

## ON THE RATE OF CONVERGENCE TO NORMALITY FOR GENERALIZED LINEAR RANK STATISTICS

MADAN L. PURI AND MUNSUP SEOH

(Received Dec. 1, 1983)

### Summary

A generalized linear rank statistic is introduced to include, as special cases, both signed as well as unsigned linear rank statistics. For this statistic, the rate of convergence to asymptotic normality is investigated. It is shown that this rate is of order  $O(N^{-1/2} \log N)$  if the score generating function  $\phi$  is twice differentiable, and it is of order  $O(N^{-1/2})$  if the second derivative of  $\phi$  satisfies Lipschitz's condition of order  $\geq 1/2$ . The results obtained extend as well as generalize most of the earlier results obtained in this direction.

### 1. Introduction

Let  $X_{N1}, X_{N2}, \dots, X_{NN}$ ,  $N \geq 1$  be independent r.v.'s (random variables) and let  $g$  be a real valued measurable function such that  $X_{Nj}^g = g(X_{Nj})$ ,  $1 \leq j \leq N$ , has a continuous c.d.f. (cumulative distribution function)  $G_{Nj}$ . We introduce the generalized linear rank statistic

$$(1.1) \quad T_N = \sum_{j=1}^N c_{Nj} a_{NR_{Nj}^g}(X_{Nj})$$

where  $\{c_{Nj}: 1 \leq j \leq N\}$  is an array of regression constants;  $\{a_{Nj}(\cdot): 1 \leq j \leq N\}$  is an array of real valued functions (called scores); and  $R_{Nj}^g$ ,  $1 \leq j \leq N$ , is the rank of  $X_{Nj}^g$  among  $\{X_{Nk}^g: 1 \leq k \leq N\}$ . We assume that the  $a_{Nj}(\cdot)$ ,  $1 \leq j \leq N$ , are generated by a known (nonconstant) function (called a score generating function)  $\phi(s, t)$ ,  $0 < s < 1$ ,  $-\infty < t < \infty$ , in either of following two ways:

$$(1.2) \quad a_{Nj}(t) = \phi(E U_{N:j}, t), \quad 1 \leq j \leq N \text{ (approximate scores)}$$

$$(1.3) \quad a_{Nj}(t) = E \phi(U_{N:j}, t), \quad 1 \leq j \leq N \text{ (exact scores)}$$

where  $U_{N:j}$  denotes the  $j$ th order statistic in a random sample of size

---

Key words and phrases: Generalized rank statistics, rate of convergence.

$N$  from the uniform distribution on  $(0, 1)$ . To exclude trivialities, we assume that  $\sum_{j=1}^N |c_{Nj}| > 0$ .

When  $g(x) = x$  and  $\phi(s, t) = \tilde{\phi}(s)$ , the statistic (1.1) reduces to the simple linear rank statistic

$$(1.4) \quad T_N = \sum_{j=1}^N c_{Nj} a_{NR_{Nj}}$$

where  $a_{Nj}$ 's are usual scores of constants and  $R_{Nj}$ ,  $1 \leq j \leq N$ , is the rank of  $X_{Nj}$  among  $\{X_{Nk} : 1 \leq k \leq N\}$ . For this statistic, the asymptotic normality was obtained by Hájek [6], [7]; the rate of convergence was investigated by Jurečková and Puri [11], Bergström and Puri [1] and Hušková [9], [10]; and Does [2] studied the rate of convergence as well as asymptotic expansions.

On the other hand, when  $g(x) = |x|$  and  $\phi(s, t) = \tilde{\phi}(s) \operatorname{sgn} t$ , where  $\operatorname{sgn} t = 1$  or  $-1$  according as  $t \geq 0$  or  $< 0$ , the statistic (1.1) reduces to the signed linear rank statistic

$$(1.5) \quad T_N = \sum_{j=1}^N c_{Nj} a_{NR_{Nj}^+} \operatorname{sgn} X_{Nj}$$

where  $R_{Nj}^+$  is the rank of  $|X_{Nj}|$  among  $\{|X_{Nk}| : 1 \leq k \leq N\}$ . For this statistic, the reader is referred to Hájek [6] and Hušková [8] (for the asymptotic normality); Puri and Wu [16] and Puri and Seoh [13] (for the rate of convergence); and Puri and Seoh [14], [15] (for the Edgeworth expansions).

The aim of this paper is to investigate the problem of the rate of convergence for the statistics (1.1) which include the signed as well as (unsigned) simple linear rank statistics. The results obtained improve or include as special cases some of the results of Jurečková and Puri [11], von Bahr [18], Bergström and Puri [1], Hušková [9], [10] and Does [3] among others.

## 2. Assumptions and main theorems

We introduce some notations. For any function  $f$ , with the domain  $D(f)$ ,  $\|f\|$  denotes the supremum norm of  $f$ , i.e.,

$$\|f\| = \sup_{x \in D(f)} |f(x)| \quad \text{or} \quad \sup_{(x, y) \in D(f)} |f(x, y)|$$

depending on the number of variables. We also use  $S^\circ$ , for a r.v.  $S$ , centered at its expectation, i.e.,  $S^\circ = S - ES$ . Noting that  $R_{Nj}^q = \sum_{k=1}^N u(X_{Nj}^q - X_{Nk}^q)$ , where  $u(x) = (1 + \operatorname{sgn} x)/2$ , we define r.v.'s  $\rho_{Nj}$  and  $\rho_{Njj}$ ,  $1 \leq j \leq N$ , as

$$(2.1) \quad \begin{aligned} \rho_{Nj} &= R_{Nj}^g / (N+1) = (N+1)^{-1} \left\{ 1 + \sum_{k \neq j}^N u(X_{Nj}^g - X_{Nk}^g) \right\}, \\ \rho_{Njj} &= E(\rho_{Nj} | X_{Nj}) = (N+1)^{-1} \left\{ 1 + \sum_{k \neq j}^N G_{Nk}(X_{Nj}^g) \right\}. \end{aligned}$$

ASSUMPTION (A<sub>l</sub>): The score generating function  $\phi(s, t)$  is  $l$ -times differentiable with respect to its first argument  $s$  such that its  $l$ th partial derivatives  $\phi_1^{(l)}(s, t) = \partial^l \phi(s, t) / (\partial s)^l$  satisfy Lipschitz's condition of order  $\delta$ ,  $0 < \delta \leq 1$ , with respect to  $s$ , i.e.,

$$(2.2) \quad \sup_{-\infty < t < \infty} |\phi_1^{(l)}(x, t) - \phi_1^{(l)}(y, t)| \leq A_l |x - y|^\delta, \quad x, y \in (0, 1),$$

for some absolute constant  $A_l$ , and  $\phi(s, t), \phi_1^{(1)}(s, t), \dots, \phi_1^{(l)}(s, t)$  are jointly measurable in  $s$  and  $t$ .

*Remark 2.1.* Assumption (A<sub>l</sub>) is satisfied with Lipschitz's condition of order  $\delta = 1/2$  if there exist  $(l+1)$ th partial derivatives which are square integrable uniformly in  $t$ , i.e.,

$$\sup_{-\infty < t < \infty} \int_{-\infty}^{\infty} |\phi_1^{(l+1)}(s, t)|^2 ds < \infty.$$

Assuming the first partial derivatives, we shall approximate  $T_N$  by

$$(2.3) \quad \hat{S}_N = \sum_{i=1}^N E(S_N | X_i) - (N-1) E S_N$$

where  $S_N$  is the first two terms of Taylor's expansion (with respect to its first argument) of  $T_N$  with approximate scores, i.e.,

$$(2.4) \quad S_N = \sum_{j=1}^N c_{Nj} \{ \phi(\rho_{Njj}, X_{Nj}) + (\rho_{Nj} - \rho_{Njj}) \phi_1^{(1)}(\rho_{Njj}, X_{Nj}) \}.$$

Let  $\Phi(x)$  denote the standard normal c.d.f. and put

$$(2.5) \quad \tau_N^2 = \text{Var } T_N, \quad \sigma_N^2 = \text{Var } S_N, \quad \hat{\sigma}_N^2 = \text{Var } \hat{S}_N, \quad \omega_3 = \hat{\sigma}_N^{-3} \sum_{j=1}^N |c_{Nj}|^3.$$

