

## Supplementary Material for “Bartlett Correction of Frequency Domain Empirical Likelihood for Time Series with Unknown Innovation Variance”

Kun Chen<sup>a</sup>, Ngai Hang Chan<sup>b</sup> and Chun Yip Yau<sup>b</sup>

<sup>a</sup> *Southwestern University of Finance and Economics*

<sup>b</sup> *The Chinese University of Hong Kong*

In this supplementary material, we give proofs of to the Lemmas 1–2 and Theorems 1–2. First, we begin by deriving the stochastic expansion of the profiled frequency domain empirical likelihood (FDEL).

### Stochastic Expansion

To prove Lemma 2 and Theorems 1–2, we need the stochastic expansion of the profiled FDEL firstly. Based on the Lagrange multiplier argument, the profiled FDEL is

$$\tilde{l}_n(\beta_0) = 4 \sum_{j=1}^n \log(1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)), \quad (1)$$

where  $\tilde{t}$  is the solution of the nonlinear equation

$$\frac{1}{n} \sum_{j=1}^n \frac{\tilde{m}_j(\beta_0)}{1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)} = 0. \quad (2)$$

To establish the stochastic expansion of the profiled FDEL  $\tilde{l}(\beta_0)$ , we first prove that the magnitude  $\tilde{t}(\beta_0) = O_p(n^{-1/2})$ . As  $w_j = n^{-1}(1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0))^{-1}$ , we have  $1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0) \geq 0$  and

$$|1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)|^{-1} \geq (1 + |\tilde{t}(\beta_0)| \max_j |\tilde{m}_j(\beta_0)|)^{-1}. \quad (3)$$

Substituting  $1/(1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)) = 1 - \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)/(1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0))$  into the (2) gives

$$\begin{aligned} 0 &= \left| \frac{1}{n} \sum_{j=1}^n \frac{\tilde{t}(\beta_0) \tilde{m}_j^2(\beta_0)}{1 + \tilde{t}(\beta_0) \tilde{m}_j(\beta_0)} - \frac{1}{n} \sum_{j=1}^n \tilde{m}_j(\beta_0) \right| \\ &\geq |\tilde{t}(\beta_0)| \frac{1}{n} \sum_{j=1}^n \frac{\tilde{m}_j^2(\beta_0)}{1 + |\tilde{t}(\beta_0)| \max_{1 \leq j \leq n} |\tilde{m}_j(\beta_0)|} - \left| \frac{1}{n} \sum_{j=1}^n \tilde{m}_j(\beta_0) \right|, \end{aligned}$$

using the (3). Therefore, we have

$$|\tilde{t}(\beta_0)| \left( \frac{1}{n} \sum_{j=1}^n \tilde{m}_j^2(\beta_0) - \max_{1 \leq j \leq n} |\tilde{m}_j(\beta_0)| |\bar{\tilde{m}}| \right) \leq |\bar{\tilde{m}}|.$$

Under **RC**, the standard argument in Owen (2001) gives  $\max_{1 \leq j \leq n} |\tilde{m}_j(\beta_0)| = o_p(n^{1/2})$ . Applying the results of Lemma 1, we have  $\bar{\tilde{m}} = O_p(n^{-1/2})$  and  $\frac{1}{n} \sum_j \tilde{m}_j^2(\beta_0) = O_p(1)$ . Hence,

$$\tilde{t}(\beta_0) = O_p(n^{-1/2}).$$

Applying Taylor's expansion to the (2), we get the stochastic expansion of  $\tilde{t}(\beta_0)$ ,

$$\tilde{t}(\beta_0) = \frac{\bar{m}}{\lambda_2} - \frac{\bar{m}\Delta_2}{\lambda_2^2} + \frac{\lambda_3\bar{m}^2}{\lambda_2^3} + \frac{\bar{m}\Delta_2^2}{\lambda_2^3} - 3\frac{\lambda_3\bar{m}^2\Delta_2}{\lambda_2^4} + 2\frac{\lambda_3^2\bar{m}^3}{\lambda_2^5} + \frac{\bar{m}^2\Delta_3}{\lambda_2^3} - \frac{\lambda_4\bar{m}^3}{\lambda_2^4} + O_p(n^{-2}). \quad (4)$$

