

ON THE UNIFORM COMPLETE CONVERGENCE OF DENSITY FUNCTION ESTIMATES

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Abstract

Let f be a uniformly continuous density function. Let W be a non-negative weight function which is continuous on its compact support $[a, b]$ and $\int_a^b W(x)dx=1$. The complete convergence of

$$\sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-X_k}{b(n)}\right) - f(s) \right|$$

to zero is obtained under varying conditions on the bandwidths $b(n)$, support or moments of f , and smoothness of W . For example, one choice of the weight function for these results is Epanechnikov's optimal function and $nb^2(n) > n^\delta$ for some $\delta > 0$. The uniform complete convergence of the mode estimate is also considered.

1. Introduction and preliminaries

The construction of a family of estimates of a density function $f(x)$ and of the mode has been studied by several people. Rosenblatt [9] considered a general class of density estimates:

$$(1.1) \quad f_n(x) = \frac{1}{nb(n)} \sum_{i=1}^n W\left(\frac{x-X_i}{b(n)}\right),$$

where X_1, \dots, X_n are i.i.d. random variables with continuous density function $f(x)$, $W(x)$ is a bounded integrable weight function such that

$$\int_{-\infty}^{\infty} W(x)dx = 1$$

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and $b(n)$ is a bandwidth that tends to zero as $n \rightarrow \infty$. Thus, the question arises as to suitable choices of $W(x)$ and $b(n)$ so that the estimate function $f_n(x)$ is optimal (in some sense). The local properties of the estimate function in (1.1) have been studied extensively (see Rosenblatt [10] for a general survey), and a global measure of deviation of the curve $f_n(x)$ from $f(x)$ by

$$(1.2) \quad \|f_n - f\|_\infty = \sup_{x \in R} |f_n(x) - f(x)|$$

has been considered. Parzen [8] showed that if the (true) underlying density function $f(x)$ is uniformly continuous then $\|f_n - f\|_\infty$ converges in probability to zero under the following conditions:

$$(P1) \quad \phi_w(t) = \int_{-\infty}^{\infty} e^{itx} W(x) dx \text{ is absolutely integrable,}$$

$$(P2) \quad \sqrt{n} b(n) \rightarrow \infty \text{ as } n \rightarrow \infty.$$

The results of Nadaraya [7], Woodroffe [15], Deheuvels [4], Devroye and Wagner [5], and Silverman [12] on uniform consistency in the strong sense are discussed in the last section for comparison with the results of this paper.

The major results of the paper give a new class of "good" weight functions (which includes the optimal function) under mild conditions on the bandwidth sequence $b(n)$ where uniform consistency of the estimate $f_n(x)$ is obtained by the complete convergence (see Stout [13]) of $\|f_n - f\|_\infty$ to zero (which implies convergence with probability one). The main tools used in obtaining these results will be the smoothness of the weight function and sub-Gaussian techniques.

Throughout this paper, attention will be restricted to a density function which is uniformly continuous (or is continuous on its compact support $[a, b]$) (see Schuster [11] for a discussion of necessity of uniform continuity) and weight functions $W(x)$ which satisfy

$$(i) \quad \int_a^b W(x) dx = 1 \text{ and}$$

$$(ii) \quad W(x) \text{ is nonnegative and continuous on } [a, b] \text{ and vanishes outside } [a, b].$$

Let U_n be a polygonal approximating function on the space of continuous functions with domain $[a, b]$, $C[a, b]$. That is,

$$g\left(a + \frac{(b-a)i}{n}\right) = [U_n(g)]\left(a + \frac{(b-a)i}{n}\right)$$

for $i=0, 1, \dots, n$ and $g \in C[a, b]$, and U_n is linear between the points $a + (b-a)i/n$ and $a + (b-a)(i+1)/n$. Recall that the modulus of continuity, $\omega_g(\delta)$, is defined by Billingsley [2] by

$$(1.3) \quad \omega_g(\gamma) = \sup_{|t-s| \leq \gamma} |g(t) - g(s)|$$

for $\gamma > 0$, $s, t \in [a, b]$, and $g \in C[a, b]$.

