

DISTRIBUTION OF WILKS' LIKELIHOOD-RATIO CRITERION IN THE COMPLEX CASE*

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(Received Sept. 8, 1969; revised June 8, 1970)

1. Summary

The null distribution of Wilks' likelihood ratio criterion, Λ , in the complex case, is obtained, and explicit expressions for the same are given for $p=2$ and 3 , where p is the number of variables. It is shown that unlike the real case the distributions derived have closed form representation for all p and for all f_2 , the hypothesis degrees of freedom. Tables of correction factors for converting chi-square percentiles to exact percentiles of a logarithmic function of Λ are provided for fourteen (p, f_2) pairs. Tables for an additional thirteen pairs can be obtained from those tabulated by interchanging p and f_2 .

2. Introduction

Let the joint density function of complex random matrices $X: p \times f_2$ and $S: p \times p$ be (see [2], [3] and [10])

$$(2.1) \quad CN(X; \mu, \Sigma) \quad CW(S; p, f_1, \Sigma)$$

where $CN(X; \mu, \Sigma)$ is the complex multivariate normal distribution defined by

$$(2.2) \quad CN(X; \mu, \Sigma) = \pi^{-pf_2} |\Sigma|^{-f_2} \exp [-\text{tr } \Sigma^{-1}(X - \mu)(\overline{X - \mu})'] .$$

$CW(S; p, f_1, \Sigma)$ is the complex Wishart distribution defined by

$$(2.3) \quad CW(S; p, f_1, \Sigma) = \{\Gamma_p(f_1)\}^{-1} |\Sigma|^{-f_1} |S|^{f_1-p} \exp (-\text{tr } \Sigma^{-1}S) ,$$

where

$$(2.4) \quad \Gamma_p(f_1) = \pi^{p(p-1)/2} \left\{ \prod_{j=1}^p \Gamma(f_1 - j + 1) \right\} ,$$

* This research was supported (in part) by the National Science Foundation under Grant Number GU-15-34. Reproduction in whole or in part is permitted for any purpose of the United States Government.

$\Sigma: p \times p$ and $S: p \times p$ are hermitian positive definite matrices and $\mu: p \times f_2$ is a complex matrix. The likelihood ratio criterion for testing $H_0: \mu(p \times f_2) = 0$ against $\mu \neq 0$ can be expressed in terms of the following criterion suggested by Wilks [9] and Pearson and Wilks [6]

$$(2.5) \quad \Lambda = |S| / |S + X\bar{X}'|$$

where the critical region is defined by $\Lambda \leq \lambda_0$ (which is parallel to the likelihood-ratio criterion for real Gaussian variables) and $P(\Lambda \leq \lambda_0 / H_0) = \alpha$.

3. Distributional properties of Λ

For purposes of notational ease, the symbol Λ will be replaced by U_{p, f_2, f_1} . Let us denote by $\beta(a, b; X)$ the density function

$$(3.1) \quad \{1/\beta(a, b)\} X^{a-1}(1-X)^{b-1}, \quad 0 \leq X \leq 1$$

of a real beta variable X . Khatri [5] has proved the following result which we state as

THEOREM 3.1. *In the complex case, when the hypothesis to be tested is true, U_{p, f_2, f_1} is distributed like $X_1 \cdots X_p$ where X_i are independently distributed real beta variables with density $\beta(f_1 - i + 1, f_2; X_i)$.*

Now if we consider (5.2.1) of Khatri [5] when $\mu=0$ (and therefore in Khatri's notation $\alpha=0$), a little algebraic manipulations give us the following theorem.

THEOREM 3.2. *In the complex case, the distribution of U_{p, f_2, f_1} is the same as that of U_{f_2, p, f_1+f_2-p} .*

This implies that without loss of generality we need consider only values of $f_2 \geq p$. The method employed in this paper relies on these two theorems.

4. The method of derivation

An immediate consequence of Theorem 3.1 is that since

$$(4.1) \quad -\log U_{p, f_2, f_1} = \sum_{i=1}^p (-\log X_i) = \sum_{i=1}^p Y_i$$

say, the distribution problem in hand can be reduced to that of a sum of independently distributed random variables. The latter distribution can be handled by taking successive convolutions provided that the procedure yields expressions which can be easily integrated at each stage. Schatzoff [8] has proved that this is in fact the case. Gupta [4], Pillai

and Gupta [7] and Schatzoff [8] have used this method to obtain the distribution of Wilks' Λ in the real case.

