

# ASYMPTOTIC CONSISTENCY OF FIXED-WIDTH SEQUENTIAL CONFIDENCE INTERVALS FOR A MULTIPLE REGRESSION FUNCTION

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## Summary

Let  $m_n(x)$  be the recursive kernel estimator of the multiple regression function  $m(x) = E[Y|X=x]$ . For given  $\alpha$  ( $0 < \alpha < 1$ ) and  $d > 0$  we define a certain class of stopping times  $N = N(\alpha, d, x)$  and take  $I_{N,d}(x) = [m_N(x) - d, m_N(x) + d]$  as a  $2d$ -width confidence interval for  $m(x)$  at a given point  $x$ . In this paper it is shown that the probability  $P\{m(x) \in I_{N,d}(x)\}$  converges to  $\alpha$  as  $d$  tends to zero.

## 1. Introduction

Let  $Z = (X, Y)$ ,  $Z_1 = (X_1, Y_1)$ ,  $\dots$ ,  $Z_n = (X_n, Y_n)$  be independent and identically distributed  $R^p \times R$ -valued random vectors on a probability space  $(\Omega, \mathcal{B}, P)$  with an unknown joint probability density function (p.d.f.)  $f^*(x, y)$  with respect to the Lebesgue measure. There have been many papers on the estimation of the nonparametric regression function  $m(x) = E[Y|X=x]$  (of  $Y$  on  $X$ ) by

$$(1.1) \quad m_n(x) = \sum_{i=1}^n W_{ni}(x) Y_i,$$

where  $W_{ni}(x) = W_{ni}(x, X_1, \dots, X_n)$  for each  $i$  ( $1 \leq i \leq n$ ) is a suitable real-valued Borel measurable function of  $x, X_1, \dots, X_n$ .

Nadaraya [8] and Watson [14] proposed the estimator (1.1) with  $p = 1$  and

$$W_{ni}(x) = K((x - X_i)/h_n) / \sum_{j=1}^n K((x - X_j)/h_n),$$

where  $K(x)$  is a suitable kernel function and  $\{h_n\}$  is a sequence of window-widths tending to zero. Later on, many authors have studied

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its asymptotic properties (see Prakasa Rao [10] for example).

When data increase we may be faced with computational burdens in processing them. To decrease these burdens Ahmad and Lin [1] proposed a recursive version of (1.1) with

$$W_{ni}(x) = h_i^{-p} K((x - X_i)/h_i) \Big/ \sum_{j=1}^n h_j^{-p} K((x - X_j)/h_j),$$

or equivalently,

$$\begin{aligned} m_0(x) &= f_0(x) \equiv 0 \\ (1.2) \quad f_n(x) &= (h_n/h_{n-1})^p f_{n-1}(x) + K((x - X_n)/h_n) \\ m_n(x) &= m_{n-1}(x) + f_n^{-1}(x) \{ Y_n - m_{n-1}(x) \} K((x - X_n)/h_n) \end{aligned}$$

and they proved some pointwise results for these estimators. Devroye and Wagner [4] considered a still simpler recursive estimator than (1.2). The author [6] proposed a class of recursive estimators  $m_n(x)$  defined in Section 2 and used in this paper, which contains (1.2) as a special case. Stone [13] has investigated the estimator (1.1) and discussed about sufficient conditions on  $\{W_{ni}(x)\}$  for  $m_n(x)$  to be consistent.

On the other hand, when one uses a recursive estimator in practical situations one may be required to terminate the computations to obtain the estimator with given accuracy. In this case the sample size is a random variable. Suppose that  $N_t$  for each  $t \in (0, \infty)$  is a stopping time. Recently, Samanta [12] has shown the asymptotic normality of  $m_{N_t}(x)$  by using the estimator (1.2).

In this paper we propose a class of stopping times  $N = N(\alpha, d, x)$  based on the idea of Chow and Robbins [3], construct a sequence of  $2d$ -width sequential confidence intervals  $I_{N,d}(x) = [m_N(x) - d, m_N(x) + d]$  for  $m(x)$  and show that the probability  $P\{m(x) \in I_{N,d}(x)\}$  converges to  $\alpha$  as  $d$  tends to zero.

In Section 2 we shall make some preparations and give several lemmas. In Section 3 we shall prove the asymptotic consistency of the sequential confidence intervals  $I_{N,d}(x)$ .

## 2. Preliminaries and several lemmas

In this section we shall make some preparations for Section 3.

Let  $K(x)$  be a given bounded p.d.f. on  $R^p$  with respect to the Lebesgue measure satisfying  $\|u\|_p^p K(u) \rightarrow 0$  as  $\|u\|_p \rightarrow \infty$ , where  $\|\cdot\|_p$  denotes the Euclidean norm on  $R^p$ . We shall impose either of the following conditions on  $K(x)$ :

$$(K1) \quad \int_{R^p} u_i K(u_1, \dots, u_p) du_1 \cdots du_p = 0 \quad \text{for all } i=1, \dots, p, \text{ and}$$

$$(K2) \quad \int_{R^p} \|u\|_p^2 K(u) du < \infty.$$

Let  $\{h_n\}$  be a nonincreasing sequence of positive numbers converging to zero, on which some of the following conditions are imposed:

$$(H1) \quad nh_n^p \uparrow \infty \quad \text{as } n \rightarrow \infty;$$

$$(H2) \quad \text{For some } a \ (0 < a \leq 1)$$

$$n^{1-2a} h_n^p \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

$$n^{1-2a} h_n^p \sum_{j=1}^n j^{2(a-1)} h_j^{-p} \rightarrow \beta \quad \text{as } n \rightarrow \infty \text{ for some positive constant } \beta,$$

