

## INEQUALITIES FOR ORDERED SUMS

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### Summary

Let  $x_i = y_i + z_i$ ,  $i = 1, \dots, n$ , and write  $x_{(1)} \leq \dots \leq x_{(n)}$ , with corresponding notation for the ordered  $y_i$  and  $z_i$ . It is shown, for example, that  $x_{(r)} \geq \max_{i=1, \dots, r} (y_{(i)} + z_{(r+1-i)})$ ,  $r = 1, \dots, n$ . Inequalities are also obtained for convex (or concave) functions of the  $x_{(i)}$ . The results lead immediately to bounds for the expected values of order statistics in nonstandard situations in terms of simpler expectations. A small numerical example illustrates the method.

### 1. Introduction

Let  $X_{(1)} \leq \dots \leq X_{(n)}$  be the order statistics formed from random variables  $X_1, \dots, X_n$ . Smith and Tong [6] have developed inequalities for convex (or concave) functions of the  $X_{(i)}$ ,  $i = 1, \dots, n$ , when  $X_i$  is expressible as a sum of two other random variables,  $X_i = Y_i + Z_i$ . The  $Y_i$  are not necessarily independent or identically distributed, nor are the  $Z_i$ . In applications the  $Z_i$  are often constants,  $Z_i = \delta_i$ ,  $i = 1, \dots, n$ . The inequalities are in terms of the ordered  $Y_i$ ,  $Z_i$ , or  $\delta_i$ , and are useful whenever, for example,  $E Y_{(i)}$  and  $E Z_{(i)}$  can be handled more easily than  $E X_{(i)}$ .

Of the order statistics only the maximum is convex (and the minimum concave). In this note we derive a simple inequality (Theorem 1) that holds for order statistics of any rank. We also strengthen one of the results of Smith and Tong [6] for convex functions of order statistics. Our results hold for any numbers  $x_i$ ,  $y_i$ ,  $z_i$  linked by  $x_i = y_i + z_i$ ,  $i = 1, \dots, n$ . Applications to bounds for the  $E X_{(i)}$  are immediate. A small numerical example illustrates our methods on a normal sample with an unidentified outlier and permits some comparisons with the inequalities in Mallows and Richter [3], Arnold and Groeneveld [1], and Nagaraja [5].

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## 2. Inequalities for ordered sums

We begin with an inequality for an ordered  $x$ -value in terms of the ordered  $y$  and  $z$  values, where  $x_i = y_i + z_i$ ,  $i = 1, \dots, n$ . It will be convenient to write the ordered  $x_i$  in either ascending or descending order:

$$(1a, 1b) \quad x_{(1)} \leq \dots \leq x_{(n)} \quad \text{or} \quad x_{[1]} \geq \dots \geq x_{[n]},$$

with corresponding notation for the ordered  $y_i$  and  $z_i$ . Then clearly

$$(2a, 2b) \quad x_{[1]} \leq y_{[1]} + z_{[1]}, \quad x_{(1)} \geq y_{(1)} + z_{(1)},$$

and, with range  $x_i = x_{[1]} - x_{(1)}$ , etc., one has

$$(3) \quad \text{range } x_i \leq \text{range } y_i + \text{range } z_i.$$

The results (2) may be generalized as follows:

**THEOREM 1.** *If  $x_i = y_i + z_i$ ,  $i = 1, \dots, n$ , then with the notation of (1) we have for  $r = 1, \dots, n$*

$$(4a) \quad x_{[r]} \leq \min_{i=1, \dots, r} (y_{[i]} + z_{[r+1-i]})$$

and

$$(4b) \quad x_{(r)} \geq \max_{i=1, \dots, r} (y_{(i)} + z_{(r+1-i)}).$$

**PROOF.** For some  $i$  in  $\{1, \dots, r\}$  and some  $j$  in  $\{1, \dots, n\}$  suppose that  $x_j > y_{[i]} + z_{[r+1-i]}$ . A necessary condition for this to hold is that either  $y_j > y_{[i]}$  or  $z_j > z_{[r+1-i]}$ . Thus at most  $(i-1) + (r-i) = r-1$  of the  $x_j$  can exceed  $y_{[i]} + z_{[r+1-i]}$ , i.e.,  $x_{[r]} \leq y_{[i]} + z_{[r+1-i]}$ ,  $i = 1, \dots, r$ , which is (4a). The proof of (4b) is similar.