Then our main theorems are as follows:

**THEOREM 2.1.** *If Assumption (A<sub>1</sub>) is satisfied, then*

$$(2.6) \quad \|P(T_N \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| \leq C_1 L_N + 2e M_N \max \{N^{-1/2}, N^{-\delta/2}\} \log N$$

where  $C_1$  is an absolute constant;  $L_N$  and  $M_N$  are defined by

$$(2.7) \quad \begin{aligned} L_N &= 4(2\|\phi\|^3 + \|\phi_1^{(1)}\|^3) \omega_3, \\ M_N &= 64e \max \{A_1, \|\phi\|, \|\phi_1^{(1)}\|\} \left( \sum_{j=1}^N c_{Nj}^2 \right)^{1/2} \hat{\sigma}_N^{-1}; \end{aligned}$$

and  $\delta$  and  $\Delta_1$  are given by (2.2). Moreover,

$$(2.8) \quad |\tau_N - \hat{\sigma}_N| \leq \hat{\sigma}_N M_N \max \{N^{-1/2}, N^{-\delta/2}\}.$$

**THEOREM 2.2.** *If the first partial derivative  $\phi_1^{(1)}$  of the score generating function  $\phi$  exists and is bounded on  $(0, 1) \times (-\infty, \infty)$ , then*

$$(2.9) \quad \|P(S_N^\circ \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| = O(\omega_\delta),$$

$$(2.10) \quad |\sigma_N - \hat{\sigma}_N| = O(\hat{\sigma}_N \omega_\delta).$$

**THEOREM 2.3.** *If Assumption  $(A_2)$  is satisfied, then*

$$(2.11) \quad \|P(T_N^\circ \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| = O(\max \{\omega_\delta, \omega_\delta^{2\beta}\}),$$

$$(2.12) \quad |\tau_N - \hat{\sigma}_N| = O(\hat{\sigma}_N \omega_\delta).$$

*Remark 2.2.* Theorem 2.1 is the improvement, as well as the generalization, of Theorem 1.2 of Bergström and Puri [1]. The latter, dealing with only approximate scores, gives the bound of order  $O(N^{-1/2} \log N)$  under stronger condition that  $\max_{1 \leq j \leq N} |c_{Nj}| = O(N^{-1/2})$  and requires bounded second derivatives, whereas we, dealing with both exact scores and approximate scores, derive the bound under much weaker assumption, namely  $\sum_{j=1}^N |c_{Nj}|^8 = O(N^{-1/2})$ , and require only the first derivatives satisfying Lipschitz's condition of order  $\delta=1$ .

*Remark 2.3.* Theorem 2.3 is the generalization of Hušková [9], [10] and also the improvement of Hušková [10]. But, in Hušková [10], Theorem B and proofs of Lemma 2, Lemma 4 and Lemma 6 are incorrect.

From now on, for the ease of convenience, we shall suppress the subscript  $N$  in  $c_{Nj}$ ,  $a_{Nj}$ ,  $R_{Nj}^q$ , etc., whenever it causes no confusion. For notational convenience we also use  $\phi_1(x, y)$ , instead of  $\phi_1^{(1)}(x, y)$ , for the first partial derivatives.

### 3. Preliminaries

In this section we prove several lemmas which are needed in the proofs of main theorems.

In view of (2.1) and (2.4) we can write

$$(3.1) \quad \rho_j - \rho_{jj} = (N+1)^{-1} \sum_{k \neq j}^N H_{jk},$$

$$(3.2) \quad S_N = \sum_{j=1}^N c_j \left\{ \phi(\rho_{jj}, X_j) + (N+1)^{-1} \sum_{k \neq j}^N H_{jk} \phi_1(\rho_{jj}, X_j) \right\},$$

where r.v.'s  $H_{jk}$ ,  $1 \leq j, k \leq N$ , are defined by

$$(3.3) \quad H_{jk} = u(X_j^g - X_k^g) - G_k(X_j^g) .$$

For each  $j$ , conditionally given  $X_j^g$ ,  $\rho_j - \rho_{jj}$  is the sum of independent r.v.'s with zero means. Thus we may apply Lemma A.1 (in the appendix) to obtain for any integer  $r \geq 1$  and  $1 \leq j \leq N$ , that

$$(3.4) \quad E(\rho_j - \rho_{jj})^{2r} \leq (4e)^r r^r N^{-r} .$$

We now estimate  $E(S_N^\circ - \hat{S}_N^\circ)^{2r}$  where  $r$  is an integer  $\geq 1$ . It follows by elementary computations that for  $1 \leq l \leq N$ ,

$$(3.5) \quad E(S_N | X_l) = c_l \phi(\rho_{ll}, X_l) + E S_N - c_l E \phi(\rho_{ll}, X_l) \\ + (N+1)^{-1} \sum_{j \neq l}^N c_j E \{H_{jl} \phi_1(\rho_{jj}, X_j) | X_l\} .$$

From (3.5), (2.3) and (3.2), we obtain

$$(3.6) \quad S_N^\circ - \hat{S}_N^\circ = (N+1)^{-1} \sum_{j=1}^N \sum_{k \neq j}^N c_j V_{jk}$$

where

$$(3.7) \quad V_{jk} = H_{jk} \phi_1(\rho_{jj}, X_j) - E \{H_{jk} \phi_1(\rho_{jj}, X_j) | X_k\} .$$

Since  $E \{V_{jk} | X_j\} = E \{V_{jk} | X_k\} = 0$  for  $j \neq k$ , an application of Lemma A.2 (in the appendix) yields that for any positive integer  $r$

$$(3.8) \quad E(S_N^\circ - \hat{S}_N^\circ)^{2r} \leq (8e \|\phi_1\|)^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r (2r)^{2r} N^{-r} .$$

*Remark 3.1.* It may be noted that the results (3.4) and (3.8) are generalizations, as well as improvements, of Lemma 2.1 and Lemma 2.2 of Bergström and Puri [1].

We now define r.v.'s  $I_k$  for  $0 \leq k \leq l$  as

$$(3.9) \quad I_k = \sum_{j=1}^N c_j \frac{(\rho_j - \rho_{jj})^k}{k!} \phi_1^{(k)}(\rho_{jj}, X_j)$$

so that  $S_N = I_0 + I_1$ . Also by Taylor's expansion for any  $x, y$  and  $t$ ,

$$\phi(x, t) = \sum_{k=0}^{l-1} \frac{(x-y)^k}{k!} \phi_1^{(k)}(y, t) + \frac{(x-y)^l}{l!} \phi_1^{(l)}(\lambda x + (1-\lambda)y, t)$$

where  $0 \leq \lambda \leq 1$  depends on  $x, y$  and  $t$ . Thus, under Assumption (A<sub>1</sub>)

$$(3.10) \quad \left| \phi(x', t) - \sum_{k=0}^l \frac{(x-y)^k}{k!} \phi_1^{(k)}(y, t) \right| \leq \frac{A_l}{l!} |x-y|^{l+s} .$$

**LEMMA 3.1.** *Let  $\tilde{T}_N$  denote the statistic (1.1) with exact scores (given by (1.3)) to distinguish it from the statistic  $T_N$  with approximate scores and let  $r > 0$  be an integer. Then under Assumption (A<sub>1</sub>)*

$$(3.11) \quad \mathbb{E}(\tilde{T}_N^\circ - T_N^\circ)^{2r} \leq (2A_1)^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r N^{-2r}$$

and under Assumption  $(A_l)$ ,  $l \geq 2$ ,

$$(3.12) \quad \mathbb{E}(\tilde{T}_N^\circ - T_N^\circ)^{2r} \leq \|\phi_1^{(2)}\|^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r N^{-r}.$$

PROOF. Put  $\tilde{a}_j(t) = \mathbb{E} \phi(U_{N:j}, t)$ ,  $1 \leq j \leq N$ . Then

$$(3.13) \quad \begin{aligned} \mathbb{E}(\tilde{T}_N^\circ - T_N^\circ)^{2r} &\leq 2^{2r} \mathbb{E}(\tilde{T}_N - T_N)^{2r} \\ &\leq 2^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r \mathbb{E} \left[ \sum_{j=1}^N \left\{ \tilde{a}_{R_j}(X_j) - \phi\left(\frac{R_j}{N+1}, X_j\right) \right\}^2 \right]^r. \end{aligned}$$

Under Assumption  $(A_1)$ , (3.10) ensures that for  $1 \leq j \leq N$  and any  $t$

$$(3.14) \quad \left| \tilde{a}_j(t) - \phi\left(\frac{j}{N+1}, t\right) \right| \leq A_1 \mathbb{E} \left| U_{N:j} - \frac{j}{N+1} \right|^{1+\delta} \leq A_1 N^{-(1+\delta)/2}.$$

(3.11) follows from (3.13) and (3.14).