Substituting the (4) into the (1) leads to the stochastic expansion of  $\tilde{l}(\beta_0)$ ,

$$\begin{aligned} \frac{1}{n}\tilde{l}_n(\beta_0) &= 2\frac{\bar{m}^2}{\lambda_2} - 2\frac{\bar{m}^2\Delta_2}{\lambda_2^2} + \frac{4}{3}\frac{\lambda_3\bar{m}^3}{\lambda_2^3} + 2\frac{\bar{m}^2\Delta_2^2}{\lambda_2^3} + \frac{4}{3}\frac{\bar{m}^3\Delta_3}{\lambda_2^3} - 4\frac{\lambda_3\bar{m}^3\Delta_2}{\lambda_2^4} \\ &\quad + 2\frac{\lambda_3^2\bar{m}^4}{\lambda_2^5} - \frac{\lambda_4\bar{m}^4}{\lambda_2^4} + O_p(n^{-5/2}). \end{aligned} \quad (5)$$

By equating the terms in the (5) to the terms of  $\tilde{S}R^2$  of the same orders, we obtain the (12).

### Proofs of Lemmas 1–2 and Theorems 1–2

In this section, we give the proofs of our lemmas and theorems.

*Proof of Lemma 1:* From the definition,

$$\begin{aligned} \lambda_4 &= \frac{1}{n} \sum_{j=1}^n \text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) \\ &\quad + \frac{4}{n} \sum_{j=1}^n \text{cum}(\tilde{m}_j(\beta_0)) \text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) \\ &\quad + \frac{3}{n} \sum_{j=1}^n \text{cum}^2(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) + \frac{6}{n} \sum_{j=1}^n \text{cum}^2(\tilde{m}_j(\beta_0)) \text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) \\ &\quad + \frac{1}{n} \sum_{j=1}^n \text{cum}^4(\tilde{m}_j(\beta_0)) + O(n^{-1}). \end{aligned} \quad (6)$$

Note that this expansion for  $\lambda_4$  is not the same as the one with known variance  $\sigma_\epsilon^2$  given in Chen *et al.* (2016) because the extra terms with  $\text{cum}(\tilde{m}_j(\beta_0))$  cannot be ignored. Based on a calculation similar to that given in Lemma 2 of Chen *et al.* (2016), we have

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) = 6\sigma_\epsilon^8 p_j(\beta_0)^4 + O(n^{-1}), \quad (7)$$

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) = 2\sigma_\epsilon^6 p_j(\beta_0)^3 + O(n^{-1}), \quad (8)$$

and

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0)) = \sigma_\epsilon^4 p_j(\beta_0)^2 + O(n^{-1}). \quad (9)$$

Substituting the equations (9) and (7)–(8) into the (6), we have

$$\lambda_4 = \frac{24}{n} \sigma_\epsilon^8 \sum_{j=1}^n p_j(\beta_0)^4 + O(n^{-1}).$$

Applying similar arguments to those used to derive  $\lambda_4$ , it can be shown that

$$\lambda_2 = \frac{2}{n} \sigma_\epsilon^4 \sum_{j=1}^n p_j(\beta_0)^2 + O(n^{-1}),$$

$$\lambda_3 = \frac{6}{n} \sigma_\epsilon^6 \sum_{j=1}^n p_j(\beta_0)^3 + O(n^{-1}).$$

