

AN ASYMPTOTIC EXPANSION FOR THE DISTRIBUTION OF  
A FUNCTION OF LATENT ROOTS OF THE NONCENTRAL  
WISHART MATRIX, WHEN  $\Omega=O(n)$

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1. Introduction

Let  $S (=XX', X=(x_{\alpha j}): p \times n)$  be a  $p \times p$  noncentral Wishart matrix having  $W(\Sigma, n; \Omega)$ . Here we shall define the matrix of noncentrality parameters  $\Omega$  by  $\Sigma^{-1}MM'/2$  instead of  $\Sigma^{-1}MM'$ , for  $M=(m_{\alpha j})=\mathbf{E}(X)$ . In the previous paper [6], we gave the asymptotic expansion of the distribution of the generalized variance under a natural assumption of  $\Omega=O(n)$ . In this paper we shall derive the asymptotic expansion for the distribution of a function of latent roots of  $S$  under the above assumption. The method adopted here is based on a perturbation expansion for latent roots and the work of Nagao [5].

2. Preliminary lemmas

So far as we are concerned with the distribution for latent roots of  $S$ , we can take that  $S$  has  $W(\Lambda=\text{diag}(\lambda_1, \dots, \lambda_p), n; \Omega)$  and further we assume  $\Omega=\Lambda^{-1}MM'/2=m\theta=m \text{ diag}(\theta_1, \dots, \theta_p)$ , where  $m=n-2\Delta$  with  $\Delta=O(1)$ .

**LEMMA 2.1.** *Let  $V=(v_{ij})=\sqrt{m}\{S/m-\Lambda(I+2\theta)\}$ , where  $m=n-2\Delta$  for a correction factor  $\Delta=O(1)$ . Then  $v=(v_{11}, \dots, v_{pp}, v_{12}, \dots, v_{p-1,p})$  converges in law to a  $p(p+1)/2$  variate normal distribution with mean zero and covariance matrix  $\Psi$ , where, by putting  $c_\alpha=1+4\theta_\alpha$ ,*

$$(2.1) \quad \Psi = 2 \begin{pmatrix} c_1 \lambda_1^2 & & & & & \\ & \ddots & & & & \\ & & c_p \lambda_p^2 & & & \\ & & & \frac{1}{4} (c_1 + c_2) \lambda_1 \lambda_2 & & 0 \\ & & & & \ddots & \\ 0 & & & & & \frac{1}{4} (c_{p-1} + c_p) \lambda_{p-1} \lambda_p \end{pmatrix}.$$

The above proof can be obtained by considering the characteristic function of  $v$ .

Jensen [2] showed a similar result above under a different assumption.

**LEMMA 2.2.** *Let  $v = (v_{11}, \dots, v_{pp}, v_{12}, \dots, v_{p-1,p})$  have a normal distribution with mean zero and covariance matrix  $\Psi = (\psi_{ij;kl})$ . Then we have*

$$(2.2) \quad E v_{ij}^2 v_{kl}^2 = 2\psi_{ij;kl}^2 + \psi_{ij;ij}\psi_{kl;kl}$$

$$(2.3) \quad E v_{ij}^2 v_{kk} v_{ll} = 2\psi_{ij;kk}\psi_{ij;ll} + \psi_{ij;ij}\psi_{kk;ll}$$

$$(2.4) \quad E v_{ii} v_{jj} v_{kk} v_{ll} = \psi_{ii;jj}\psi_{kk;ll} + \psi_{ii;kk}\psi_{jj;ll} + \psi_{ii;ll}\psi_{jj;kk} .$$

### 3. Derivation of asymptotic expansion

Let  $l_1 > \dots > l_p$  be latent roots of  $S/m$  and  $\lambda_1 \geq \dots \geq \lambda_p$  be latent roots of  $\Sigma$ . Put  $S/m = A(I + 2\theta) + (1/\sqrt{m})V$ . If  $\tau_\alpha = \lambda_\alpha(1 + 2\theta_\alpha)$  is simple, the perturbation expansion (see for example Wigner [8], p. 40) for the  $\alpha$ th largest root  $l_\alpha$  of  $S/m$  can be expressed as follows:

$$(3.1) \quad l_\alpha = \tau_\alpha + \frac{1}{\sqrt{m}} v_{\alpha\alpha} + \frac{1}{m} \sum \tau_{\alpha\beta}^* v_{\alpha\beta}^2 + \frac{1}{m\sqrt{m}} \left\{ \sum \tau_{\alpha\beta}^* \tau_{\alpha\gamma}^* v_{\alpha\beta} v_{\alpha\gamma} v_{\gamma\beta} - v_{\alpha\alpha} \sum \tau_{\alpha\beta}^{*2} v_{\alpha\beta}^2 \right\} + O_p(m^{-2}) ,$$

where the summation  $\sum$  stands for the summation with respect to all subscripts appearing in each term except for index  $\alpha$  and

$$(3.2) \quad \tau_{\alpha\beta}^* = \begin{cases} (\tau_\alpha - \tau_\beta)^{-1} & \text{if } \alpha \neq \beta \\ 0 & \text{if } \alpha = \beta . \end{cases}$$

It is noted that the method due to a perturbation expansion was used by Lawley [4], Sugiura [7], Fujikoshi [1] and Konishi [3].

