# ASYMPTOTIC FORMULAS FOR THE NON-NULL DISTRIBUTIONS OF THREE STATISTICS FOR MULTIVARIATE LINEAR HYPOTHESIS\*

#### YASUNORI FUJIKOSHI

(Received Nov. 16, 1972; revised Jan. 24, 1974)

### Summary

Asymptotic formulas for the distributions of the likelihood ratio statistic, Hotelling's statistic and Pillai's statistic for multivariate linear hypothesis are derived under the assumption of  $n_e=ne$ ,  $n_h=nh$  (e>0, h>0, e+h=1) and  $\Omega=O(n)$ , where  $n_e$  and  $n_h$  are the degrees of freedom for the error and for the hypothesis, respectively and  $\Omega$  is the non-centrality matrix. New asymptotic formulas are given in terms of normal distribution function and its derivatives up to the order  $n^{-1}$ . We give also some numerical results of our asymptotic approximations.

# 1. Introduction

Let  $S_e$  and  $S_h$  denote the independent  $p \times p$  matrices with the central Wishart distribution  $W_p(n_e, \Sigma)$  and the non-central Wishart distribution  $W_p(n_h, \Sigma, \Omega)$ , respectively. The likelihood ratio statistic, Hotelling's statistic and Pillai's statistic for multivariate linear hypothesis are expressed by  $W = -\{n_e + (n_h - p - 1)/2\} \log |S_e(S_h + S_e)^{-1}|$ ,  $T_0^2 = n_e \operatorname{tr} S_h S_e^{-1}$  and  $V = (n_e + n_h) \operatorname{tr} S_h (S_h + S_e)^{-1}$ , respectively. Then  $n_e$  and  $n_h$  mean the degrees of freedom for the error and the hypothesis. The asymptotic expansions for the non-null distributions of W,  $T_0^2$  and V with respect to  $n_e$  up to the order  $n_e^{-2}$ , assuming that  $n_h$  is a fixed number and  $\Omega$  is a fixed matrix, are available in the literature (Sugiura and Fujikoshi [11], Siotani [9], Fujikoshi [3], Muirhead [6] and Lee [4]). Sugiura [12] has obtained the asymptotic expansions of the non-null distributions of these statistics under the assumption that  $\Omega = O(n_e)$  and  $n_h$  is a fixed number.

In this paper we study the asymptotic distributions of these statistics in the situation that  $n_h$  is also relatively large. We derive the asymptotic expansions for the distributions of W,  $T_0^2$  and V assuming that  $n_e=ne$ ,  $n_h=nh$  (e>0, h>0, e+h=1) and  $\Omega=n\Theta$  where  $\Theta$  is a fixed

<sup>\*</sup> This research was partially supported by the Sakko-Kai Foundation,

matrix. New asymptotic formulas are given in terms of normal distribution function and its derivatives up to the order  $n^{-1}$  by using the similar line as in Sugiura [12]. Here we note that it is realistic to assume that  $n_h$  is relatively large, in some multivariate linear hypothesis, e.g., in testing the interactions in multivariate multi-way classification design with relatively large levels and small cell frequencies. In Section 4 we also attempt to test numerically our asymptotic approximations. In the following we may assume  $\Sigma = I$  and  $\Omega = n\Theta = n$  diag  $(\theta_1, \theta_2, \dots, \theta_p)$  since we treat the distributions of invariant tests for multivariate linear hypothesis.

#### 2. Preliminary lemmas

Let  $T_e$  and  $T_h$  be the statistics defined by

$$(2.1) T_e = \sqrt{m} \left( S_e / m - \mu e I \right) \text{and} T_h = \sqrt{m} \left\{ S_h / m - \mu (h I + 2\Theta) \right\} ,$$

for  $m = \mu^{-1}n - r$  where  $\mu$  and r are fixed numbers. Then, by the same technique as in Fujikoshi [2], it is easily seen that the statistics  $T_{\epsilon}$  and  $T_{h}$  converge in law to p(p+1)/2 variate normal distribution with mean 0 as m tends to infinity. Put

(2.2) 
$$\Delta_{\beta}^{\alpha} = \mu \operatorname{tr} \left[ eA^{\alpha} \partial_{e}^{\beta} + \left\{ hI + 2(\alpha + \beta)\Theta \right\} B^{\alpha} \partial_{h}^{\beta} \right]$$

for any given diagonal matrices A and B where  $\partial_e$  and  $\partial_h$  are defined by

(2.3) 
$$\partial_e = \left(\frac{1}{2}(1+\delta_{ij})\partial/\partial\lambda_{ij}^{(e)}\right)$$
 and  $\partial_h = \left(\frac{1}{2}(1+\delta_{ij})\partial/\partial\lambda_{ij}^{(h)}\right)$ 

for any symmetric matrices  $\Lambda_e = (\lambda_{ij}^{(e)})$  and  $\Lambda_h = (\lambda_{ij}^{(h)})$  of order p. We write  $\Delta_0^a$  as  $d_a$  which does not depend on the operators  $\partial_e$  and  $\partial_h$ . We need the following lemma similar to Sugiura [12].