Next, we consider the orders of  $\text{Var}(\Delta_k)$ . By Chebyshev’s inequality, for any  $\delta > 0$ ,  $P(|\Delta_k| \geq \delta) \leq E(\Delta_k^2)/\delta^2$ , and as  $E(\Delta_k) = 0$ , the orders of  $\Delta_k$  can be directly obtained by computing the orders of  $\text{Var}(\Delta_k)$ . The case  $k = 1$  is omitted because it is simpler. For  $k = 2$ ,

$$\begin{aligned} \text{Var} \left( \frac{1}{n} \sum_{j=1}^n \tilde{m}_j(\beta_0)^2 \right) &= \frac{1}{n^2} \sum_{j=1}^n \sum_{k=1}^n [\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0), \tilde{m}_k(\beta_0)) \\ &\quad + 2\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0))\text{cum}(\tilde{m}_k(\beta_0)) \\ &\quad + 2\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0), \tilde{m}_k(\beta_0))\text{cum}(\tilde{m}_j(\beta_0)) \\ &\quad + 4\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0))\text{cum}(\tilde{m}_j(\beta_0))\text{cum}(\tilde{m}_k(\beta_0)) \\ &\quad + 2\text{cum}^2(\tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0))] + O(n^{-1}). \end{aligned} \quad (10)$$

For the double summation in the (10), the terms corresponding to  $j = k$  can be collected as

$$\frac{1}{n^2} \sum_{j=1}^n \text{Var}(\tilde{m}_j(\beta_0)) = \frac{20}{n^2} \sigma_\epsilon^8 \sum_{j=1}^n p_j(\beta_0)^4 + O(n^{-2}), \quad (11)$$

using the equations (9) and (7)–(9).

By applying an argument similar to that used in the second equality of Lemma 2 in Chen *et al.* (2016), for  $j \neq k$ , we have

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0)) = \frac{1}{n} \kappa_{\epsilon,4} p_j(\beta_0) p_k(\beta_0) + O(n^{-2}), \quad (12)$$

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0)) = \frac{2}{n} \kappa_{\epsilon,4} \sigma_\epsilon^2 p_j(\beta_0)^2 p_k(\beta_0) + O(n^{-2}), \quad (13)$$

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0), \tilde{m}_k(\beta_0)) = \frac{2}{n} \kappa_{\epsilon,4} \sigma_\epsilon^2 p_j(\beta_0) p_k(\beta_0)^2 + O(n^{-2}), \quad (14)$$

and

$$\text{cum}(\tilde{m}_j(\beta_0), \tilde{m}_j(\beta_0), \tilde{m}_k(\beta_0), \tilde{m}_k(\beta_0)) = \frac{4}{n} \kappa_{\epsilon,4} \sigma_\epsilon^4 p_j(\beta_0)^2 p_k(\beta_0)^2 + O(n^{-2}), \quad (15)$$

where the  $O(n^{-2})$  terms are uniform in  $j, k$ .

Substituting the equations (9) and (11)–(15) into the (10) yields

$$\text{Var}(\Delta_2) = \frac{20}{n^2} \sigma_\epsilon^8 \sum_{j=1}^n p_j(\beta_0)^4 + \frac{16}{n^3} \kappa_{\epsilon,4} \sigma_\epsilon^4 \sum_{j \neq k} p_j(\beta_0)^2 p_k(\beta_0)^2 + O(n^{-2}).$$

By applying similar but more tedious calculations to  $\Delta_3$  and  $\Delta_4$ , we have

$$\text{Var}(\Delta_3) = \frac{K_3}{n^2} \sigma_\epsilon^{12} \sum_{j=1}^n p_j(\beta_0)^6 + \frac{K'_3}{n^3} \kappa_{\epsilon,4} \sigma_\epsilon^8 \sum_{j \neq k} p_j(\beta_0)^3 p_k(\beta_0)^3 + O(n^{-2}),$$

$$\text{Var}(\Delta_4) = \frac{K_4}{n^2} \sigma_\epsilon^{16} \sum_{j=1}^n p_j(\beta_0)^8 + \frac{K'_4}{n^3} \kappa_{\epsilon,4} \sigma_\epsilon^{12} \sum_{j \neq k} p_j(\beta_0)^4 p_k(\beta_0)^4 + O(n^{-2}),$$

where  $K_3, K_4, K'_3$  and  $K'_4$  are some constants. This completes the proof of Lemma 1.

