



Correction to: On the strong universal consistency of local averaging regression estimates

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There is a gap at the end of the proof of Theorem 1, since there the application of the conditional McDiarmid inequality yields

$$J_n - \mathbf{E}\{J_n | X_1, \dots, X_n\} \rightarrow 0 \quad a.s.,$$

where $J_n = \int \left| \sum_{i=1}^n W_{n,i}(x) \cdot (Y_i - m(X_i)) \right| \mu(dx)$, and not yet the assertion

$$J_n \rightarrow 0 \quad a.s.$$

in the last step of the proof of Theorem 1.

This gap can be filled by adding into assumption (A3) the second condition

$$\sum_{i=1}^n \int |W_{n,i}(x)|^2 \mu(dx) \rightarrow 0 \quad a.s. \quad (29)$$

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Using this condition together with $|Y| \leq L$ a.s., it is easy to see that one has

$$\mathbf{E}\{J_n|X_1, \dots, X_n\} \rightarrow 0 \text{ a.s.},$$

which is still needed to obtain the assertion.

In order to verify (29) in the applications of Theorem 1, for kernel estimation in the context of Lemma 6 one notices that, up to some constant factor, the left-hand side of (29) is majorized by

$$\int \frac{1}{1 + \sum_{i=1}^n I_{S_{r_1}}\left(\frac{x-X_i}{h_n}\right)} \mu(dx),$$

which can be treated similarly to the verification of (A4) in Lemma 6. The verification of (29) for partitioning estimation in the context of Lemma 9 is analogous.

Details

Last part of the proof of Theorem 1. It remains to show

$$J_n \cdot I_{B_n} \rightarrow 0 \text{ a.s.}$$

Application of the conditional McDiarmid inequality as in the proof of Theorem 1 yields

$$J_n \cdot I_{B_n} - \mathbf{E}\{J_n \cdot I_{B_n}|X_1, \dots, X_n\} \rightarrow 0 \text{ a.s.}$$

Hence, it suffices to show

$$\mathbf{E}\{J_n|X_1, \dots, X_n\} \rightarrow 0 \text{ a.s.} \quad (30)$$

By the inequality of Jensen, the independence of the data and $|Y| \leq L$ a.s., we get

$$\begin{aligned} & (\mathbf{E}\{J_n|X_1, \dots, X_n\})^2 \\ & \leq \mathbf{E}\{J_n^2|X_1, \dots, X_n\} \\ & \leq \mathbf{E} \left\{ \int \left| \sum_{i=1}^n W_{n,i}(x) \cdot (Y_i - m(X_i)) \right|^2 \mu(dx) \middle| X_1, \dots, X_n \right\} \\ & = \mathbf{E} \left\{ \left| \sum_{i=1}^n W_{n,i}(X) \cdot (Y_i - m(X_i)) \right|^2 \middle| X_1, \dots, X_n \right\} \\ & = \mathbf{E} \left\{ \mathbf{E} \left\{ \left| \sum_{i=1}^n W_{n,i}(X) \cdot (Y_i - m(X_i)) \right|^2 \middle| X, X_1, \dots, X_n \right\} \middle| X_1, \dots, X_n \right\} \end{aligned}$$

$$\begin{aligned}
 &= \mathbf{E} \left\{ \sum_{i=1}^n W_{n,i}(X)^2 \cdot \mathbf{E} \left\{ (Y_i - m(X_i))^2 \middle| X, X_1, \dots, X_n \right\} \middle| X_1, \dots, X_n \right\} \\
 &\leq 4L^2 \cdot \mathbf{E} \left\{ \sum_{i=1}^n W_{n,i}(X)^2 \middle| X_1, \dots, X_n \right\} \\
 &= 4L^2 \cdot \sum_{i=1}^n \int |W_{n,i}(x)|^2 \mu(dx).
 \end{aligned}$$

Thus, (30) follows from (29).

Proof of (29) in the context of Lemma 6. On the one hand, we have

$$\sum_{i=1}^n W_{n,i}(x)^2 = \frac{\sum_{i=1}^n K\left(\frac{x-X_i}{h_n}\right)^2}{\left(\sum_{j=1}^n K\left(\frac{x-X_j}{h_n}\right)\right)^2} \leq 1.$$

On the other hand, it holds

$$\begin{aligned}
 \sum_{i=1}^n W_{n,i}(x)^2 &\leq c_2 \cdot \frac{\sum_{i=1}^n K\left(\frac{x-X_i}{h_n}\right)}{\left(\sum_{j=1}^n K\left(\frac{x-X_j}{h_n}\right)\right)^2} \cdot I_{\left\{\sum_{j=1}^n K\left(\frac{x-X_j}{h_n}\right) > 0\right\}} \\
 &\leq c_2 \cdot \frac{1}{\sum_{j=1}^n K\left(\frac{x-X_j}{h_n}\right)}.
 \end{aligned}$$