$$n^{3/2-3a} h_n^{3p/2} \sum_{j=1}^n j^{3(a-1)} h_j^{-2p} \rightarrow 0 \quad \text{as } n \rightarrow \infty, \text{ and}$$

$$n^{1/2-a} h_n^{p/2} \sum_{j=1}^n j^{a-1} h_j^2 \rightarrow 0 \quad \text{as } n \rightarrow \infty;$$

$$(H3) \quad \text{For any } \varepsilon \ (>0) \text{ there exists a positive constant } \delta \text{ such that } |n/m-1| < \delta \text{ implies } |h_n/h_m-1| < \varepsilon;$$

$$(H4) \quad \sum_{n=1}^{\infty} (n^2 h_n^p)^{-1} < \infty;$$

$$(H5) \quad n^{1+\gamma} h_n^{p+4} \rightarrow 0 \quad \text{as } n \rightarrow \infty \text{ for some positive constant } \gamma.$$

Throughout this paper we use the following class of recursive estimators  $m_n(x)$  of the regression function  $m(x)$ , which is proposed by the author [6]:

$$m_0(x) \equiv 0, \quad f_0(x) \equiv c \quad \text{for an arbitrary positive constant } c$$

$$f_n(x) = f_{n-1}(x) + a_n \{K_n(x, X_n) - f_{n-1}(x)\}$$

$$m_n(x) = m_{n-1}(x) + a_n G(f_n(x)) \{Y_n - m_{n-1}(x)\} K_n(x, X_n)$$

for each  $n \geq 1$ , where

$$(2.1) \quad a_n = a/n \quad \text{with } 0 < a \leq 1,$$

$$G(y) = \begin{cases} y^{-1} & \text{if } y > 0 \\ 0 & \text{otherwise,} \end{cases}$$

and

$$K_n(x, s) = h_n^{-p} K((x-s)/h_n) \quad \text{for } x, s \in R^p.$$

In this paper, if in some term its denominator is less than or equal to zero we define the value of the term to be zero. Let

$$f(x) = \int_R f^*(x, y) dy, \quad q(x) = \int_R y f^*(x, y) dy, \quad g(x) = \int_R y^2 f^*(x, y) dy,$$

$$\phi(x) = \int_R y^4 f^*(x, y) dy \quad \text{and} \quad v(x) = (g(x)/f(x)) - (q(x)/f(x))^2 \quad (\geq 0).$$

Clearly,  $v(x) > 0$  is equivalent to  $f(x) > 0$  and  $v(x) > 0$ . Throughout this paper we assume that  $f(x)$ ,  $q(x)$ ,  $g(x)$  and  $\phi(x)$  are all finite on  $R^p$  and write  $m(x) = q(x)/f(x)$ . Also, let  $C$ ,  $C_1$ ,  $C_2$ ,  $\dots$  denote appropriate positive constants.

Define sequences  $\{q_n(x)\}$ ,  $\{g_n(x)\}$  and  $\{v_n(x)\}$  as follows:

$$\begin{aligned} q_0(x) &= g_0(x) \equiv 0, \\ q_n(x) &= q_{n-1}(x) + a_n \{Q_n(x, Z_n) - q_{n-1}(x)\} \\ g_n(x) &= g_{n-1}(x) + a_n \{G_n(x, Z_n) - g_{n-1}(x)\} \\ v_n(x) &= (g_n(x)/f_n(x)) - (q_n(x)/f_n(x))^2 \end{aligned}$$

for each  $n \geq 1$ , where for  $x \in R^p$ ,  $z = (u, y) \in R^p \times R$  and  $n \geq 1$

$$Q_n(x, z) = y K_n(x, u) \quad \text{and} \quad G_n(x, z) = y^2 K_n(x, u).$$

For  $a_n$  in (2.1) set

$$\gamma_0 = \gamma_1 = 1, \quad \gamma_n = \prod_{j=2}^n (1 - a_j) \quad \text{for } n \geq 2 \quad \text{and}$$

$$\beta_{mn} = \begin{cases} \prod_{j=m+1}^n (1 - a_j) & \text{if } n > m \geq 0 \\ 1 & \text{if } n = m \geq 0. \end{cases}$$

Obviously,  $\gamma_n \downarrow 0$  as  $n \rightarrow \infty$ ,  $\gamma_n > 0$  for all  $n \geq 0$  and

$$(2.2) \quad \beta_{mn} = \gamma_m^{-1} \gamma_n \quad \text{if } n \geq m \geq 1.$$

It is known that

$$(2.3) \quad \beta_{mn} \sim m^a n^{-a} \quad \text{as } n \geq m \rightarrow \infty, \quad \text{and}$$

$$(2.4) \quad C_1 n^{-a} \leq \gamma_n \leq C_2 n^{-a} \quad \text{for all } n \geq 1,$$

where " $\phi_n \sim \psi_n$  as  $n \rightarrow \infty$ " means  $\lim_{n \rightarrow \infty} \phi_n / \psi_n = 1$ .

*Remark 2.1.* We can write  $f_n(x)$ ,  $q_n(x)$  and  $g_n(x)$  as follows:

$$f_n(x) = \sum_{j=1}^n a_j \beta_{jn} K_j(x, X_j) + \beta_{0n} c,$$

$$q_n(x) = \sum_{j=1}^n a_j \beta_{jn} Q_j(x, Z_j) \quad \text{and} \quad g_n(x) = \sum_{j=1}^n a_j \beta_{jn} G_j(x, Z_j),$$

where  $\sum_{j=m}^n (\cdot) = 0$  if  $n < m$ .