Consider the density of X_i ,

$$(4.2) \quad \beta(f_1-i+1, f_2; X_i) = K_i X_i^{f_1-i} (1-X_i)^{f_2-1}, \quad 0 \leq X_i \leq 1, \quad f_1 \geq i,$$

where

$$(4.3) \quad K_i = [1/\beta(f_1-i+1, f_2)] = \Gamma(f_1+f_2-i+1)/\{\Gamma(f_1-i+1)\Gamma(f_2)\}.$$

Using the binomial theorem and transforming $Y_i = -\log X_i$, the density of Y_i is given by,

$$(4.4) \quad K_i \sum_{l=0}^{f_2-1} (-1)^l \binom{f_2-1}{l} e^{-(f_1+l-i+1)Y_i}, \quad Y_i > 0.$$

Note that in this case, unlike the real case, an application of the binomial theorem gives a finite series for all values of f_2 .

Now making use of (2.4) of Schatzoff [8], the density of $Y_1 + \dots + Y_p$ can be obtained by means of successive convolutions, as given in the next section.

5. Exact distribution of U_{p, f_2, f_1} in the complex case

In this section, we consider the density and the c.d.f. of U_{p, f_2, f_1} for $p=2$ and 3. We also give the general form of the density and the c.d.f. for any p and general f_1 and f_2 .

Case (i). $p=2$. We have $U_{2, f_2, f_1} = X_1 X_2$, and hence

$$(5.1) \quad -\log U_{2, f_2, f_1} = Y_1 + Y_2.$$

Now using the method of Section 4, we get the density of U_{2, f_2, f_1} in the following finite series form:

$$(5.2) \quad K_1 K_2 U^{f_1-2} \left[\sum_{m=1}^{f_2-1} \binom{f_2-1}{m} \binom{f_2-1}{m-1} U^m \log U \right. \\ \left. + \sum_{l=0}^{f_2-1} \sum_{\substack{m=0 \\ l+m-1}}^{f_2-1} \frac{(-1)^{l+m}}{m-l-1} \binom{f_2-1}{l} \binom{f_2-1}{m} (U^{l+1} - U^m) \right], \\ 0 < U \leq 1,$$

where K_1 and K_2 are given by (4.3) for $i=1$ and 2 respectively. The c.d.f. of U_{2, f_2, f_1} is obtained from (5.2) by integrating between $(0, u)$, $0 < u \leq 1$, as follows

$$(5.3) \quad P[U_{2, f_2, f_1} \leq u] \\ = K_1 K_2 u^{f_1-1} \left[- \sum_{m=1}^{f_2-1} \binom{f_2-1}{m} \binom{f_2-1}{m-1} \frac{u^m}{(f_1+m-1)^2} (1 - (f_1+m-1) \log u) \right]$$

$$+ \sum_{l=0}^{f_2-1} \sum_{\substack{m=0 \\ l \neq m-1}}^{f_2-1} \frac{(-1)^{l+m}}{m-l-1} \binom{f_2-1}{l} \binom{f_2-1}{m} \\ \cdot \left(\frac{u^{l+1}}{f_1+l} - \frac{u^m}{f_1+m-1} \right) \Big] .$$

It is interesting to note the special cases for $f_2=2$ and 3. For $f_2=2$, we get the density and the c.d.f. of U_{2,f_2,f_1} respectively as follows:

$$(5.4) \quad f_1^2(f_1^2-1)U^{f_1-1}[\log U + ((U^{-1}-U)/2)] , \quad 0 \leq U \leq 1 ,$$

and

$$(5.5) \quad P[U_{2,2,f_1} \leq u] = u^{f_1}[-(f_1^2-1)(1-f_1 \log u) \\ + (f_1^2((f_1+1)u^{-1}-(f_1-1)u))/2] .$$

Similarly for $f_2=3$, we obtain the density and the c.d.f. of U_{2,f_2,f_1} respectively as follows:

$$(5.6) \quad \{f_1^2(f_1+1)^2(f_1-1)(f_1+2)U^{f_1-2}/12\} \\ \cdot \{6U(1+U) \log U + 1 + 9U - 9U^2 - U^3\} , \quad 0 \leq U \leq 1 ,$$

and

$$(5.7) \quad P[U_{2,3,f_1} \leq u] \\ = [f_1^2(f_1+1)^2(f_1-1)(f_1+2)u^{f_1-1}/12] \\ \cdot \left[\frac{1}{f_1-1} + \frac{9u}{f_1} - \frac{9u^2}{f_1+1} - \frac{u^3}{f_1+2} - \frac{6u}{f_1^2}(1-f_1 \log u) \right. \\ \left. - \frac{6u^2\{1-(f_1+1) \log u\}}{(f_1+1)^2} \right] .$$