LEMMA 1. Let  $S_e$  and  $S_h$  have  $W_p(n_e, I)$  and  $W_p(n_h, I, \Omega)$  respectively. Suppose  $f(\Lambda_e, \Lambda_h)$  is an analytic function of two positive definite matrices  $\Lambda_e$  and  $\Lambda_h$ . Then the following asymptotic formula holds:

$$\begin{split} (2.4) \qquad & \text{E}\left[f(S_e/m,\,S_h/m)\,\text{etr}\,\{\phi(A\,T_e+B\,T_h)\}\right] \\ = & \text{etr}\,(-d_2t^2)[1+m^{-1/2}\phi\,\{\Gamma_1+4d_3\phi^2/3\}+m^{-1}\,\{\Gamma_2+(\Gamma_3+\Gamma_1^2/2)\phi^2\\ & +2(d_4+2d_3\Gamma_1/3)\phi^4+8d_3^2\phi^8/9\}+m^{-3/2}\phi\,\{\Gamma_4+\Gamma_1\Gamma_2+(\Gamma_5+\Gamma_1\Gamma_3\\ & +4d_3\Gamma_2/3+\Gamma_1^3/6)\phi^2+2(8d_5/5+2d_3\Gamma_3/3+d_4\Gamma_1+d_3\Gamma_1^2/3)\phi^4\\ & +8(d_3d_4/3+d_3^2\Gamma_1/9)\phi^6+32d_3^3\phi^8/81\}+m^{-2}\,\{(\varDelta_2^0)^2/2+2(\varDelta_1^1)^2\varDelta_2^0\phi^2\\ & +2(\varDelta_1^1)^4\phi^4/3+the\,\,lower\,\,order\,\,derivatives\}\\ & +O(m^{-5/2})]f(\varLambda_e,\,\varLambda_h)|_{\varLambda_e=\mu eI,\varLambda_h=\mu(hI+2\Theta)}\;, \end{split}$$

where  $\phi = it$   $(i = \sqrt{-1})$  and  $\Gamma_i$  are defined by

$$\begin{array}{ll} (2.5) & \varGamma_{1} = 2\varDelta_{1}^{1} + rd_{1} \; , & \varGamma_{2} = \varDelta_{2}^{0} + r\varDelta_{1}^{0} \; , \\ & \varGamma_{3} = 4\varDelta_{1}^{2} + rd_{2} \; , & \varGamma_{4} = 4\mu \; \mathrm{tr} \; \{eA\partial_{e}^{2} + (hI + 4\Theta)B\partial_{h}^{2}\} + 8\mu \; \mathrm{tr} \; \Theta\partial_{h}B\partial_{h} + 2r\varDelta_{1}^{1} \; , \\ & \varGamma_{5} = 8\varDelta_{3}^{3} + 4rd_{3}/3 \; . \end{array}$$

PROOF. By expanding f around  $S_e/m = \mu eI$ ,  $S_h/m = \mu(hI + 2\Theta)$  in Taylor's series and evaluating the expectations with respect to  $S_e$  and  $S_h$  by using the formula (Anderson [1])

(2.6) 
$$E \left[ \text{etr} (iTS_h) \right] = |I - 2iT|^{-n_h/2} \text{etr} \left\{ 2i\Omega T (I - 2iT)^{-1} \right\}$$

for any symmetric matrix  $T = ((1 + \delta_{ij})t_{ij}/2)$ , we can express the left-hand side of (2.4) as follows:

$$(2.7) \quad \exp\left\{-m^{-1/2}\phi d_{1}-\varDelta_{1}^{0}\right\} |I-2m^{-1/2}(\phi A+m^{-1/2}\partial_{e})|^{-\mu\epsilon(m+r)/2} \\ \cdot |I-2m^{-1/2}(\phi B+m^{-1/2}\partial_{h})|^{-\mu h(m+r)/2} \operatorname{etr}\left[2\mu(m^{1/2}+rm^{-1/2})\Theta\right] \\ \cdot (\phi B+m^{-1/2}\partial_{h})\left\{I-2m^{-1/2}(\phi B+m^{-1/2}\partial_{h})\right\}^{-1} \\ \cdot f(\varLambda_{e},\varLambda_{h})|_{\varLambda_{e}=\mu\epsilon I,\varLambda_{h}=\mu(hI+2\Theta)}.$$

Applying the asymptotic formulas,  $-\log |I-n^{-1}A| = \sum_{\alpha=1}^{l} n^{-\alpha} \operatorname{tr} A^{\alpha}/\alpha + O(n^{-(l+1)})$  and  $(I-n^{-1}A)^{-1} = \sum_{\alpha=1}^{l} n^{-\alpha}A^{\alpha} + O(n^{-(l+1)})$  which hold for large n, to (2.7), we obtain (2.4).

Let  $G[h(T_{\epsilon}, T_{b})]$  be an abbreviated notation for

(2.8) 
$$e^{d_2t^2} E_{T_e,T_h} [h(T_e, T_h) \text{ etr } \{\phi(AT_e + BT_h)\}]$$
.

The following lemma is useful for our asymptotic expansions.

