

EXTENSION OF THE INEQUALITY FOR THE VARIANCE OF  
 AN ESTIMATOR BY BAYESIAN PROCESS

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1. Introduction

Let  $X_1, X_2, \dots$  be a sequence of random variables identically distributed, whose distribution function depends on the parametric vector  $\theta \in \Omega_p \subset R^p$ . Let  $L_j(x/\theta)$  and  $G(\theta)$ , defined over  $\Omega_p$  be the conditional likelihood function for

$$x = (x_1, x_2, \dots, x_j), \quad j = 1, 2, \dots$$

and the distribution of  $\theta$ , respectively. Let us consider the same sequential process described by Wolfowitz [6] and also the existence of a estimating function  $g_{N,h}(x, \theta_h)$  for  $\theta_h$ , with  $1 \leq h \leq p$ , defined by Godambe [4].

2. Regularity conditions

We now postulate the following regularity conditions to be satisfied by  $L_j(x/\theta)$  and  $g_{N,h}(x, \theta_h)$

- i)  $\theta_h \in \Omega \subset R$  for  $1 \leq h \leq p$ .
- ii) There exists a set  $A_j, j=1, 2, \dots$ , independent of  $\theta_h, \Pr(A_j) = 0$ , such that for all

$$(x_1, x_2, \dots, x_j) \notin A_j, \quad \text{we have that}$$

$$\frac{\partial^r}{\partial \theta_h^r} L_j(x/\theta)$$

exists for  $h=1, 2, \dots, p$  and  $r=1, 2, \dots, k$ . Now if  $L_j(x/\theta) = 0$ , we define

$$H_{j,h}^r(x/\theta) = 0, \quad \text{that is,}$$

$H_{j,h}^r(x/\theta)$  is defined completely for all  $\theta_h \in \Omega$  ( $1 \leq h \leq p$ ) and for almost all  $x \in R_j$ , where  $H_{j,h}^r(x/\theta) = (1/L_j(x/\theta)) (\partial^r / \partial \theta_h^r) L_j(x/\theta)$ .

- iii)  $E_\theta E [g_{N,h}(X, \theta_h) / \theta] = 0$  for  $1 \leq h \leq p$  and all  $\theta \in \Omega_p$ .

iv) There exists a set  $B_j$ ,  $j=1, 2, \dots$ , independent of  $\theta_h$  ( $1 \leq h \leq p$ ),  $\Pr(B_j)=0$ , such that

$$\frac{\partial^r}{\partial \theta_h^r} g_{j,h}(x, \theta_h)$$

exists for all

$$(x_1, x_2, \dots, x_j) \notin B_j \quad \text{and} \quad r=1, 2, \dots, k.$$

v) For  $j=1, 2, \dots, h$  and  $r=1, 2, \dots, k$ , let

$$T_j^r(x_1, x_2, \dots, x_j)$$

be non-negative and  $L$ -measurable functions.

a) There exists a set  $C_j$ ,  $j=1, 2, \dots$  independent of

$$\theta_h \quad (1 \leq h \leq p), \quad \Pr(C_j)=0, \quad \text{such that,}$$

for all  $(x_1, x_2, \dots, x_j) \notin C_j$ , we have

$$\left| \frac{\partial^r}{\partial \theta_h^r} g_{j,h}(x, \theta_h) L_j(x/\theta) \right| < T_j^r(x_1, x_2, \dots, x_j).$$

$$\text{b) } \int_{R_j} T_j^r(x_1, x_2, \dots, x_j) \prod_{i=1}^j dx_i < \infty.$$

vi) Let us define for  $1 \leq h \leq p$  and  $j=1, 2, \dots$ ,

$$t_j(\theta_h) = \int_{R_j} g_{j,h}(x, \theta_h) L_j(x/\theta) \prod_{i=1}^j dx_i,$$

and let us also postulate the uniform convergence of the following series

$$\sum_{j=1}^{\infty} \frac{d^r}{d\theta_h^r} t_j(\theta_h) \quad \text{for } r=1, 2, \dots, k.$$

vii) For each  $j=1, 2, \dots$ , there exist functions

$$S_j^r(x_1, x_2, \dots, x_j) \quad \text{for } r=1, 2, \dots, k,$$

such that, when  $g_{j,h}(x, \theta_h)$  and

$$T_j^r(x_1, x_2, \dots, x_j)$$

are replaced by unity and  $S_j^r(x_1, x_2, \dots, x_j)$ , respectively, the conditions v) and vi) remain valid.

