ASYMPTOTIC FORMULAS FOR THE HYPERGEOMETRIC FUNCTION ${}_2F_1$ OF MATRIX ARGUMENT, USEFUL IN MULTIVARIATE ANALYSIS

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(Received Aug. 30, 1971; revised Oct. 18, 1973)

Summary

Let S_i have the Wishart distribution $W_p(\Sigma_i, n_i)$ for i=1, 2. An asymptotic expansion of the distribution of $-2\log\lambda = -2\log\left[\prod_{\alpha=1}^2 |S_\alpha/n_\alpha|^{n_\alpha/2}\cdot|(S_1+S_2)/n|^{-n/2}\right]$ for large $n=n_1+n_2$ is derived, when $\Sigma_1\Sigma_2^{-1}=I+n^{-1/2}\theta$, based on an asymptotic solution of the system of partial differential equations for the hypergeometric function ${}_2F_1$, obtained recently by Muirhead [2]. Another asymptotic formula is also applied to the distributions of $-2\log\lambda$ and $-\log|S_2(S_1+S_2)^{-1}|$ under fixed $\Sigma_1\Sigma_2^{-1}$, which gives the earlier results by Nagao [4]. Some useful asymptotic formulas for ${}_1F_1$ were investigated by Sugiura [7].

1. Preliminaries

Based on a random sample of size n_i+1 from p-variate normal population $N(\mu_i, \Sigma_i)$ for i=1 and 2, the modified likelihood ratio statistic for testing $\Sigma_1 = \Sigma_2$ against $\Sigma_1 \neq \Sigma_2$ can be expressed as

(1.1)
$$\lambda = |S_1/n_1|^{n_1/2} |S_2/n_2|^{n_2/2} / |(S_1 + S_2)/n|^{n/2},$$

where S_1 has the Wishart distribution $W_p(\Gamma, n_1)$ and S_2 has $W_p(I, n_2)$ independently with $n = n_1 + n_2$. The diagonal matrix Γ has the characteristic roots of $\Sigma_1 \Sigma_2^{-1}$ on the diagonal. The moments of λ was given by Sugiura [5] as

(1.2)
$$\mathbb{E}\left[\lambda^{h}\right] = \left(\frac{n^{n}}{n_{1}^{n_{1}}n_{2}^{n_{2}}}\right)^{ph/2} \frac{\Gamma_{p}(n/2)}{\Gamma_{p}(n_{1}/2)\Gamma_{p}(n_{2}/2)} \frac{\Gamma_{p}(n_{1}(1+h)/2)\Gamma_{p}(n_{2}(1+h)/2)}{\Gamma_{p}(n(1+h)/2)} \\ \cdot |\Gamma|^{n_{1}h/2} \cdot {}_{2}F_{1}(nh/2, n_{1}(1+h)/2; n(1+h)/2; I-\Gamma) ,$$

which can be derived by starting the joint distribution of (S_1, S_2) and using the formula

$$\operatorname{etr}\left[\frac{1}{2}(I-\Gamma^{-1})S_{1}\right] = \sum_{k=0}^{\infty} \Sigma_{(\epsilon)}C_{\epsilon}\left(\frac{1}{2}(I-\Gamma^{-1})S_{1}\right) / k! ,$$

together with (12) and (22) in Constantine [1]. Kummer transformation formula $_2F_1(a_1, a_2; b; Z) = |I-Z|^{-a_2} _2F_1(b-a_1, a_2; b; -Z(I-Z)^{-1})$ is also used in the final expression.

Put $m_{\alpha} = \rho n_{\alpha}$ for $\alpha = 1$, 2 and the correction factor $\rho = 1 - (2p^2 + 3p - 1)(n_1^{-1} + n_2^{-1} - n^{-1})/\{6(p+1)\}$ with $m = m_1 + m_2$. Asymptotic expansion of the distribution of $-2\rho \log \lambda$ for large m and fixed $\rho_{\alpha} = m_{\alpha}/m > 0$, when $\Gamma = I + m^{-1}\theta$, was obtained by Sugiura [5] in terms of χ^2 -distributions and when Γ is fixed, by Nagao [4] in terms of normal distribution function and its derivatives, in a more general case, namely for k-sample problem. When $\Gamma = I + m^{-1/2}\theta$, however, both methods are not available and the present approach by system of partial differential equations for ${}_2F_1$ by Muirhead [3] is essentially useful.