On the other hand, under Assumption  $(A_l)$ ,  $l \geq 2$ , there are second partial derivatives, bounded, such that

$$\left| \tilde{a}_j(t) - \phi\left(\frac{j}{N+1}, t\right) \right| \leq \frac{\|\phi_1^{(2)}\|}{2} \mathbb{E} \left( U_{N:j} - \frac{j}{N+1} \right)^2 \leq \frac{\|\phi_1^{(2)}\|}{2N}.$$

This inequality and (3.13) ensure (3.12) and the proof is complete.

LEMMA 3.2. Let  $T_N$  be the statistic (1.1) with approximate scores (given by (1.2)) and suppose that Assumption  $(A_l)$ ,  $l \geq 1$ , is satisfied. Then for any integer  $r > 0$ ,

$$\mathbb{E} \left( T_N^\circ - \sum_{k=0}^l I_k^\circ \right)^{2r} \leq \left( \frac{2A_l}{l!} \right)^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r (4e)^{r(l+1)} (r(l+1))^{r(l+1)} N^{-r(l+\delta-1)}$$

where  $I_k$ ,  $1 \leq k \leq l$ , are defined by (3.9).

PROOF. It follows from (3.10) and Hölder's inequality that

$$\begin{aligned} \mathbb{E} \left( T_N^\circ - \sum_{k=0}^l I_k^\circ \right)^{2r} &\leq 2^{2r} \mathbb{E} \left( \sum_{j=1}^N c_j \left\{ \phi(\rho_j, X_j) - \sum_{k=0}^l \frac{(\rho_j - \rho_{jj})^k}{k!} \phi_1^{(k)}(\rho_{jj}, X_j) \right\} \right)^{2r} \\ &\leq 2^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r \left( \frac{A_l}{l!} \right)^{2r} N^{r-1} \sum_{j=1}^N \mathbb{E} |\rho_j - \rho_{jj}|^{2r(l+\delta)} \end{aligned}$$

which, together with (3.4), completes the proof.

Define for any positive integer  $r$

$$(3.15) \quad \omega_r = \sum_{j=1}^N |c_j^*|^r$$

where  $c_j^* = \hat{\sigma}_N^{-1} c_j$ ,  $1 \leq j \leq N$ . Also consider the r.v.'s  $\hat{S}_N^{(j)}$ ,  $1 \leq j \leq N$ , defined as

$$(3.16) \quad \hat{S}_N^{(j)} = c_j^* \{ \phi(\rho_{jj}, X_j) - \mathbb{E} \phi(\rho_{jj}, X_j) \} \\ + (N+1)^{-1} \sum_{k \neq j}^N c_k^* \mathbb{E} \{ [u(X_k^g - X_j^g) - G_j(X_k^g)] \phi_1(\rho_{kk}, X_k) | X_j \} .$$

Then, in view of (2.3) and (3.5), we have

$$(3.17) \quad \hat{\sigma}_N^{-1} \hat{S}_N^{\circ} = \sum_{j=1}^N \hat{S}_N^{(j)} .$$

It is easy to check that for  $1 \leq j \leq N$

$$(3.18) \quad \mathbb{E} \hat{S}_N^{(j)} = 0, \quad \sum_{j=1}^N \mathbb{E} (\hat{S}_N^{(j)})^2 = 1, \\ \mathbb{E} (\hat{S}_N^{(j)})^2 \leq 2 \{ (c_j^*)^2 \|\phi\|^2 + N^{-1} \omega_2 \|\phi_1\|^2 \}, \\ \mathbb{E} |\hat{S}_N^{(j)}|^3 \leq 4 \{ 2 |c_j^*|^3 \|\phi\|^3 + N^{-1} \omega_3 \|\phi_1\|^3 \} .$$

As  $\hat{\sigma}_N^{-1} \hat{S}_N^{\circ}$  is the sum of independent r.v.'s with zero means and satisfying  $\sum_{j=1}^N \mathbb{E} |\hat{S}_N^{(j)}|^3 \leq L_N$ , ( $L_N$  defined by (2.7)), Lemma V.2.1 and Theorem V.2.3 of Petrov [12] ensure

LEMMA 3.3. For  $|t| \leq (4L_N)^{-1}$

$$(3.19) \quad |\mathbb{E} \exp(it \hat{\sigma}_N^{-1} \hat{S}_N^{\circ}) - \exp(-t^2/2)| \leq 16L_N |t|^3 \exp(-t^2/3) .$$

Moreover, there is an absolute constant  $C_1$  such that

$$(3.20) \quad \|\mathbb{P}(\hat{\sigma}_N^{-1} \hat{S}_N^{\circ} \leq \cdot) - \Phi(\cdot)\| \leq C_1 L_N .$$

In the rest of this section, we assume that

$$(3.21) \quad \omega_3 = \sum_{j=1}^N |c_j^*|^3 \leq 1, \quad \max \{ \|\phi\|, \|\phi_1\|, \|\phi_1^{(2)}\| \} \leq 1/6 .$$

Thus we have

$$(3.22) \quad \mathbb{E} (\hat{S}_N^{(j)})^2 \leq ((c_j^*)^2 + N^{-1} \omega_2) / 18, \quad \mathbb{E} |\hat{S}_N^{(j)}|^3 \leq (2 |c_j^*|^3 + N^{-1} \omega_3) / 54 .$$

Since  $1 = \sum_{j=1}^N \mathbb{E} (\hat{S}_N^{(j)})^2 \leq 9^{-1} \omega_2$ , Hölder's inequality yields

$$(3.23) \quad N^{-1/2} \leq 27^{-1} \omega_3 .$$

The inequality (3.23) plays an important role in the following proofs.

Let  $\mu_j(t)$  denote the characteristic function of  $\hat{S}_N^{(j)}$ , i.e.,

$$(3.24) \quad \mu_j(t) = \mathbb{E} \exp(it \hat{S}_N^{(j)}) .$$

Then it follows from Petrov ([12], p. 110) that

$$|\mu_j(t)| \leq \exp \left\{ -\frac{t^2}{2} \mathbb{E} (\hat{S}_N^{(j)})^2 + \frac{2}{3} |t|^3 \mathbb{E} |\hat{S}_N^{(j)}|^3 \right\}$$

which implies that for  $|t| \leq \omega_3^{-1}$

$$(3.25) \quad \prod_{j \in J} |\mu_j(t)| \leq \exp \left\{ -\frac{t^2}{2} + \frac{t^2}{2} \sum_{j \in J} \mathbb{E} (\hat{S}_N^{(j)})^2 + \frac{2}{3} |t|^3 \sum_{j \in J} \mathbb{E} |\hat{S}_N^{(j)}|^3 \right\} \\ \leq \exp \left( -\frac{4t^2}{9} + \frac{t^2}{18} \#J \max_{1 \leq j \leq N} |c_j^*| \right)$$

where  $J$  is a subset of  $\{1, 2, \dots, N\}$  and  $\#J$  is its cardinality.

