

ON THE REDUCTION TO A COMPLETE CLASS IN MULTIPLE DECISION PROBLEMS (2)

MASAKATSU MURAKAMI

(Received Apr. 26, 1977; revised Mar. 4, 1978)

Summary

As a criterion for the reduction to a complete class of decision rule in case where actions, samples and states are finite in number, "regret-relief ratio" criterion and "incremental loss-gain ratio" criterion were introduced in 2-state of nature case [2]. In this paper, "generalized regret-relief ratio" criterion in k -state of nature case is introduced as an extension of "regret-relief ratio" criterion and its usefulness is shown with an example.

1. Introduction

Consider a following decision problem. Let $L(\theta, a)$ be a loss function incurred if an action a is taken when the state of nature is θ and let x be an information about to guess the true state of nature. In this paper, we have assumed only the case the state space Θ contains k points, the action space A contains n points and the sample space contains m points. Hence, $\Theta = \{\theta_1, \dots, \theta_k\}$, $A = \{a_1, \dots, a_n\}$ and $X = \{x_1, \dots, x_m\}$. For each θ , there is a corresponding cumulative distribution function $F(x|\theta)$, which represents the distribution of X when the true state is θ .

A non-randomized decision rule $d \in D$ and a randomized decision rule $\delta \in D^*$ are defined respectively as follows,

$$d(x) = a \quad \text{and} \quad \delta = \sum_{i=1}^t \pi_i d_i$$

where $\pi_i \geq 0$ for all i and $\sum_{i=1}^t \pi_i = 1$. For any d and δ , the risk can be defined by

$$R(\theta, d) = E_{\theta} L(\theta, d(x)) = \int L(\theta, d(x)) dF(x|\theta) = \sum_{i=1}^m L(\theta, d(x_i)) f(x_i|\theta)$$

and

$$R(\theta, \delta) = \sum_{i=1}^t \pi_i R(\theta, d_i) .$$

Let $w_i \in W$ be a prior probability of θ_i . If the prior distribution is known, one may use the minimal risk criterion for selecting the best decision rule requires to minimize the expected risk defined by

$$r(\delta) = E_w R(\theta, \delta) = \sum_{s=1}^k w_s R(\theta_s, \delta) .$$

However, if W is not known, a reasonable way in choosing a decision rule is to use the relation of natural ordering. A decision rule δ is said to be better than a decision rule δ^* if $R(\theta, \delta) \leq R(\theta, \delta^*)$ for all $\theta \in \Theta$ and $R(\theta, \delta) < R(\theta, \delta^*)$ for at least one $\theta \in \Theta$. (Natural ordering)

In such a case we say that δ dominates δ^* . A class C of decision rules, $C \subset D^*$, is said to be complete, if given any rule $\delta \in D^*$ not in C , there exists a rule $\delta_0 \in C$ that is better than δ .

For the reduction to the complete class of decision rule, we have introduced a new criterion, that is, "regret-relief ratio criterion," in 2-state of nature case, [2]. The basic idea of this criterion is following; Let $l(\theta, a_k)$ be regret of action a_k in θ which is caused by failing to take the best action in θ and let $\bar{l}(\theta, a_k)$ be relief of action a_k in θ which is given by avoiding the worst action in θ . Hence,

$$l(\theta, a_k) = L(\theta, a_k) - \min_{a \in A} L(\theta, a) , \quad \text{and}$$

$$\bar{l}(\theta, a_k) = \max_{a \in A} L(\theta, a) - L(\theta, a_k) .$$

A reasonable decision maker may not take action a_k if $\bar{l}(\theta, a_k) < l(\theta, a_k)$. In this case, regret-relief ratio $r(\theta, a_k)$ is greater than 1,

$$r(\theta, a_k) = \frac{l(\theta, a_k)}{\bar{l}(\theta, a_k)} > 1 .$$

In 2-state of nature case, this idea is summarized as "regret-relief ratio criterion" (Theorem 2 in [2]).

The purpose of the paper is to extend this idea from 2-state of nature case to k -state of nature case.