For any real-valued function  $\theta$  on  $R^p$  let  $C(\theta)$  be the set of continuity points of  $\theta$  and  $\|\theta\|_\infty = \sup \{|\theta(x)| : x \in R^p\}$ . For any fixed  $x \in R^p$  and  $n \geq 1$  let

$$U_n^{(0)} = K_n(x, X_n) - E[K_n(x, X_n)], \quad U_n^{(1)} = Q_n(x, Z_n) - E[Q_n(x, Z_n)],$$

$$W_n = a_n \gamma_n^{-1} (U_n^{(0)}, U_n^{(1)})', \quad B_n = (nh_n^p)^{1/2} \gamma_n \sum_{j=1}^n W_j \quad \text{and}$$

$$B_n^* = (nh_n^p)^{1/2} (f_n(x) - f(x), q_n(x) - q(x))',$$

where the prime denotes transpose.

We shall summarize some results obtained in [5] and [6].

LEMMA 2.1 ([6]). *Let  $\{d_n\}$  be a sequence of positive numbers converging to zero. Let  $k(x)$  be a bounded, integrable, real-valued Borel measurable function on  $R^p$  satisfying  $\|u\|_p^p |k(u)| \rightarrow 0$  as  $\|u\|_p \rightarrow \infty$ . Let  $\theta(x)$  be an integrable, real-valued Borel measurable function on  $R^p$ . Then, for each point  $x \in C(\theta)$  we have*

$$\int_{R^p} d_n^{-p} k((x-u)/d_n) \theta(u) du \rightarrow \theta(x) \int_{R^p} k(u) du \quad \text{as } n \rightarrow \infty \quad \text{and}$$

$$\sup_{n \geq 1} \int_{R^p} d_n^{-p} |k((x-u)/d_n)| |\theta(u)| du \leq C,$$

where  $C$  may depend on  $x$ .

LEMMA 2.2 ([6]). *Assume  $E[Y^2] < \infty$ . Let (H4) be satisfied, and let  $x$  be a point with  $f(x) > 0$ , belonging to the set  $C(f) \cap C(q) \cap C(g)$ . Then we have*

$$\lim_{n \rightarrow \infty} f_n(x) = f(x) \quad \text{with probability one (w.p.1.)},$$

$$\lim_{n \rightarrow \infty} q_n(x) = q(x) \quad \text{w.p.1.}, \quad \text{and} \quad \lim_{n \rightarrow \infty} m_n(x) = m(x) \quad \text{w.p.1.}$$

LEMMA 2.3 ([5]). *Let  $\{y_n\}$  be a sequence of random variables on a probability space  $(\Omega, \mathcal{F}, P)$ . Suppose that there exists a null set  $A$  such that for each  $\omega \in A^c$*

$$y_n(\omega) \geq 0 \quad \text{for all } n \geq 1, \quad \lim_{n \rightarrow \infty} y_n(\omega) = 1 \quad \text{and}$$

$$y_n(\omega) > 0 \quad \text{for all } n \geq m \quad \text{if } y_m(\omega) > 0 \quad \text{for some } m = m(\omega).$$

Let  $\{b(n)\}$  be a sequence of positive numbers satisfying

$$\lim_{n \rightarrow \infty} b(n) = \infty \quad \text{and} \quad \lim_{n \rightarrow \infty} b(n)/b(n-1) = 1.$$

For each  $t \in (0, \infty)$  define  $N_t$  as the smallest integer  $n \geq 1$  such that  $b(n)/t \geq y_n > 0$ . Then, we have

$$\begin{aligned} P\{N_t < +\infty\} &= 1 \quad \text{for each } t \in (0, \infty), \quad P\{N_t \uparrow \infty \text{ as } t \rightarrow \infty\} = 1 \quad \text{and} \\ P\{b(N_t)/t \rightarrow 1 \text{ as } t \rightarrow \infty\} &= 1. \end{aligned}$$

The following lemma gives the asymptotic normality of  $B_n$ .

**LEMMA 2.4.** Assume  $E[Y^4] < \infty$ . Let (H2) be satisfied. Consider a point  $x \in C(f) \cap C(q) \cap C(g)$  with  $v(x) > 0$ . If either  $x \in C(\phi)$  or  $\|\phi\|_\infty < \infty$  holds, then we have

$$B_n \xrightarrow{L} N(0, \Gamma) \quad \text{as } n \rightarrow \infty \text{ (in law),}$$

where the covariance matrix  $\Gamma = \Gamma(x)$  is given by

$$\Gamma = a^2 \beta \int_{\mathbb{R}^p} K^2(u) du \begin{pmatrix} f(x) & q(x) \\ q(x) & g(x) \end{pmatrix}.$$

**PROOF.** By Hölder's inequality we get that for each  $j \geq 1$

$$\begin{aligned} E[|U_j^{(1)}|^8] &\leq 8 E[|Q_j(x, Z_j)|^8] \\ &= 8 \int_{\mathbb{R}^p \times \mathbb{R}} \{|y|^8 K_j^2(x, u) (f^*(u, y))^{3/4}\} \{K_j(x, u) (f^*(u, y))^{1/4}\} du dy \\ &\leq 8 \left[ \int_{\mathbb{R}^p \times \mathbb{R}} y^4 \{K_j(x, u)\}^{8/3} f^*(u, y) du dy \right]^{3/4} \\ &\quad \times \left[ \int_{\mathbb{R}^p \times \mathbb{R}} K_j^4(x, u) f^*(u, y) du dy \right]^{1/4} = 8 h_j^{-2p} \times I_1 \times I_2, \end{aligned}$$

where

$$\begin{aligned} I_1 &= \left[ \int_{\mathbb{R}^p} h_j^{-p} \{K((x-u)/h_j)\}^{8/3} \phi(u) du \right]^{3/4} \quad \text{and} \\ I_2 &= \left[ \int_{\mathbb{R}^p} h_j^{-p} K^4((x-u)/h_j) f(u) du \right]^{1/4}. \end{aligned}$$