Another asymptotic formula for ${}_2F_1$, based on the differential equations gives the asymptotic expansions of the distributions of $-2\rho \log \lambda$ and $-\log |S_2(S_1+S_2)^{-1}|$ under fixed Γ obtained in Nagao [4] by different technique. The characteristic function of $-\sqrt{n} \log |S_2(S_1+S_2)^{-1}|$ can be written as

(1.3)
$$\Gamma_{p}\left(\frac{1}{2}n_{2}-\sqrt{n}it\right)\Gamma_{p}\left(\frac{1}{2}n\right)/\left[\Gamma_{p}\left(\frac{1}{2}n_{2}\right)\Gamma_{p}\left(\frac{1}{2}n-\sqrt{n}it\right)\right]$$

$$\cdot {}_{2}F_{1}\left(-\sqrt{n}it,\frac{1}{2}n_{1};\frac{1}{2}n-\sqrt{n}it;I-\Gamma\right),$$

by the similar argument as for (1.2), which was remarked by Sugiura [6]. This statistic can be used to test $\Sigma_1 = \Sigma_2$ against $\Sigma_1 \Sigma_2^{-1} \ge I$ for two normal populations.

2. General asymptotic formulas for ${}_2F_1$

Muirhead [2] has proved that the hypergeometric function ${}_{2}F_{1}(a, b; c; R)$ for $R = \text{diag}(r_{1}, \dots, r_{p})$ can be characterized as the unique solution of the partial differential equation

$$(2.1) r_{i}(1-r_{i})\frac{\partial^{2}F}{\partial r_{i}^{2}} + \left[c - \frac{1}{2}(p-1) - \left\{a + b + 1 - \frac{1}{2}(p-1)\right\}r_{i} + \frac{1}{2}\sum_{j=2}^{p}\frac{r_{i}(1-r_{j})}{r_{i}-r_{j}}\right]\frac{\partial F}{\partial r_{i}} - \frac{1}{2}\sum_{j=2}^{p}\frac{r_{j}(1-r_{j})}{r_{i}-r_{j}}\frac{\partial F}{\partial r_{j}} = abF,$$

subject to the condition that F is symmetric with respect to r_i and F is analytic at R=0 with F(0)=1.

Generalizing the characteristic function of $-2\rho \log \lambda$ obtained by putting $h = -2\rho it$ in (1.2) under $\Gamma = I + m^{-1/2}\theta$, we shall consider G(R) =

 $\log \left[{}_{2}F_{1}(\alpha n; \beta_{1}n + \beta_{0}; \gamma_{1}n + \gamma_{0}; n^{-1/2}R) \right]$ for $\gamma_{1} \neq 0$, which is the unique solution for

$$(2.2) r_{1}(1-r_{1}/\sqrt{n})\left\{\frac{\partial^{2}G}{\partial r_{1}^{2}}+\left(\frac{\partial G}{\partial r_{1}}\right)^{2}\right\}+\left[r_{1}n-(\alpha+\beta_{1})\sqrt{n}\,r_{1}-\frac{1}{2}\,(p-1)+r_{0}\right] \\ -\left(\beta_{0}+1-\frac{1}{2}\,(p-1)\right)r_{1}/\sqrt{n}+\frac{1}{2}\,\sum_{j=2}^{p}\,\frac{r_{1}}{r_{1}-r_{j}} \\ -\frac{1}{2}\,\sum_{j=2}^{p}\,\frac{r_{1}^{2}}{r_{1}-r_{j}}\,\frac{1}{\sqrt{n}}\,\right]\frac{\partial G}{\partial r_{1}}-\frac{1}{2}\,\sum_{j=2}^{p}\,\frac{r_{j}}{r_{1}-r_{j}}\,\frac{\partial G}{\partial r_{j}} \\ +\frac{1}{2}\,\sum_{j=2}^{p}\,\frac{r_{j}^{2}}{r_{1}-r_{j}}\,\frac{\partial G}{\partial r_{j}}\,\frac{1}{\sqrt{n}}=\alpha\beta_{1}n\sqrt{n}+\alpha\beta_{0}\sqrt{n}\;,$$