2. Generalized regret-relief ratio criterion

A new notation is introduced to simplify our discussion. In case that $\Theta = \{\theta_1, \dots, \theta_k\}$ and $A = \{a_1, \dots, a_n\}$, let $a_{j(s)}$ be any action with $\min_{a \in A} L(\theta_s, a)$. For instance, in 3-state and 5-action case like

$$L(\theta_1, a_1) < L(\theta_1, a_2) < L(\theta_1, a_3) < L(\theta_1, a_4) < L(\theta_1, a_5)$$

$$L(\theta_2, a_1) > L(\theta_2, a_2) > L(\theta_2, a_3) < L(\theta_2, a_4) < L(\theta_2, a_5)$$

$$L(\theta_3, a_1) > L(\theta_3, a_2) > L(\theta_3, a_3) > L(\theta_3, a_4) > L(\theta_3, a_5)$$

one would know

$$a_{j(1)} = a_1, \quad a_{j(2)} = a_3 \quad \text{and} \quad a_{j(3)} = a_5.$$

Let us define generalized regret $l^*(\theta, a)$, generalized relief $\bar{l}^*(\theta, a)$ and generalized regret-relief ratio $\gamma^*(\theta, a)$ of action a as follows.

DEFINITION 1 (Generalized regret). The magnitude of regret caused by failing to take an action $a_{j(s)}$ which satisfies

$$L(\theta, a_j) > L(\theta, a_{j(s)})$$

is called generalized regret $l_s^*(\theta, a_j)$ of action a_j and defined by

$$(2.1) \quad l_s^*(\theta, a_j) = L(\theta, a_j) - L(\theta, a_{j(s)}).$$

DEFINITION 2 (Generalized relief). The magnitude of relief given by avoiding an action $a_{j(s)}$ which satisfies

$$L(\theta, a_j) < L(\theta, a_{j(s)})$$

is called generalized relief $\bar{l}_s^*(\theta, a_j)$ of action a_j and defined by

$$(2.2) \quad \bar{l}_s^*(\theta, a_j) = L(\theta, a_{j(s)}) - L(\theta, a_j).$$

DEFINITION 3 (Generalized regret-relief ratio). The generalized regret-relief ratio $\gamma_{s,t}^*(\theta, a_j)$ of action a_j is defined by

$$(2.3) \quad \gamma_{s,t}^*(\theta, a_j) = \frac{l_s^*(\theta, a_j)}{\bar{l}_t^*(\theta, a_j)}.$$

Using generalized regret-relief ratio, we have obtained the following theorem.

THEOREM. In the case that $\Theta = \{\theta_1, \dots, \theta_k\}$, $X = \{x_1, \dots, x_m\}$ and $A = \{a_1, \dots, a_n\}$, assume that $f(x_i|\theta) > 0$ ($i=1, \dots, m$; $\theta \in \Theta$) and that there exist integers $1=n(1) < \dots < n(k)=n$ such that

$$(2.4) \quad \begin{aligned} L(\theta_s, a_h) &> L(\theta_s, a_{h'}) & h < h' < n(s) \quad \text{and} \\ L(\theta_s, a_h) &< L(\theta_s, a_{h'}) & n(s) < h < h'. \end{aligned}$$

Then, for each a_j ($\neq a_{j(s)}$, $s=1, \dots, k$), the following holds. If the condition

$$(2.5) \quad \gamma_{t,t+1}^*(\theta_\alpha, a_j) \cdot \gamma_{t,t+1}^*(\theta_\beta, a_j) > 1$$

is satisfied, for integers t , α and β such that

$$j(t) < j < j(t+1) ,$$

$$(2.6) \quad \gamma_{t,t+1}^*(\theta_\alpha, a_j) = \min_{s \leq t} \{ \gamma_{t,t+1}^*(\theta_s, a_j) \} ,$$

and

$$(2.7) \quad \gamma_{t,t+1}^*(\theta_\beta, a_j) = \min_{s > t} \{ \gamma_{t,t+1}^*(\theta_s, a_j) \} .$$

then any non-randomized decision rule which takes action a_j is dominated by some randomized decision rule.