*Proof of Lemma 2:* To derive  $\kappa_1(\sqrt{n}\tilde{S}\tilde{R})$ , first note that

$$\text{cum}(\bar{m}) = \frac{1}{n^2} \sum_{j=1}^n \frac{p_j(\beta_0)}{g_j(\beta_0)} b_I(\omega_j) + O(n^{-2}), \quad (16)$$

$$\text{cum}(\bar{m}^2) = \text{cum}(\bar{m}, \bar{m}) + (\text{cum}(\bar{m}))^2 = \frac{1}{n} \frac{1}{2} \lambda_2 + O(n^{-2}), \quad (17)$$

and

$$\begin{aligned} \text{cum}(\bar{m}\Delta_2) &= \text{cum}(\bar{m}, \Delta_2) \\ &= \frac{1}{n^2} \sum_{j=1}^n \frac{p_j(\beta_0)^3}{g_j(\beta_0)^3} [\text{cum}(I_n(\omega_j), I_n(\omega_j), I_n(\omega_j)) \\ &\quad + 2\text{cum}(I_n(\omega_j), I_n(\omega_j))\text{cum}(I_n(\omega_j))] + O(n^{-2}) \\ &= \frac{4\sigma_\epsilon^6}{n^2} \sum_{j=1}^n p_j(\beta_0)^3 + O(n^{-2}) = \frac{2}{3} \frac{\lambda_3}{n} + O(n^{-2}). \end{aligned} \quad (18)$$

By using the equations (16)–(18), we have

$$\begin{aligned} \kappa_1(\sqrt{n}\tilde{S}\tilde{R}) &= \text{cum}(\sqrt{n}(\tilde{R}_1 + \tilde{R}_2 + \tilde{R}_3 + \tilde{R}_n)) \\ &= \sqrt{n} \text{cum} \left\{ \sqrt{2} \frac{\bar{m}}{\lambda_2^{1/2}} + \frac{\sqrt{2}}{3} \frac{\lambda_3 \bar{m}^2}{\lambda_2^{5/2}} - \frac{\sqrt{2}}{2} \frac{\bar{m}\Delta_2}{\lambda_2^{3/2}} \right\} + O(n^{-2}) \\ &= \frac{1}{n^{3/2}} \frac{\sqrt{2}}{\lambda_2^{1/2}} \sum_{j=1}^n \frac{b_I(\omega_j)}{g_j(\beta_0)} p_j(\beta_0) - \frac{1}{n^{1/2}} \frac{\sqrt{2}}{6} \frac{\lambda_3}{\lambda_2^{3/2}} + O(n^{-3/2}). \end{aligned}$$

Hence, the (13) is established. For  $\kappa_2(\sqrt{n}\tilde{S}\tilde{R})$ , consider

$$\begin{aligned} \kappa_2(\sqrt{n}\tilde{S}\tilde{R}) &= \text{cum}(\sqrt{n}\tilde{S}\tilde{R}, \sqrt{n}\tilde{S}\tilde{R}) \\ &= n \left\{ \text{cum}(\tilde{R}_1, \tilde{R}_1) + 2\text{cum}(\tilde{R}_1, \tilde{R}_2) + 2\text{cum}(\tilde{R}_1, \tilde{R}_3) + \text{cum}(\tilde{R}_2, \tilde{R}_2) \right\} \\ &\quad + O(n^{-2}). \end{aligned} \quad (19)$$