Now we shall estimate  $\mathbb{E} \exp(it\hat{\sigma}_N^{-1}S_N^\circ)$  and  $\mathbb{E} \exp(it\hat{\sigma}_N^{-1}(S_N^\circ + I_2^\circ))$ . To this end we need the following two lemmas:

Recalling that the r.v.'s  $H_{jk}$ ,  $1 \leq j, k \leq N$  are given by (3.3), we define

$$(3.26) \quad I_{21} = (2(N+1)^2)^{-1} \sum_{(j,k) \neq} \sum_{\neq} c_j H_{jk}^2 \phi_1^{(2)}(\rho_{jj}, X_j), \\ I_{22} = (2(N+1)^2)^{-1} \sum_{(j,k,l) \neq} \sum_{\neq} c_j H_{jk} H_{jl} \phi_1^{(2)}(\rho_{jj}, X_j),$$

where  $(j, k) \neq$  and  $(j, k, l) \neq$  mean that summations are taken for distinct indices  $j, k$  and  $l$  over the set  $\{1, 2, \dots, N\}$ . Then it follows from (3.1), (3.9) and (3.26) that

$$(3.27) \quad I_2^\circ = I_{21}^\circ + I_{22}^\circ.$$

LEMMA 3.4. *Suppose the condition (3.21) is satisfied and let  $I_{21}$  and  $I_{22}$  be defined by (3.26). Then for all  $N \geq 3$  we have*

$$(3.28) \quad \mathbb{E} (\hat{\sigma}_N^{-1} I_{21}^\circ)^2 \leq 3^{-12} \omega_3^4,$$

$$(3.29) \quad \mathbb{E} (\hat{\sigma}_N^{-1} I_{22}^\circ)^{2r} \leq \left( \frac{2}{9} r \omega_3 \right)^{2r},$$

where  $\omega_3$  is given by (3.15) and  $r \geq 0$  is an integer.

PROOF. Set  $\tilde{H}_{jk} = H_{jk}^2 \phi_1^{(2)}(\rho_{jj}, X_j) - \mathbb{E} \{H_{jk}^2 \phi_1^{(2)}(\rho_{jj}, X_j)\}$ ,  $1 \leq j, k \leq N$ , and then direct expansion yields that

$$(3.30) \quad 4(N+1)^4 \mathbb{E} (\hat{\sigma}_N^{-1} I_{21}^\circ)^2 = \sum_{(j,k) \neq} \sum_{\neq} \mathbb{E} \{ (c_j^*)^2 \tilde{H}_{jk}^2 + c_j^* c_k^* \tilde{H}_{jk} \tilde{H}_{kj} \} \\ + \sum_{(j,k,l) \neq} \sum_{\neq} \mathbb{E} \{ (c_j^*)^2 \tilde{H}_{jk} \tilde{H}_{jl} + c_j^* c_k^* \tilde{H}_{jk} \tilde{H}_{kl} \\ + c_j^* c_l^* \tilde{H}_{jk} \tilde{H}_{lj} + c_j^* c_l^* \tilde{H}_{jk} \tilde{H}_{lk} \} \\ + \sum_{(j,k,l,m) \neq} \sum_{\neq} \mathbb{E} \{ c_j^* c_l^* \tilde{H}_{jk} \tilde{H}_{lm} \}.$$

Since  $(\mathbb{E} \tilde{H}_{jk}^2)^{1/2} \leq \|\phi_1^{(2)}\| \leq 1/6$ ,  $1 \leq j, k \leq N$ , and expectations in the last

sum of (3.30) vanish, we have

$$\mathbb{E}(\hat{\sigma}_N^{-1}I_{21}^\circ)^2 \leq (144N^4)^{-1} \{N\omega_2 + \omega_1^2 + N^2\omega_2 + 3N\omega_1^2\} \leq 3^{-12}\omega_1^4$$

where the second inequality follows by (3.23) and facts

$$(3.31) \quad \omega_1 \leq N^{2/3}\omega_3^{1/3}, \quad \omega_2 \leq N^{1/3}\omega_3^{2/3}.$$

To prove (3.29), we set  $\hat{\sigma}_N^{-1}I_{22}^\circ = (2(N+1))^{-1} \sum_{j=1}^N c_j^* \hat{H}_j$ , where

$$\hat{H}_j = (N+1)^{-1} \sum_{k \neq j}^N \sum_{l \neq j, k}^N H_{jk} H_{jl} \phi_1^{(2)}(\rho_{jl}, X_j).$$

Then, for any fixed  $j$ ,  $1 \leq j \leq N$ , conditionally given  $X_j$ , summands of  $\hat{H}_j$  satisfy the assumptions of Lemma A.2. Thus an application of this lemma yields that for  $1 \leq j \leq N$ ,

$$\mathbb{E} \hat{H}_j^{2r} \leq \mathbb{E}[\mathbb{E}(\hat{H}_j^{2r} | X_j)] \leq (4e \|\phi_1^{(2)}\|)^{2r} (2r)^{2r}$$

which, together with (3.23), (3.31) and Hölder's inequality, ensures (3.29) to complete the proof.

Finally we need the following lemma, the proof of which is given in the appendix.

**LEMMA 3.5.** *Suppose the condition (3.21) is satisfied and let  $\hat{S}_N$ ,  $S_N$  and  $I_{22}$  be defined by (2.3), (2.4) and (3.26) respectively. Then there is an absolute constant  $K$  such that for  $|t| \leq \omega_3^{-1}$ , any  $N \geq 3$  and a positive integer  $r$  where  $r \leq N/3$ ,*

$$(3.32) \quad |\mathbb{E} \exp(it\hat{\sigma}_N^{-1}\hat{S}_N^\circ) \{\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ)\}^r| \\ \leq K^r r^{2r} \omega_3 (1+|t|+t^2)^r \exp\left(-\frac{4t^2}{9} + \frac{t^2 r}{9} \max_{1 \leq j \leq N} |c_j^*|\right),$$

$$(3.33) \quad |\mathbb{E} \exp(it\hat{\sigma}_N^{-1}\hat{S}_N^\circ) \{\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ)\}^r| \\ \leq K^r r^{2r} \omega_3 (1+|t|+t^2)^r \exp\left(-\frac{4t^2}{9} + \frac{t^2 r}{6} \max_{1 \leq j \leq N} |c_j^*|\right).$$

Furthermore, if  $\omega_3 \leq e^{-2}$  and  $2r \leq \log \omega_3^{-1}$ , then

$$(3.34) \quad |\mathbb{E} \exp(it\hat{\sigma}_N^{-1}\hat{S}_N^\circ) \{\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ)\}^r| \\ \leq K^r r^{2r} \omega_3 (1+|t|+t^2)^r \exp(-t^2/3),$$

$$(3.35) \quad |\mathbb{E} \exp(it\hat{\sigma}_N^{-1}\hat{S}_N^\circ) \{\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ)\}^r| \\ \leq K^r r^{2r} \omega_3 (1+|t|+t^2)^r \exp(-t^2/4).$$

#### 4. Proofs of main theorems

We now prove the main theorems stated in the Section 2.

PROOF OF THEOREM 2.1. By standard arguments, we have for any  $\varepsilon > 0$

$$(4.1) \quad \begin{aligned} P(T_N^\circ \leq \hat{\sigma}_N x) &\leq P(\hat{S}_N^\circ \leq \hat{\sigma}_N(x + \varepsilon)) + P(|T_N^\circ - \hat{S}_N^\circ| \geq \hat{\sigma}_N \varepsilon), \\ P(T_N^\circ \leq \hat{\sigma}_N x) &\geq P(\hat{S}_N^\circ \leq \hat{\sigma}_N(x - \varepsilon)) - P(|T_N^\circ - \hat{S}_N^\circ| \geq \hat{\sigma}_N \varepsilon). \end{aligned}$$

Recalling the well known inequalities

$$\Phi(x + \varepsilon) \leq \Phi(x) + \varepsilon, \quad \Phi(x - \varepsilon) \geq \Phi(x) - \varepsilon,$$

we obtain, using (3.20), (4.1) and Chebychev's inequality, that for any  $\varepsilon > 0$  and positive integer  $r$ ,

$$(4.2) \quad \|P(T_N^\circ \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| \leq C_1 L_N + \varepsilon + (\hat{\sigma}_N \varepsilon)^{-2r} E(T_N^\circ - \hat{S}_N^\circ)^{2r}$$

both for the approximate as well as exact scores.