PROOF. Since any non-randomized decision rule $d \in D$ is defined by assigning $d(x_i) = a^i$ ($i = 1, 2, \dots, m$), we write a non-randomized decision rule d in the form of an ordered m -tuple

$$d \equiv (a^1, \dots, a^m) .$$

For $d = (a^1, \dots, a^m)$, i ($1 \leq i \leq m$) and $a_j \in A$, we define a new non-randomized decision rule

$$d^{(i)} * a_j = (a^1, \dots, a^{i-1}, \underset{\substack{\uparrow \\ \text{ith place}}}{a_j}, a^{i+1}, \dots, a^m) .$$

That is, a non-randomized decision rule with a_j in i th place and remaining $a^1, \dots, a^{i-1}, a^{i+1}, \dots, a^m$ unchanged from d . We have to show, for each i ($i = 1, \dots, m$), that some randomized decision rule $\delta^{(i)}$ which is some mixture of $d^{(i)} * a_{j(t)}$ and $d^{(i)} * a_{j(t+1)}$ dominates $d^{(i)} * a_j$. Note that $a_{j(t)}$ and $a_{j(t+1)}$ are the best actions when θ_t and θ_{t+1} are true respectively. First, there exists a number q ($0 < q < 1$) which satisfies the following equation;

$$(2.8) \quad R(\theta_\alpha, d^{(i)} * a_j) = (1-q) \cdot R(\theta_\alpha, d^{(i)} * a_{j(t)}) + q \cdot R(\theta_\alpha, d^{(i)} * a_{j(t+1)}) .$$

For let q be

$$\begin{aligned} q &= \frac{R(\theta_\alpha, d^{(i)} * a_j) - R(\theta_\alpha, d^{(i)} * a_{j(t)})}{R(\theta_\alpha, d^{(i)} * a_{j(t+1)}) - R(\theta_\alpha, d^{(i)} * a_{j(t)})} \\ &= \frac{f(x_i | \theta_\alpha) \{ L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)}) \}}{f(x_i | \theta_\alpha) \{ L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)}) \}} \\ &= \frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)})} . \end{aligned}$$

Since $L(\theta_\alpha, a_{j(t)}) < L(\theta_\alpha, a_j) < L(\theta_\alpha, a_{j(t+1)})$, q satisfies (2.8) and $0 < q < 1$.

Suppose $\delta^{(i)}$ is the randomized decision rule of $d^{(i)} * a_{j(t)}$ and $d^{(i)} * a_{j(t+1)}$ in the ratio of $(1-q) : q$. Then

$$\begin{aligned} R(\theta_\alpha, \delta^{(i)}) &= (1-q) \cdot R(\theta_\alpha, d^{(i)} * a_{j(t)}) + q \cdot R(\theta_\alpha, d^{(i)} * a_{j(t+1)}) \\ &= R(\theta_\alpha, d^{(i)} * a_j) \end{aligned}$$

and, for $s \neq \alpha$

$$\begin{aligned} R(\theta_s, \delta^{(i)}) &= (1-q) \cdot R(\theta_s, d^{(i)} * a_{j(t)}) + q \cdot R(\theta_s, d^{(i)} * a_{j(t+1)}) \\ &= \frac{1}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)})} \\ &\quad \cdot [\{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)\} R(\theta_s, d^{(i)} * a_{j(t)}) \\ &\quad + \{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})\} R(\theta_s, d^{(i)} * a_{j(t+1)})]. \end{aligned}$$

Hence

$$\begin{aligned} (2.9) \quad R(\theta_s, d^{(i)} * a_j) - R(\theta_s, \delta^{(i)}) &= \frac{1}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)})} [\{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})\} \\ &\quad \cdot \{R(\theta_s, d^{(i)} * a_j) - R(\theta_s, d^{(i)} * a_{j(t+1)})\} - \{L(\theta_\alpha, a_{j(t+1)}) \\ &\quad - L(\theta_\alpha, a_j)\} \{R(\theta_s, d^{(i)} * a_{j(t)}) - R(\theta_s, d^{(i)} * a_j)\}] \\ &= \frac{f(x_i | \theta_s)}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)})} [\{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})\} \\ &\quad \cdot \{L(\theta_s, a_j) - L(\theta_s, a_{j(t+1)})\} - \{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)\} \\ &\quad \cdot \{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)\}] \\ &= \frac{f(x_i | \theta_s)}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)})} \{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)\} \\ &\quad \cdot \{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)\} \\ &\quad \cdot \left\{ \frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} \cdot \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t+1)})}{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)} - 1 \right\}. \end{aligned}$$

i) In case $s \leq t$.