Using the equations (7)–(9) and (12)–(15), we have

$$\begin{aligned} \text{cum}(\tilde{R}_1, \tilde{R}_1) &= \text{cum} \left( \sqrt{2} \frac{\bar{m}}{\lambda_2^{1/2}}, \sqrt{2} \frac{\bar{m}}{\lambda_2^{1/2}} \right) = \frac{1}{n} + \frac{4\sigma_\epsilon^2}{\lambda_2} \frac{1}{n^3} \sum_{j=1}^n \frac{p_j(\beta_0)^2}{g_j(\beta_0)} b_I(\omega_j) \\ &\quad + \frac{2}{\lambda_2} \frac{1}{n^2} \sum_{j \neq k} \frac{p_j(\beta_0)p_k(\beta_0)}{g_j(\beta_0)g_k(\beta_0)} \text{cum}(I_n(\omega_j), I_n(\omega_k)) + O(n^{-3}), \end{aligned} \quad (20)$$

$$\begin{aligned} \text{cum}(\tilde{R}_1, \tilde{R}_2) &= \frac{2}{3} \frac{\lambda_3}{\lambda_2^3} \text{cum}(\bar{m}, \bar{m}^2) - \frac{1}{\lambda_2^2} \text{cum}(\bar{m}, \bar{m}\Delta_2) \\ &= \frac{1}{n^2} \left\{ \frac{2}{9} \frac{\lambda_3^2}{\lambda_2^3} - \frac{1}{2} \frac{\lambda_4}{\lambda_2^2} - 2 \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-3}), \end{aligned} \quad (21)$$

$$\begin{aligned}
\text{cum}(\tilde{R}_1, \tilde{R}_3) &= \frac{3}{4} \frac{1}{\lambda_2^3} \text{cum}(\bar{m}, \bar{m} \Delta_2^2) + \frac{2}{3} \frac{1}{\lambda_2^3} \text{cum}(\bar{m}, \bar{m}^2 \Delta_3) - \frac{5}{3} \frac{\lambda_3}{\lambda_2^4} \text{cum}(\bar{m}, \bar{m}^2 \Delta_2) \\
&\quad + \frac{8}{9} \frac{\lambda_3^2}{\lambda_2^5} \text{cum}(\bar{m}, \bar{m}^3) - \frac{1}{2} \frac{\lambda_4}{\lambda_2^4} \text{cum}(\bar{m}, \bar{m}^3) \\
&= \frac{1}{n^2} \left\{ \frac{11}{16} \frac{\lambda_4}{\lambda_2^2} - \frac{1}{3} \frac{\lambda_3^2}{\lambda_2^3} + \frac{3}{2} \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-3}),
\end{aligned} \tag{22}$$

and

$$\begin{aligned}
\text{cum}(\tilde{R}_2, \tilde{R}_2) &= \frac{2}{9} \frac{\lambda_3^2}{\lambda_2^5} \text{cum}(\bar{m}^2, \bar{m}^2) - \frac{2}{3} \frac{\lambda_3}{\lambda_2^4} \text{cum}(\bar{m} \Delta_2, \bar{m}^2) + \frac{1}{2} \frac{1}{\lambda_2^3} \text{cum}(\bar{m} \Delta_2, \bar{m} \Delta_2) \\
&= \frac{1}{n^2} \left\{ -\frac{1}{9} \frac{\lambda_3^2}{\lambda_2^3} + \frac{5}{24} \frac{\lambda_4}{\lambda_2^2} + \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-3}).
\end{aligned} \tag{23}$$