As  $x \in C(f)$ , it follows from Lemma 2.1 that  $I_2 \leq C_1$ . If  $x \in C(\phi)$  then by Lemma 2.1 we get  $I_1 \leq C_2$ , and if  $\|\phi\|_\infty < \infty$  then we have

$$I_1 \leq \{\|\phi\|_\infty \|K\|_\infty^{5/3}\}^{3/4} < \infty.$$

In any case,  $I_1 \times I_2$  is bounded by a constant and therefore, we obtain

$$(2.5) \quad E[|U_j^{(1)}|^8] \leq C_3 h_j^{-2p} \quad \text{for all } j \geq 1.$$

Replacing (4.17) in [6] by (2.5) we can use the same arguments as in

the proof of Lemma 4.1 of [6], to complete the proof of Lemma 2.4.

LEMMA 2.5. Let  $\{h_n\}$  be a nonincreasing sequence of positive numbers converging to zero and satisfy (H1) and (H2). Let  $\{V_n\}$  be a sequence of independent random variables on a probability space  $(\Omega, \mathcal{F}, P)$  satisfying

$$E[V_n] = 0 \quad \text{and} \quad h_n^2 E[V_n^2] \leq C \quad \text{for all } n \geq 1.$$

For any  $d \in (0, \infty)$  let  $N(d)$  be a positive integer-valued random variable on  $(\Omega, \mathcal{F}, P)$  and  $n(d)$  a positive integer with  $\lim_{d \rightarrow 0} n(d) = \infty$ . Set

$$T_n = \sum_{j=1}^n a_j \beta_{jn} V_j \quad \text{for each } n \geq 1,$$

where  $a_n$  is given in (2.1). If

$$N(d)/n(d) \xrightarrow{P} 1 \quad \text{as } d \rightarrow 0,$$

then we have

$$(N(d)h_{N(d)}^2)^{1/2}(T_{N(d)} - T_{n(d)}) \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0.$$

PROOF. For simplicity let  $N = N(d)$  and  $n = n(d)$ . Let any positive numbers  $\varepsilon$  and  $\xi$  be fixed. For any  $\rho (> 0)$  set

$$M_1 = [(1-\rho)n] \quad \text{and} \quad M_2 = [(1+\rho)n],$$

where  $[b]$  denotes the largest integer not greater than  $b$ . Since as  $\rho \rightarrow 0$

$$2\rho/(1-\rho) \rightarrow 0 \quad \text{and} \quad \{(1-\rho)/(1+\rho)\}^a \rightarrow 1,$$

there exists a positive constant  $\rho = \rho(\varepsilon, \xi) < 1/2$  such that

$$(2.6) \quad 2\rho/(1-\rho) < \varepsilon^2 \xi / 2 \quad \text{and} \quad (1 - \{(1-\rho)/(1+\rho)\}^a)^2 < \varepsilon^2 \xi / 2.$$

As  $\lim_{d \rightarrow 0} n = \infty$  we get  $M_i \rightarrow \infty$  as  $d \rightarrow 0$  for  $i=1, 2$  and  $M_1/M_2 \sim (1-\rho)/(1+\rho)$  as  $d \rightarrow 0$ . Hence by (2.3) and (2.6) we have

$$(2.7) \quad (1 - \beta_{M_1, M_2})^2 < \varepsilon^2 \xi \quad \text{for } d \text{ sufficiently small.}$$

Also, since  $(M_2 - M_1)/M_1 \sim 2\rho/(1-\rho)$  as  $d \rightarrow 0$ , it follows from (2.6) that

$$(2.8) \quad (M_2 - M_1)/M_1 < \varepsilon^2 \xi \quad \text{for } d \text{ sufficiently small.}$$

It is clear that

$$(2.9) \quad M_2/M_1 \leq C_1 \quad \text{for all } d > 0.$$

Now, we shall prove the lemma. By assumption we get

$$(2.10) \quad P\{|N/n-1|\geq \rho\} < \xi \quad \text{for } d \text{ sufficiently small.}$$

Put

$$I = P\{(Nh_N^p)^{1/2}|T_N - T_n| \geq \varepsilon, |N-n| < \rho n\}.$$

Then

$$\begin{aligned} (2.11) \quad I &\leq P\{(ih_i^p)^{1/2}|T_i - T_n| \geq \varepsilon \text{ for some } i \in (M_1, M_2)\} \\ &\leq P\{(ih_i^p)^{1/2}(|T_i - T_{M_1}| + |T_n - T_{M_1}|) \geq \varepsilon \text{ for some } i \in (M_1, M_2)\} \\ &\leq P\left\{\max_{M_1 < i \leq M_2} (ih_i^p)^{1/2} \times 2 \max_{M_1 < i \leq M_2} |T_i - T_{M_1}| \geq \varepsilon\right\} \\ &= P\left\{\max_{M_1 < i \leq M_2} (ih_i^p)^{1/2} \times \max_{M_1 < i \leq M_2} |T_i - T_{M_1}| \geq \varepsilon/2\right\}. \end{aligned}$$