To complete the proof it remains to estimate the last quantity in (4.2). For the statistic  $T_N$  with approximate scores, Lemma 3.2 ensures that for any positive integer  $r$ ,

$$(4.3) \quad E(T_N^\circ - S_N^\circ)^{2r} \leq (2A_1)^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r (4e)^{2r} (2r)^{2r} N^{-\delta r}.$$

For exact scores, it follows from (3.11), (4.3) and Minkowski's inequality that for any positive integer  $r$ ,

$$(4.4) \quad E(T_N^\circ - S_N^\circ)^{2r} \leq (16A_1 e)^{2r} \left( \sum_{j=1}^N c_j^2 \right)^r (2r)^{2r} N^{-\delta r}$$

and this inequality holds for approximate scores as well. Now set

$$B_N^{(r)} = \left[ 32e(A_1 + \|\phi_1\|) \left( \sum_{j=1}^N c_j^2 \right)^{1/2} r \max\{N^{-\delta/2}, N^{-1/2}\} \right]^{2r}.$$

It then follows from (3.8) and (4.4) that for any positive integer  $r$ ,

$$(4.5) \quad E(T_N^\circ - \hat{S}_N^\circ)^{2r} \leq B_N^{(r)}.$$

Substitution of (4.5) into (4.2) yields

$$(4.6) \quad \|P(T_N^\circ \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| \leq C_1 L_N + \varepsilon + (\hat{\sigma}_N \varepsilon)^{-2r} B_N^{(r)}.$$

Next we choose  $\varepsilon$  such that  $\varepsilon = (\hat{\sigma}_N \varepsilon)^{-2r} B_N^{(r)}$ , i.e.,

$$(4.7) \quad \begin{aligned} \varepsilon &= \left[ 32e(A_1 + \|\phi_1\|) \left( \sum_{j=1}^N c_j^2 \right)^{1/2} \hat{\sigma}_N^{-1} r \max\{N^{-\delta/2}, N^{-1/2}\} \right]^{2r/(2r+1)} \\ &\leq [M_N r \max\{N^{-\delta/2}, N^{-1/2}\}]^{2r/(2r+1)} \end{aligned}$$

where  $M_N$  is given by (2.7). In view of (3.18) we have

$$(4.8) \quad 1 = \text{Var} (\hat{\sigma}_N^{-1} \hat{S}_N) \leq 2(\|\phi\|^2 + \|\phi_1\|^2) \sum_{j=1}^N (c_j^*)^2 \leq M_N^2$$

and hence for any positive integer  $r$ ,

$$(4.9) \quad 1 \leq M_N^{2r/(2r+1)} \leq M_N.$$

Taking  $r = [\log N]$ , where  $[\cdot]$  denotes the integer part, we have

$$(4.10) \quad (\max \{N^{-1/2}, N^{-\delta/2}\})^{2r/(2r+1)} \leq e \max \{N^{-1/2}, N^{-\delta/2}\}.$$

Finally (2.6) follows by combining (4.6) through (4.10), and (2.8) follows by (4.5). The proof of Theorem 2.1 is complete.

To prove Theorems 2.2 and 2.3, we shall invoke Esseen's smoothing lemma (see e.g. Feller [5], p. 538), which implies that for all  $\omega > 0$

$$(4.11) \quad |\mathbb{P}(S \leq \cdot) - \Phi(\cdot)| \leq \frac{1}{\pi} \int_{|t| \leq \omega} |t|^{-1} |\eta(t) - e^{-t^2/2}| dt + O(\omega^{-1})$$

where  $S$  is a r.v. such that  $\mathbb{E} S = 0$  and  $\eta(t) = \mathbb{E} \exp(itS)$ .

Note that both Theorems 2.2 and 2.3 are trivial if  $\omega_3 \geq e^{-2}$ . Thus we consider the case when

$$(4.12) \quad \omega_3 < e^{-2}.$$

Furthermore, we may, without loss of generality, assume that

$$(4.13) \quad \max \{\|\phi\|, \|\phi_1\|, \|\phi_1^{(2)}\|\} \leq 1/6.$$

PROOF OF THEOREM 2.2. It follows by (3.19) and (4.11) that, in order to prove (2.9), it suffices to show that for some  $\varepsilon_1$ ,  $0 < \varepsilon_1 \leq 1$ ,

$$(4.14) \quad \int_{|t| \leq \varepsilon_1 \omega_3^{-1}} |t|^{-1} |\mathbb{E} \exp(it\hat{\sigma}_N^{-1} S_N^\circ) - \mathbb{E} \exp(it\hat{\sigma}_N^{-1} \hat{S}_N^\circ)| dt = O(\omega_3).$$

We now estimate the integral (4.14). For any positive integer  $k$ , Taylor's expansion yields

$$(4.15) \quad \begin{aligned} & |t|^{-1} |\mathbb{E} \exp(it\hat{\sigma}_N^{-1} S_N^\circ) - \mathbb{E} \exp(it\hat{\sigma}_N^{-1} \hat{S}_N^\circ)| \\ & \leq \sum_{r=1}^{2k-1} \frac{|t|^{r-1}}{r!} |\mathbb{E} [\{\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ)\}^r \exp(it\hat{\sigma}_N^{-1} \hat{S}_N^\circ)]| \\ & \quad + O\left(\frac{|t|^{2k-1}}{(2k)!} \mathbb{E} (\hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ))^{2k}\right). \end{aligned}$$

By taking  $k=1$  and using (3.8), (3.31) and (3.34), we have for any  $\varepsilon_1 > 0$

$$(4.16) \quad \int_{|t| \leq \varepsilon_1 \omega_3^{-1/2}} |t|^{-1} |\mathbb{E} \exp(it\hat{\sigma}_N^{-1} S_N^\circ) - \mathbb{E} \exp(it\hat{\sigma}_N^{-1} \hat{S}_N^\circ)| = O(\omega_3).$$

We next take  $2k = [\log \omega_3^{-1}]$ . Then (3.34) insures that for any  $\varepsilon_1$ ,  $0 < \varepsilon_1 \leq 1$ ,

$$(4.17) \quad \int_{\varepsilon_1 \omega_3^{-1/2} \leq |t| \leq \varepsilon_1 \omega_3^{-1}} \sum_{r=1}^{2k-1} \frac{|t|^{r-1}}{r!} |\mathbb{E} \{ \hat{\sigma}_N^{-1} (S_N^\circ - \hat{S}_N^\circ) \}^r \exp(it \hat{\sigma}_N^{-1} \hat{S}_N^\circ)| dt \\ \leq \omega_3 \exp\left(-\frac{\varepsilon_1^2 \omega_3^{-1}}{3}\right) (2k)^{4k} \sum_{r=1}^{2k-1} \frac{K^r}{r!} \int_{|t| \leq \omega_3^{-1}} |t|^{r-1} (1+|t|+t^2)^r dt \\ \leq \omega_3 \exp\left(-\frac{\varepsilon_1^2 \omega_3^{-1}}{3}\right) (2k)^{4k} \omega_3^{-6k} e^{3K} = O(\omega_3).$$

It follows from (3.8), (3.23) and (3.31) that

$$(4.18) \quad \int_{|t| \leq \varepsilon_1 \omega_3^{-1}} \frac{|t|^{2k-1}}{2k!} \mathbb{E} (\hat{\sigma}_N^{-1} (S_N^\circ - \hat{S}_N^\circ))^{2k} dt \\ \leq \frac{(8e \|\phi_1\|)^{2k} (2k)^{2k} \omega_3^{2k}}{(2k)!} (\varepsilon_1 \omega_3^{-1})^{2k} \\ \leq (2e\varepsilon_1)^{2k} (2k)^{2k} [(2k)!]^{-1} \leq (2e^2 \varepsilon_1)^{2k}$$

where the last inequality follows by Stirling's formula (see, e.g. Feller [4]). By taking  $\varepsilon_1 = 1/(2e^2)$  and combining (4.15), (4.17) and (4.18), we obtain that

$$(4.19) \quad \int_{\varepsilon_1 \omega_3^{-1/2} \leq |t| \leq \varepsilon_1 \omega_3^{-1}} |t|^{-1} |\mathbb{E} \exp(it \hat{\sigma}_N^{-1} S_N^\circ) - \mathbb{E} \exp(it \hat{\sigma}_N^{-1} \hat{S}_N^\circ)| dt = O(\omega_3)$$

which, together with (4.16), ensures (4.14). Clearly (2.10) follows from (3.8), (3.23) and (3.31). The proof follows.