DEFINITION (Chow [3]). A random variable X is said to be *sub-Gaussian* if there exists $\alpha \geq 0$ such that

$$(1.4) \quad E[\exp(tX)] \leq \exp\left(\frac{\alpha^2 t^2}{2}\right) \quad \text{for all } t \in R.$$

If X is sub-Gaussian, then let

$$\tau(X) = \inf \{ \alpha \geq 0 : \text{Inequality (1.4) holds} \}.$$

Some basic properties on sub-Gaussian random variables include:

1. If $P[|X| \leq K] = 1$ and $E X = 0$, then

$$(1.5) \quad E[\exp(tX)] \leq \exp(K^2 t^2).$$

2. If $\tau(X) = \alpha$, then

$$(1.6) \quad P[|X| \geq \lambda] \leq 2 \exp(-\lambda^2/2\alpha^2).$$

3. The sum of two independent sub-Gaussian random variables is sub-Gaussian.

Finally, a sequence of random variables $\{X_n\}$ is said to *converge completely* to a random variable X if

$$(1.7) \quad \sum_{n=1}^{\infty} P[|X_n - X| > \varepsilon] < \infty$$

for each $\varepsilon > 0$. Thus, complete convergence implies convergence with probability one by Boole's inequality.

2. Main results

In this section the complete convergence of $\|f_n - f\|_{\infty}$ to zero is obtained under conditions on the modulus of continuity of the weight function $W(x)$ and the rate of convergence to zero by the bandwidth $b(n)$. Also, the uniform consistency of the mode estimate is obtained in this setting. The uniform consistency of the estimate $f_n(x)$ (in the complete sense) is accomplished by two lemmas.

LEMMA 1. (i) If $nb^2(n) > n^{\delta}$ for some $\delta > 0$ and (ii) $\int |x|^p f(x) dx < \infty$ for some $p > 0$, and (iii) $\omega_w((2b - 2a)/n^r b(n)) = o(b(n))$ for some integer $r > 1/p$, then

$$(2.1) \quad \sup_{-\infty < s < \infty} \left| f_n(s) - \frac{1}{b(n)} E W\left(\frac{s - X_1}{b(n)}\right) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$ where $f_n(s)$ is defined in (1.1).

PROOF. From (ii) it follows that

$$(2.2) \quad \sum_{n=1}^{\infty} P[|X_1|^p > n] < \infty.$$

Thus, define

$$Y_n = X_n I_{[|X_n| \leq n^{1/p}]}$$

for each n . By (2.2) $P[X_n \neq Y_n \text{ i.o.}] = 0$ since

$$(2.3) \quad \sum_{n=1}^{\infty} P[Y_n \neq X_n] = \sum_{n=1}^{\infty} P[|X_n| > n^{1/p}] < \infty.$$

Next,

$$(2.4) \quad \begin{aligned} & \sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-X_k}{b(n)}\right) - \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) \right| \\ & \leq \sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n \left[W\left(\frac{s-X_k}{b(n)}\right) - W\left(\frac{s-Y_k}{b(n)}\right) \right] \right| \\ & \quad + \sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-Y_k}{b(n)}\right) - \frac{1}{nb(n)} \sum_{k=1}^n E W\left(\frac{s-Y_k}{b(n)}\right) \right| \\ & \quad + \sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n E W\left(\frac{s-Y_k}{b(n)}\right) - \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) \right|. \end{aligned}$$

The first and third terms of the right-hand side of Inequality (2.4) converge to 0 completely by (2.3) and the boundedness of W .