By (2.2) and the definition of  $T_n$  we have that for  $i \in (M_1, M_2]$

$$\begin{aligned} |T_i - T_{M_1}| &\leq \left| \sum_{j=1}^{M_1} a_j(\beta_{ji} - \beta_{jM_1})V_j \right| + \left| \sum_{j=M_1+1}^i a_j\beta_{ji}V_i \right| \\ &\leq (\gamma_{M_1} - \gamma_i) \left| \sum_{j=1}^{M_1} a_j\gamma_j^{-1}V_j \right| + \gamma_i \left| \sum_{j=M_1+1}^i a_j\gamma_j^{-1}V_j \right|, \end{aligned}$$

which, together with (2.11) and the monotonicity of  $h_n$ ,  $nh_n^p$  and  $\gamma_n$ , implies that

$$\begin{aligned} (2.12) \quad I &\leq P\left\{(M_2h_{M_1}^p)^{1/2}(\gamma_{M_1} - \gamma_{M_2}) \left| \sum_{j=1}^{M_1} a_j\gamma_j^{-1}V_j \right| \right. \\ &\quad \left. + (M_2h_{M_2}^p)^{1/2} \max_{M_1 < i \leq M_2} \gamma_i \left| \sum_{j=M_1+1}^i a_j\gamma_j^{-1}V_j \right| \geq \varepsilon/2\right\} \leq J_1 + J_2, \end{aligned}$$

where

$$\begin{aligned} J_1 &= P\left\{(M_2h_{M_1}^p)^{1/2}(\gamma_{M_1} - \gamma_{M_2}) \left| \sum_{j=1}^{M_1} a_j\gamma_j^{-1}V_j \right| \geq \varepsilon/4\right\} \quad \text{and} \\ J_2 &= P\left\{(M_2h_{M_2}^p)^{1/2} \max_{M_1 < i \leq M_2} \gamma_i \left| \sum_{j=M_1+1}^i a_j\gamma_j^{-1}V_j \right| \geq \varepsilon/4\right\}. \end{aligned}$$

From (H2), (2.4) and assumption we get

$$ih_i^p \sum_{j=1}^i a_j^2 \beta_{ji}^2 E[V_j^2] \leq C_2 i^{1-2a} h_i^p \sum_{j=1}^i j^{2(a-1)} h_j^{-p} \leq C_3 \quad \text{for all } i \geq 1.$$

Hence by Chebychev's inequality, (2.2), (2.7) and (2.9) we obtain

$$(2.13) \quad J_1 \leq C_4 \varepsilon^{-2} (1 - \beta_{M_1 M_2})^2 (M_2/M_1) M_1 h_{M_1}^p \sum_{j=1}^{M_1} a_j^2 \beta_{jM_1}^2 E[V_j^2] \leq C_5 \xi$$

for  $d$  sufficiently small. From assumption, the Hájek-Rényi inequality (see Petrov [9], page 51), the monotonicity of  $h_n$ , (2.8) and (2.9) we have that for  $d$  sufficiently small



$$\begin{aligned}
 (2.14) \quad J_2 &\leq C_6 \varepsilon^{-2} M_2 h_{M_2}^2 \sum_{j=M_1+1}^{M_2} \alpha_j^2 E[V_j^2] \leq C_7 \varepsilon^{-2} M_2 \sum_{j=M_1+1}^{M_2} j^{-2} \\
 &\leq C_7 \varepsilon^{-2} M_2 M_1^{-2} (M_2 - M_1) \leq C_8 \xi.
 \end{aligned}$$

Combining (2.10), (2.12), (2.13) and (2.14) we obtain

$$P\{(Nh_N^p)^{1/2}|T_N - T_n| \geq \varepsilon\} \leq I + P\{|N - n| \geq \rho n\} \leq C_9 \xi$$

for  $d$  sufficiently small, which concludes the proof of Lemma 2.5.

### 3. Main results

In this section we shall propose sequential confidence intervals  $I_{n,d}(x)$  for  $m(x)$  with prescribed width  $2d$  and coverage probability  $\alpha$ . Then the asymptotic consistency of these confidence intervals will be shown, that is,  $P\{m(x) \in I_{N(d),d}(x)\} \rightarrow \alpha$  as  $d \rightarrow 0$ , where  $\{N(d)\}$  is the class of stopping times defined below.

Let any  $\alpha$  ( $0 < \alpha < 1$ ) be given. Define  $D = D(\alpha) > 0$  by  $\Phi(D) - \Phi(-D) = \alpha$ , where  $\Phi(\cdot)$  is the standard normal distribution function. Let  $d$  be any positive number, and let any  $x \in R^p$  be fixed. We define the stopping time  $N(d) = N(\alpha, d, x)$  as the smallest integer  $n \geq 1$  such that  $(D^2 B)^{-1} d^2 n h_n^2 \geq v_n(x)/f_n(x) > 0$ , where

$$(3.1) \quad B = \alpha^2 \beta \int_{R^p} K^2(u) du (> 0) \quad \text{with } \beta \text{ being given in (H2).}$$

Define the confidence interval  $I_{n,d}(x)$  as

$$I_{n,d}(x) = [m_n(x) - d, m_n(x) + d].$$

Also, define  $n(d) = n(\alpha, d, x)$  as the smallest integer  $n \geq 1$  such that  $(D^2 B)^{-1} d^2 n h_n^2 \geq v(x)/f(x) > 0$ .

Let  $\sigma^2(x) = Bv(x)/f(x)$  with  $B$  being given in (3.1).

The following lemma states the asymptotic properties of  $N(d)$ .