**PROOF OF THEOREM 2.3.** By standard arguments, to prove (2.11), it suffices to show that

$$(4.20) \quad \mathbb{P}(|T_N^\circ - S_N^\circ - I_2^\circ + I_{21}^\circ| \geq \hat{\sigma}_N \omega_3) = O(\max\{\omega_3, \omega_3^{2\delta}\})$$

and that

$$(4.21) \quad \|\mathbb{P}(S_N^\circ + I_{22}^\circ \leq \hat{\sigma}_N \cdot) - \Phi(\cdot)\| = O(\max\{\omega_3, \omega_3^{2\delta}\}).$$

Applying Lemma 3.2 with  $r=1$  and  $l=2$  and then using (3.23) and (3.31), we obtain that

$$\mathbb{E}(\hat{\sigma}_N^{-1}(T_N^\circ - S_N^\circ - I_2^\circ))^2 = O(\omega_3^{2+2\delta})$$

which, together with (3.28) and Chebychev's inequality, ensures (4.20). Similarly, like the proof of Theorem 2.2, (4.21) follows from (3.8), (3.29) and (3.35). Since Assumption  $(A_1)$  is satisfied with  $\delta=1$  under Assumption  $(A_2)$ , (2.8), (3.23) and (3.31) ensure (2.12). The proof is complete.

5. Appendix

PROOF OF LEMMA 3.5. Since  $\log \omega_3^{-1} \leq 2(\max_{1 \leq j \leq N} |c_j^*|)^{-1}$  when  $\omega_3 \leq 1$ , (3.34) and (3.35) are immediate consequences of (3.32) and (3.33) respectively. Also the proof of (3.32) is very similar to that of (3.33). Thus we omit it and prove only (3.33).

Recalling that  $H_{jk}$  and  $V_{jk}$ ,  $1 \leq j, k \leq N$ , are given by (3.3) and (3.7) respectively, we define r.v.'s for  $1 \leq j, k, l \leq N$

$$(A.1) \quad \tilde{V}_{jkl} = c_j^* \left\{ \frac{N+1}{N-2} V_{jk} + (1/2) H_{jk} H_{jl} \phi_1^{(2)}(\rho_{jj}, X_j) \right\}.$$

Then it follows from (3.6) and (3.26) that

$$(A.2) \quad \hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ) = (N+1)^{-2} \sum_{(j,k,l) \neq} \sum_{\neq} \tilde{V}_{jkl}.$$

Note that for any distinct indices  $j, k$  and  $l$

$$(A.3) \quad E(\tilde{V}_{jkl} | X_j) = E(\tilde{V}_{jkl} | X_k) = E(\tilde{V}_{jkl} | X_l) = 0.$$

In view of (3.16) and (3.17),  $\hat{\sigma}_N^{-1} \hat{S}_N^\circ$  is a sum of independent r.v.'s. This fact and (A.2) ensure that

$$(A.4) \quad E \{ \exp(it \hat{\sigma}_N^{-1} \hat{S}_N^\circ) \hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ) \} \\ = (N+1)^{-2} \sum_{(j,k,l) \neq} \sum_{\neq} \left( \prod_{\nu \neq j,k,l} \mu_\nu(t) \right) E \{ \hat{V}_{jkl} \exp(it(\hat{S}_N^{(j)} + \hat{S}_N^{(k)} + \hat{S}_N^{(l)})) \}$$

where  $\mu_\nu(t)$  is the characteristic function of the r.v.  $\hat{S}_N^{(\nu)}$ , defined by (3.24). It follows by Lemma XV.4.1 of Feller [5] and (A.3) that

$$(A.5) \quad E \{ \tilde{V}_{jkl} \exp(it(\hat{S}_N^{(j)} + \hat{S}_N^{(k)} + \hat{S}_N^{(l)})) \} \\ = O(t^2 |c_j^*| \{ (c_j^*)^2 + (c_k^*)^2 + (c_l^*)^2 + N^{-1} \omega_2 \}).$$

Hence for  $r=1$  (3.33) follows from (3.25), (A.4) and (A.5).

We next consider the expansion of  $\{ \hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ) \}^r$ . Expanding it directly, by the method leading to (3.30), we find that

$$(A.6) \quad \{ \hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ) \}^r = (N+1)^{-2r} \left( \sum_{(j,k,l) \neq} \sum_{\neq} \tilde{V}_{jkl} \right)^r \\ = (N+1)^{-2r} \sum_{b=3}^{3r} \sum_{\nu} q_b^{(r)}$$

where  $q_b^{(r)}$  is of the form

$$(A.7) \quad q_b^{(r)} = \sum_{(j_1, j_2, \dots, j_b) \neq} \tilde{V}_{\partial \partial \partial_1} \tilde{V}_{\partial \partial \partial_2} \dots \tilde{V}_{\partial \partial \partial_r}$$

and each " $\partial$ " is one of indices  $j_1, j_2, \dots, j_b$ . Thus we have

$$(A.8) \quad |E \exp(it \hat{\sigma}_N^{-1} \hat{S}_N^\circ) \{ \hat{\sigma}_N^{-1}(S_N^\circ - \hat{S}_N^\circ + I_{22}^\circ) \}^r|$$

$$\begin{aligned}
&= \left| (N+1)^{-2r} \sum_{b=3}^{3r} \left( \prod_{j \neq j_1, \dots, j_b} \mu_j(t) \right) \sum_{\nu} \mathbb{E} \left\{ q_{b\nu}^{(r)} \prod_{k=1}^b \exp(it\hat{S}_N^{(j_k)}) \right\} \right| \\
&\leq \exp \left( -\frac{4t^2}{9} + \frac{t^2}{6} r \max_{1 \leq j \leq N} |c_j^*| \right) \\
&\quad \times \sum_{b=3}^{3r} \sum_{\nu} N^{-2r} \left| \mathbb{E} \left\{ q_{b\nu}^{(r)} \prod_{k=1}^b \exp(it\hat{S}_N^{(j_k)}) \right\} \right|
\end{aligned}$$

in view of (3.25).

Let  $Q_b^{(r)}$ ,  $3 \leq b \leq 3r$ , be the cardinality of the collection of all different terms  $q_{b\nu}^{(r)}$  and put  $\Omega_r = \sum_{b=3}^{3r} Q_b^{(r)}$ , which is the total number of terms  $q_{b\nu}^{(r)}$  in the expansion (A.6). Then clearly  $\Omega_1=1$ . To estimate  $\Omega_r$ , we first note the recursive relation

$$\Omega_r \leq \Omega_{r-1} \{(\alpha^2 - \alpha + 1) + \alpha^2 + \alpha + 1\} \leq 2\alpha^3 \Omega_{r-1}$$

where  $\alpha=3(r-1)$ . Thus there is a constant  $K_1$  (independent of  $N$  and  $r$ ) such that for any positive integer  $r$ ,

$$(A.9) \quad \Omega_r \leq (54)^{r-1} ((r-1)!)^3 \leq K_1^r r^{3r}$$

where the second inequality follows by Stirling's formula (see e.g. Feller [4]).

To complete the proof, it remains to show that for any positive integer  $r$  and  $b$ ,  $3 \leq b \leq 3r$ ,

$$(A.10) \quad N^{-2r} \left| \mathbb{E} q_{b\nu}^{(r)} \prod_{k=1}^b \exp(it\hat{S}_N^{(j_k)}) \right| \leq K_2^r \omega_s (1 + |t| + t^2)^r$$

where  $K_2$  is an absolute constant. We shall prove this by induction on  $r$ . Note that (A.10) holds for  $r=1$  in view of (A.5). We now suppose that (A.10) holds for  $r \leq m-1$  and let  $r=m \geq 2$ .

If  $b \leq 2m=2r$ , then (A.10) is trivial. Thus we consider only terms  $q_{b\nu}^{(m)}$ ,  $b \geq 2m+1$ . Pick any  $s$ ,  $1 \leq s \leq m$ , and let  $b=2m+s$ . Let  $\gamma$  denote the number of indices  $j$ 's (in (A.7)) which appear at least twice as a subscript of one of  $\tilde{V}$ -terms and then we have that  $3m \geq 2\gamma + 2m + s - \gamma$ , i.e.

$$(A.11) \quad 0 \leq \gamma \leq m - s.$$

Hence at least  $m+2s$  indices among  $j_1, j_2, \dots, j_{2m+s}$  appear exactly once. If one of  $\tilde{V}$ -terms has three subscripts appeared only once, then the expectation in (A.10) is split into expectations of two groups to ensure (A.10) immediately by the induction hypothesis. This is the case when  $s > m/2$  or  $\gamma < s$ .