Using the compact support $[a, b]$ for W ,

$$\frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-Y_k}{b(n)}\right) = 0 \quad \text{unless } a \leq \frac{s-Y_k}{b(n)} \leq b$$

for some $1 \leq k \leq n$ or unless $-n^{1/p} + ab(n) \leq s \leq n^{1/p} + bb(n)$. Since $b(n) \rightarrow 0$, the $\sup_{-\infty < s < \infty}$ in the second term of (2.4) need only be taken over $[-n^{1/p} + a, n^{1/p} + b]$. Let $\delta_n = 2(b-a)/n^r b(n)$ and let $t_i = -n^{1/p} + a + (i(2n + (b-a))/n^{2r})$ for $1 \leq i \leq n^{2r}$. Hence, $t_i - t_{i-1} = (2n^{1/p} + (b-a))/n^{2r} \leq 2(b-a)/n^r$ for n large enough. Let $\tilde{W}_k(s) = W(s - Y_k/b(n)) - E W(s - Y_k/b(n))$ for each $k = 1, \dots, n$. Thus, $E \tilde{W}_k(s) = 0$ for each $s \in [a, b]$ and each k . Furthermore,

$$(2.5) \quad \begin{aligned} \omega_{\tilde{W}_k}(\delta_n) &= \sup_{|t-s| \leq \delta_n} |\tilde{W}_k(s) - \tilde{W}_k(t)| \\ &\leq \sup_{|t-s| \leq \delta_n} \left| W\left(s - \frac{Y_k}{b(n)}\right) - W\left(t - \frac{Y_k}{b(n)}\right) \right| \\ &\quad + \sup_{|t-s| \leq \delta_n} \left| E W\left(s - \frac{Y_k}{b(n)}\right) - E W\left(t - \frac{Y_k}{b(n)}\right) \right| \\ &\leq 2\omega_W(\delta_n). \end{aligned}$$

Hence, $\omega_{\tilde{W}_k}(\delta_n) \leq 2\omega_W(\delta_n) = o(b(n))$ for each k from condition (iii). For $\varepsilon > 0$ let

$$(2.6) \quad A_n = \left[\sup_{-n^{1/p} + a \leq s \leq n^{1/p} + b} \left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{s}{b(n)} \right) \right| > \varepsilon \right] \\ = \left[\max_{1 \leq i \leq n^{2r}} \sup_{s \in I_i} \left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{s}{b(n)} \right) \right| > \varepsilon \right]$$

where $I_i = [t_{i-1}, t_i]$. Hence,

$$(2.7) \quad A_n \subset \left[\max_{1 \leq i \leq n^{2r}} \left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right| \right. \\ \left. + \max_{1 \leq i \leq n^{2r}} \sup_{s \in I_i} \left| \frac{1}{nb(n)} \sum_{k=1}^n \left[\tilde{W}_k \left(\frac{s}{b(n)} \right) - \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right] \right| > \varepsilon \right].$$

However,

$$(2.8) \quad \max_{1 \leq i \leq n^{2r}} \sup_{s \in I_i} \left| \frac{1}{nb(n)} \sum_{k=1}^n \left[\tilde{W}_k \left(\frac{s}{b(n)} \right) - \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right] \right| \leq \frac{2}{b(n)} \omega_W(\delta_n).$$

Since $\omega_W(\delta_n) = o(b(n))$ by condition (iii), there exists $N(r)$ such that

$$A_n \subset \left[\max_{1 \leq i \leq n^{2r}} \left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right| > \frac{\varepsilon}{2} \right]$$

for all $n \geq N(r)$. Using the basic properties of sub-Gaussian random variables ($\{\tilde{W}_k(t_i/b(n)) : k=1, 2, \dots\}$ for each i), for each $n \geq N(r)$

$$(2.9) \quad P(A_n) \leq P \left[\max_{1 \leq i \leq n^{2r}} \left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right| > \frac{\varepsilon}{2} \right] \\ \leq \sum_{i=1}^{n^{2r}} P \left[\left| \frac{1}{nb(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_i}{b(n)} \right) \right| > \frac{\varepsilon}{2} \right] \\ \leq n^{2r} 2 \exp [-\varepsilon^2/64 \|W\|_\infty^2 B_n]$$