subject to G(0)=0 and G is symmetric for R. Putting $G(R)=\sum_{k=-1}^{\infty}Q_k(R)n^{-k/2}$ with $Q_k(0)=0$ as in Sugiura [7], we can successively determine $Q_k(R)$ from (2.2). Equating the term of order $n\sqrt{n}$ in (2.2) yields $\gamma_1\partial Q_{-1}/\partial r_1=\alpha\beta_1$, which implies $Q_{-1}=(\alpha\beta_1/\gamma_1)\sum_{j=1}^p r_j$. We shall write $\sigma_j=\operatorname{tr} R^j$ and put $\xi_1=\alpha+\beta_1-\alpha\beta_1/\gamma_1$, $\xi_2=\alpha+\beta_1-2\alpha\beta_1/\gamma_1$ for abbreviation. Then equating the term of order n, \sqrt{n} and 1 in (2.2) yields, after some computation:

$$\begin{split} Q_0 &= \frac{1}{2} (\alpha \beta_1 / \gamma_1^2) \xi_1 \sigma_2 \;, \\ (2.3) \qquad Q_1 &= \frac{1}{3} \sigma_3 (\alpha \beta_1 / \gamma_1^3) (\alpha \beta_1 + \xi_1 \xi_2) + (\alpha / \gamma_1) (\beta_0 - \beta_1 \gamma_0 / \gamma_1) \sigma_1 \;, \\ Q_2 &= \frac{1}{4} (\alpha \beta_1 / \gamma_1^4) \sigma_4 \{ \alpha \beta_1 \xi_1 (2 - \xi_1 / \gamma_1) + \xi_2 (\alpha \beta_1 + \xi_1 \xi_2) \} \\ &\qquad \qquad + \frac{1}{2} (\alpha / \gamma_1^2) \sigma_2 \Big\{ - (\beta_1 / \gamma_1) \Big(\gamma_0 + \frac{1}{2} \Big) \xi_1 + (\beta_0 - \gamma_0 \beta_1 / \gamma_1) \xi_2 + \beta_1 \Big(\beta_0 + \frac{1}{2} \Big) \Big\} \\ &\qquad \qquad + \frac{1}{4} (\alpha \beta_1 / \gamma_1^2) \sigma_1^2 (1 - \xi_1 / \gamma_1) \;. \end{split}$$

Hence we have:

THEOREM 2.1. An asymptotic formula of $_2F_1(\alpha n, \beta_1 n + \beta_0; \gamma_1 n + \gamma_0; n^{-1/2}R)$ for large value of n, when $\gamma_1 \neq 0$ is given by

$$(2.4) \qquad \exp\left[\sqrt{n}\,(\alpha\beta_1/\gamma_1)\sigma_1 + \frac{1}{2}\,(\alpha\beta_1/\gamma_1^2)\xi_1\sigma_2 + n^{-1/2}Q_1 + n^{-1}Q_2 + O(n^{-3/2})\right],$$

where $\sigma_j = \operatorname{tr} R^j$, $\xi_1 = \alpha + \beta_1 - \alpha \beta_1 / \gamma_1$ and Q_1 , Q_2 are given by (2.3) with $\xi_2 = \xi_1 - \alpha \beta_1 / \gamma_1$.

Generalizing the characteristic function (1.3), an asymptotic formula

for ${}_2F_1(\alpha\sqrt{n}, \beta_2n+\beta_1\sqrt{n}+\beta_0; \gamma_2n+\gamma_1\sqrt{n}+\gamma_0; R)$, when $\gamma_2\neq 0$ should be investigated. As in Muirhead [3], transforming the variables R to $W=I-(I-(\beta_2/\gamma_2)R)^{-1}$ and putting $F=|I-W|^{\alpha\sqrt{n}}G$, we can rewrite the partial differential equation (2.1) for the function G(W). Further the transformation $H=\log G$ is useful to simplify the right-hand side of the equation. Denoting $a=\alpha\sqrt{n}$, $b=\beta_2n+\beta_1\sqrt{n}+\beta_0$ and $c=\gamma_2n+\gamma_1\sqrt{n}+\gamma_0$ for abbreviation, the differential equation for H(W) is expressed by