Next we consider the remaining case when

$$(A.12) \quad 1 \leq s \leq m/2, \quad s \leq \gamma \leq m-s$$

and none of  $\tilde{V}$ -terms in (A.7) satisfy that every one of its subscripts appears only once. Since at least  $m+2s$  indices among  $j_1, j_2, \dots, j_{2m+s}$  appear exactly once, we must have at least  $2s$   $\tilde{V}$ -terms, of which exactly two subscripts appear only once. Thus, by rearranging the order of summations, we can write the typical form of  $q_b^{(\gamma)} \prod_{k=1}^b \exp(it\hat{S}_N^{(j,k)})$  as follows:

$$(A.14) \quad \begin{aligned} & N^{-2m} (\sum_{k_1} \sum_{l_1} \dots \sum_{k_\alpha} \sum_{l_\alpha} \sum_{m_1} \sum_{m_\beta} \sum_{j_1} \dots \sum_{j_\gamma}) \\ & \times \tilde{V}_{\partial k_1 l_1} \dots \tilde{V}_{\partial k_j l_j} \tilde{V}_{k_{j+1} \partial l_{j+1}} \dots \tilde{V}_{k_\alpha l_\alpha \partial} \quad (A) \\ & \times \tilde{V}_{\partial \partial m_1} \dots \tilde{V}_{\partial \partial m_\varepsilon} \tilde{V}_{\partial m_{\varepsilon+1} \partial} \dots \tilde{V}_{m_\beta \partial \partial} \quad (B) \\ & \times \tilde{V}_{\partial \partial \partial_1} \dots \tilde{V}_{\partial \partial \partial_\eta} \quad (C) \\ & \times \prod_{\nu=1}^{\alpha} \exp(it(\hat{S}_N^{(k_\nu)} + \hat{S}_N^{(l_\nu)})) \prod_{\nu=1}^{\beta} \exp(it\hat{S}_N^{(m_\nu)}) \prod_{\nu=1}^{\gamma} \exp(it\hat{S}_N^{(j_\nu)}) \end{aligned}$$

where summations are taken over distinct indices; indices  $k$ 's,  $l$ 's and  $m$ 's appear only once;  $j$ 's appear at least twice among places marked by " $\partial$ ";  $\alpha, \beta, \gamma, \delta, \varepsilon$  and  $\eta$  are nonnegative integers such that

$$(A.15) \quad \begin{aligned} & 1 \leq s \leq \alpha/2 \leq m/2, \quad s \leq \gamma \leq m-s, \\ & 0 \leq \delta \leq \alpha \leq m, \quad 0 \leq \varepsilon \leq \beta \leq m, \\ & 2\alpha + \beta + \gamma = 2m + s, \quad \alpha + \beta + \gamma = m. \end{aligned}$$

We next consider the conditional expectation of (A.14) given r.v.'s  $X_{j_1}, X_{j_2}, \dots, X_{j_\gamma}$ . To this end, applying Lemma XV.4.1 of Feller [5] and then taking conditional expectations, we obtain that for distinct indices  $j, k$  and  $l$

$$(A.16) \quad \begin{aligned} & \sum_{(k,l) \neq} \sum E(\tilde{V}_{jkl} \tilde{e}_{kl} | X_j) \leq \sum_{(k,l) \neq} \sum |t| \{|c_k^*| + |c_l^*| + N^{-1}\omega_1\} \leq 3|t|N^{5/3}\omega_3^{1/3}, \\ & \sum_{(k,l) \neq} \sum E(\tilde{V}_{kji} \tilde{e}_{kl} | X_j) \leq \sum_{(k,l) \neq} \sum |t| |c_k^*| \{|c_k^*| + |c_l^*| + N^{-1}\omega_1\} \leq 3|t|N^{4/3}\omega_3^{2/3}, \\ & \sum_{(k,l) \neq} \sum E(V_{klj} \tilde{e}_{kl} | X_j) \leq 3|t|N^{4/3}\omega_3^{2/3} \end{aligned}$$

where  $\tilde{e}_{kl} = \exp(it(\hat{S}_N^{(k)} + \hat{S}_N^{(l)}))$ .

We finally split the proof into two cases when  $\alpha = m$  or  $\alpha < m$ . We first prove (A.10) when  $\alpha = m$  and then when  $\alpha < m$ .

When  $\alpha = m$ , we have only group (A) in (A.14). If  $s \geq 2$ , then the expectation of (A.14) is split into two groups and the induction hypothesis ensures (A.10) immediately. If  $s = 1$  and  $\delta = 0$ , then it follows by

substitution of (A.16) into (A.14) that the expectation of (A.14) is bounded by

$$N^{-2m}(3|t|N^{4/3}\omega_3^{2/3})^m N \leq (3|t|)^m \omega_3.$$

If  $s=1$  and  $\delta>0$  then the expectation is bounded by

$$N^{-2m}(3|t|N^{5/3}\omega_3^{1/3})^m \omega_1 \leq (3|t|)^m \omega_3.$$

This completes the proof for the case when  $\alpha=m$ .

We now suppose that  $\alpha<m$ . If some of indices  $j$ 's appear only in group (A) of (A.14), then we can again split the expectation into two groups. Thus it remains to show (A.10) when all indices  $j$ 's in group (A) appear again in group (B) or (C). Note that  $\beta+\gamma$  is the total number of indices  $m$ 's and  $j$ 's which appear in group (B) or (C) and that the number of indices  $m$ 's and  $j$ 's, which appear as one of the second or the third subscripts of  $\tilde{V}$ -terms in group (B) or (C), is at most  $2(\gamma+\beta)$ . Moreover,

$$\beta+\gamma-2(\gamma+\beta)=\gamma-2(m-\alpha-\beta)-\beta=s$$

in view of (A.15). Thus at least  $s$  indices among  $m_1, m_2, \dots, m_\beta, j_1, j_2, \dots, j_r$  must appear as the first subscript of one of  $\tilde{V}$ -terms in group (B) or (C). Utilizing this fact and substituting (A.16) into (A.14), we find that the expectation of (A.14) is bounded by

$$\begin{aligned} & N^{-2m}(3|t|N^{5/3}\omega_3^{1/3})^s(3|t|N^{4/3}\omega_3^{2/3})^{\alpha-\delta}\omega_1^s N^{2m+s-(2\alpha+s)} \\ & \leq (3|t|)^\alpha N^{-2\alpha+\delta/3+4\alpha/3+2s/3}\omega_3^{2\alpha/3-\delta/3+s/3} \\ & \leq (3(1+|t|))^m N^{-(\alpha-2s)/3}\omega_3^{(\alpha+s)/3} \leq (3(1+|t|))^m \omega_3 \end{aligned}$$

because  $2s\leq\alpha$  and  $\delta\leq\alpha$ . This completes the proof of the case when  $\alpha<m$ . The proof follows.

Let  $\{Y_j\}_{j=1}^\infty$  be a sequence of independent r.v.'s and  $\{d_j\}_{j=1}^\infty$  a sequence of real numbers. Then we have the following lemmas.

LEMMA A.1. *Let  $Z_j, j\geq 1$ , be r.v.'s of the form  $Z_j=g_j(Y_1, Y_2, \dots, Y_j)$  such that for  $j\geq 2$*

$$(A.17) \quad \mathbf{E}(Z_j|Y_1, Y_2, \dots, Y_{j-1})=0.$$

*If the sequence  $\{d_j\}$  is non-increasing in absolute values, then for any positive integers  $r$  and  $l$*

$$(A.18) \quad \mathbf{E}\left(\sum_{j=1}^l d_j Z_j\right)^{2r} \leq (4e)^r \left(\sum_{j=1}^l d_j^2\right)^r r^r m_l^{(r)}$$

where  $m_l^{(r)} = \max_{1\leq j\leq l} \mathbf{E} Z_j^{2r}$ .