Case (ii). $p=3$. For $p=3$, $-\log U_{3,f_2,f_1}=Y_1+Y_2+Y_3$ and the density of U_{3,f_2,f_1} is given by the following finite series:

$$(5.8) \quad \left(\prod_{i=1}^3 K_i \right) U^{f_1-2} \left[\frac{1}{2} \sum_{m=1}^{f_2-2} (-U)^m \mathcal{J}(m+1, m, m-1) (\log U)^2 \right. \\ \left. - \sum_{\substack{m=1 \\ m \neq n-1}}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^n U^m}{(n-m-1)^2} \mathcal{J}(m, m-1, n) \{U^{n-m-1} - (n-m-1) \log U - 1\} \right. \\ \left. - \sum_{l=0}^{f_2-3} \sum_{\substack{m=0 \\ l \neq m-1}}^{f_2-1} \frac{(-1)^m U^{l+1}}{(m-l-1)} \mathcal{J}(l+2, l, m) \log U \right. \\ \left. + \sum_{\substack{l=0 \\ l \neq m-1, n-2}}^{f_2-1} \sum_{m=0}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^{l+m+n}}{(m-l-1)(n-l-2)} \mathcal{J}(l, m, n) (U^{l+1} - U^{n-1}) \right. \\ \left. - \sum_{l=0}^{f_2-1} \sum_{\substack{m=0 \\ l \neq m-1}}^{f_2-2} \frac{(-1)^l U^m}{(m-l-1)} \mathcal{J}(m+1, m, l) \log U \right]$$

$$-\sum_{l=0}^{f_2-1} \sum_{\substack{m=0 \\ m \neq l+1, n-1}}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^{l+m+n}}{(m-l-1)(n-m-1)} \mathcal{A}(l, m, n) (U^m - U^{n-1}) \Big], \\ 0 \leqq U \leqq 1,$$

where $\mathcal{A}(l, m, n) = \binom{f_2-1}{l} \binom{f_2-1}{m} \binom{f_2-1}{n}$, and K_i are given by (4.3). The c.d.f. of U_{3, f_2, f_1} can be easily obtained by straightforward integration of (5.8) between $(0, u)$, $0 < u \leqq 1$ and is given by

$$(5.9) \quad P[U_{3, f_2, f_1} \leqq u] \\ = \left(\prod_{i=1}^3 K_i \right) u^{f_1-1} \left[\sum_{m=1}^{f_2-1} \frac{(-u)^m}{2(f_1+m-1)^3} \mathcal{A}(m+1, m, m-1) \right. \\ \cdot \{ (f_1+m-1)^2 (\log u)^2 - (f_1+m-1) \log u + 2 \} \\ - \sum_{\substack{m=1 \\ m \neq n-1}}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^n u^n}{(f_1+m-1)(n-m-1)^2} \mathcal{A}(n, m, m-1) \\ \cdot \left\{ \frac{(f_1+m-1)u^{n-m-1}}{f_1+n-2} + \frac{n-m-1}{f_1+m-1} (1 - (f_1+m-1) \log u) - 1 \right\} \\ + \sum_{\substack{l=0 \\ l \neq m-1}}^{f_2-3} \sum_{m=0}^{f_2-1} \frac{(-1)^m u^{l+1}}{(m-l-1)(f_1+l)^2} \mathcal{A}(m, l+2, l) \{ 1 - (f_1+l) \log u \} \\ + \sum_{\substack{l=0 \\ l \neq m-1, n-2}}^{f_2-1} \sum_{m=0}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^{l+m+n}}{(m-l-1)(n-l-2)} \mathcal{A}(l, m, n) \left(\frac{u^{l+1}}{f_1+l} - \frac{u^{n-1}}{f_1+n-2} \right) \\ + \sum_{\substack{l=0 \\ l \neq m-1}}^{f_2-1} \sum_{m=0}^{f_2-2} \frac{(-1)^l u^m}{(m-l-1)(f_1+m-1)^2} \mathcal{A}(l, m, m+1) \\ \cdot \{ 1 - (f_1+m-1) \log u \} \\ - \sum_{\substack{l=0 \\ l \neq m+1, n-1}}^{f_2-1} \sum_{m=0}^{f_2-1} \sum_{n=0}^{f_2-1} \frac{(-1)^{l+m+n}}{(m-l-1)(n-m-1)} \mathcal{A}(l, m, n) \\ \cdot \left(\frac{u^m}{f_1+m-1} - \frac{u^{n-1}}{f_1+n-2} \right) \Big].$$