Substituting the equations (20)–(23) into the (19) yields

$$\begin{aligned}
\kappa_2(\sqrt{n} \tilde{S}R) &= 1 + \frac{1}{n} \left\{ -\frac{1}{3} \frac{\lambda_3^2}{\lambda_2^3} + \frac{7}{12} \frac{\lambda_4}{\lambda_2^2} + \frac{4\sigma_\epsilon^2}{\lambda_2} \frac{1}{n} \sum_{j=1}^n \frac{p_j(\beta_0)^2}{g_j(\beta_0)} b_I(\omega_j) \right. \\
&\quad \left. + \frac{2}{\lambda_2} \sum_{j \neq k} \frac{p_j(\beta_0) p_k(\beta_0)}{g_j(\beta_0) g_k(\beta_0)} \text{cum}(I_n(\omega_j), I_n(\omega_k)) \right\} + O(n^{-2}).
\end{aligned}$$

For  $\kappa_3(\sqrt{n} \tilde{S}R)$ , note that

$$\begin{aligned}
\kappa_3(\sqrt{n} \tilde{S}R) &= \text{cum}(\sqrt{n} \tilde{S}R, \sqrt{n} \tilde{S}R, \sqrt{n} \tilde{S}R) \\
&= n^{3/2} \{ \text{cum}(\tilde{R}_1, \tilde{R}_1, \tilde{R}_1) + 3 \text{cum}(\tilde{R}_1, \tilde{R}_1, \tilde{R}_2) \} + O(n^{-3/2}).
\end{aligned} \tag{24}$$

Using the equations (7)–(9) and (12)–(15), we have

$$\text{cum}(\tilde{R}_1, \tilde{R}_1, \tilde{R}_1) = \frac{1}{n^2} \sqrt{2} \frac{\lambda_3}{\lambda_2^{3/2}} + O(n^{-3}). \tag{25}$$

For the second term on the right-hand side of the (24), it can be shown that

$$3 \text{cum}(\tilde{R}_1, \tilde{R}_1, \tilde{R}_2) = 2\sqrt{2} \frac{\lambda_3}{\lambda_2^{7/2}} \text{cum}(\bar{m}, \bar{m}, \bar{m}^2) - 3\sqrt{2} \frac{1}{\lambda_2^{5/2}} \text{cum}(\bar{m}, \bar{m}, \bar{m} \Delta_2).$$

Again using the equations (7)–(9) and (12)–(15), we have

$$\begin{aligned}
\text{cum}(\bar{m}, \bar{m}, \bar{m}^2) &= \frac{\lambda_2^2}{2n^2} + O(n^{-3}), \\
\text{cum}(\bar{m}, \bar{m}, \bar{m} \Delta_2) &= \frac{2}{3n^2} \lambda_2 \lambda_3 + O(n^{-3}).
\end{aligned}$$

Thus, we have

$$3 \text{cum}(\tilde{R}_1, \tilde{R}_1, \tilde{R}_2) = -\frac{\sqrt{2}}{n^2} \frac{\lambda_3}{\lambda_2^{3/2}} + O(n^{-3}). \tag{26}$$

Substituting the equations (25) and (26) into the (24), we have  $\kappa_3(\sqrt{n}\tilde{S}R) = O(n^{-3/2})$ .

For  $\kappa_4(\sqrt{n}\tilde{S}R)$ , we have

$$\begin{aligned}\kappa_4(\sqrt{n}\tilde{S}R) &= \text{cum}(\sqrt{n}\tilde{S}R, \sqrt{n}\tilde{S}R, \sqrt{n}\tilde{S}R, \sqrt{n}\tilde{S}R) \\ &= \text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1) + 4\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_2) \\ &\quad + 4\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_3) + 6\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_2, \sqrt{n}\tilde{R}_2) \\ &\quad + O(n^{-2}).\end{aligned}\tag{27}$$