Now we consider an analytic function  $f(l_1, \dots, l_p)$  of  $l_1, \dots, l_p$ . Then by using (3.1), we have

$$(3.3) \quad f(l_1, \dots, l_p) = f(\tau_1, \dots, \tau_p) + \frac{1}{\sqrt{m}} q_0 + \frac{1}{m} q_1 + \frac{1}{m\sqrt{m}} q_2 + O_p(m^{-2}) ,$$

where

$$q_0 = \sum v_{\alpha\alpha} f_\alpha$$

$$(3.4) \quad q_1 = \sum \tau_{\alpha\beta}^* v_{\alpha\beta}^2 f_\alpha + \frac{1}{2} \sum v_{\alpha\alpha} v_{\beta\beta} f_{\alpha\beta}$$

$$q_2 = \sum \tau_{\alpha\beta}^* \tau_{\alpha\gamma}^* v_{\alpha\beta} v_{\beta\gamma} v_{\gamma\alpha} f_\alpha - \sum \tau_{\alpha\beta}^{*2} v_{\alpha\alpha} v_{\alpha\beta}^2 f_\alpha$$

$$+ \frac{1}{2} (\sum \tau_{\beta\gamma}^* v_{\alpha\alpha} v_{\beta\gamma}^2 + \sum \tau_{\alpha\gamma}^* v_{\beta\beta} v_{\alpha\gamma}^2) f_{\alpha\beta} + \frac{1}{6} \sum v_{\alpha\alpha} v_{\beta\beta} v_{\gamma\gamma} f_{\alpha\beta\gamma} .$$

The symbol  $\sum$  in (3.4) stands for the summation with respect to all subscripts appearing in each term and from now on we will use this symbol. Also  $f_\alpha$ ,  $f_{\alpha\beta}$  and  $f_{\alpha\beta\gamma}$  mean

$$(3.5) \quad \begin{aligned} f_\alpha &= \frac{\partial}{\partial \tau_\alpha} f(\tau_1, \dots, \tau_p), & f_{\alpha\beta} &= \frac{\partial^2}{\partial \tau_\alpha \partial \tau_\beta} f(\tau_1, \dots, \tau_p), \\ f_{\alpha\beta\gamma} &= \frac{\partial^3}{\partial \tau_\alpha \partial \tau_\beta \partial \tau_\gamma} f(\tau_1, \dots, \tau_p). \end{aligned}$$

Then from (3.3), the statistic  $\sqrt{m}\{f(l_1, \dots, l_p) - f(\tau_1, \dots, \tau_p)\}$  converges in law to a normal distribution with mean zero and variance  $\tau^2 = 2 \operatorname{tr}(I + 4\theta)(\Lambda A)^2$  with  $A = \operatorname{diag}(f_1, \dots, f_p)$ . Thus the characteristic function of  $\sqrt{m}\{f(l_1, \dots, l_p) - f(\tau_1, \dots, \tau_p)\}/\tau$  is given by

$$(3.6) \quad \phi(t) = E \left[ \exp \left( \frac{it}{\tau} q_0 \right) \left\{ 1 + \frac{1}{\sqrt{m}} \left( \frac{it}{\tau} \right) q_1 + \frac{1}{m} \left[ \left( \frac{it}{\tau} \right) q_2 + \frac{1}{2} \left( \frac{it}{\tau} \right)^2 q_1^2 \right] \right\} \right] + O(m^{-3/2}).$$