**LEMMA 3.1.** Assume  $E[Y^4] < \infty$ . Let (H1), (H3) and (H4) be satisfied. Consider a point  $x \in C(f) \cap C(g) \cap C(g)$  with  $v(x) > 0$ . If either  $x \in C(\psi)$  or  $\|\psi\|_\infty < \infty$  holds then we have

$$P\{N(d) < +\infty\} = 1 \quad \text{for each } d > 0, \quad P\{N(d) \uparrow \infty \text{ as } d \rightarrow 0\} = 1,$$

$$N(d) h_{N(d)}^2 d^2 / (D^2 \sigma^2(x)) \rightarrow 1 \quad \text{as } d \rightarrow 0 \text{ w.p.1., and}$$

$$N(d) h_{N(d)}^2 / (n(d) h_{n(d)}^2) \rightarrow 1 \quad \text{as } d \rightarrow 0 \text{ w.p.1..}$$

**PROOF.** From the definition of  $N(d)$  we get

$$(3.2) \quad N(d) = \text{smallest integer } n \geq 1 \text{ such that } b(n)/t \geq y_n > 0,$$

where

$$y_n = \frac{v_n(x)/f_n(x)}{v(x)/f(x)}, \quad b(n) = \frac{nh_n^2}{v(x)/f(x)} \quad \text{and} \quad t = D^2 B d^{-2}.$$

Clearly,

$$(3.3) \quad b(n) > 0 \quad \text{for all } n \geq 1, \quad \lim_{n \rightarrow \infty} b(n) = \infty \quad \text{and} \quad \lim_{n \rightarrow \infty} b(n)/b(n-1) = 1.$$

We shall show that

$$(3.4) \quad v_n(x) \geq 0 \quad \text{for all } n \geq 1 \text{ on } \mathcal{Q}$$

and that for any fixed  $\omega \in \mathcal{Q}$

$$(3.5) \quad v_n(x) > 0 \quad \text{for all } n \geq m \quad \text{if } v_m(x) > 0 \quad \text{for some } m = m(\omega) \geq 1.$$

For simplicity we omit  $\omega$ . By the definition of  $v_n(x)$  it suffices to consider the case where  $f_n(x) > 0$ .

That  $v_n(x) \geq (\text{resp. } >) 0$  for all  $n \geq 1$  is equivalent to

$$(3.6) \quad g_n(x)f_n(x) - q_n^2(x) \geq (\text{resp. } >) 0 \quad \text{for all } n \geq 1.$$

Let any  $\omega \in \mathcal{Q}$  be fixed. First we shall prove (3.4). By the definitions of  $f_n(x)$ ,  $q_n(x)$  and  $g_n(x)$  we have

$$(3.7) \quad A_{n+1} = (1 - a_n)^2 A_n + a_n(1 - a_n) D_n \quad \text{for each } n \geq 1,$$

where

$$A_n = g_{n-1}(x)f_{n-1}(x) - q_{n-1}^2(x) \quad \text{and}$$

$$D_n = G_n(x, z_n)f_{n-1}(x) + K_n(x, X_n)g_{n-1}(x) - 2Q_n(x, Z_n)q_{n-1}(x).$$

By Remark 2.1 we get

$$(3.8) \quad D_n = \sum_{j=1}^{n-1} a_j \beta_{j, n-1} \{G_n(x, Z_n)K_j(x, X_j) + K_n(x, X_n)G_j(x, Z_j) - 2Q_n(x, Z_n)Q_j(x, Z_j)\} + G_n(x, Z_n)\beta_{0, n-1}c$$

for each  $n \geq 1$ . From the definitions of  $K_n(x, X_n)$ ,  $Q_n(x, Z_n)$  and  $G_n(x, Z_n)$  we have

$$G_n(x, Z_n)\beta_{0, n-1}c \geq 0 \quad \text{for each } n \geq 1$$

and

$$G_n(x, Z_n)K_j(x, X_j) + K_n(x, X_n)G_j(x, Z_j) - 2Q_n(x, Z_n)Q_j(x, Z_j) \geq 0$$

for each  $j = 1, \dots, n-1$  with  $n \geq 2$ ,

which, together with (3.8), yields that

$$(3.9) \quad D_n \geq 0 \quad \text{for each } n \geq 1.$$

It follows from (3.7) and (3.9) that

$$(3.10) \quad A_{n+1} \geq (1 - a_n)^2 A_n \quad \text{for each } n \geq 1.$$

As  $A_1 = 0$ , by (3.10) and induction we have  $A_n \geq 0$  for each  $n \geq 1$ , which, together with (3.6), gives (3.4).

Next we shall prove (3.5). Assume that  $v_m(x) > 0$  for some  $m \geq 1$ . From (3.6) we get  $A_{m+1} > 0$ . Suppose that  $A_n > 0$  for some  $n > m + 1$ . Then by (2.1) and (3.10) we have  $A_{n+1} > 0$ . Hence by induction we obtain  $A_n > 0$  for all  $n \geq m + 1$ , which is equivalent to  $v_n(x) > 0$  for all  $n \geq m$ . Thus (3.5) was proved.