Using the equations (7)–(9) and (12)–(15), it can be shown that

$$\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1) = \frac{1}{n} \left\{ \frac{\lambda_4}{\lambda_2^2} + 12 \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-2}),\tag{28}$$

$$\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_2) = \frac{1}{n} \left\{ -\frac{3}{2} \frac{\lambda_4}{\lambda_2^2} - 6 \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-2}),\tag{29}$$

$$\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_3) = \frac{1}{n} \frac{3}{4} \frac{\lambda_4}{\lambda_2^2} + O(n^{-2}),\tag{30}$$

and

$$\text{cum}(\sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_1, \sqrt{n}\tilde{R}_2, \sqrt{n}\tilde{R}_2) = \frac{1}{n} \left\{ \frac{1}{3} \frac{\lambda_4}{\lambda_2^2} + 2 \frac{\kappa_{\epsilon,4}}{\sigma_\epsilon^4} \right\} + O(n^{-2}).\tag{31}$$

Substituting the equations (28)–(31) into the (27), we have  $\kappa_4(\sqrt{n}\tilde{S}R) = O(n^{-2})$ . This completes the proof of Lemma 2.

*Proof of Theorem 1:* The proof is standard and can be found in Chan *et al.* (2014).  $\square$

*Proof of Theorem 2:* Based on the cumulant expansions in Lemma 2,

$$\begin{aligned}\mathbb{E}(\tilde{l}(\beta_0)) = \mathbb{E}(n\tilde{S}R^2) &= \kappa_2(\sqrt{n}\tilde{S}R) + \kappa_1(\sqrt{n}\tilde{S}R)^2 = 1 + \frac{\tilde{k}_{1,1}^2 + \tilde{k}_{2,2}}{n} + O(n^{-2}) \\ &= 1 + \frac{\tilde{b}}{n} + O(n^{-2}).\end{aligned}$$

Next, we derive the coverage error of the Bartlett-corrected confidence interval based on the Edgeworth expansion of  $\sqrt{n}\tilde{S}R^* := \sqrt{n}\tilde{S}R(1 - \tilde{b}/2n)$ . Specifically,

$$\begin{aligned}P(\tilde{l}(\beta_0) \leq (1 + \tilde{b}/n)\chi_{1,1-\alpha}^2) &= P((\sqrt{n}\tilde{S}R^*)^2 + O_p(n^{-2}) \leq \chi_{1,1-\alpha}^2) \\ &= \int_{-\sqrt{\chi_{1,1-\alpha}^2}}^{\sqrt{\chi_{1,1-\alpha}^2}} \phi(x) dx + \int_{-\sqrt{\chi_{1,1-\alpha}^2}}^{\sqrt{\chi_{1,1-\alpha}^2}} \left\{ \frac{\tilde{r}_1^*(x)}{\sqrt{n}} + \frac{\tilde{r}_2^*(x)}{n} + \frac{\tilde{r}_3^*(x)}{n^{3/2}} \right\} \phi(x) dx + O(n^{-2}) \\ &= 1 - \alpha + O(n^{-2}),\end{aligned}\tag{32}$$

where

$$\begin{aligned}\tilde{r}_1^*(x) &= \sqrt{n} \left\{ \kappa_1(\sqrt{n}\tilde{S}R^*)x + \frac{1}{6} \kappa_3(\sqrt{n}\tilde{S}R^*)(x^3 - 3x) \right\}, \\ \tilde{r}_2^*(x) &= \frac{n}{2} \left\{ \kappa_2(\sqrt{n}\tilde{S}R^*) - 1 + \kappa_1(\sqrt{n}\tilde{S}R^*) \right\} (x^2 - 1).\end{aligned}$$

The last equality in the (32) holds because the integral of  $\tilde{r}_2^*(x)$  equals  $O(n^{-1})$  based on the equations (13) and (14). Moreover, the integrals of terms with order  $n^{-1/2}$  and  $n^{-3/2}$  are equal to zero because of the oddness of polynomials  $\tilde{r}_1^*(x)$  and  $\tilde{r}_3^*(x)$ . This completes the proof of Theorem 2.