Consequently,

$$\begin{aligned}
 \sum_{i=1}^n W_{n,i}(x)^2 &\leq \min \left\{ 1, c_2 \cdot \frac{1}{\sum_{j=1}^n K\left(\frac{x-X_j}{h_n}\right)} \right\} \\
 &\leq \min \left\{ 1, \frac{c_2}{c_1} \cdot \frac{1}{\sum_{j=1}^n I_{S_{r_1}}\left(\frac{x-X_j}{h_n}\right)} \right\} \\
 &\leq \max \left\{ 1, \frac{c_2}{c_1} \right\} \cdot \min \left\{ 1, \frac{1}{\sum_{j=1}^n I_{S_{r_1}}\left(\frac{x-X_j}{h_n}\right)} \right\} \\
 &\leq \max \left\{ 1, \frac{c_2}{c_1} \right\} \cdot \frac{2}{1 + \sum_{j=1}^n I_{S_{r_1}}\left(\frac{x-X_j}{h_n}\right)}.
 \end{aligned}$$

Hence, it suffices to show

$$W_n := \int \frac{1}{1 + \sum_{j=1}^n I_{S_{r_1}}\left(\frac{x-X_j}{h_n}\right)} \mu(dx) \rightarrow 0 \quad a.s. \tag{31}$$

For any bounded sphere S around 0, by Lemma 2a and by assumption (9), we get

$$\begin{aligned} & \mathbf{E} \left\{ \int_S \frac{1}{1 + \sum_{j=1}^n I_{S_{r_1}} \left(\frac{x - X_j}{h_n} \right)} \mu(dx) \right\} \\ &= \int_S \mathbf{E} \left\{ \frac{1}{1 + \sum_{j=1}^n I_{S_{r_1}} \left(\frac{x - X_j}{h_n} \right)} \right\} \mu(dx) \\ &\leq \int_S \frac{1}{n \cdot \mu(x + h_n \cdot S_{r_1})} \mu(dx) \\ &\leq \frac{const}{n \cdot h_n^d} \rightarrow 0 \quad (n \rightarrow \infty), \end{aligned}$$

where the last inequality holds because of equation (5.1) in Györfi et al. (2002). Thus, it suffices to show

$$W_n - \mathbf{E}\{W_n\} \rightarrow 0 \quad a.s. \tag{32}$$

Analogously to the proof of (A4), with X'_1, X_1, \dots, X_n independent and identically distributed and

$$W'_n := \int \frac{1}{1 + I_{S_{r_1}} \left(\frac{x - X'_1}{h_n} \right) + \sum_{j=2}^n I_{S_{r_1}} \left(\frac{x - X_j}{h_n} \right)} \mu(dx),$$

by Lemma 4.2 in Kohler et al. (2003), one has

$$\mathbf{E}\{|W_n - \mathbf{E}\{W_n\}|^4\} \leq c_{11} \cdot n^2 \cdot \mathbf{E}\{(W_n - W'_n)^4\} \quad (n \in \mathbb{N}).$$

Furthermore, by the second part of Lemma 5 one gets

$$\begin{aligned} & \mathbf{E}\{|W_n - W'_n|^4\} \\ &\leq 16 \cdot \mathbf{E} \left\{ \left(\int \frac{I_{S_{r_1}} \left(\frac{x - X_1}{h_n} \right)}{\left(1 + \sum_{j=2}^n I_{S_{r_1}} \left(\frac{x - X_j}{h_n} \right) \right)^2} \mu(dx) \right)^4 \right\} \\ &\leq 16 \cdot \mathbf{E} \left\{ \left(\int \frac{I_{S_{r_1}} \left(\frac{x - X_1}{h_n} \right)}{1 + \sum_{j=2}^n I_{S_{r_1}} \left(\frac{x - X_j}{h_n} \right)} \mu(dx) \right)^4 \right\} \\ &\leq \frac{const}{n^4}. \end{aligned}$$

From these relations, one obtains (32) by the Borel–Cantelli lemma and the Markov inequality.

Proof of (29) in the context of Lemma 9. Analogously to above it suffices to show

$$V_n := \int \frac{1}{1 + \sum_{j=1}^n I_{A_{\mathcal{P}_n}(x)}(X_j)} \mu(dx) \rightarrow 0 \quad a.s.$$

For any bounded sphere S around zero, by assumption (12) we get

$$\int_S \frac{1}{n \cdot \mu(A_{\mathcal{P}_n}(x))} \mu(dx) \rightarrow 0 \quad (n \rightarrow \infty),$$

from which by Lemma 2a we can conclude analogously to above

$$\mathbf{E}V_n \rightarrow 0 \quad (n \rightarrow \infty).$$

Hence, it suffices to show

$$V_n - \mathbf{E}\{V_n\} \rightarrow 0 \quad a.s.,$$

which follows analogously to above from the second part of Lemma 7.