where $\|W\|_\infty = \sup_s |W(s)|$ and $B_n = \sum_{k=1}^n (1/nb(n))^2 = 1/nb^2(n)$. To obtain the complete convergence in (2.1), consider

$$(2.10) \quad \sum_{n=1}^{\infty} P(A_n) = \sum_{n=1}^{N(r)} P(A_n) + \sum_{n=N(r)+1}^{\infty} P(A_n) \\ \leq N(r) + \sum_{n=N(r)+1}^{\infty} 2n^{2r} \exp \left(\frac{-\varepsilon^2 nb^2(n)}{64 \|W\|_\infty^2} \right) \\ \leq N(r) + \sum_{n=N(r)+1}^{\infty} 2n^{2r} \exp(-cn^\delta)$$

where $c = \varepsilon^2/64 \|W\|_\infty^2$. Thus, the series in (2.10) converges by the integral test.

LEMMA 2. *If the underlying density, f , is uniformly continuous, then*

$$(2.11) \quad \sup_{-\infty < s < \infty} \left| \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) - f(s) \right| \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

PROOF. Since f is uniformly continuous given $\varepsilon > 0$ there exists $\delta > 0$ such that $|f(x) - f(x')| < \varepsilon$ whenever $|x - x'| < \delta$. Let N be sufficiently large so that $|b(n)y| < \delta$ for all $n \geq N$ and $y \in [a, b]$. Since $W(y) = 0$ for $y \notin [a, b]$,

$$(2.12) \quad \begin{aligned} & \left| \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) - f(s) \right| \\ &= \left| \frac{1}{b(n)} \int_{-\infty}^{\infty} W\left(\frac{s-x}{b(n)}\right) f(x) dx - f(s) \right| \\ &= \left| \int_{-\infty}^{\infty} W(y) f(s-b(n)y) dy - f(s) \right| \\ &= \left| \int_a^b W(y) [f(s-b(n)y) - f(s)] dy \right| \\ &< \varepsilon \int_a^b W(y) dy = \varepsilon \end{aligned}$$

uniformly in s for all $n \geq N$. Hence,

$$\sup_{-\infty < s < \infty} \left| \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) - f(s) \right| \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Thus, the proof of Theorem 1 is immediate from Lemmas 1 and 2 since for each $\varepsilon > 0$

$$(2.13) \quad \begin{aligned} & P \left[\sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-X_k}{b(n)}\right) - f(s) \right| > \varepsilon \right] \\ &\leq P \left[\sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n \left(W\left(\frac{s-X_k}{b(n)}\right) - E W\left(\frac{s-X_1}{b(n)}\right) \right) \right| > \frac{\varepsilon}{2} \right] \\ &\quad + P \left[\sup_{-\infty < s < \infty} \left| \frac{1}{b(n)} E W\left(\frac{s-X_1}{b(n)}\right) - f(s) \right| > \frac{\varepsilon}{2} \right] \end{aligned}$$

and each of the terms in (2.13) is a convergent series in n . All of the conditions will be stated in Theorem 1 for easy reference.

THEOREM 1. *Let $\{X_n\}$ be independent random variables with the same density function $f(s)$ which is uniformly continuous. Let $W(x)$ be a nonnegative weight function which is continuous on its compact support and integrates to 1. If*

- (a) $nb^2(n) > n^\delta$ for some $\delta > 0$,
- (b) $\int |x|^p f(x) dx < \infty$ for some $p > 0$, then

$$(c) \quad \omega_W\left(\frac{2b-2a}{n^r b(n)}\right) = o(b(n)) \text{ for some integer } r > \frac{1}{p}, \text{ then } \sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{i=1}^n W\left(\frac{s-X_i}{b(n)}\right) - f(s) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$.