Since

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} = \min_{s \leq t} \left\{ \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t)})}{L(\theta_s, a_{j(t+1)}) - L(\theta_s, a_j)} \right\}$$

then for $s \leq t$ ($s \neq \alpha$), we have

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} < \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t)})}{L(\theta_s, a_{j(t+1)}) - L(\theta_s, a_j)}.$$

Both sides are positive. So

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} \cdot \frac{L(\theta_s, a_{j(t+1)}) - L(\theta_s, a_j)}{L(\theta_s, a_j) - L(\theta_s, a_{j(t)})} < 1,$$

and hence

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} \cdot \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t+1)})}{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)} - 1 < 0.$$

Since $f(x_i|\theta_s) > 0$, $L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_{j(t)}) > 0$, $L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j) > 0$ and $L(\theta_s, a_{j(t)}) - L(\theta_s, a_j) < 0$. In (2.9), we get for $s \leq t$

$$R(\theta_s, d^{(i)} * a_j) - R(\theta_s, \delta^{(i)}) > 0.$$

ii) In case $s > t$.

From the condition

$$\gamma_{t,t+1}^*(\theta_\alpha, a_j) \cdot \gamma_{t,t+1}^*(\theta_\beta, a_j) > 1$$

we get

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} \cdot \frac{L(\theta_\beta, a_j) - L(\theta_\beta, a_{j(t+1)})}{L(\theta_\beta, a_{j(t)}) - L(\theta_\beta, a_j)} - 1 > 0.$$

Since

$$\frac{L(\theta_\beta, a_j) - L(\theta_\beta, a_{j(t+1)})}{L(\theta_\beta, a_{j(t)}) - L(\theta_\beta, a_j)} = \min_{s > t} \left\{ \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t+1)})}{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)} \right\},$$

for $s > t$, we have

$$\frac{L(\theta_\alpha, a_j) - L(\theta_\alpha, a_{j(t)})}{L(\theta_\alpha, a_{j(t+1)}) - L(\theta_\alpha, a_j)} \cdot \frac{L(\theta_s, a_j) - L(\theta_s, a_{j(t+1)})}{L(\theta_s, a_{j(t)}) - L(\theta_s, a_j)} - 1 > 0.$$

Because $L(\theta_s, a_{j(t)}) - L(\theta_s, a_j) > 0$, we get for $s > t$

$$R(\theta_s, d^{(i)} * a_j) - R(\theta_s, \delta^{(i)}) > 0.$$

Hence

$$R(\theta_\alpha, d^{(i)} * a_j) = R(\theta_\alpha, \delta^{(i)}) \quad \text{for some } \alpha$$

and

$$R(\theta_s, d^{(i)} * a_j) > R(\theta_s, \delta^{(i)}) \quad \text{for } s \neq \alpha,$$

which completes the proof.

3. Some example

Example. Consider a decision problem of 4-state, m -sample and 8-action. A loss function is given by Table 1 and $f(x_i|\theta_s) > 0$ for $i =$

$1, \dots, m, s=1, \dots, 4.$

Table 1. $L(\theta, a)$

	a_1	a_2	a_3	a_4	a_5	a_6	a_7	a_8
θ_1	0	2	3	5	9	10	13	14
θ_2	4	3	2	0	3	4	7	10
θ_3	8	7	6	3	1	0	3	6
θ_4	9	8	7	5	4	3	2	0

In this problem, the number of all possible non-randomized decision rules is 8^m . Since $a_{j(1)}=a_1$, $a_{j(2)}=a_4$, $a_{j(3)}=a_6$ and $a_{j(4)}=a_8$, the assumption of Theorem is satisfied. For action a_2 , generalized regret-relief ratios are

$$\begin{aligned}
 \gamma_{1,2}^*(\theta_\alpha, a_2) &= \gamma_{1,2}^*(\theta_1, a_2) = \frac{2}{3} \\
 \gamma_{1,2}^*(\theta_\beta, a_2) &= \min_{s>1} \{\gamma_{1,2}^*(\theta_s, a_2)\} \\
 &= \min \{\gamma_{1,2}^*(\theta_2, a_2), \gamma_{1,2}^*(\theta_3, a_2), \gamma_{1,2}^*(\theta_4, a_2)\} \\
 &= \min \left\{ \frac{3}{1}, \frac{4}{1}, \frac{3}{1} \right\} = 3.
 \end{aligned}$$

Hence,

$$\gamma_{1,2}^*(\theta_\alpha, a_2) \cdot \gamma_{1,2}^*(\theta_\beta, a_2) = \frac{2}{3} \cdot 3 = 2 > 1.$$

By Theorem, we know that the non-randomized decision rules which take action a_2 for any x_i ($i=1, \dots, m$) are dominated.