*Comment 1.* Equations (4a) and (4b) may equivalently be stated as

$$\max_{i=1, \dots, r} (y_{(i)} + z_{(r+1-i)}) \leq x_{(r)} \leq \min_{j=1, \dots, n+1-r} (y_{(n+1-j)} + z_{(r-1+j)}).$$

*Comment 2.* The bounds in (4) are attainable. In (4a) equality in the form  $x_{[r]} = y_{[i]} + z_{[r+1-i]}$  is achieved if and only if there exists  $j$  such that  $y_j = y_{[i]}$  and  $z_j = z_{[r+1-i]}$ . For example,  $x_{[1]} = y_{[1]} + z_{[1]}$  if there exists  $j$  such that  $y_j = y_{[1]}$  and  $z_j = z_{[1]}$ ; in words, equality is achieved if  $y_j$  and  $z_j$  are the largest or, in case of ties, one of the largest  $y$ 's and  $z$ 's, respectively.

Frequently the sum  $S_k = \sum_{i=1}^k x_{[i]}$  of the  $k$  largest  $x_i$  is of interest,  $k = 1, \dots, n$ . In obtaining inequalities for  $S_k$  one can do better than merely add the first  $k$  inequalities of (4). Thus, it is obvious that

$$(5a) \quad S_k \leq \sum_{i=1}^k (y_{[i]} + z_{[i]}) .$$

To deal with inequalities in the other direction note that from (4b)

$$S_1 \equiv x_{(n)} \geq \max_{i=1, \dots, n} (y_{(i)} + z_{(n+1-i)}) .$$

Let  $x'_{(i)}$ ,  $i=1, \dots, n$  denote the  $n$  sums  $y_{(i)} + z_{(n+1-i)}$  arranged in ascending order of magnitude. Then  $x'_{(n)}$  is an attainable lower bound for  $x_{(n)}$ . If  $x_{(n)} = x'_{(n)}$ , with  $x'_{(n)} = y_{(h)} + z_{(n+1-h)}$  (say),  $h=1, \dots, n$ , we note that  $x'_{(n-1)}$  is an attainable lower bound for  $x_{(n-1)}$ , where here

$$x'_{(n-1)} = \max_{\substack{j=1, \dots, n \\ j \neq i}} (y_{(j)} + z_{(n+1-j)}) .$$

Repeating the process we have the sharp inequality

$$(5b) \quad S_k = \sum_{i=n+1-k}^n x'_{(i)} .$$

The results (5) may be extended from  $S_k$  to a convex linear function  $l$  of ordered  $x_i$ , viz.  $l = \sum_{i=1}^n c_i x_{(i)}$  with  $c_1 \leq \dots \leq c_n$ . To see this, note that

$$l = c_1 S_n + (c_2 - c_1) S_{n-1} + \dots + (c_n - c_{n-1}) S_1$$

so that from (5) we have easily for  $c_1 \geq 0$

$$(6) \quad \sum_{i=1}^n c_i x'_{(i)} \leq l \leq \sum_{i=1}^n c_i (y_{(i)} + z_{(i)}) .$$

If  $c_m < 0$ ,  $c_{m+1} \geq 0$ ,  $m=1, \dots, n$ , equation (6) continues to hold since  $l$  may be split into

$$\sum_{i=1}^m c_i x_{(i)} + \sum_{i=m+1}^n c_i x_{(i)}$$

and the lower end counterparts of (5) applied to the first sum.

The right hand inequality of (6) is essentially the same as that in Theorem 2.1 of Smith and Tong [6], where majorization arguments (e.g., Marshall and Olkin [4]) are used. Our lower bound in (6) is, however, superior to theirs, viz.  $\sum c_i (y_{(i)} + z_{[i]})$ .

Clearly it is also possible to generalize (3) to

$$\sum_{i=1}^m d_i (x_{[i]} - x_{(i)}) \leq \sum_{i=1}^m d_i (y_{[i]} - y_{(i)}) + \sum_{i=1}^m d_i (z_{[i]} - z_{(i)}) ,$$

where  $d_1 \geq \dots \geq d_m$  and  $m=n/2$  or  $(n+1)/2$  according as  $n$  is even or odd.