Thus we must calculate each term in (3.6). At first the first term is given by

$$(3.7) \quad \begin{aligned} E \left[ \exp \left[ \frac{it}{\tau} q_0 \right] \right] &= \operatorname{etr} \left[ -\frac{it}{\tau} \sqrt{m} \Lambda \Lambda (I + 2\theta) \right] \left| I - \frac{2}{\sqrt{m}} \left( \frac{it}{\tau} \right) \Lambda \Lambda \right|^{-n/2} \\ &\times \operatorname{etr} \left[ \frac{1}{2} \left( I - \frac{2}{\sqrt{m}} \left( \frac{it}{\tau} \right) \Lambda \Lambda \right)^{-1} \Lambda^{-1} M M' - \frac{1}{2} \Lambda^{-1} M M' \right] \\ &= \exp \left[ \frac{(it)^2}{2} \right] \left\{ 1 + \frac{1}{\sqrt{m}} \left[ 2\Lambda \left( \frac{it}{\tau} \right) \operatorname{tr} \Lambda A + \frac{4}{3} \left( \frac{it}{\tau} \right)^3 \operatorname{tr}(I + 6\theta)(\Lambda A)^3 \right] \right. \\ &\quad + \frac{1}{m} \left[ \left( \frac{it}{\tau} \right)^2 \{2\Lambda^2(\operatorname{tr} \Lambda A)^2 + 2\Lambda \operatorname{tr}(\Lambda A)^2\} + \left( \frac{it}{\tau} \right)^4 \right. \\ &\quad \times \left. \left\{ 2 \operatorname{tr}(I + 8\theta)(\Lambda A)^4 + \frac{8}{3} \Lambda \operatorname{tr} \Lambda A \operatorname{tr}(I + 6\theta)(\Lambda A)^3 \right\} \right. \\ &\quad \left. + \frac{8}{9} \left( \frac{it}{\tau} \right)^6 [\operatorname{tr}(I + 6\theta)(\Lambda A)^3]^2 \right] \right\} + O(m^{-3/2}). \end{aligned}$$

For the second term, expressing  $q_0$  and  $q_1$  as functions of  $x_{\alpha j}$  ( $\alpha = 1, \dots, p$ ,  $j = 1, \dots, n$ ), after some algebraic manipulation, it can be expressed as an expectation of a function of new variables  $Y = (y_{\alpha j})$  ( $\alpha = 1, \dots, p$ ,  $j = 1, \dots, n$ ), the columns of which are statistically independent and are distributed as normal with mean  $(I - (2/\sqrt{m})(it/\tau)\Lambda A)^{-1}M$  and covariance

matrix  $(A^{-1} - (2/\sqrt{m})(it/\tau)A)^{-1}$ . By noting  $\sum_{j=1}^n m_{\alpha j}m_{\beta j} = 2m\lambda_\alpha\theta_\alpha\delta_{\alpha\beta}$ , where the symbol  $\delta_{\alpha\beta}$  stands for the Kronecker delta, we have

$$(3.8) \quad \begin{aligned} & E\left[\left(\frac{it}{\tau}\right)q_1 \exp\left[\frac{it}{\tau}q_0\right]\right] \\ &= \exp\left[\frac{(it)^2}{2}\right]\left\{\left(\frac{it}{\tau}\right)\left[\frac{1}{2}\sum \tau_{\alpha\beta}^*\lambda_\alpha\lambda_\beta(c_\alpha+c_\beta)f_\alpha + \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha}\right]\right. \\ & \quad \left.+ 2\left(\frac{it}{\tau}\right)^3 \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} + \frac{1}{\sqrt{m}} \sum_{\alpha=1}^3 g_{2\alpha} \left(\frac{it}{\tau}\right)^{2\alpha}\right\} + O(m^{-1}), \end{aligned}$$

where  $c_\alpha = 1 + 4\theta_\alpha$  and

$$(3.9) \quad \begin{aligned} g_2 &= 2A(\text{tr } A) \left\{ \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \right\} \\ & \quad + \sum \tau_{\alpha\beta}^* \lambda_\alpha^2 \lambda_\beta (2c_\alpha + c_\beta - 1) f_\alpha^2 + \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta^2 (c_\alpha + 2c_\beta - 1) f_\alpha f_\beta \\ & \quad + 2 \sum \lambda_\alpha^3 (3c_\alpha - 1) f_\alpha f_{\alpha\alpha} + 4A \sum \lambda_\alpha^2 \lambda_\beta c_\alpha f_\alpha f_{\alpha\beta}, \\ g_4 &= \frac{4}{3} [\text{tr } (I + 6\theta)(A)A^3] \left\{ \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \right\} \\ & \quad + 4A(\text{tr } A) \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} + 4 \sum \lambda_\alpha^3 \lambda_\beta^2 (3c_\alpha - 1) c_\beta f_\alpha^2 f_\beta f_{\alpha\beta}, \\ g_6 &= \frac{8}{3} [\text{tr } (I + 6\theta)(A)A^3] \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta}. \end{aligned}$$