$$(2.5) \qquad (-\beta_{2}/\gamma_{2})\{1-(1-\gamma_{2}/\beta_{2})w_{1}\}\left[a(a+1)-2(a+1)(1-w_{1})\frac{\partial H}{\partial w_{1}}\right] \\ + (1-w_{1})^{2}\left\{\frac{\partial^{2} H}{\partial w_{1}^{2}}+\left(\frac{\partial H}{\partial w_{1}}\right)^{2}\right\}-(\beta_{2}/\gamma_{2})\left[\{c-(p-1)/2\}(1-w_{1})\right] \\ - \{a+b+1-(p-1)/2\}(\gamma_{2}/\beta_{2})w_{1}+\frac{1}{2}\sum_{j=2}^{p}\frac{w_{1}(1-w_{j})}{w_{1}-w_{j}} \\ \cdot \{1-(1-\gamma_{2}/\beta_{2})w_{1}\}\left[-a+(1-w_{1})\frac{\partial H}{\partial w_{1}}\right]+\frac{1}{2}\frac{\beta_{2}}{\gamma_{2}} \\ \cdot \sum_{j=2}^{p}\frac{(1-w_{1})(1-w_{j})w_{j}}{w_{1}-w_{j}}\left\{1-\left(1-\frac{\gamma_{2}}{\beta_{2}}\right)w_{j}\right\}\left\{-\frac{a}{1-w_{j}}+\frac{\partial H}{\partial w_{j}}\right\} \\ = ab.$$

Putting $H(W) = \sum_{k=0}^{\infty} Q_k(w) n^{-k/2}$ with $Q_k(0) = 0$ for $k = 0, 1, \dots$, and equating the term of order n, \sqrt{n} , and 1 in (2.5), we can get:

THEOREM 2.2. An asymptotic formula for ${}_{2}F_{1}(\alpha\sqrt{n}; \beta_{2}n+\beta_{1}\sqrt{n}+\beta_{0}; \gamma_{2}n+\gamma_{1}\sqrt{n}+\gamma_{0}; R)$ for large value of n, when $\gamma_{2}\neq 0$ is given by

$$(2.6) \qquad |I-W|^{lpha\sqrt{n}} \exp\left[-lpha\delta_1\sigma_1 + rac{1}{2}lpha^2\delta_2\sigma_2 + n^{-1/2}Q_1 + n^{-1}Q_2 + O(n^{-3/2})
ight],$$

where $\sigma_{j} = \text{tr } W^{j}$, $W = I - (I - (\beta_{2}/\gamma_{2})R)^{-1}$ and $\delta_{1} = \beta_{1}/\beta_{2} - \gamma_{1}/\gamma_{2}$ $\delta_{2} = \beta_{2}^{-1} - \gamma_{2}^{-1}$. The coefficients Q_{1} and Q_{2} are given by, using $\delta_{0} = \beta_{0}/\beta_{2} - \gamma_{0}/\gamma_{2}$,

$$Q_{1} = \alpha \sigma_{1}(-\delta_{0} + \gamma_{1}\delta_{1}/\gamma_{2}) + \frac{1}{2}\alpha \sigma_{2}\left\{\delta_{2}\left(\frac{1}{2} - \alpha \gamma_{1}/\gamma_{2}\right) + \delta_{1}(\delta_{1} + \alpha \delta_{2} - \alpha/\gamma_{2})\right\} + \frac{1}{3}\alpha^{2}\delta_{2}\sigma_{3}\left\{-3\delta_{1} + \alpha(\gamma_{2}^{-1} - \delta_{2})\right\} + \frac{1}{2}\alpha^{3}\sigma_{4}\delta_{2}^{2} + \frac{1}{4}\alpha\delta_{2}\sigma_{1}^{2},$$

$$Q_{2} = (\alpha/\gamma_{2})\sigma_{1}[\gamma_{0}\delta_{1} + \gamma_{1}(\delta_{0} - \gamma_{1}\delta_{1}/\gamma_{2})] + \frac{1}{2}\alpha\sigma_{2}\left[-(\alpha/\gamma_{2})\left\{\delta_{1}^{2} + (1 + \gamma_{0})\delta_{2}\right\} + \frac{1}{2}\delta_{1}(2\delta_{0} + \delta_{2} - \gamma_{2}^{-1}) - \frac{1}{2}(\gamma_{1}/\gamma_{2})\delta_{2}(1 - \alpha\gamma_{1}^{-1} - 2\alpha\gamma_{1}\gamma_{2}^{-1}) + (\alpha\delta_{2} + \delta_{1} - \alpha/\gamma_{2})(\delta_{0} - 2\delta_{1}\gamma_{1}/\gamma_{2})\right] + \frac{1}{3}\sigma_{3}\left[\alpha\delta_{1}\delta_{2}\left(-\frac{3}{2} + 6\alpha\gamma_{1}\gamma_{2}^{-1} + 6\alpha\gamma_{1}\gamma_{2}^{-1})\right]$$