Case (iii). General p. The density and the distribution function of U_{p, f_2, f_1} for $p > 3$ becomes too involved for presentation as well as for computational purposes. However, the above analysis forms the basis of the following theorems.

THEOREM 5.1. *The probability density function of U_{p, f_2, f_1} , in the complex case, is of the form*

$$(5.10) \quad \mathcal{A}(U) = \left(\prod_{i=1}^p K_i \right) \sum_{j=1}^m C_j U^{f_1-l_j} (-\log U)^{k_j},$$

where K_i is defined by (4.3) and the constants C_j and the integers m, l_j, k_j are determined from p, f_2 and f_1 .

PROOF. The theorem can be easily proved by induction. The proof parallels the proof of Theorem 2.3 of Schatzoff [8] in the real case and hence the same has been omitted.

When $p=1$, $\beta(U)$ is the density function of a beta random variable with parameters f_2 and f_1 . This is readily seen to be of the required form where $m=f_2-1$, $l_j=1-j$, $k_j=0$ and $C_j=(-1)^j \binom{f_2-1}{j}$.

The corresponding distribution function is obtained by integrating (5.10) between $(0, u)$, $0 < u \leq 1$, as indicated by

THEOREM 5.2. *The distribution function of U_{p,f_2,f_1} , in the complex case, is of the form*

$$(5.11) \quad P(U_{p,f_2,f_1} \leq u) = \left(\prod_{i=1}^p K_i \right) \sum_{j=1}^m \left\{ C_j u^{f_1-l_j} \sum_{r_j=1}^{k_j+1} \frac{(-\log u)^{k_j-r_j+1} k_j!}{(f_1-l_j)^{r_j} (k_j-r_j+1)!} \right\},$$

$$= \left(\prod_{i=1}^p K_i \right) \sum_{h=1}^M \{ d_h u^{f_1-l_h} (-\log u)^{s_h} (f_1-l_h)^{-r_h} \},$$

where r_j , M , d_h and s_h are also constants and can be determined from p , f_2 and f_1 .

Note that, unlike the real case, here the formulae for the density and the distribution function are in the closed form for all values of p and f_2 , even or odd. Theorem 5.1 does not indicate explicitly how to find the values of the constants m , C_j , l_j and k_j , a task which is by no means easy for large values of p or f_2 . However, the theorem provides a basis for a recursive algorithm which can be programmed and run on the computer to obtain the density and the corresponding distribution functions at successive stages of the convolution process (see Schatzoff [8]).

It should be noted that (5.11) is a finite series of weighted chi-square distributions. From a practical standpoint (5.11) makes possible for the first time the calculation of exact percentage points of Wilks' criterion in the complex case for high dimensionality and small sample size. As noted by Goodman [3], that for every distributional result of classical multivariate Gaussian statistical analysis obtainable in closed (explicit) form, the counterpart analysis for complex Gaussian is also obtainable in closed (explicit) form with necessary changes. The present analysis is the counterpart analysis of the real case, obtained by Consul [1], Gupta [4], Pillai and Gupta [7] and Schatzoff [8].

6. Computation of percentage points

As per Theorem 3.1 the distribution of U_{p,f_2,f_1} , when $p=1$, is $\beta(f_1, f_2; U)$ and hence the percentage points can be obtained by entering a

variance ratio table with

$$F = \frac{1-U}{U} \frac{f_1}{f_2}, \quad \nu_1 = 2f_2, \quad \nu_2 = 2f_1,$$

or by entering directly with U the percentage point table of the beta variable (e.g. Table 16 of Biometrika Tables for Statisticians) with the same degrees of freedom. Also, according to Theorem 3.2, when $f_2=1$, $U_{p,1,f_1}$ (which has the same distribution as U_{1,p,f_1+1-p}) is distributed as $\beta(f_1+1-p, p; U)$ and we use the same tables with degrees of freedom $\nu_1=2p$, $\nu_2=2(f_1+1-p)$.