*Proof of Lemma 3:* It is sufficient to show that  $\hat{k}_{1,1} = \tilde{k}_{1,1} + O_p(n^{-1/2})$  and  $\hat{k}_{2,2} = \tilde{k}_{2,2} + O_p(n^{-1/2})$ . Note that the convergences hold point-wise on the interior of the compact parameter space  $\Theta$ . Under **RC**, the  $\sqrt{n}$ -consistency of  $\hat{\beta}$  implies  $\hat{\lambda}_k = \lambda_k + O_p(n^{-1/2})$ ,  $k = 2, 3, 4$ . Thus,  $\hat{k}_{2,2} = \tilde{k}_{2,2} + O_p(n^{-1/2})$ . For  $\hat{k}_{1,1}$ , we argue that  $\hat{b}_I(\omega) = b_I(\omega) + O_p(n^{-1/2})$ , where the error is uniform in  $\omega \in \Pi$ . We consider the decomposition  $|b_I(\omega) - \hat{b}_I(\omega)| \leq T_1(\omega) + T_2(\omega)$ , where

$$T_1(\omega) = \left| b_I(\omega) + \frac{1}{2\pi} \sum_{u=-\infty}^{\infty} |u| \gamma_{\hat{\theta}}(u) \exp(iu\omega) \right|,$$

$$T_2(\omega) = \left| -\frac{1}{2\pi} \sum_{u=-\infty}^{\infty} |u| \gamma_{\hat{\theta}}(u) \exp(iu\omega) + \frac{1}{2\pi} \sum_{u=-[n/2]+1}^{[n/2]+1} |u| \gamma_{\hat{\theta}}(u) \exp(iu\omega) \right|.$$

Next, we prove that  $\sup_{\omega \in \Pi} T_1(\omega) = O_p(n^{-1/2})$ . By applying the Taylor expansion, we have

$$\begin{aligned} \sup_{\omega \in \Pi} T_1(\omega) &\leq \frac{1}{2\pi} \sup_{\omega \in \Pi} \sum_{u=-\infty}^{\infty} |u| |\gamma_{\hat{\theta}}(u) - \gamma_{\theta_0}(u)| |\exp(iu\omega)| \\ &\leq \frac{1}{2\pi} \sum_{u=-\infty}^{\infty} |u| |\gamma_{\hat{\theta}}(u) - \gamma_{\theta_0}(u)| \\ &= \frac{1}{2\pi} \sum_{u=-\infty}^{\infty} |u| |(\hat{\theta} - \theta_0)' \frac{\partial}{\partial \theta} \gamma_{\theta^*}(u)| \\ &\leq \frac{1}{2\pi} \|\hat{\theta} - \theta_0\| \sum_{u=-\infty}^{\infty} |u| \left\| \frac{\partial}{\partial \theta} \gamma_{\theta^*}(u) \right\| \\ &= O_p(n^{-1/2}), \end{aligned}$$

where  $\|\theta^* - \theta_0\| \leq \|\hat{\theta} - \theta_0\|$  and the last equality follows from Regularity Condition (a) and  $\|\hat{\theta} - \theta_0\| = O_p(n^{-1/2})$ . As  $\sup_{\omega \in \Pi} T_2(\omega) \leq (2\pi)^{-1} \sum_{u=[n/2]+1}^{\infty} |u| |\gamma_{\hat{\theta}}(u)| = o_p(n^{-1/2})$ , it follows that  $\hat{b}_I(\omega) = b_I(\omega) + O_p(n^{-1/2})$ , where the error term is uniform in  $\omega$ . Hence,  $\hat{k}_{1,1} = \tilde{k}_{1,1} + O_p(n^{-1/2})$ . Under Condition (b), by using the similar arguments as above, we can show that  $\widehat{\text{cum}}(I_n(\omega_j), I_n(\omega_k)) = \text{cum}(I_n(\omega_j), I_n(\omega_k)) + O_p(n^{-1/2})$  and thus  $\hat{k}_{2,2} = \tilde{k}_{2,2} + O_p(n^{-1/2})$ .

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