viii) The covariance matrix  $\|\lambda_{rs}\|$  of  $H_{N,h}^r(x/\theta)$ ,  $r=1, 2, \dots, k$  exists and is non-singular.

3. Theorem

Under the above conditions, we can show that

$$(3.1) \quad E_{\theta} E [g_{N,h}^2(X, \theta_h)/\theta] \geq \sum_{r,s=1}^k \lambda^{rs} E_{\theta} E [W_{N,h}^{r,0}(X, \theta_h)/\theta] \cdot E_{\theta} E [W_{N,h}^{s,0}(X, \theta_h)/\theta] ,$$

where  $W_{N,h}^{k,i}(x, \theta_h) = H_{N,h}^k(x/\theta)(\partial^i/\partial\theta_h^i)g_{N,h}(x, \theta_h)$  and  $\lambda^{rs}$  is the element of the  $r$ th line and  $s$ th column of the inverse matrix of  $\|\lambda_{rs}\|$ .

PROOF. Let

$$(3.2) \quad D = g_{N,h}(x, \theta_h) - \sum_{r=1}^k \alpha_r H_{N,h}^r(x/\theta)$$

where  $\alpha_r$ 's are independent of  $X$  and  $\theta$  for  $r=1, 2, \dots, k$ . According to the sequential procedure developed by Wolfowitz [6], we can write

$$(3.3) \quad \sum_{j=1}^{\infty} \int_{R_j} L_j(x/\theta) \prod_{i=1}^j dx_i = 1 .$$

Now taking the  $r$ th derivative with respect to  $\theta_h$  ( $1 \leq h \leq p$ ) in both sides of (3.3), we get by virtue of condition vii)

$$\sum_{j=1}^{\infty} \int_{R_j} H_{j,h}^r(x/\theta) L_j(x/\theta) \prod_{i=1}^j dx_i = 0 ,$$

that is,

$$(3.4) \quad E [H_{N,h}^r(X/\theta)/\theta] = 0 \quad \text{for } r=1, 2, \dots, k .$$

From iii) and (3.4), it follows that

$$E_{\theta} E (D/\theta) = 0 .$$

So

$$(3.5) \quad \text{Var} (D) = E_{\theta} E (D^2/\theta) = E_{\theta} E \left\{ \left[ g_{N,h}(X, \theta_h) - \sum_{r=1}^k \alpha_r H_{N,h}^r(X/\theta) \right]^2 / \theta \right\} .$$

Equation (3.5) can be written like

$$(3.6) \quad \text{Var} (D) = E_{\theta} E [g_{N,h}^2(X, \theta_h)/\theta] - 2 \sum_{r=1}^k \alpha_r A^r + \sum_{r,s=1}^k \alpha_r \alpha_s \lambda_{rs} ,$$

where

$$\lambda_{rs} = E_{\theta} E [H_{N,h}^r(X/\theta) \cdot H_{N,h}^s(X/\theta)/\theta] = \text{Cov} [H_{N,h}^r(X/\theta), H_{N,h}^s(X/\theta)]$$

and

$$A^r = E_{\theta} E [W_{N,h}^{r,0}(X, \theta_h)/\theta] = \text{Cov} [g_{N,h}(X, \theta_h), H_{N,h}^r(X/\theta)] .$$