For $p=2$, (5.3) is programmed to compute the values of U_{2,f_2,f_1} on CDC 6400 to a minimum accuracy of five significant digits based on three arguments (f_2, f_1, α) where α is the lower probability level. For large values of f_1 (≥ 40) we used the approximation (5.3.2) of Khatri [5] neglecting terms of order $O(m^{-3})$. These values were then used to obtain the correction factor for converting chi-square percentiles with $2pf_2$ degrees of freedom to the exact percentiles of $-(2f_1+f_2-p) \log U_{p,f_2,f_1}$. Finally tabulation of the correction factors,

$$C = [\text{percentiles of } -(2f_1+f_2-p) \log U_{p,f_2,f_1}] / [\text{percentile of } \chi^2_{2pf_2}]$$

was made for $\alpha=0.1, 0.05, 0.025, 0.01, 0.005$; $f_2=2(1)10(2)20$ and $f_1=2(1)10(2)20, 24, 30, 40, 60, 120, \infty$. These factors are given to three decimal places although they were obtained to an accuracy of four decimals. The correction factors are presented in Table 1. Schatzoff [8] has shown that the linear interpolation in these tables gives reasonably accurate approximations to C . Pillai and Gupta [7] have described some of the uses of such tables in the real case. The computations in Table 1 were carried on CDC 6400 of the University of Arizona's Computer Center.

The author wishes to thank Professor K. C. S. Pillai of Purdue University for calling this subject to his attention.

Table 1. Chi-square adjustments to Wilks' criterion U , in the complex case.
Factor C for lower percentiles of U (upper percentiles of χ^2), $p=2$.

α f_1	$f_2=2$					$f_2=3$				
	0.100	0.050	0.025	0.010	0.005	0.100	0.050	0.025	0.010	0.005
2	1.200	1.223	1.244	1.271	1.290	1.261	1.289	1.315	1.349	1.374
3	1.074	1.081	1.088	1.097	1.104	1.105	1.115	1.124	1.135	1.143
4	1.039	1.043	1.046	1.051	1.054	1.059	1.064	1.068	1.074	1.078
5	1.024	1.027	1.029	1.032	1.034	1.038	1.041	1.044	1.047	1.050
6	1.017	1.018	1.020	1.022	1.023	1.026	1.029	1.031	1.033	1.035
7	1.012	1.013	1.014	1.016	1.017	1.020	1.021	1.023	1.024	1.026
8	1.009	1.010	1.011	1.012	1.013	1.015	1.016	1.017	1.019	1.020
9	1.007	1.008	1.009	1.009	1.010	1.012	1.013	1.014	1.015	1.016
10	1.006	1.006	1.007	1.008	1.008	1.010	1.011	1.011	1.012	1.013
12	1.004	1.004	1.005	1.005	1.006	1.007	1.007	1.008	1.008	1.009
14	1.003	1.003	1.004	1.004	1.004	1.005	1.005	1.006	1.006	1.007
16	1.002	1.002	1.003	1.003	1.003	1.004	1.004	1.004	1.005	1.005
18	1.002	1.002	1.002	1.002	1.002	1.003	1.003	1.004	1.004	1.004
20	1.001	1.001	1.002	1.002	1.002	1.002	1.003	1.003	1.003	1.003
24	1.001	1.001	1.001	1.001	1.001	1.002	1.002	1.002	1.002	1.002
30	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.002
40	1.000	1.000	1.000	1.000	1.000	1.001	1.001	1.001	1.001	1.001
60	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
120	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi_{2p}^2 f_2$	13.3616	15.5073	17.5346	20.0902	21.9550	18.5494	21.0261	23.3367	26.2170	28.2995
f_2	$f_2=4$					$f_2=5$				
	2	1.319	1.351	1.382	1.422	1.450	1.373	1.409	1.444	1.488
3	1.138	1.149	1.160	1.173	1.183	1.170	1.183	1.195	1.211	1.222
4	1.080	1.086	1.092	1.099	1.104	1.102	1.109	1.116	1.124	1.130
5	1.053	1.057	1.060	1.065	1.068	1.069	1.074	1.078	1.083	1.087
6	1.038	1.040	1.043	1.046	1.048	1.050	1.053	1.056	1.060	1.063
7	1.029	1.030	1.032	1.035	1.036	1.038	1.041	1.043	1.046	1.048
8	1.022	1.024	1.025	1.027	1.028	1.030	1.032	1.034	1.036	1.038
9	1.018	1.019	1.020	1.022	1.023	1.024	1.026	1.027	1.029	1.030
10	1.015	1.016	1.017	1.018	1.019	1.020	1.021	1.023	1.024	1.025
12	1.010	1.011	1.012	1.013	1.013	1.014	1.015	1.016	1.017	1.018
14	1.008	1.008	1.009	1.009	1.010	1.011	1.012	1.012	1.013	1.014
16	1.006	1.006	1.007	1.007	1.008	1.008	1.009	1.010	1.010	1.011
18	1.005	1.005	1.005	1.006	1.006	1.007	1.007	1.008	1.008	1.008
20	1.004	1.004	1.004	1.005	1.005	1.006	1.006	1.006	1.007	1.007
24	1.003	1.003	1.003	1.003	1.003	1.004	1.004	1.004	1.005	1.005
30	1.002	1.002	1.002	1.002	1.002	1.003	1.003	1.003	1.003	1.003
40	1.001	1.001	1.001	1.001	1.001	1.002	1.002	1.002	1.002	1.002
60	1.000	1.000	1.000	1.001	1.001	1.001	1.001	1.001	1.001	1.001
120	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi_{3p}^2 f_2$	23.5418	26.2962	28.8454	31.9999	34.2672	28.4120	31.4104	34.1696	37.5662	39.9968