Next we shall show that

$$(3.11) \quad g_n(x) \rightarrow g(x) \quad \text{as } n \rightarrow \infty \text{ w.p.1.}$$

Since Lemma 2.1, the definition of  $G_n(x, Z_n)$  and  $x \in C(g)$  give  $E[G_n(x, Z_n)] \rightarrow g(x)$  as  $n \rightarrow \infty$ , it follows from Remark 2.1 and the Toeplitz lemma (see Loève [7], page 238) that

$$(3.12) \quad E[g_n(x, Z_n)] \rightarrow g(x) \quad \text{as } n \rightarrow \infty.$$

Lemma 2.1 and (H4) give

$$\sum_{n=1}^{\infty} a_n^2 E[G_n^2(x, Z_n)] < \infty,$$

which, together with Kolmogorov's convergence theorem, Kronecker's lemma, Remark 2.1 and (2.2), implies that

$$(3.13) \quad g_n(x) - E[g_n(x)] \rightarrow 0 \quad \text{as } n \rightarrow \infty \text{ w.p.1.}$$

Thus, according to (3.12) and (3.13) we obtain (3.11). Lemma 2.2 and (3.11) give

$$(3.14) \quad v_n(x) \rightarrow v(x) \quad \text{as } n \rightarrow \infty \text{ w.p.1.}$$

From Lemma 2.2, (3.4), (3.5), (3.14) and the property of  $f_n(x)$ , there exists a null set  $A$  such that for each  $\omega \in A^c$

$$y_n(\omega) \geq 0 \quad \text{for all } n \geq 1, \quad \lim_{n \rightarrow \infty} y_n(\omega) = 1 \quad \text{and}$$

$$y_n(\omega) > 0 \quad \text{for all } n \geq m \quad \text{if } y_m(\omega) > 0 \quad \text{for some } m = m(\omega) \geq 1,$$

which, together with (3.2) and (3.3), permits us to apply Lemma 2.3 to obtain the first three assertions of Lemma 3.1. Replacing  $N(d)$ ,  $v_n(x)$  and  $f_n(x)$  by  $n(d)$ ,  $v(x)$  and  $f(x)$ , respectively, we have that as  $d \rightarrow 0$   $n(d) \rightarrow \infty$  and  $n(d)h_{n(d)}^2 d^2 / (D^2 \sigma^2(x)) \rightarrow 1$ , which, together with the third assertion, implies the last assertion. This completes the proof.

*Remark 3.1.* By use of Lemmas 2.2 and 3.1, and Theorem 1 of Richter [11] we have that under all the conditions of Lemma 3.1

$$m_{N(d)}(x) \rightarrow m(x) \quad \text{as } d \rightarrow 0 \text{ w.p.1.}$$

The following theorem is one of main theorems.

**THEOREM 3.1.** Assume  $E[Y^4] < \infty$ . Let (K1), (K2) and (H1)~(H5) be satisfied. Suppose that there exist bounded, continuous second partial derivatives  $\partial^2 f(x)/\partial x_i \partial x_j$  and  $\partial^2 q(x)/\partial x_i \partial x_j$  on  $R^p$  for  $i, j=1, \dots, p$ . Consider a point  $x \in C(g)$  with  $v(x) > 0$ . Assume  $x \in C(\phi)$  or  $\|\phi\|_\infty < \infty$ . If

$$(3.15) \quad N(d)/n(d) \xrightarrow{P} 1 \quad \text{as } d \rightarrow 0$$

then we obtain

$$(N(d)h_{N(d)}^2)^{1/2}(m_{N(d)}(x) - m(x)) \xrightarrow{L} N(0, \sigma^2(x)) \quad \text{as } d \rightarrow 0.$$

**PROOF.** For simplicity put  $N=N(d)$  and  $n=n(d)$ . It follows from Lemma 3.1 that  $n \rightarrow \infty$  as  $d \rightarrow 0$ . First we shall show that

$$(3.16) \quad B_N \xrightarrow{L} N(0, \Gamma) \quad \text{as } d \rightarrow 0,$$

where  $\Gamma$  is given in Lemma 2.4. From Lemma 2.4 and the Cramér-Wold theorem (see Billingsley [2], page 49) we get

$$(3.17) \quad D'B_n \xrightarrow{L} N(0, D'\Gamma D) \quad \text{as } d \rightarrow 0 \text{ for any } D' \in R^2.$$

Since

$$D'B_N = D'B_n + (D'B_N - D'B_n) \quad \text{for any } D' \in R^2,$$

to prove (3.16) it suffices from (3.17) and the Cramér-Wold theorem to show that

$$(3.18) \quad D'B_N - D'B_n \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0 \text{ for any } D' \in R^2.$$

Let any  $D' = (d_0, d_1) \in R^2$  be fixed. For  $i \geq 1$  set

$$S_i^{(t)} = \sum_{j=1}^i \alpha_j \beta_{ji} U_j^{(t)} \quad \text{for } t=0, 1.$$

It is clear that

$$(3.19) \quad \begin{aligned} D'B_N - D'B_n &= \sum_{t=0}^1 d_t (Nh_N^2)^{1/2} (S_N^{(t)} - S_n^{(t)}) \\ &\quad + \{(Nh_N^2/(nh_n^2))^{1/2} - 1\} D'B_n. \end{aligned}$$

Put

$$\xi^{(t)}(x) = (f(x))^{1-t} (g(x))^t \quad \text{for } t=0, 1.$$

It follows from assumption and Lemma 2.1 that for  $t=0, 1$

$$h_i^2 E [(U_i^{(t)})^2] \leq \int_{R^p} h_i^{-p} K^2((x-u)/h_i) \xi^{(t)}(u) du \leq C_1 \quad \text{for all } i \geq 1.$$

Thus by use of Lemma 2.5 we have

$$(3.20) \quad \sum_{i=0}^1 d_i (N h_N^p)^{1/2} (S_N^{(i)} - S_n^{(i)}) \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0.$$