LEMMA 2. Let R, S, W and Z be any fixed diagonal matrices of order p. Then the following identities hold:

$$egin{align*} G[1] = &1 + m^{-1/2}\phi(rd_1 + 4\phi^2d_3/3) + m^{-1}\phi^2\{r(d_2 + rd_1^2/2) \ &+ 2\phi^2(d_4 + 2rd_1d_3/3) + 8\phi^4d_3^2/9\} + O(m^{-3/2}) \;, \ G[\operatorname{tr} RT_eST_e](\mu e)^{-1} \ &= a(R,S) + 4\phi^2\mu e\alpha + m^{-1/2}\phi[2c(R,S;A) + r\{d_1a(R,S) \ &+ 4\mu e \operatorname{tr} RSA\} + 4\phi^2\{d_3a(R,S)/3 + 4\mu e \operatorname{tr} RSA^3 + rd_1\mu e\alpha\} \ &+ 16\phi^4\mu ed_3lpha/3] + O(m^{-1}) \;, \ G[\operatorname{tr} RT_eST_h](\mu^2 e)^{-1} \ &= 4\phi^2\beta + 2m^{-1/2}\phi[r\operatorname{tr} RS(AL + BM) \ &+ 2\phi^2\{2\operatorname{tr} RSAB(AM + BN) + rd_1\beta\} + 8\phi^4d_3\beta/3] + O(m^{-1}) \;, \ G[\operatorname{tr} RT_hST_h]\mu^{-1} \ &= c(R,S;M)/2 + 4\phi^2\mu\gamma + m^{-1/2}\phi[2\{c(R,S;BM)\}] \end{split}$$

$$\begin{split} &+2a(R\theta,SB)+2a(RB,S\theta)\}+r\{d_{i}c(R,S;M)/2\\ &+4\mu \operatorname{tr} RSBLM\}+4\phi^{2}\{d_{i}c(R,S;M)/6\\ &+4\mu \operatorname{tr} RSB^{3}MN+rd_{i}\mu_{i}\gamma\}+16\phi^{4}\mu d_{i}\gamma/3]+O(m^{-1})\;,\\ G[\operatorname{tr} RT_{c}ST_{c}WT_{c}](\mu e)^{-2}\\ &=2\phi b(R,S,W;A)+8\phi^{3}\mu e \operatorname{tr} RSWA^{5}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{c}WT_{h}](\mu^{2}e)^{-1}\\ &=2\phi a(RWBM,S)+8\phi^{3}\mu e \operatorname{tr} RSWA^{2}BM+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{h}WT_{h}](\mu^{2}e)^{-1}\\ &=2\phi c(RSA,W;M)+8\phi^{3}\mu \operatorname{tr} RSWAB^{2}M^{2}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{h}ST_{h}WT_{h}]\mu^{-2}\\ &=\phi\{b(R,S,W;BM^{2})+c(RM,SM,WM;BM^{-1})\}\\ &+8\phi^{3}\mu \operatorname{tr} RSWB^{3}M^{3}+O(m^{-1/2})\;,\\ G[(\operatorname{tr} RT_{c}ST_{c})^{2}](\mu e)^{-2}\\ &=a(R,S)^{2}+2a(R^{2},S^{2})+2a(RS,RS)+8\phi^{3}\mu e\{4\operatorname{tr} (RSA)^{2}\\ &+\alpha a(R,S)\}+16\phi^{4}(\mu e\alpha)^{2}+O(m^{-1/2})\;,\\ G[\operatorname{tr} T_{c}^{2}\operatorname{tr} (RT_{c})^{2}](\mu e)^{-2}\\ &=(p^{2}+p+4)a(R,R)+4\phi^{2}\mu e\{(p^{2}+p+8)\operatorname{tr} (RA)^{2}\\ &+a(R,R)\operatorname{tr} A^{2}\}+16\phi^{4}(\mu e)^{2}\operatorname{tr} A^{2}\operatorname{tr} (RA)^{2}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{c}\operatorname{tr} WT_{c}ZT_{h}](\mu^{3}e^{3})^{-1}\\ &=4\phi^{2}\{4\operatorname{tr} RSWZABM+\tilde{\beta}a(R,S)\}+16\phi^{4}\mu e\alpha\tilde{\beta}+O(m^{-1/2})\;,\\ G[(\operatorname{tr} RT_{c}ST_{h})^{2}](\mu^{2}e)^{-1}\\ &=c(R,S,R,S;M)+8\phi^{2}\mu \operatorname{tr} R^{3}S^{3}(eA^{2}+B^{2}M)M\\ &+16\phi^{4}\mu^{2}e\beta^{2}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{h}\operatorname{tr} WT_{c}ZT_{h}](\mu^{2}e)^{-1}\\ &=c(R,S,W,Z;M)+8\phi^{2}\mu \operatorname{tr} RSWZ(eA^{2}+B^{2}M)M\\ &+16\phi^{4}\mu^{2}e\beta\tilde{\beta}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{c}\operatorname{tr} WT_{c}ZT_{b}](\mu^{2}e)^{-1}\\ &=c(R,S,W,Z;M)+8\phi^{2}\mu \operatorname{tr} RSWZ(eA^{2}+B^{2}M)M\\ &+16\phi^{4}\mu^{2}e\beta\tilde{\beta}+O(m^{-1/2})\;,\\ G[\operatorname{tr} RT_{c}ST_{c}\operatorname{tr} WT_{c}Z$$

$$egin{aligned} G[\operatorname{tr} RT_eST_h\operatorname{tr} WT_hST_h](\mu^3e)^{-1} \ =& 2\phi^2 \{8\operatorname{tr} RSWZABM^2 + eta c(W,Z;M)\} + 16\phi^4\mueta ilde{\gamma} + O(m^{-1/2}) \;, \end{aligned}$$