The case of the density function having compact support is discussed in Taylor and Cheng [14] and being discontinuous off of its support is not entirely excluded in Theorem 1. The following steps indicate modifications which allows the theory to include a large class of density functions (for example, the uniform densities).

Step 1. For an unknown density function which is continuous only on $[a, b]$ and vanishes outside $[a, b]$, there is no change in Lemma 1.

Step 2. In Lemma 2 it is easy to verify that

$$\sup_{a+b(n)C \leq s \leq b-b(n)C} \left| E \frac{1}{b(n)} W\left(\frac{s-X_1}{b(n)}\right) - f(s) \right| \rightarrow 0$$

as $n \rightarrow \infty$ where $C = \max\{|a|, |b|\}$.

Step 3. Combining steps 1 and 2, for each $\varepsilon > 0$

$$\sup_{a+b(n)C \leq s \leq b-b(n)C} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-X_k}{b(n)}\right) - f(s) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$.

Hence, the complete convergence of the maximal deviation of the density estimate holds on arbitrary closed intervals inside of $[a, b]$. Similar consideration was also given in Woodroffe [15].

In Lemma 1 the modulus of continuity was used only to replace $f_n(s)$ by a polygonal approximation. Thus, the following corollary can be obtained with basically the same proof.

COROLLARY 1. *Let the density function $f(s)$ be as stated in Theorem 1. Let $W(x)$ be a nonnegative weight function which has compact support and integrates to 1. If*

(a) $nb^3(n) > n^\delta$ for some $\delta > 0$, and

(b') $\sup_{-n^{1/p+a} \leq s \leq n^{1/p+b}} |W(s) - U_{n^{2r}}(W)(s)| = o(b(n))$, then

$$\sup_{-\infty < s < \infty} \left| \frac{1}{nb(n)} \sum_{k=1}^n W\left(\frac{s-X_k}{b(n)}\right) - f(s) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$.

The condition $nb^3(n) > n^\delta$ need not hold for all n but only eventually. Also, the condition can be stated as

$$(a') \quad \int_d^\infty x^r \exp(-cxb^2(x))dx < \infty$$

for some $d > 0$ where $b(x)$ is a function which generates the bandwidths $b(1), b(2), \dots$ and c is a constant.

In considering mode estimates, assume that the continuous density function $f(s)$ has a unique mode θ , that is,

$$f(\theta) = \max_{-\infty < s < \infty} f(s).$$

The sample mode θ_n is also assumed to uniquely satisfy

$$f_n(\theta_n) = \max_{-\infty < s < \infty} f_n(s) \quad \text{for each } n.$$

THEOREM 2. *If the regularity conditions of Theorem 1 or Corollary 1 (or condition (b')) are satisfied, then*

$$|\theta_n - \theta| \rightarrow 0$$

completely as $n \rightarrow \infty$.

PROOF. Since $f(s)$ is uniformly continuous and has a unique mode θ , for $\varepsilon > 0$ there exists $\eta > 0$ such that $|x - \theta| \geq \varepsilon$ implies that $|f(\theta) - f(x)| \geq \eta$. Thus, it suffices to show that $f(\theta_n) \rightarrow f(\theta)$ completely. But,

$$\begin{aligned} (2.14) \quad |f(\theta_n) - f(\theta)| &\leq |f(\theta_n) - f_n(\theta_n)| + |f_n(\theta_n) - f(\theta)| \\ &\leq \sup_{-\infty < s < \infty} |f(s) - f_n(s)| + \left| \max_{-\infty < s < \infty} f_n(s) - \max_{-\infty < s < \infty} f(s) \right| \\ &\leq 2 \sup_{-\infty < s < \infty} |f_n(s) - f(s)| \end{aligned}$$

pointwise for each n . From (2.14) and the complete convergence of $\|f_n - f\|_\infty$, it follows that $|\theta_n - \theta| \rightarrow 0$ completely.