$$\begin{split} &+2\alpha^{2}\gamma_{2}^{-1}\Big)+(\gamma_{2}^{-1}-\delta_{2})\Big(\delta_{1}^{2}+\frac{1}{2}\,\delta_{2}\Big)\alpha^{2}-3\alpha^{2}\delta_{0}\delta_{2}-\alpha^{2}\delta_{2}(\delta_{2}-\gamma_{2}^{-1})\\ &\quad\cdot(1-2\alpha\gamma_{1}\gamma_{2}^{-1})-\alpha\delta_{1}(\delta_{1}+\alpha\delta_{2}-\alpha/\gamma_{2})^{2}\Big]\\ &+\frac{1}{4}\,\sigma_{4}\alpha^{2}\delta_{2}\Big[\delta_{1}^{2}+\delta_{2}\Big(\frac{5}{2}-\alpha^{2}\gamma_{2}^{-1}-4\alpha\gamma_{1}\gamma_{2}^{-1}\Big)-2\alpha\delta_{1}(\gamma_{2}^{-1}-\delta_{2})\\ &\quad+(\alpha\delta_{2}+\delta_{1}-\alpha\gamma_{2}^{-1})(5\delta_{1}+\alpha\delta_{2}-\alpha\gamma_{2}^{-1})\Big]+\sigma_{5}\alpha^{3}\delta_{2}^{2}[-2\delta_{1}+\alpha\gamma_{2}^{-1}-\alpha\delta_{2}]\\ &+\frac{5}{6}\,\alpha^{4}\delta_{2}^{3}\sigma_{6}+\frac{1}{4}\,\alpha\sigma_{1}^{2}[\delta_{1}(\delta_{2}-\gamma_{2}^{-1})-\gamma_{2}^{-1}\delta_{2}(\gamma_{1}+\alpha)]\\ &-\frac{1}{2}\,\alpha\delta_{2}\sigma_{1}\sigma_{2}[\delta_{1}+\alpha\delta_{2}-\alpha\gamma_{2}^{-1}]+\frac{1}{2}\,\alpha^{2}\delta_{2}^{2}\sigma_{1}\sigma_{3}+\frac{1}{8}\,\alpha^{2}\delta_{2}^{2}\sigma_{2}^{2}\;. \end{split}$$

3. Asymptotic expansions

The characteristic function of the likelihood ratio statistic $-2\rho \log \lambda$ under $\Gamma = I + m^{-1/2}\theta$ is expressed by (1.2) as

$$(3.1) \qquad \left(\frac{m^m}{m_1^{m_1}m_2^{m_2}}\right)^{-pit} \frac{\Gamma_p(m/2+\varDelta)\Gamma_p(m_1(1-2it)/2+\varDelta_1)\Gamma_p(m_2(1-2it)/2+\varDelta_2)}{\Gamma_p(m_1/2+\varDelta_1)\Gamma_p(m_2/2+\varDelta_2)\Gamma_p(m(1-2it)/2+\varDelta)} \\ \qquad \qquad \cdot |I+m^{-1/2}\theta|^{-m_1tt} \cdot {}_2F_1(-mit, \, m_1(1-2it)/2+\varDelta_1; \\ \qquad \qquad m(1-2it)/2+\varDelta; \, -m^{-1/2}\theta) \ ,$$

where $\Delta_i = \rho_i \Delta = (n_i - m_i)/2$ and $\rho_i = n_i/n$. By Stirling's formula, the first factor in (3.1), namely, the product of gamma functions can be evaluated as