For the third term, Lemmas 2.1 and 2.2 are useful to calculate them. By Lemma 2.1, the third term can be rewritten as an expectation of a function of  $p(p+1)/2$  new variables  $W = (w_{11}, \dots, w_{pp}, w_{12}, \dots, w_{p-1,p})$  distributed as normal with mean  $(2(it/\tau)\lambda_1^2 c_1 f_1, \dots, 2(it/\tau)\lambda_p^2 c_p f_p, 0, \dots, 0)$  and covariance matrix  $\Psi$  in (2.1). Thus we have

$$(3.10) \quad \begin{aligned} & E\left[\left(\frac{it}{\tau}\right)q_2 \exp\left[\frac{it}{\tau}q_0\right]\right] \\ &= \exp\left[\frac{(it)^2}{2}\right]\left\{\left(\frac{it}{\tau}\right)^2 \left[ \sum \tau_{\beta i}^* \lambda_\beta^2 \lambda_i c_\alpha (c_\alpha + c_\beta) f_\alpha f_{\alpha\beta} \right.\right. \\ & \quad \left.+ 2 \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\beta f_{\alpha\beta} + \sum \tau_{\alpha\beta}^* \lambda_\alpha^2 \lambda_\beta c_\alpha (c_\alpha + c_\beta) f_\alpha (f_\beta - f_\alpha) \right] \\ & \quad \left.+ \left(\frac{it}{\tau}\right)^4 \frac{4}{3} \sum \lambda_\alpha^2 \lambda_\beta^2 \lambda_i^2 c_\alpha c_\beta c_i f_\alpha f_\beta f_i f_{\alpha\beta i} \right\} + O(m^{-1/2}) \end{aligned}$$

and

$$(3.11) \quad E\left[\frac{1}{2}\left(\frac{it}{\tau}\right)^2 q_1^2 \exp\left[\frac{it}{\tau}q_0\right]\right] = \exp\left[\frac{(it)^2}{2}\right] \sum_{\alpha=1}^3 g'_{2\alpha} \left(\frac{it}{\tau}\right)^{2\alpha} + O(m^{-1/2}),$$

where

$$\begin{aligned}
g'_2 &= \frac{1}{2} \left\{ \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \right\}^2 \\
&\quad + \frac{1}{4} \sum \tau_{\alpha\beta}^{*2} \lambda_\alpha^2 \lambda_\beta^2 (c_\alpha + c_\beta)^2 f_\alpha (f_\alpha - f_\beta) + \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_{\alpha\beta}^2 , \\
(3.12) \quad g'_4 &= \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} \\
&\quad + 2 \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} + 4 \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta c_\gamma f_\beta f_\gamma f_{\alpha\beta} f_{\alpha\gamma} , \\
g'_6 &= 2 \left\{ \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} \right\}^2 .
\end{aligned}$$

Therefore the characteristic function  $\phi(t)$  in (3.6) is given by

$$\begin{aligned}
(3.13) \quad \phi(t) &= \exp \left[ \frac{(it)^2}{2} \right] \left\{ 1 + \frac{1}{\sqrt{m}} \left[ \left( \frac{it}{\tau} \right) \left\{ \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \right. \right. \right. \\
&\quad + \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + 2A \sum \lambda_\alpha f_\alpha \Big\} \\
&\quad \left. \left. \left. + \left( \frac{it}{\tau} \right)^3 \left\{ \frac{4}{3} \operatorname{tr} (I + 6\theta)(\Lambda A)^3 + 2 \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} \right\} \right] \right. \\
&\quad \left. + \frac{1}{m} \sum_{\alpha=1}^3 \left( \frac{it}{\tau} \right)^{2\alpha} h_{2\alpha} \right\} + O(m^{-3/2}) ,
\end{aligned}$$