3. Comparisons and useful weight functions

Brief comments and comparisons of these results with existing results will be listed in this section. Also, some useful weight functions will be considered.

Nadaraya [7] had the weaker bandwidth condition, $\sum_{n=1}^{\infty} \exp(-rnb^2(n)) < \infty$ for each $r > 0$, but required W to be of bounded variation. Woodroffe [15] also considered weight functions with compact support. In addition, his conditions included: $W \in \text{LIP}(\beta)$, $0 < \beta \leq 1$, and $b(n)^{-r} = o(n)$ with $n = o(b(n)^{-\delta})$, $1 < r < \delta$. It will be shown that $W \in \text{LIP}(\beta)$, $\beta > 0$ or $\beta < -1$, is sufficient for the smoothness condition of Theorem 1(c).

Deheuvels [4] also used the p th moment condition of Theorem 1(b) in relating necessary and sufficient conditions on the bandwidth sequence

(namely, $1/b(n)=o(n/\log n)$) and the almost sure convergence of $\sup_s |f_n(s) - f(s)|$ for a Riemann integrable weight function $W(x, y)$. In the nearest neighbor method, Devroye and Wagner [5] required only the condition $nb(n)/\log n \rightarrow \infty$ and $b(n)/n \rightarrow 0$ in obtaining almost sure convergence for bounded weight functions having compact support. Finally, Silverman [12] required the bandwidth condition of $b(n) \rightarrow 0$ and $\log n/nb(n) \rightarrow 0$ in obtaining almost sure convergence when W is uniformly continuous and of bounded variation.

For the results of this paper, the weight function $W(x)$ needed to be continuous on its compact support and satisfy a smoothness condition ((b) or (b')). However, it need not be of bounded variation even if it satisfies a Lipschitz condition (which yields (b) or (b')). Some useful weight functions for these results will now be listed. First, Epanechnikov's [6] optimal weight function

$$W(x) = \begin{cases} \frac{3}{4\sqrt{5}} \left(1 - \frac{x^2}{5}\right) & \text{if } |x| \leq \sqrt{5} \\ 0 & \text{otherwise} \end{cases}$$

can be used. In this case, let $a = -5$ and $b = 5$. Then $|W(x) - W(y)| \leq C|x - y|$ and $\omega_w((2b - 2a)/b(n)n^r) \leq C'(1/b(n)n^r)$ for constants C and C' . Thus, condition (c) is easily satisfied.

If the weight function W satisfies a Lipschitz condition of order α , then $|W(x) - W(y)| < M|x - y|^\alpha$ and

$$\omega_w\left(\frac{1}{b(n)n^r}\right) \leq M\left(\frac{1}{b(n)n^r}\right)^\alpha$$

for some $M > 0$. Thus, the bandwidth sequence $b(n)$ must be chosen so that

$$(3.1) \quad b^{\alpha+1}(n)n^{\alpha r} \rightarrow \infty \quad \text{as } n \rightarrow \infty \text{ for some } r > 0.$$

Case 1. $-1 \leq \alpha \leq 0$. No bandwidth exists for (3.1).

Case 2. $\alpha > 0$. If $b(n) = n^{-p}$ for some $p > 0$, then r is an integer $\geq 2p(1+\alpha)/\alpha$. Then, $b^{1+\alpha}(n)n^{\alpha r} \geq n^{p(1+\alpha)} \rightarrow \infty$, and (3.1) is satisfied.

Case 3. $\alpha < -1$. Again, if $b(n) = n^{-p}$ for some $p > 0$, then r is an integer $\geq -p(\alpha+1)/2\alpha$. Then $b^{1+\alpha}(n)n^{\alpha r} \geq n^{-p(1+\alpha)/2} \rightarrow \infty$, and (3.1) is satisfied.

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