p =number of variates; f_1 =error degrees of freedom; f_2 =hypothesis degrees of freedom; $C=[\text{percentile for } -(2f_1+f_2-p) \log_e U]/(\text{percentile for } \chi_{2p}^2 f_2)$.

Table 1. (Continued)

$\alpha \setminus f_1$	$f_2=6$					$f_2=7$				
	0.100	0.050	0.025	0.010	0.005	0.100	0.050	0.025	0.010	0.005
2	1.423	1.462	1.500	1.549	1.585	1.468	1.511	1.552	1.605	1.644
3	1.201	1.216	1.229	1.247	1.259	1.231	1.247	1.262	1.281	1.295
4	1.124	1.132	1.140	1.149	1.156	1.145	1.154	1.162	1.173	1.181
5	1.085	1.091	1.096	1.102	1.106	1.102	1.108	1.113	1.120	1.125
6	1.063	1.067	1.070	1.074	1.078	1.076	1.080	1.084	1.089	1.093
7	1.048	1.051	1.054	1.057	1.060	1.059	1.062	1.065	1.069	1.072
8	1.039	1.041	1.043	1.045	1.047	1.047	1.050	1.052	1.055	1.057
9	1.032	1.033	1.035	1.037	1.038	1.039	1.041	1.043	1.045	1.047
10	1.026	1.028	1.029	1.031	1.032	1.033	1.034	1.036	1.038	1.039
12	1.019	1.020	1.021	1.022	1.023	1.024	1.025	1.026	1.028	1.029
14	1.014	1.015	1.016	1.017	1.018	1.018	1.019	1.020	1.021	1.022
16	1.011	1.012	1.012	1.013	1.014	1.014	1.015	1.016	1.017	1.017
18	1.009	1.010	1.010	1.011	1.011	1.012	1.012	1.013	1.014	1.014
20	1.008	1.008	1.008	1.009	1.009	1.010	1.010	1.011	1.011	1.012
24	1.005	1.006	1.006	1.006	1.007	1.007	1.007	1.008	1.008	1.008
30	1.004	1.004	1.004	1.004	1.004	1.005	1.005	1.005	1.005	1.005
40	1.002	1.002	1.002	1.002	1.002	1.003	1.003	1.003	1.003	1.003
60	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.001	1.001
120	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi^2_{2p} f_2$	33.1963	36.4151	39.3641	42.9798	45.5585	37.9159	41.3372	44.4607	48.2782	50.9933
	$f_2=8$					$f_2=9$				
	0.100	0.050	0.025	0.010	0.005	0.100	0.050	0.025	0.010	0.005
2	1.511	1.557	1.600	1.657	1.698	1.551	1.599	1.645	1.704	1.748
3	1.259	1.277	1.293	1.314	1.328	1.286	1.305	1.323	1.345	1.361
4	1.166	1.176	1.186	1.197	1.205	1.187	1.198	1.208	1.220	1.229
5	1.118	1.125	1.131	1.138	1.144	1.134	1.141	1.148	1.156	1.162
6	1.089	1.094	1.098	1.104	1.108	1.102	1.107	1.112	1.118	1.122
7	1.070	1.074	1.077	1.081	1.084	1.081	1.085	1.088	1.093	1.096
8	1.056	1.059	1.062	1.065	1.068	1.066	1.069	1.072	1.075	1.078
9	1.047	1.049	1.051	1.054	1.058	1.054	1.057	1.060	1.062	1.064
10	1.039	1.041	1.043	1.045	1.047	1.046	1.048	1.050	1.053	1.054
12	1.029	1.030	1.032	1.033	1.034	1.034	1.036	1.037	1.039	1.040
14	1.022	1.023	1.024	1.026	1.026	1.026	1.028	1.029	1.030	1.031
16	1.018	1.018	1.019	1.020	1.021	1.021	1.022	1.023	1.024	1.025
18	1.014	1.015	1.016	1.016	1.017	1.017	1.018	1.019	1.020	1.020
20	1.012	1.012	1.013	1.014	1.014	1.014	1.015	1.016	1.016	1.017
24	1.009	1.009	1.009	1.010	1.010	1.010	1.011	1.011	1.012	1.012
30	1.006	1.006	1.006	1.007	1.007	1.007	1.007	1.008	1.008	1.008
40	1.003	1.004	1.004	1.004	1.004	1.004	1.004	1.004	1.005	1.005
60	1.001	1.002	1.002	1.002	1.002	1.002	1.002	1.002	1.002	1.002
120	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.001	1.001	1.001
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi^2_{2p} f_2$	42.5847	46.1943	49.4804	53.4858	56.3281	47.2122	50.9985	54.4373	58.6192	61.5812