(3.2) The first factor =
$$(1-2it)^{-f/2} + O(m^{-2})$$

for f = p(p+1)/2. The second factor can easily be written by

$$(3.3) \quad \log |I+m^{-1/2}\theta|^{-m_1it} = \rho_1 it \Big(-\sqrt{m} \, \sigma_1 + \frac{1}{2} \, \sigma_2 - \frac{1}{3} \, m^{-1/2} \sigma_3 + \frac{1}{4} \, m^{-1} \sigma_4 \Big) \\ + O(m^{-8/2}) \; ,$$

where $\sigma_j = \text{tr } \theta^j$. Finally putting n = m, $\alpha = -it$, $\beta_1 = \rho_1(1-2it)/2$, $\beta_0 = \Delta_1$, $\gamma_1 = (1-2it)/2$, $\gamma_0 = \Delta$ and $R = -\theta$ in Theorem 2.1, we can evaluate the third factor in (3.1) as, denoting (1-2it) by T,

$$(3.4) \qquad \log \, {}_{2}F_{1} = \sqrt{m} \, \rho_{1}it\sigma_{1} + \frac{1}{4} \, \rho_{1}\sigma_{2}(-2it-\rho_{2}+\rho_{2}T^{-1}) - \frac{1}{6} \, \rho_{1}\sigma_{3}m^{-1/2} \\ \qquad \cdot \, [-2it-\rho_{2}-\rho_{2}^{2}+3\rho_{2}^{2}T^{-1}+\rho_{2}(1-2\rho_{2})T^{-2}] \\ \qquad + m^{-1} \Big[\frac{1}{8} \, \rho_{1}\sigma_{4}\{-2it-\rho_{2}-\rho_{2}^{2}-\rho_{2}^{3}+6\rho_{2}^{3}T^{-1}+2\rho_{2}^{2}(3-5\rho_{2})T^{-2} \Big] \Big]$$

$$egin{aligned} &+
ho_2 (1 - 5
ho_2 + 5
ho_2^2) T^{-3} \} + rac{1}{4}
ho_1
ho_2 \sigma_2 \{ -2 \varDelta + (1 + 4 \varDelta) T^{-1} \ &- (1 + 2 \varDelta) T^{-2} \} + rac{1}{4}
ho_1
ho_2 \sigma_1^2 (T^{-1} - T^{-2}) \, \Big] + O(m^{-3/2}) \; . \end{aligned}$$

Combined with (3.2), (3.3) and (3.4), we can get the asymptotic expansion of the distribution of $-2\rho \log \lambda$ in terms of noncentral χ^2 -distributions with noncentrality $\delta = \rho_1 \rho_2 \sigma_2/4$ as

(3.5)
$$P(-2\rho \log \lambda < x) = P_f + \frac{1}{6} \rho_1 \rho_2 \sigma_3 m^{-1/2} \{ (1+\rho_2) P_f - 3\rho_2 P_{f+2} - (1-2\rho_2) P_{f+4} \} + m^{-1} \sum_{k=0}^{4} h_k P_{f+2k} + O(m^{-3/2}) ,$$

where $P_f = P(\chi_f^2(\delta) < x)$ for the noncentral χ^2 -variate and

$$egin{aligned} h_0 = & rac{1}{8}
ho_1
ho_2 \Big\{ -\sigma_4 (1 +
ho_2 +
ho_2^2) - 4 \Delta \sigma_2 + rac{1}{9}
ho_1
ho_2 \sigma_3^2 (1 +
ho_2)^2 \Big\} \;, \ h_1 = & rac{1}{8}
ho_1
ho_2 \Big\{ 6
ho_2^2 \sigma_4 + 2 (1 + 4 \Delta) \sigma_2 + 2 \sigma_1^2 - rac{2}{3}
ho_1
ho_2^2 (1 +
ho_2) \sigma_3^2 \Big\} \;, \end{aligned}$$

$$(3.6) \qquad h_2 = \frac{1}{8} \rho_1 \rho_2 \Big\{ 2\rho_2 (3 - 5\rho_2) \sigma_4 - 2(2\varDelta + 1) \sigma_2 - 2\sigma_1^2 + \frac{1}{9} \rho_1 \rho_2 \sigma_3^2 (13\rho_2^2 + 2\rho_2 - 2) \Big\} ,$$

$$h_3 = \frac{1}{8} \rho_1 \rho_2 \Big\{ (1 - 5\rho_2 + 5\rho_2^2) \sigma_4 + \frac{2}{3} \rho_1 \rho_2^2 (1 - 2\rho_2) \sigma_3^2 \Big\} ,$$

$$h_4 = \frac{1}{72} \rho_1^2 \rho_2^2 \sigma_3^2 (1 - 2\rho_2)^2 .$$

Next we shall consider the distribution of $-2\rho m^{-1/2} \log \lambda$ under fixed Γ , the characteristic function of which can be expressed from (1.2) as