Let any  $\varepsilon (>0)$  be fixed. From (3.15) and (H3) we get that for  $\delta$  in (H3)

$$P \{ |h_N/h_n - 1| \geq \varepsilon \} \leq P \{ |N/n - 1| \geq \delta \} \rightarrow 0 \quad \text{as } d \rightarrow 0,$$

which implies that

$$h_N/h_n \xrightarrow{P} 1 \quad \text{as } d \rightarrow 0.$$

Therefore, (3.15) gives

$$\{N h_N^p / (n h_n^p)\}^{1/2} - 1 \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0,$$

which, together with (3.17), yields that

$$(3.21) \quad \{(N h_N^p / (n h_n^p))^{1/2} - 1\} D' B_n \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0.$$

From (3.19), (3.20) and (3.21) we obtain (3.18). In the proof of Theorem 4.1 of [6] it was proved, under the assumptions of Theorem 3.1, that  $\|B_i^* - B_i\|_2 \rightarrow 0$  as  $i \rightarrow \infty$  on  $\Omega$ . Hence by Lemma 3.1 we get  $\|B_N^* - B_N\|_2 \rightarrow 0$  as  $d \rightarrow 0$  w.p.1., which, together with (3.16), implies that

$$(3.22) \quad B_N^* \xrightarrow{L} N(0, \Gamma) \quad \text{as } d \rightarrow 0.$$

Define a function  $T(u, v)$  on  $R^2$  as

$$T(u, v) = \begin{cases} v/u & \text{if } u \neq 0 \\ 0 & \text{otherwise.} \end{cases}$$

Let  $L' = (-q(x)/f^2(x), f^{-1}(x))$ . By the Taylor theorem we get

$$(3.23) \quad (N h_N^p)^{1/2} \{T(f_N(x), q_N(x)) - T(f(x), q(x))\} = L' B_N^* + \varepsilon_N \|B_N^*\|_2$$

on  $[N < +\infty]$ ,

where

$$\varepsilon_i \rightarrow 0 \quad \text{if } \|(f_i(x), q_i(x))' - (f(x), q(x))'\|_2 \rightarrow 0.$$

According to Lemma 2.1 of [6] we have  $m_N(x) = q_N(x)/f_N(x)$  on  $[N < +\infty]$ . Hence by the definitions of  $T(u, v)$  and  $N$  we obtain

$$(3.24) \quad T(f_N(x), q_N(x)) = m_N(x) \quad \text{on } [N < +\infty].$$

Since  $m(x) = q(x)/f(x)$  it follows from (3.23) and (3.24) that

$$(3.25) \quad (Nh_N^p)^{1/2}(m_N(x) - m(x)) = L'B_N^* + \varepsilon_N \|B_N^*\|_2 \quad \text{on } [N < +\infty].$$

Lemma 2.2 gives  $\varepsilon_i \rightarrow 0$  as  $i \rightarrow \infty$  w.p.1., which, together with Lemma 3.1, yields that

$$(3.26) \quad \varepsilon_N \rightarrow 0 \quad \text{as } d \rightarrow 0 \text{ w.p.1.}$$

Combining (3.22) and (3.26) we get

$$(3.27) \quad \varepsilon_N \|B_N^*\|_2 \xrightarrow{P} 0 \quad \text{as } d \rightarrow 0.$$

Therefore, according to (3.22), (3.25), (3.27) and Lemma 3.1 we obtain

$$(Nh_N^p)^{1/2}(m_N(x) - m(x)) \xrightarrow{L} N(0, L'\Gamma L) \quad \text{as } d \rightarrow 0,$$

which concludes the proof of Theorem 3.1.

We are now in the position to give our main result.

**THEOREM 3.2.** *Under all the conditions of Theorem 3.1 we have*

$$P \{m(x) \in I_{N(d),d}(x)\} \rightarrow \alpha \quad \text{as } d \rightarrow 0.$$

**PROOF.** Put  $N = N(d)$ . By Lemma 3.1 and Theorem 3.1 we have

$$\begin{aligned} Dd^{-1}(m_N(x) - m(x)) \\ = (D^2\sigma^2(x)/(Nh_N^p d^2))^{1/2} (Nh_N^p/\sigma^2(x))^{1/2} (m_N(x) - m(x)) \xrightarrow{L} N(0, 1) \\ \text{as } d \rightarrow 0. \end{aligned}$$

Thus we obtain

$$P \{m(x) \in I_{N,d}(x)\} = P \{|Dd^{-1}(m_N(x) - m(x))| \leq D\} \rightarrow \Phi(D) - \Phi(-D) = \alpha \quad \text{as } d \rightarrow 0.$$

This completes the proof.

**COROLLARY 3.1.** *Assume  $E[Y^4] < \infty$ . Let (K1) and (K2) be satisfied, and let  $\|\phi\|_\infty < \infty$ . Suppose that there exist bounded, continuous second partial derivatives  $\partial^2 f(x)/\partial x_i \partial x_j$  and  $\partial^2 q(x)/\partial x_i \partial x_j$  on  $R^p$  for  $i, j = 1, \dots, p$  and that  $g(x)$  is continuous on  $R^p$ . Set*

$$h_n = n^{-r/p} \quad \text{with } p/(p+4) < r < 1.$$

*Let  $a$  in (2.1) satisfy  $1 \geq a > (1-r)/2$ . Then, for each point  $x$  with  $v(x) > 0$  we obtain*

$$P \{m(x) \in I_{N(d),d}(x)\} \rightarrow \alpha \quad \text{as } d \rightarrow 0.$$

PROOF. We can easily verify (H1)~(H5) with  $\beta=(2a+r-1)^{-1}$ . Lemma 3.1 gives (3.15). Thus, since all the conditions of Theorem 3.1 are fulfilled, we obtain Corollary 3.1 by Theorem 3.2. This completes the proof.

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