where  $h_2$ ,  $h_4$  and  $h_6$  are given by

$$\begin{aligned}
h_2 &= \frac{1}{2} \left\{ \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} + \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + 2A \sum \lambda_\alpha f_\alpha \right\}^2 \\
&\quad + \sum \tau_{\alpha\beta}^* \lambda_\alpha^2 \lambda_\beta^2 (2c_\alpha + c_\beta - 1) f_\alpha^2 + \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta^2 (c_\alpha + 2c_\beta - 1) f_\alpha f_\beta \\
&\quad + \sum \tau_{\beta\gamma}^* \lambda_\alpha^2 \lambda_\beta \lambda_\gamma (c_\beta + c_\gamma) f_\alpha f_{\alpha\beta} + \frac{1}{4} \sum \tau_{\alpha\beta}^{*2} \lambda_\alpha^2 \lambda_\beta^2 (c_\alpha + c_\beta)^2 \\
&\quad \times f_\alpha (f_\alpha - f_\beta) + \sum \tau_{\alpha\beta}^{*2} \lambda_\alpha^3 \lambda_\beta^3 c_\alpha (c_\alpha + c_\beta) f_\alpha (f_\beta - f_\alpha) \\
&\quad + \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta (f_{\alpha\beta}^2 + 2f_\beta f_{\alpha\beta}) + 2 \sum \lambda_\alpha^3 (3c_\alpha - 1) f_\alpha f_{\alpha\alpha} \\
&\quad + 4A \sum \lambda_\alpha^2 \lambda_\beta c_\alpha f_\alpha f_{\alpha\beta} + 2A \sum \lambda_\alpha^2 f_\alpha^2 , \\
(3.14) \quad h_4 &= 2 \operatorname{tr} (I + 8\theta)(\Lambda A)^4 + \frac{4}{3} [\operatorname{tr} (I + 6\theta)(\Lambda A)^3] \left\{ \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \right. \\
&\quad \left. + \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha + 2A \sum \lambda_\alpha f_\alpha \right\} + \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha \\
&\quad \times \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} + 4 \sum \lambda_\alpha^3 \lambda_\beta^2 (3c_\alpha - 1) c_\beta f_\alpha^2 f_\beta f_{\alpha\beta} \\
&\quad + \frac{4}{3} \sum \lambda_\alpha^2 \lambda_\beta^2 \lambda_\gamma^2 c_\alpha c_\beta c_\gamma f_\beta f_\gamma (f_\alpha f_{\alpha\beta\gamma} + 3f_{\alpha\beta} f_{\alpha\gamma}) \\
&\quad + 2 \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} + 4A \sum \lambda_\alpha f_\alpha \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} , \\
h_6 &= \frac{8}{9} \left\{ \operatorname{tr} (I + 6\theta)(\Lambda A)^3 + \frac{3}{2} \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} \right\}^2 .
\end{aligned}$$

Thus we have the following theorem :

**THEOREM 3.1.** *Let  $S$  be a noncentral Wishart matrix having  $W(\Sigma, n; \Omega)$ , and  $l_1 > \dots > l_p$ , and  $\lambda_1 \geq \dots \geq \lambda_p$  be latent roots of  $S/m$  and  $\Sigma$ , respectively. For any analytic function  $f(l_1, \dots, l_p)$ , if  $\tau_\alpha = \lambda_\alpha(1+2\theta_\alpha)$  is simple, then the distribution function of  $f^* = \sqrt{m}\{f(l_1, \dots, l_p) - f(\tau_1, \dots, \tau_p)\}/\tau$  can be expanded for large  $m=n-2A$  under  $\Omega=m\theta$  as follows:*

$$(3.15) \quad P(f^* \leq x) = \Phi(x) - \frac{1}{\sqrt{m}} \left\{ \left[ \sum \lambda_\alpha^2 c_\alpha f_{\alpha\alpha} + \frac{1}{2} \sum \tau_{\alpha\beta}^* \lambda_\alpha \lambda_\beta (c_\alpha + c_\beta) f_\alpha \right. \right. \\ \left. \left. + 2A \sum \lambda_\alpha f_\alpha \right] \Phi^{(1)}(x)/\tau + \left[ \frac{4}{3} \operatorname{tr}(I+6\theta)(AA)^3 \right. \right. \\ \left. \left. + 2 \sum \lambda_\alpha^2 \lambda_\beta^2 c_\alpha c_\beta f_\alpha f_\beta f_{\alpha\beta} \right] \Phi^{(3)}(x)/\tau^3 \right\} \\ + \frac{1}{m} \sum_{\alpha=1}^3 h_{2\alpha} \Phi^{(2\alpha)}(x)/\tau^{2\alpha} + O(m^{-3/2}),$$

where  $\Phi^{(j)}(x)$  stands for the  $j$ th derivative of the standard normal distribution function  $\Phi(x)$ , the symbol  $\sum$  means the summation with respect to all subscripts appearing in each term and the coefficients  $h_{2\alpha}$  with  $c_\alpha = 1+4\theta_\alpha$  are given by (3.14). The asymptotic variance is  $\tau^2 = 2 \operatorname{tr}(I+4\theta) \cdot (AA)^2$  with  $A = \operatorname{diag}(\lambda_1, \dots, \lambda_p)$  and  $A = \operatorname{diag}(f_1, \dots, f_p)$  and  $f_\alpha, f_{\alpha\beta}$  and  $f_{\alpha\beta\gamma}$  are given by (3.5).

In particular, in case  $\theta=0$ , the above result agrees with Konishi [3] after some calculation.

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