$$(3.7) \qquad \left(\frac{m^{m}}{m_{1}^{m_{1}}m_{2}^{m_{2}}}\right)^{-itp/\sqrt{m}} \\ \cdot \frac{\Gamma_{p}(m/2+\varDelta)\Gamma_{p}(m_{1}/2-\rho_{1}it\sqrt{m}+\varDelta_{1})\Gamma_{p}(m_{2}/2-\rho_{2}it\sqrt{m}+\varDelta_{2})}{\Gamma_{p}(m_{1}/2+\varDelta_{1})\Gamma_{p}(m_{2}/2+\varDelta_{2})\Gamma_{p}(m/2-\sqrt{m}it+\varDelta)} \\ \cdot |\Gamma|^{-\rho_{1}it\sqrt{m}}{}_{2}F_{1}(-\sqrt{m}it,\ m_{1}/2-\rho_{1}it\sqrt{m}+\varDelta_{1}; \\ m/2-it\sqrt{m}+\varDelta;\ I-\Gamma) \ .$$

By Stirling's formula, the first factor in (3.7) can be evaluated as

$$(3.8) \qquad \exp\left[\frac{1}{2}p(p+1)itm^{-1/2} + \frac{1}{2}p(p+1)(it)^2m^{-1} + O(m^{-3/2})\right].$$

Putting n=m, $\alpha=-it$, $\beta_2=\rho_1/2$, $\beta_1=-\rho_1it$, $\beta_0=\Delta_1$, $\gamma_2=1/2$, $\gamma_1=-it$, $\gamma_0=\Delta_1$ and $R=I-\Gamma$ in Theorem 2.2, we can get an asymptotic formula for

the second factor in (3.7) with respect to $W=I-\tilde{\Gamma}^{-1}$ for $\tilde{\Gamma}=\rho_1\Gamma+\rho_2I$, which gives the following asymptotic expansion under fixed Γ in terms of the standard normal distribution function $\Phi(x)$ and the derivatives:

$$(3.9) \quad \mathrm{P}\left(-2\rho m^{-1/2}\log\lambda + \sqrt{m}\log\left(|\varGamma|^{\rho_1}/|\tilde{\varGamma}|\right) < \tau_1 x\right)$$

$$= \Phi(x) + m^{-1/2} \left\{ g_1 \Phi^{(1)}(x) / \tau_1 + g_3 \Phi^{(3)}(x) / \tau_1^3 \right\} + m^{-1} \sum_{k=1}^3 h_{2k} \Phi^{(2k)}(x) / \tau_1^{2k} + O(m^{-3/2}) \right\},$$

where
$$au_1^2 = 2(\rho_2/\rho_1)\sigma_2$$
 for $\sigma_j = \text{tr } W^j = \text{tr } (I - \tilde{\Gamma}^{-1})^j$ and $g_1 = \frac{1}{2}(\rho_2/\rho_1)(\sigma_2 + \sigma_1^2) - \frac{1}{2}p(p+1)$, $g_3 = 2(\rho_2/\rho_1)\left\{-\sigma_2 + \frac{2}{2}(1 - \rho_2/\rho_1)\sigma_3 + (\rho_2/\rho_1)\sigma_4\right\}$,

$$(3.10) \quad h_2 = \frac{1}{2} p(p+1) - 2(\rho_2/\rho_1)(1+\Delta)\sigma_2 + 2(\rho_2/\rho_1)(1-\rho_2/\rho_1)\sigma_3 + \frac{5}{2} (\rho_2/\rho_1)^2 \sigma_4 \\ - 2(\rho_2/\rho_1)\sigma_1^2 + 2(\rho_2/\rho_1)(1-\rho_2/\rho_1)\sigma_1\sigma_2 + 2(\rho_2/\rho_1)^2 \sigma_1\sigma_3 \\ + \frac{1}{2} (\rho_2/\rho_1)^2 \sigma_2^2 + \frac{1}{2} g_1^2 ,$$

$$\begin{split} h_4 &= 4(\rho_2/\rho_1)\sigma_2 - \frac{16}{3}(\rho_2/\rho_1)(1-\rho_2/\rho_1)\sigma_3 + 2(\rho_2/\rho_1)\{1-7\rho_2/\rho_1 + (\rho_2/\rho_1)^2\}\sigma_4 \\ &+ 8(\rho_2/\rho_1)^2(1-\rho_2/\rho_1)\sigma_5 + \frac{20}{3}(\rho_2/\rho_1)^3\sigma_6 + g_1g_3 \;, \end{split}$$

$$h_6 = \frac{1}{2} g_3^2$$
.