Finally, generalizations to the case  $x_i = \sum_{j=1}^p x_{ij}$ ,  $i=1, \dots, n$  and  $p$  a

positive integer ( $p > 1$ ), are straightforward. For example, in obvious notation, (4a) generalizes to

$$x_{[r]} \leq \min \sum_{j=1}^p x_{[i_j]j},$$

where the minimum is taken over all positive integers  $i_j$  for which  $\sum_{i=1}^p i_j = r + p - 1$ .

### 3. Bounds for the expectations of order statistics in nonstandard situations

Let  $X_i = Y_i + Z_i$ ,  $i = 1, \dots, n$ , where the  $Y_i$  and  $Z_i$  are random variables. Then provided only that  $E Y_i$  and  $E Z_i$  exist,  $i = 1, \dots, n$ , previous results can immediately be converted into corresponding inequalities between expectations. For example, (4b) becomes

$$E X_{(r)} \geq E \max_{i=1, \dots, r} (Y_{(i)} + Z_{(r+1-i)}) \geq \max_{i=1, \dots, r} (E Y_{(i)} + E Z_{(r+1-i)}),$$

the last step following from Jensen's inequality.

It should be noted that no assumptions of independence and common distributions are needed in the above but some simplifying assumptions will usually be invoked in applications. A case of special interest occurs when the  $Y_i$  are iid and the  $Z_i$  are constants, say  $Z_i = \delta_i$ ,  $i = 1, \dots, n$ . We consider an example of this kind. Other applications are given in Smith and Tong [6].

*Sample with one outlier.* Let  $Y_i$ ,  $i = 1, \dots, n$ , be iid,  $Z_i = \delta > 0$  for some unknown value of  $i$  and  $Z_i = 0$  otherwise. Then the  $X_i$  are a sample with an unidentified outlier.

From (4) we have for  $r = 1, \dots, n-1$

$$(7a) \quad E Y_{(r)} \leq E X_{(r)} \leq \min (E Y_{(r+1)}, E Y_{(r)} + \delta),$$

and, for  $r = n$ ,

$$(7b) \quad \max (E Y_{(n)}, E Y_{(1)} + \delta) \leq E X_{(n)} \leq E Y_{(n)} + \delta.$$

As a small numerical illustration we take the  $Y_i$  to be independent normal  $N(0, 1)$ ,  $n = 5$ , and  $\delta = 2$ . Results are presented in Table 1. In this case it is possible to obtain exact values for  $E X_{(r)}$  (David et al. [2]) but the bounds can easily be calculated for general patterns of the  $\delta_i$  in the case of any distribution for which  $E Y_{(r)}$ , the expected values of the order statistics, are available.

In our example the bounds for  $E X_{(5)}$  are rather wide and may be

Table 1. Expected value,  $E X_{(r)}$ , of  $r$ -th order statistic in a standard normal sample of size 5 in the presence of an outlier with mean  $\delta=2$ 

$r$	$E Y_{(r)}$	$E X_{(r)}$		
		lower bound(7)	exact value	upper bound(7)
1	-1.1630	-1.1630	-1.0316	-0.4950
2	-0.4950	-0.4950	-0.3054	0
3	0	0	0.2698	0.4950
4	0.4950	0.4950	0.9167	1.1630
5	1.1630	1.1630	2.1504	3.1630

improved as follows :

- (a)  $E X_{(5)} \geq 2$  since  $E X_{(n)} \geq \delta$ , a result that follows, for example, by a majorization argument (Marshall and Olkin [4], p. 348).
- (b)  $E X_{(5)} \leq 2.8$  since  $E X_{(n)} \leq E \bar{X} + (n-1)(E S^2/n)^{1/2}$ , where  $(n-1)S^2 = \sum (X_i - \bar{X})^2$  (Nagaraja [5]);  $E \bar{X} = \delta/n$  and  $E S^2 = 1 + \delta^2/n$ .

It will be clear that while methods (a) and (b) lead to better bounds for  $E X_{(5)}$  in the present instance, they will not necessarily do so. Both these approaches are confined to  $X_{(n)}$  or other convex functions such as  $\sum_{i=n+1-k}^n X_{(i)}$  (with corresponding results for lower extremes). For order statistics that are not extremes our methods tend to give much better bounds than other approaches. Note that, as a result of the outlier, the expected value of the sample median has increased from 0 to 0.2698 with bounds  $[0, 0.4950]$ ; corresponding figures for the trimmed mean based on the three central order statistics are 0.2937 and  $[0, 0.5527]$ .

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