 $=a(R, S)c(W, Z; M)/2+2\phi^2\mu\{e\alpha c(W, Z; M)+2a(R, S)\tilde{\gamma}\}$ 

 $+16\phi^4\mu^2e\alpha\tilde{\gamma}+O(m^{-1/2})$ ,

$$G[(\operatorname{tr} RT_hST_h)^2]\mu^{-2} \ = \{c(R,S;M)^2/2 + a(R^2M^2,S^2) + a(R^2,S^2M^2)\}/2 \ + a(R^2M,S^2M) + a(RSM,RSM) + a(RSM^2,RS) \ + 4\phi^2\mu\{8\operatorname{tr} R^2S^2B^2M^3 + \gamma c(R,S;M)\} + 16\phi^4(\mu\gamma)^2 + O(m^{-1/2}),$$

where  $\alpha = \operatorname{tr} RSA^2$ ,  $\beta = \operatorname{tr} RSABM$ ,  $\tilde{\beta} = \operatorname{tr} WZABM$ ,  $\gamma = \operatorname{tr} RSB^2M^2$ ,  $\tilde{\gamma} = \operatorname{tr} WZB^2M^2$ ,  $L = hI + 2\Theta$ ,  $M = hI + 4\Theta$ ,  $N = hI + 6\Theta$ ,  $a(R, S) = \operatorname{tr} RS + \operatorname{tr} R + \operatorname{tr} S$ ,  $b(R, S, W; Q) = 3\operatorname{tr} RSWQ + \operatorname{tr} RSQ\operatorname{tr} W + \operatorname{tr} SWQ\operatorname{tr} R + \operatorname{tr} RWQ\operatorname{tr} S$ , c(R, S; Q) = c(R, S, I, I; Q) and  $c(R, S, W, Z; Q) = \{4\operatorname{tr} RSWZQ + \operatorname{tr} RZQ + \operatorname{tr} SW + \operatorname{tr} RWQ\operatorname{tr} SZ + \operatorname{tr} SZQ\operatorname{tr} RW + \operatorname{tr} SWQ\operatorname{tr} RZ\}/2$ .

PROOF. The first identity is obtained by putting  $f(\Lambda_e, \Lambda_h) = 1$  in Lemma 1. Putting  $f(\Lambda_e, \Lambda_h) = \operatorname{tr} R(\Lambda_e - \mu e I) S(\Lambda_e - \mu e I)$  in Lemma 1, we can reduce  $G[\operatorname{tr} RT_e ST_e](\mu e)^{-1}$  as follows:

(2.10) 
$$[\operatorname{tr} \partial_e^2 + 2\phi^2 \mu e (\operatorname{tr} A \partial_e)^2 + m^{-1/2} \phi \{ 4 \operatorname{tr} A \partial_e^2 + r d_1 \operatorname{tr} \partial_e^2 + 2r \mu e \operatorname{tr} A \partial_e \operatorname{tr} \partial_e + 2\phi^2 (2d_3 \operatorname{tr} \partial_e^2/3 + 4 \mu e \operatorname{tr} A \partial_e \operatorname{tr} A^2 \partial_e + r d_1 \mu e (\operatorname{tr} A \partial_e)^2 ) \\ + 8\phi^4 d_3 \mu e (\operatorname{tr} A \partial_e)^2 \} + O(m^{-1}) ] \operatorname{tr} R \Lambda_e S \Lambda_e .$$

Let  $E_{ij}$  be the  $p \times p$  matrix defined by  $(1/2)(1+\delta_{ij})(\partial/\partial \lambda_{ij}^{(e)})\Lambda_e$ . We have

(2.11) 
$$\operatorname{tr} \Theta \partial_e^2 \operatorname{tr} R \Lambda_e S \Lambda_e = 2 \sum_{i,j=1}^p \theta_i \operatorname{tr} R E_{ij} S E_{ij} = c(R, S; \Theta)/2$$
,

(2.12) 
$$\operatorname{tr} \Theta \partial_e \operatorname{tr} \Gamma \partial_e \operatorname{tr} R \Lambda_e S \Lambda_e = 2 \sum_{i,j=1}^p \theta_i \gamma_j \operatorname{tr} R E_{ii} S E_{jj} = 2 \operatorname{tr} R \Theta S \Gamma$$
,

where  $\theta = \text{diag}(\theta_1, \theta_2, \dots, \theta_p)$  and  $\Gamma = \text{diag}(\gamma_1, \gamma_2, \dots, \gamma_p)$ . Hence we obtain the second identity. Similarly we can derive the other identities by using Lemma 1 and Lemma 2.2 in Sugiura [12].

