

NOTE ON K -POINT SEPARATION MEASUREMENT

RINYA TAKAHASHI

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Summary

Glick [1] introduced the notion of a separation measurement and showed that for a set f_1, \dots, f_K of densities, $s_K^*(f_1, f_2, \dots, f_K) = 2 \left[\int \max \{f_1, f_2, \dots, f_K\} - 1 \right]$ is a K -point separation measurement. This notion is some generalization of Matusita's distance (affinity) of densities f_1, f_2, \dots, f_K , and its interesting applications were shown in Matusita [2], [3]. In this paper we give some statistical remarks on a separation measurement.

If f_1 and f_2 are densities, then the L_1 norm satisfies that

$$\begin{aligned} \|f_1 - f_2\| &= \int |f_1 - f_2| \\ (1) \quad &= 2 \left[\int \max \{f_1, f_2\} - 1 \right] \\ (2) \quad &= 2 \left[1 - \int \min \{f_1, f_2\} \right] \\ (3) \quad &= \int \max \{f_1, f_2\} - \int \min \{f_1, f_2\} . \end{aligned}$$

Therefore, Glick's s_K^* is a natural extension of the L_1 norm in the sense (1). We show that extensions of the L_1 norm in the sense (2) and (3) are K -point separation measurements, too.

DEFINITION (Glick). A symmetric function s will be called a K -point separation measurement ($K \geq 2$) for a subset S in a vector space with norm $\|\cdot\|$ if, for any elements $a_1, a_2, \dots, a_K \in S$, the function s satisfies

$$\max_{i < j} \|a_i - a_j\| \leq s(a_1, a_2, \dots, a_K) \leq \sum_{i < j} \|a_i - a_j\| .$$

We assume that f_1, f_2, \dots, f_K are densities with respect to some σ -finite measure μ .

THEOREM. The function s_*^K given by

$$s_*^K(f_1, f_2, \dots, f_K) = 2 \left[1 - \int \min \{f_1, f_2, \dots, f_K\} \right]$$

is a K -point separation measurement for the probability densities in $L_1[\mu]$.

PROOF. From the definition, we obtain

$$\begin{aligned} s_*^K(f_1, f_2, \dots, f_K) &= 2 \left[1 - \int \min \{f_1, f_2, \dots, f_K\} \right] \\ &= \max_{i < j} 2 \left[1 - \int \min \{f_i, f_j\} \right] \\ &= \max_{i < j} \|f_i - f_j\|. \end{aligned}$$

This completes the proof.

If we define $s_K(f_1, f_2, \dots, f_K) = \int \max \{f_1, f_2, \dots, f_K\} - \int \min \{f_1, f_2, \dots, f_K\} = s_*^K(f_1, f_2, \dots, f_K)/2 + s_*^K(f_1, f_2, \dots, f_K)/2$, then we get the following corollary.

COROLLARY. The function $s_K(f_1, f_2, \dots, f_K)$ is a K -point separation measurement.

Remarks.

(i) By using the same technique as in Glick [1] we can obtain an upper bound for s_*^K , that is

$$s_*^K(f_1, f_2, \dots, f_K) \leq \frac{2}{K} \sum_{i < j} \|f_i - f_j\|.$$

(ii) Glick's s^* and our s_* have the following relationship:

$$\begin{aligned} s_K^*(f_1, f_2, \dots, f_K) &= \sum_{i_1 < i_2} s_*^2(f_{i_1}, f_{i_2}) - \sum_{i_1 < i_2 < i_3} s_*^3(f_{i_1}, f_{i_2}, f_{i_3}) \\ &\quad + \dots + (-1)^K s_*^K(f_1, f_2, \dots, f_K), \end{aligned}$$

which follows from the identity:

$$\begin{aligned} \int \max \{f_1, f_2, \dots, f_K\} &= K - \sum_{i_1 < i_2} \int \min \{f_{i_1}, f_{i_2}\} \\ &\quad + \sum_{i_1 < i_2 < i_3} \int \min \{f_{i_1}, f_{i_2}, f_{i_3}\} \\ &\quad - \dots + (-1)^K \int \min \{f_1, f_2, \dots, f_K\}. \end{aligned}$$

(iii) In contrast with the interpretation of s_K^* given in Glick [1] the

function s_*^K may be interpreted as follows: a least favourable classification rule has the unconditional probability of incorrect classification given by $1 - \int \min \{g_1, g_2, \dots, g_K\} = s_*^K(g_1, g_2, \dots, g_K)/2$, where $g_j(x) = g_j f_j(x)$, g_j standing for the prior probability of f_j .

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OSAKA UNIVERSITY

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