Similarly as a_2 , for a_3, a_5, a_7 we get the following result. For a_3 , we get

$$\begin{aligned}
 \gamma_{1,2}^*(\theta_\alpha, a_3) &= \gamma_{1,2}^*(\theta_1, a_3) = \frac{3}{2}, \\
 \gamma_{1,2}^*(\theta_\beta, a_3) &= \min_{s>1} \{\gamma_{1,2}^*(\theta_s, a_3)\} \\
 &= \min \{\gamma_{1,2}^*(\theta_2, a_3), \gamma_{1,2}^*(\theta_3, a_3), \gamma_{1,2}^*(\theta_4, a_3)\} \\
 &= \min \left\{ \frac{2}{2}, \frac{3}{2}, \frac{2}{2} \right\} = 1
 \end{aligned}$$

and

$$\gamma_{1,2}^*(\theta_\alpha, a_3) \cdot \gamma_{1,2}^*(\theta_\beta, a_3) = \frac{3}{2} \cdot 1 = \frac{3}{2} > 1.$$

For a_5 , we get

$$\begin{aligned}\gamma_{2,3}^*(\theta_\alpha, a_5) &= \min \{ \gamma_{2,3}^*(\theta_1, a_5), \gamma_{2,3}^*(\theta_2, a_5) \} \\ &= \min \left\{ \frac{4}{1}, \frac{3}{1} \right\} = 3 ,\end{aligned}$$

$$\begin{aligned}\gamma_{2,3}^*(\theta_\beta, a_5) &= \min \{ \gamma_{2,3}^*(\theta_3, a_5), \gamma_{2,3}^*(\theta_4, a_5) \} \\ &= \min \left\{ \frac{1}{2}, \frac{1}{1} \right\} = \frac{1}{2} ,\end{aligned}$$

and

$$\gamma_{2,3}^*(\theta_\alpha, a_5) \cdot \gamma_{2,3}^*(\theta_\beta, a_5) = 3 \cdot \frac{1}{2} = \frac{3}{2} > 1 .$$

For a_7 , we get

$$\begin{aligned}\gamma_{3,4}^*(\theta_\alpha, a_7) &= \min \{ \gamma_{3,4}^*(\theta_1, a_7), \gamma_{3,4}^*(\theta_2, a_7), \gamma_{3,4}^*(\theta_3, a_7) \} \\ &= \min \left\{ \frac{3}{1}, \frac{3}{3}, \frac{3}{3} \right\} = 1 ,\end{aligned}$$

$$\gamma_{3,4}^*(\theta_\beta, a_7) = \gamma_{3,4}^*(\theta_4, a_7) = 2 ,$$

and

$$\gamma_{3,4}^*(\theta_\alpha, a_7) \cdot \gamma_{3,4}^*(\theta_\beta, a_7) = 1 \cdot 2 = 2 > 1 .$$

Therefore, by Theorem, all of the non-randomized decision rules which take action a_3 , a_5 or a_7 for any x_i ($i=1, \dots, m$) are dominated. Hence the number of non-randomized decision rule in the complete class is reduced from 8^m to 4^m .

We conclude that for the problem of this example, generalized regret-relief ratio criterion is very effective in the reduction to a complete class.

Acknowledgement

The author would like to thank Dr. N. Matsubara, of University of Tsukuba, for his frequent, stimulating and helpful discussion. He also wishes to thank the referee for his kind comments.

THE INSTITUTE OF STATISTICAL MATHEMATICS

REFERENCES

- [1] Ferguson, T. S. (1967). *Mathematical Statistics: A Decision Theoretic Approach*, Academic Press, New York.
- [2] Murakami, M. (1976). On the reduction to a complete class in multiple decision problems, *Ann. Inst. Statist. Math.*, 28, 145-165.