Table 1. (Continued)

$\frac{\alpha}{f_1}$	$f_2=10$					$f_2=12$				
	0.100	0.050	0.025	0.010	0.005	0.100	0.050	0.025	0.010	0.005
2	1.588	1.638	1.687	1.749	1.795	1.655	1.710	1.763	1.830	1.880
3	1.312	1.332	1.351	1.374	1.391	1.361	1.383	1.404	1.429	1.448
4	1.207	1.218	1.229	1.242	1.252	1.244	1.257	1.269	1.284	1.295
5	1.150	1.158	1.165	1.174	1.180	1.181	1.190	1.198	1.208	1.215
6	1.115	1.121	1.126	1.132	1.137	1.141	1.147	1.153	1.160	1.166
7	1.092	1.096	1.100	1.105	1.108	1.113	1.118	1.123	1.128	1.132
8	1.075	1.078	1.082	1.086	1.088	1.093	1.098	1.101	1.106	1.109
9	1.062	1.065	1.068	1.071	1.074	1.079	1.082	1.085	1.089	1.092
10	1.053	1.056	1.058	1.060	1.062	1.067	1.070	1.073	1.076	1.078
12	1.040	1.042	1.043	1.045	1.046	1.051	1.053	1.055	1.057	1.059
14	1.031	1.032	1.034	1.035	1.036	1.040	1.042	1.043	1.045	1.046
16	1.025	1.026	1.027	1.028	1.029	1.032	1.034	1.035	1.036	1.037
18	1.020	1.021	1.022	1.023	1.024	1.027	1.028	1.029	1.030	1.031
20	1.017	1.018	1.018	1.019	1.020	1.022	1.023	1.024	1.025	1.026
24	1.012	1.013	1.013	1.014	1.014	1.016	1.017	1.018	1.018	1.019
30	1.008	1.009	1.009	1.009	1.010	1.011	1.012	1.012	1.012	1.013
40	1.005	1.005	1.005	1.006	1.006	1.007	1.007	1.007	1.008	1.008
60	1.002	1.002	1.002	1.003	1.003	1.003	1.003	1.003	1.004	1.004
120	1.001	1.001	1.001	1.001	1.001	1.000	1.000	1.000	1.001	1.001
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi^2_{2pf_2}$	51.8050	55.7585	59.3417	63.6907	66.7659	60.9066	65.1708	69.0226	73.6826	76.9688
$f_2=14$						$f_2=16$				
2	1.715	1.774	1.830	1.902	1.956	1.769	1.831	1.891	1.967	2.024
3	1.405	1.429	1.452	1.480	1.500	1.447	1.472	1.496	1.526	1.548
4	1.280	1.294	1.307	1.324	1.335	1.313	1.329	1.343	1.360	1.373
5	1.210	1.220	1.229	1.240	1.248	1.238	1.248	1.258	1.270	1.279
6	1.165	1.172	1.179	1.187	1.193	1.189	1.197	1.204	1.213	1.220
7	1.134	1.140	1.145	1.152	1.156	1.155	1.161	1.167	1.174	1.179
8	1.112	1.116	1.121	1.126	1.129	1.130	1.135	1.140	1.145	1.149
9	1.095	1.099	1.102	1.106	1.109	1.111	1.115	1.119	1.124	1.127
10	1.082	1.085	1.088	1.091	1.094	1.096	1.100	1.103	1.107	1.110
12	1.062	1.065	1.067	1.070	1.072	1.074	1.077	1.079	1.082	1.084
14	1.050	1.051	1.053	1.055	1.057	1.059	1.061	1.063	1.066	1.067
16	1.040	1.042	1.043	1.045	1.046	1.048	1.050	1.052	1.054	1.055
18	1.033	1.035	1.036	1.037	1.038	1.040	1.042	1.043	1.045	1.046
20	1.028	1.029	1.030	1.031	1.032	1.034	1.036	1.037	1.038	1.039
24	1.021	1.022	1.022	1.023	1.024	1.026	1.026	1.027	1.028	1.029
30	1.014	1.015	1.015	1.016	1.016	1.018	1.018	1.019	1.020	1.020
40	1.009	1.009	1.009	1.010	1.010	1.011	1.011	1.012	1.012	1.012
60	1.004	1.004	1.004	1.005	1.005	1.005	1.005	1.006	1.006	1.006
120	1.002	1.002	1.002	1.002	1.003	1.003	1.003	1.003	1.003	1.003
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
$\chi^2_{2pf_2}$	69.9185	74.4683	78.5672	83.5134	86.9938	78.8596	83.6753	88.0041	93.2169	96.8781