We propose now to determine the values of  $\alpha_r$  for  $r=1, 2, \dots, k$  that minimize (3.6). Condition viii) says that  $\|\lambda_{rs}\|$  is non-singular; then these values of  $\alpha_r$  ( $r=1, 2, \dots, k$ ) are given by

$$(3.7) \quad \alpha_{rs} = \sum_{s=1}^k \lambda^{rs} A^s .$$

Taking the values of  $\alpha_r$  ( $r=1, 2, \dots, k$ ) given by (3.7), into (3.6), we get

$$(3.8) \quad \text{Var}(D) = E_{\theta} E [g_{N,h}^2(X, \theta_h)/\theta] - \sum_{r,s=1}^k \lambda^{rs} A^r A^s .$$

Since  $\text{Var}(D)$  is non-negative, from (3.8) it follows that

$$(3.9) \quad E_{\theta} E [g_{N,h}^2(X, \theta_h)/\theta] \geq \sum_{r,s=1}^k \lambda^{rs} A^r A^s .$$

Obviously, by the definition of  $A^r$ , equation (3.9) can be written as

$$E_{\theta} E [g_{N,h}^2(X, \theta_h)/\theta] \geq \sum_{r,s=1}^k \lambda^{rs} E_{\theta} E [W_{N,h}^{r,0}(X, \theta_h)/\theta] \cdot E_{\theta} E [W_{N,h}^{s,0}(X, \theta_h)/\theta] ,$$

which is the result expressed by (3.1).

#### 4. Application

From inequality (3.1), for  $1 \leq h \leq s$ , let

$$g_{N,h}(X, \theta_h) = \partial_{N,h}(X) - \theta_h - E [\Psi_h(\theta)] + E(\theta_h)$$

be, such that

$$E [\partial_{N,h}(X)] = \Psi_h(\theta) .$$

Thus, we get

$$(4.1) \quad E_{\theta} E \{[\partial_{N,h}(X) - \theta_h]^2/\theta\} \\ \geq E_{\theta}^2 [\Psi_h(\theta) - \theta_h] + \sum_{r,s=1}^k \lambda^{rs} E_{\theta} \left[ \frac{\partial^r}{\partial \theta_h^r} \Psi_h(\theta) \right] E_{\theta} \left[ \frac{\partial^s}{\partial \theta_h^s} \Psi_h(\theta) \right] .$$

When  $N$  is fixed and equal to  $n$ , (4.1) can be written as

$$(4.2) \quad E_{\theta} E \{[\partial_{n,h}(X) - \theta_h]^2/\theta\} \\ \geq E_{\theta}^2 [\Psi_h(\theta) - \theta_h] + \sum_{r,s=1}^k \lambda^{rs} E_{\theta} \left[ \frac{\partial^r}{\partial \theta_h^r} \Psi_h(\theta) \right] E_{\theta} \left[ \frac{\partial^s}{\partial \theta_h^s} \Psi_h(\theta) \right]$$

which is the inequality obtained by Gart [3]. Now, when  $\theta$  is constant and equal to  $\theta$  for all  $1 \leq h \leq p$ , then (4.1) becomes

$$(4.3) \quad E \{[\partial_N(X) - \theta]^2 / \theta\} - E^2 \{\partial_N(X) - \theta\} \geq \sum_{r,s=1}^k \lambda^{rs} \frac{d^r}{d\theta^r} \Psi(\theta) \frac{d^s}{d\theta^s} \Psi(\theta) .$$

From (4.3) it follows that

$$(4.4) \quad \text{Var} [\partial_N(X)] \geq \sum_{r,s=1}^k \lambda^{rs} \frac{d^r}{d\theta^r} \Psi(\theta) \frac{d^s}{d\theta^s} \Psi(\theta) ,$$

which is the inequality obtained by Seth [5].

Finally, from (4.3) we get inequality proved by Wolfowitz [6] and from (3.1) inequalities obtained by Campos [2], which contains the Bhattacharyya [1] and Godambe [4] inequalities, setting appropriate conditions.

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