LEMMA A.2. Let  $\tilde{V}_{jk}$  be r.v.'s of the form  $\tilde{V}_{jk} = g_{jk}(Y_j, Y_k)$ ,  $1 \leq j, k < \infty$ , such that for any  $j$  and  $k$ ,  $j \neq k$ ,

$$(A.19) \quad E(\tilde{V}_{jk}|Y_j) = E(\tilde{V}_{jk}|Y_k) = 0.$$

Then for any positive integers  $l$  and  $r$

$$(A.20) \quad E \hat{V}_l^{2r} \leq (4e)^{2r} \left( \sum_{j=1}^l d_j^2 \right)^r (2r)^{2r} l^{-r} \tilde{m}_l^{(r)}$$

where

$$\hat{V}_l = \frac{1}{l} \sum_{j=1}^l \sum_{k \neq j}^l d_j \tilde{V}_{jk} \quad \text{and} \quad \tilde{m}_l^{(r)} = \max_{\substack{1 \leq j, k \leq l \\ j \neq k}} E \tilde{V}_{jk}^{2r}.$$

PROOF OF LEMMA A.1. It follows by Hölder's inequality that

$$\left( \sum_{j=1}^l d_j Z_j \right)^{2r} \leq \left( \sum_{j=1}^l d_j^2 \right)^r \left( \sum_{j=1}^l Z_j^2 \right)^r \leq \left( \sum_{j=1}^l d_j^2 \right)^r l^{r-1} \sum_{j=1}^l Z_j^{2r}$$

which insures (A.18) when  $r \geq l$ .

For  $r \leq l$ , we prove (A.18) by induction on  $l$  with  $r$  fixed. Define  $V_l = \sum_{j=1}^l d_j Z_j$  and suppose that (A.18) is true for  $l = n \geq r$ . Then it follows by (A.17) and Hölder's inequality that

$$(A.21) \quad E V_{n+1}^{2r} \leq E V_n^{2r} + \sum_{v=2}^{2r} \binom{2r}{v} (E V_n^{2r})^{(2r-v)/2r} E (d_{n+1}^{2r} Z_{n+1}^{2r})^{v/2r} \\ \leq (4e)^r \left( \sum_{j=1}^n d_j^2 \right)^r r^r \left[ 1 + \frac{4r^2 d_{n+1}^2}{4er \sum_{j=1}^n d_j^2} \left( 1 + \frac{|d_{n+1}|}{\left( 4er \sum_{j=1}^n d_j^2 \right)^{1/2}} \right)^{2r-2} \right] m_{n+1}^{(r)}.$$

Since  $r \leq n$  and  $|d_j|$ 's are decreasing, we have

$$1 + |d_{n+1}| \left( 4er \sum_{j=1}^n d_j^2 \right)^{-1/2} \leq 1 + (2r)^{-1} \leq e^{1/(2r)}$$

which, together with (A.21), ensures (A.18) for  $l = n + 1$  to complete the proof.

PROOF OF LEMMA A.2. Since the assumption and the conclusion of this lemma are invariant under simultaneous permutation of  $d_j$ 's and  $Y_j$ 's, we may, without loss of generality, assume that  $|d_1| \geq |d_2| \geq \dots \geq |d_l|$ .

Define  $Z_1 = \tilde{Z}_1 = 0$  and

$$Z_j = \sum_{k=1}^{j-1} V_{jk}, \quad \tilde{Z}_j = \sum_{k=1}^{j-1} d_k V_{kj}, \quad 2 \leq j \leq l$$

so that  $\hat{V}_l = \frac{1}{l} \left( \sum_{j=1}^l d_j Z_j + \sum_{j=1}^l \tilde{Z}_j \right)$ . It follows by (A.19) that collections

$\{Z_j\}_{j=1}^l$  and  $\{\tilde{Z}_j\}_{j=1}^l$  satisfy the assumptions of Lemma A.1 and hence

$$(A.22) \quad \begin{aligned} \mathbb{E} \hat{V}_l^{2r} &\leq l^{-2r} 2^{2r-1} \left\{ \mathbb{E} \left( \sum_{j=1}^l d_j Z_j \right)^{2r} + \mathbb{E} \left( \sum_{j=1}^l \tilde{Z}_j \right)^{2r} \right\} \\ &\leq l^{-2r} 2^{2r-1} (4e)^r r^r \left\{ \left( \sum_{j=1}^l d_j^2 \right)^r \max_{1 \leq j \leq l} \mathbb{E} Z_j^{2r} + l^r \max_{1 \leq j \leq l} \mathbb{E} \tilde{Z}_j^{2r} \right\}. \end{aligned}$$

Because of (A.19) and the definition of  $V_{jk}$ 's, it follows that for each  $j$ , conditionally given  $Y_j$ ,  $Z_j$  and  $\tilde{Z}_j$  are sums of independent r.v.'s with zero means. Thus we may again apply Lemma A.1 to obtain that for  $1 \leq j \leq l$

$$(A.23) \quad \begin{aligned} \mathbb{E} Z_j^{2r} &\leq (4e)^r (j-1)^r r^r \max_{\substack{1 \leq k \leq l \\ k \neq j}} \mathbb{E} V_{jk}^{2r} \\ \mathbb{E} \tilde{Z}_j^{2r} &\leq (4e)^r \left( \sum_{k=1}^{j-1} d_k^2 \right)^r r^r \max_{\substack{1 \leq k \leq l \\ k \neq j}} \mathbb{E} V_{kj}^{2r}. \end{aligned}$$

The proof follows by combining (A.22) and (A.23).

INDIANA UNIVERSITY  
WRIGHT STATE UNIVERSITY

#### REFERENCES

- [1] Bergström, H. and Puri, M. L. (1977). Convergence and remainder terms in linear rank statistics, *Ann. Statist.*, **5**, 671-680.
- [2] Does, R. J. M. M. (1982a). *Higher order asymptotics for simple linear rank statistics*, Mathematisch Centrum, Amsterdam.
- [3] Does, R. J. M. M. (1982b). Berry-Esséen theorem for simple linear rank statistics under the null hypothesis, *Ann. Prob.*, **10**, 982-991.
- [4] Feller, W. (1968). *An Introduction to Probability Theory and Its Applications*, Vol. I, 3rd edition, Wiley, New York.
- [5] Feller, W. (1971). *An Introduction to Probability Theory and Its Applications*, Vol. II, 2nd edition, Wiley, New York.
- [6] Hájek, J. (1962). Asymptotically most powerfull rank order tests, *Ann. Math. Statist.*, **33**, 1129-1147.
- [7] Hájek, J. (1968). Asymptotic normality of simple linear rank statistics under alternatives, *Ann. Math. Statist.*, **39**, 325-346.
- [8] Hušková, M. (1970). Asymptotic distribution of simple linear rank statistics for testing symmetry, *Zeit. Wahrscheinlichkeitsth.*, **14**, 308-322.
- [9] Hušková, M. (1977). The rate of convergence of simple linear rank statistics under hypothesis and alternatives, *Ann. Statist.*, **5**, 658-670.
- [10] Hušková, M. (1979). The rate of convergence of simple linear rank statistics under alternatives, *Contributions to Statistics*, (ed. J. Jurečková), Academic, Publishing House, Prague, 99-108.
- [11] Jurečková, J. and Puri, M. L. (1975). Order of normal approximation for rank test statistics distribution, *Ann. Prob.*, **3**, 526-533.
- [12] Petrov, V. V. (1975). *Sums of independent random variables*, Springer-Verlag, Berlin.
- [13] Puri, M. L. and Seoh, M. (1985). Berry-Esséen theorems for signed linear rank statistics with regression constants, *In Colloquium on Limit Theorems in Probability and*

- Statistics*, Veszprem, Hungary (ed. P. Révész), 875-906.
- [14] Puri, M. L. and Seoh, M. (1984a). Edgeworth expansions for signed linear rank statistics with regression constants, *J. Statist. Plann. Inf.*, **10**, 137-149.
  - [15] Puri, M. L. and Seoh, M. (1984b). Edgeworth expansions for signed linear rank statistics under near location alternatives, *J. Statist. Plann. Inf.*, **10**, 289-309.
  - [16] Puri, M. L. and Wu, T. J. (1983). The order of normal approximation for signed linear rank statistics, *Teor. Veroyat. Primen.* (to appear).
  - [17] Ruymgaart, F. M. (1980). A unified approach to the asymptotic distribution theory of certain midrank statistics, *Lecture notes in Mathematics*, Springer Verlag, **821**, 1-18.
  - [18] von Bahr, B. (1976). Remainder term estimate in a combinatorial limit theorem, *Zeit Wahrscheinlichkeitsth.*, **35**, 131-139.