Table 1. (Continued)

α f_1	$f_2=18$					$f_2=20$				
	0.100	0.050	0.025	0.010	0.005	0.100	0.050	0.025	0.010	0.005
2	1.819	1.884	1.946	2.026	2.085	1.864	1.932	1.997	2.080	2.142
3	1.485	1.512	1.538	1.569	1.593	1.520	1.549	1.576	1.610	1.634
4	1.344	1.361	1.376	1.395	1.409	1.374	1.392	1.408	1.428	1.442
5	1.264	1.276	1.286	1.299	1.308	1.290	1.302	1.313	1.327	1.337
6	1.212	1.220	1.228	1.238	1.245	1.234	1.243	1.252	1.262	1.269
7	1.175	1.182	1.188	1.196	1.201	1.194	1.202	1.208	1.216	1.222
8	1.148	1.153	1.158	1.164	1.169	1.165	1.171	1.176	1.183	1.188
9	1.127	1.131	1.136	1.141	1.144	1.142	1.147	1.152	1.157	1.161
10	1.110	1.114	1.118	1.122	1.125	1.124	1.128	1.132	1.137	1.140
12	1.086	1.089	1.091	1.095	1.097	1.097	1.101	1.104	1.107	1.110
14	1.069	1.071	1.073	1.076	1.078	1.079	1.081	1.084	1.086	1.088
16	1.057	1.059	1.060	1.062	1.064	1.065	1.067	1.069	1.072	1.073
18	1.048	1.049	1.051	1.052	1.054	1.055	1.057	1.058	1.060	1.062
20	1.041	1.042	1.043	1.045	1.046	1.047	1.048	1.050	1.051	1.052
24	1.030	1.032	1.032	1.034	1.034	1.036	1.037	1.038	1.039	1.040
30	1.021	1.022	1.023	1.023	1.024	1.025	1.026	1.026	1.027	1.028
40	1.013	1.014	1.014	1.014	1.015	1.016	1.016	1.016	1.017	1.017
60	1.012	1.012	1.013	1.013	1.013	1.014	1.015	1.015	1.015	1.015
120	1.004	1.004	1.004	1.004	1.004	1.004	1.004	1.004	1.005	1.005
∞	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
χ^2_{2p, f_2}	87.7433	92.8083	97.3532	102.816	106.648	96.5782	101.879	106.629	112.329	116.321

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