### 3. Derivation of the asymptotic expansions

We shall obtain the asymptotic expansions for the distributions of (i)  $W=-m\log|S_e(S_h+S_e)^{-1}|$ , (ii)  $T_o^2=m\operatorname{tr} S_hS_e^{-1}$  and (iii)  $V=m\operatorname{tr} S_h(S_h+S_e)^{-1}$  with respect to  $m=\mu^{-1}n-r$  assuming  $\Omega=n\Theta$  where  $\Theta$  is a fixed matrix. In derivation of expansions we assume only that the correction factors  $\mu$  and r are fixed numbers. The expansions for the distributions for traditional statistics are obtained by choosing (i)  $\mu^{-1}=(1+e)/2$ , r=(p+1)/2, (ii)  $\mu^{-1}=e$ , r=0 and (iii)  $\mu^{-1}=1$ , r=0, respectively.

#### 3.1. Expansion for the distribution of W

From (2.1)  $S_e$  and  $S_h + S_e$  are expressed in terms of  $T_e$  and  $T_h$  as  $m\mu e\{I + m^{-1/2}(\mu e)^{-1}T_e\}$  and  $m\mu (I + 2\Theta)\{I + m^{-1/2}\mu^{-1}(I + 2\Theta)^{-1}(T_h + T_e)\}$  respectively. Hence we can express the characteristic function of  $\{W - m \cdot \log |e^{-1}(I + 2\Theta)|\}/\sqrt{m}$  as follows:

(3.1) 
$$\text{E}\left[\text{etr}\left\{\phi(AT_e + BT_h)\right\}\left\{1 + m^{-1/2}\phi q_1(T) + m^{-1}\phi(q_2(T) + \phi q_1(T)^2/2)\right\}\right] + O(m^{-3/2}) ,$$

where  $A = B - (\mu e)^{-1}I$ ,  $B = \mu^{-1}(I + 2\Theta)^{-1}$  and

(3.2) 
$$q_1(T) = \frac{1}{2} [(\mu e)^{-2} \operatorname{tr} T_e^2 - \operatorname{tr} \{B(T_e + T_h)\}^2] ,$$

$$q_2(T) = -\frac{1}{3} [(\mu e)^{-3} \operatorname{tr} T_e^3 - \operatorname{tr} \{B(T_e + T_h)\}^3] .$$

With the help of Lemma 2 and noting  $d_1=0$ , we can simplify (3.1), getting

$$\exp(-\tau^2 t^2/2)[1+m^{-1/2}\{\phi v_1+\phi^3 v_3\}+m^{-1}\sum_{\alpha=1}^3\phi^{2\alpha}w_{2\alpha}+O(m^{-3/2})]$$
,

where  $\tau^2 = 2\mu^{-1}(p/e - s_2)$  for  $s_i = \operatorname{tr}(I + 2\Theta)^{-j}$  and

$$(3.3) \qquad v_1 = (2\mu)^{-1} \{ p(p+1)/e - 2(p+1)s_1 + s_1^2 + s_2 \} \ ,$$
 
$$v_3 = 2(3\mu^2)^{-1} \{ p/e^2 + 2s_3 - 3s_4 \} \ ,$$
 
$$w_2 = v_1^2/2 + (2\mu^2)^{-1} \{ p(p+1)/e^2 + 2(p+1)s_2 - 4s_1s_2 - 4(p+2)s_3 + s_2^2 + 4s_1s_3 + 5s_4 \} + r\mu^{-1} \{ -p/e + s_2 \} \ ,$$
 
$$w_4 = v_1v_3 + 2(3\mu^3)^{-1} \{ p/e^3 - 3s_4 + 12s_5 - 10s_6 \} \ ,$$
 
$$w_6 = v_3^2/2 \ .$$

Inverting this characteristic function, we have,

(3.4) 
$$P(\{W-m \log |e^{-1}(I+2\Theta)|\}/(\tau\sqrt{m}) < x)$$

$$= \Phi(x) - m^{-1/2} \{v_1 \Phi^{(1)}(x)/\tau + v_3 \Phi^{(3)}(x)/\tau^3\}$$

$$+ m^{-1} \sum_{\alpha=1}^{3} w_{2\alpha} \Phi^{(2\alpha)}(x)/\tau^{2\alpha} + O(m^{-3/2}),$$

where  $\Phi^{(j)}(x)$  denotes the jth derivative of the standard normal distribution function  $\Phi(x)$ . The coefficients  $v_j$  and  $w_j$  are given by (3.3).