After some computation, we can verify that the above result agrees with the Theorem 8.1 in Nagao [4] for k=2.

Finally for the characteristic function of $-\sqrt{n} \log |S_2(S_1+S_2)^{-1}|$ given in (1.3), the first factor can be reduced to

$$(3.11) \qquad -\sqrt{n} \, itp \log \rho_2 - pt^2 \rho_1/\rho_2 + 2pitn^{-1/2} \\ \cdot \left\{ \frac{1}{3} \, \rho_1 [(1+\rho_2)/\rho_2^2] (it)^2 + \frac{1}{4} (p+1)\rho_1/\rho_2 \right\} + p(it)^2 n^{-1} \\ \cdot \left\{ \frac{2}{3} \, (\rho_1/\rho_2) (it)^2 (1+\rho_2^{-1}+\rho_2^{-2}) + \frac{1}{2} \, (p+1)\rho_1 (1+\rho_2)/\rho_2^2 \right\} \\ + O(n^{-3/2}) \; .$$

The second factor can be expanded by Theorem 2.2, which implies

(3.12)
$$P(-\sqrt{n} \log |S_2(S_1+S_2)^{-1}| - \sqrt{n} \log |\rho_2 \tilde{\Gamma}| < x\tau_2)$$

$$=$$
 $\Phi(x)+n^{-1/2}\{g_1\Phi^{(1)}(x)/ au_2+g_3\Phi^{(3)}(x)/ au_2^3\}+n^{-1}\sum\limits_{k=1}^3h_{2k}\Phi^{(2k)}(x)/ au_2^{2k} \ +O(n^{-8/2})$,

where
$$au_2^2 = 2(2\sigma_1 + (\rho_2/\rho_1)\sigma_2 + p\rho_1/\rho_2)$$
 for $\sigma_j = \operatorname{tr} (I - \tilde{\Gamma}^{-1})^j$ with $\tilde{\Gamma} = \rho_1 \Gamma + \rho_2 I$ and $g_1 = -\frac{1}{2} (\rho_1/\rho_2) p(p+1) + \frac{1}{2} (\rho_2/\rho_1) (\sigma_1^2 + \sigma_2)$,
$$g_3 = -\frac{2}{3} (\rho_1/\rho_2) p(2 + \rho_1/\rho_2) - 4\sigma_1 + 4(1 - \rho_2/\rho_1)\sigma_2 + \frac{4}{3} (\rho_2/\rho_1) (4 - \rho_2/\rho_1)\sigma_3 + 2(\rho_2/\rho_1)^2\sigma_4$$
,
$$h_2 = \frac{1}{2} p(p+1) (\rho_1/\rho_2) (2 + \rho_1/\rho_2) + (1 - 3\rho_2/\rho_1)\sigma_2 + 2(\rho_2/\rho_1) (2 - \rho_2/\rho_1)\sigma_3 + \frac{5}{2} (\rho_2/\rho_1)^2\sigma_4 + (1 - 3\rho_2/\rho_1)\sigma_1^2 + 2(\rho_2/\rho_1) (2 - \rho_2/\rho_1)\sigma_1\sigma_2 + 2(\rho_2/\rho_1)^2\sigma_1\sigma_3 + \frac{1}{2} (\rho_2/\rho_1)^2\sigma_2^2 + \frac{1}{2} g_1^2$$
,
$$h_4 = \frac{2}{3} p(\rho_1/\rho_2) (1 + \rho_2^{-1} + \rho_2^{-2}) + 8\sigma_1 - 4(5 - 3\rho_2/\rho_1)\sigma_2 + \frac{8}{3} \{5 - 15\rho_2/\rho_1 + 3(\rho_2/\rho_1)^2\}\sigma_3 + (\rho_2/\rho_1) \{30 - 30\rho_2/\rho_1 + 2(\rho_2/\rho_1)^2\}\sigma_4 + 8(\rho_2/\rho_1)^2 (3 - \rho_2/\rho_1)\sigma_5 + \frac{20}{3} (\rho_2/\rho_1)^3\sigma_6 + g_1g_3$$
,
$$h_6 = \frac{1}{2} g_3^2$$
.

Though this result is the same as the Theorem 11.1 in Nagao [4], our expression is somewhat simpler.

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