The null distribution of W is obtained by putting  $\Theta = 0$  in (3.4) and the result is

(3.5) 
$$P(\{W+mp \log e\}/(\tilde{\tau}\sqrt{m}) < x)$$

$$= \Phi(x) - m^{-1/2} \{\tilde{v}_1 \Phi^{(1)}(x)/\tilde{\tau} + \tilde{v}_3 \Phi^{(3)}(x)/\tilde{\tau}^3\}$$

$$+ m^{-1} \sum_{\alpha=1}^{3} \tilde{w}_{2\alpha} \Phi^{(2\alpha)}(x)/\tilde{\tau}^{2\alpha} + O(m^{-3/2}) ,$$

where  $\tilde{\tau} = 2ph(\mu e)^{-1}$  and

$$\begin{split} \tilde{v}_1 &= p(p+1)h(2\mu e)^{-1} \;, \qquad \tilde{v}_3 = 2ph(1+e)(\mu e)^{-2}/3 \;, \\ \tilde{w}_2 &= p(p+1)h\{p(p+1) + 4(1+e)\}(\mu e)^{-2}/8 - rph(\mu e)^{-1} \;, \\ \tilde{w}_4 &= ph\{p(p+1)(1+e)h + 2(1+e+e^2)\}(\mu e)^{-3}/3 \;, \\ \tilde{w}_4 &= \tilde{v}_3^3/2 \;. \end{split}$$

3.2. Expansion for the distribution of  $T_0^2$ 

The characteristic function of  $\{T_0^2 - me^{-1} \operatorname{tr}(hI + 2\Theta)\}/\sqrt{m}$  can be expressed by (3.1) for  $A = -(\mu e^2)^{-1}(hI + 2\Theta)$ ,  $B = (\mu e)^{-1}I$  and

(3.7) 
$$q_1(T) = -(\mu e)^{-2} \{ \operatorname{tr} T_h T_e + \mu e \operatorname{tr} A T_e^2 \} ,$$
$$q_2(T) = (\mu e)^{-3} \{ \operatorname{tr} T_h T_e^2 + \mu e \operatorname{tr} A T_e^3 \} .$$

By using Lemma 2, we have after inversion and arranging each term with respect to  $t_i = \text{tr} (I + 2\Theta)^j$ 

(3.8) 
$$P\left(\left\{T_{0}^{2}-me^{-1}\operatorname{tr}\left(hI+2\Theta\right)\right\}/(\tau\sqrt{m})< x\right) \\ = \Phi(x)-m^{-1/2}\left\{v_{1}\Phi^{(1)}(x)/\tau+v_{3}\Phi^{(3)}(x)/\tau^{3}\right\} \\ +m^{-1}\sum_{\alpha=1}^{3}w_{2\alpha}\Phi^{(2\alpha)}(x)/\tau^{2\alpha}+O(m^{-3/2}),$$

where  $\tau^2 = 2(\mu e^3)^{-1} \{-pe + t_2\}$  and

$$(3.9) \qquad v_{1} = (\mu e^{2})^{-1}(p+1)(-pe+t_{1}) ,$$

$$v_{3} = 4(3\mu^{2}e^{5})^{-1}\{pe^{2} - 3et_{1} + 2t_{3}\} ,$$

$$w_{2} = v_{1}^{2}/2 + (\mu^{2}e^{4})^{-1}\{p(p+1)e(e-3) - 2(p+1)et_{1} + t_{1}^{2} + (3p+4)t_{2}\}$$

$$+ r(\mu e^{3})^{-1}\{pe - t_{2}\} ,$$

$$w_{4} = v_{1}v_{3} + 2(\mu^{3}e^{7})^{-1}\{pe^{2}(2-e) + 4e^{2}t_{1} - 10et_{2} + 5t_{4}\} ,$$

$$w_{6} = v_{3}^{2}/2 .$$

Similarly we have under  $\Omega = 0$ ,

(3.10) 
$$P\left( \{ T_0^2 - mphe^{-1} \} / (\tilde{\tau}\sqrt{m}) < x \right)$$

$$= \varPhi(x) - m^{-1/2} \{ \tilde{v}_1 \varPhi^{(1)}(x) / \tilde{\tau} + \tilde{v}_3 \varPhi^{(3)}(x) / \tilde{\tau}^3 \}$$

$$+ m^{-1} \sum_{\alpha=1}^{3} \tilde{w}_{2\alpha} \varPhi^{(2\alpha)}(x) / \tilde{\tau}^{2\alpha} + O(m^{-3/2}) ,$$

where  $\tilde{\tau}^2 = 2ph(\mu e^3)^{-1}$  and

$$(3.11) \quad \tilde{v}_1 = p(p+1)h(\mu e^2)^{-1} , \qquad \tilde{v}_3 = 4ph(2-e)(3\mu^2 e^5)^{-1} ,$$

$$\tilde{w}_2 = p(p+1)\{(p^2+p+8)h^2/2+3he\}(\mu^2 e^4)^{-1} - rph(\mu e^3)^{-1} ,$$

$$\tilde{w}_4 = 2ph\{2p(p+1)h(2-e)/3+e^2-5e+5\}(\mu^3 e^7)^{-1} ,$$

$$\tilde{w}_6 = \tilde{v}_3^2/2$$
.

3.3. Expansion for the distribution of V

The characteristic function of  $\{V-m \text{ tr } (hI+2\Theta)(I+2\Theta)^{-1}\}/\sqrt{m}$  can be expressed by (3.1) for  $A=\mu^{-1}\{eC^2-C\}$ ,  $B=\mu^{-1}eC^2$  and

(3.12) 
$$q_1(T) = -\mu^{-1} \operatorname{tr} (AT_e + BT_h) C(T_e + T_h) ,$$

$$q_2(T) = \mu^{-2} \operatorname{tr} (AT_e + BT_h) \{C(T_e + T_h)\}^2 ,$$

where  $C=(I+2\Theta)^{-1}$ . By the same technique as in the case of W and  $T_0^2$  we have

(3.13) 
$$P(\{V-m \text{ tr } (hI+2\Theta)(I+2\Theta)^{-1}\}/(\tau\sqrt{m}) < x)$$

$$= \Phi(x) - m^{-1/2} \{v_1 \Phi^{(1)}(x)/\tau + v_3 \Phi^{(3)}(x)/\tau^3\}$$

$$+ m^{-1} \sum_{\tau=1}^{3} w_{2\sigma} \Phi^{(2\sigma)}(x)/\tau^{2\sigma} + O(m^{-3/2})$$

and in a particular case  $\Omega = 0$ ,

(3.14) 
$$P(\{V-mph\}/(\tilde{\tau}\sqrt{m}) < x)$$

$$= \Phi(x) - m^{-1/2} \tilde{v}_{\delta} \Phi^{(8)}(x)/\tilde{\tau}^{3} + m^{-1} \sum_{s=1}^{3} \tilde{w}_{2s} \Phi^{(2s)}(x)/\tilde{\tau}^{2s} + O(m^{-8/2}),$$

where  $\tau^2 = 2e\mu^{-1}\{s_2 - es_4\}$  for  $s_j = \text{tr } C^j$ ,  $\tilde{\tau}^2 = 2phe\mu^{-1}$  and

$$\begin{aligned} (3.15) \quad & v_1\!=\!e\mu^{-1}\{-(p\!+\!1)s_2\!+\!s_1s_2\!+\!s_3\} \ , \\ & v_3\!=\!(4/3)e\mu^{-2}\{-s_3\!+\!3es_5\!+\!e^2s_6\!-\!3e^2s_7\} \ , \\ & w_2\!=\!v_1^2/2\!+\!e\mu^{-2}[es_1^2/2\!-\!es_1s_2\!+\!2(p\!+\!1)s_3\!-\!2s_1s_3\!-\!s_2^2 \\ & \quad +\{(p\!+\!1)e\!-\!3\}s_4\!-\!es_2s_3\!-\!4e(p\!+\!2)s_5\!-\!2es_1s_4\!+\!4es_1s_5\!+\!3es_2s_4 \\ & \quad +es_3^2/2\!+\!8es_6]\!-\!re\mu^{-1}(s_2\!-\!es_4) \ , \\ & w_4\!=\!v_1v_3\!+\!2e\mu^{-3}[s_4\!-\!6es_6\!-\!4e^2s_7\!-\!e^2(e\!-\!14)s_8\!+\!8e^3s_9\!-\!12e^3s_{10}] \ , \\ & w_4\!=\!v_3^2/2 \ , \end{aligned}$$

$$\begin{array}{ll} (3.16) & \tilde{v}_{3}\!=\!(4/3)phe\mu^{-2}(e\!-\!h) \; , \\ & \tilde{w}_{2}\!=\!-phe\mu^{-2}(p\!+\!1)\!-\!rphe\mu^{-1} \; , \\ & \tilde{w}_{4}\!=\!2phe\mu^{-3}(e^{2}\!+\!h^{2}\!-\!3he) \; , \qquad \tilde{w}_{8}\!=\!\tilde{v}_{3}^{2}\!/\!2 \; . \end{array}$$

Sugiura [10] and Nagao [7] have given the expansions similar to (3.5), (3.10) and (3.14), which are obtained by putting  $\mu=1$ , r=0 in (3.5), (3.10) and (3.14), respectively, in connection with the expansions for the distributions of test criteria for equality of two covariance matrices.

# 4. Numerical accuracy of the approximations

In a special case of p=2 Pillai and Jayachandran [8] have computed the exact 5% points of  $\lambda_1 = |S_e(S_h + S_e)^{-1}|^{1/2}$ ,  $\lambda_2 = \operatorname{tr} S_h S_e^{-1}$  and  $\lambda_3 = \operatorname{tr} S_h (S_h + S_e)^{-1}$  $+S_{\epsilon}$ )<sup>-1</sup> and their powers under certain alternatives for some  $n_{\epsilon}$  and  $n_{h}$ . Hence it is possible to test the accuracies of our asymptotic results. Table 1 gives the numerical comparison between the exact and approximate 5% points. Our asymptotic percentage points were computed from the formulas (c.f. Sugiura [10]) obtained by applying the general inverse expansion formula to (3.5), (3.10) and (3.14) with the correction factors such that (i)  $\mu^{-1}=(1+e)/2$ , r=(p+1)/2 for W, (ii)  $\mu^{-1}=e$ , r=p+1 for  $T_0^2$  and (iii)  $\mu^{-1}=1$ , r=0 for V. The values in the brackets ( ) for  $\lambda_2$  mean the percentage points when the traditional correction factors  $\mu^{-1}=e$ , r=0 were used. Table 1 shows that to choose  $\mu^{-1}=e$ , r=p+1 for the correction factors of  $T_0^2$  can be recommended rather than the traditional correction factors. Table 2 gives the numerical comparison between the exact and approximate powers when p=2. Our approximate powers were computed from (3.4), (3.8) and (3.13) by choos-

Table 1	Comparison of approximations to the upper 5%	points of
	$\lambda_1$ , $\lambda_2$ and $\lambda_3$ for $p=2$	

$n_e$	$n_h$	$\lambda_1$			$\lambda_2$	$\lambda_8$	
		exact	approx.	exact	approx.	exact	approx.
13	7	.4460	.4458	2.880	2.851(2.614)	1.039	1.042
	13	.3194	.3196	4.941	4.874(4.475)	1.035	1.035
33	13	.5959	.5959	1.424	1.419(1.399)	.7863	.7871

Table 2 Comparison of approximations to the powers of W,  $T_0^2$  and V for  $p{=}2$  and significance level,  $\alpha{=}0.05$ 

ne		$\omega_1$	ω <sub>2</sub>	W		$T_0^2$		$\dot{m{V}}$	
	$n_h$			exact	approx.	exact	approx.	exact	approx.
13	7	0	1	.0826	.0843	.081	.091	.0826	.0831
		.5	.5	.0839	.0857	.081	.089	.0858	.0863
	13	0	1	.0694	.0704	.068	.055	.0698	.0700
		.5	.5	.0699	.0710	.068	.054	.0710	.0712
33	13	.5	.5	.0784	.0787	.077	.081	.0788	.0785
		0	1.5	.0941	.0942	.094	.099	.0934	.0931
63	7	0	4	.3242	.3242	.330	.332	.3167	.3159
		2.5	2.5	.4241	.4241	.419	.422	.4279	.4273
	13	2	2	.231	.2328	.230	.233	.234	.2344
		0	5	.284	.2852	.292	.294	.265*	.2753

<sup>(\*</sup> It seems that this value is incorrect.)

ing the same correction factors as in Table 1, and using the exact significant points given in [8]. Exact powers for  $n_e=63$  and  $n_h=7$  in Table 2 were extracted from Lee [5]. From Tables 1 and 2 we can see that our approximations except for  $T_0^2$  are good still in the case of the small values of  $n_h$  and  $\omega_j$ . Further, it may be pointed that our approximations are excellent when  $n_e$  and  $n_h$  are both large and  $\omega_j$  are large.

# Acknowledgement

I wish to thank Professor N. Sugiura for pointing out an error in my original manuscript of  $\Gamma_4$  in E.Q. (2.5). This necessitated a correction of  $w_2$  in E.Q. (3.3) and (3.15).

KOBE UNIVERSITY

#### REFERENCES

- [1] Anderson, T. W. (1946). The non-central Wishart distribution and certain problems of multivariate statistics, *Ann. Math. Statist.*, 17, 409-431.
- [2] Fujikoshi, Y. (1968). Asymptotic expansion of the distribution of the generalized variance in the non-central case, J. Sci. Hiroshima Univ., Ser. A-I, 32, 293-299.
- [3] Fujikoshi, Y. (1970). Asymptotic expansions of the distributions of the test statistics in multivariate analysis, J. Sci. Hiroshima Univ., Ser. A-I, 34, 73-144.
- [4] Lee, Y. S. (1971). Distribution of the canonical correlations and asymptotic expansions for the distributions of certain independence test statistics, Ann. Math. Statist., 42, 526-537.
- [5] Lee. Y. S. (1971). Asymptotic formulae for the distribution of a multivariate test statistic: power comparisons of certain multivariate tests, *Biometrika*, 58, 647-651.
- [6] Muirhead, R. J. (1972). The asymptotic noncentral distribution Hotelling's generalized  $T_0^2$ , Ann. Math. Statist., 43, 1671-1677.
- [7] Nagao, H. (1970). Asymptotic expansions of some test criteria for homogeneity of variances and covariance matrices from normal populations, J. Sci. Hiroshima Univ., Ser. A-I, 34, 153-247.
- [8] Pillai, K. C. S. and Jayachandran, K. (1967). Power comparisons of tests of two multivariate hypotheses based on four criteria, *Biometrika*, 54, 195-210.
- [9] Siotani, M. (1971). An asymptotic expansion of the non-null distribution of Hotelling's generalized T<sub>0</sub><sup>2</sup>-statistic, Ann. Math. Statist., 42, 560-571.
- [10] Sugiura, N. (1969). Asymptotic expansions of the distributions of  $|S_1S_2^{-1}|$  and  $|S_2(S_1 + S_2)^{-1}|$  for testing the equality of two covariance matrices, *Mimeo Series*, No. 653, Univ. of North Carolina.
- [11] Sugiura, N. and Fujikoshi, Y. (1969). Asymptotic expansions of the non-null distributions of the likelihood ratio criteria for multivariate linear hypothesis and independence, Ann. Math. Statist., 40, 942-952.
- [12] Sugiura, N. (1973). Further asymptotic formulas for the non-null distributions of three statistics for multivariate linear hypothesis, *Ann. Inst. Statist. Math.*, **25**, 153-163.