A QUEUEING SYSTEM WITH SEVERAL TYPES OF CUSTOMERS

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1. Introduction

Consider a single-server queueing system in which

- (i) there are r types of customers
- (ii) the arrivals are in batches such that each batch contains K_1 customers of the first type, K_2 customers of the second type, \cdots , K_r customers of the rth type with the total service times X_1, X_2, \cdots , X_r respectively, where $\{K_1, K_2, \cdots, K_r, X_1, X_2, \cdots, X_r\}$ is a random vector
- (iii) the structures $\{K_1, K_2, \dots, K_r, X_1, X_2, \dots, X_r\}$ for the different batches are mutually independent
- (iv) the interarrival times of the batches are independent and identically distributed with distribution function A(t)
- (v) the batch structures are independent of the interarrival times.

In the present paper the joint behaviour of the number of customers of each type served during a busy period and their total service time is studied. Furthermore, the explicit properties of the special cases with

$$dA(t) = \frac{\lambda^m}{(m-1)!} \exp(-\lambda t) t^{m-1} dt \qquad (t>0)$$

and

$$dA(t) = \exp(-\lambda t)\lambda dt$$
 $(t>0)$

are given.

Prior to the present work Welch [4] and the author [3] had studied the joint behaviour of the number of customers of each type served during a busy period and its length for the special case of the present queueing system with

$$dA(t) = \exp\left(-\sum_{i=1}^{r} \lambda_{i} t\right) \left(\sum_{i=1}^{r} \lambda_{i}\right) dt$$

$$P(K_{j}=1, K_{i}=0 \ (i(\neq j)=1, 2, \cdots, r)) = \frac{\lambda_{j}}{\sum_{i=1}^{r} \lambda_{i}} \qquad (j=1, 2, \cdots, r).$$

2. General queueing system

Let T_i $(i=1, 2, \dots, r)$ be the total service time of the *i*th type in a busy period and N_i $(i=1, 2, \dots, r)$ and N be the number of customers of the *i*th type and the number of batches served during it respectively. Then, defining $E(Q_1; Q_2) = E(Q_1 | Q_2) P(Q_2)$ if $P(Q_2) > 0$ and = 0 otherwise, we have for suitable θ_i and z_i $(i=1, 2, \dots, r)$

$$(1) \qquad E\left(\prod_{i=1}^{r} e^{-\theta_{i}T_{i}}z_{i}^{N_{i}}; N=n\right)$$

$$= \begin{cases} \int \cdots \int dU(y_{1}; \cdot)dU(y_{2}; \cdot) \cdots dU(y_{n}; \cdot)dA(v_{1})dA(v_{2}) \cdots dA(v_{n}) \\ y_{1}+y_{2}+\cdots+y_{i}-(v_{1}+v_{2}+\cdots+v_{i}) \geq 0 \qquad (l=1, 2, \cdots, n-1) \\ y_{1}+y_{2}+\cdots+y_{n}-(v_{1}+v_{2}+\cdots+v_{n}) < 0 \end{cases}$$

$$= \begin{cases} (G(\cdot))^{n} \int \cdots \int dB(y_{1}; \cdot)dB(y_{2}; \cdot) \cdots dB(y_{n}; \cdot)dA(v_{1})dA(v_{2}) \cdots dA(v_{n}) \\ y_{1}+y_{2}+\cdots+y_{i}-(v_{1}+v_{2}+\cdots+v_{i}) \geq 0 \qquad (l=1, 2, \cdots, n-1) \\ y_{1}+y_{2}+\cdots+y_{n}-(v_{1}+v_{2}+\cdots+v_{n}) < 0 \\ (G(\cdot))^{n}P(N^{*}=n) \qquad (n=1, 2, \cdots), \end{cases}$$

where $U(y; \cdot)$ denotes

$$U(y; z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r) = E\left(\prod_{i=1}^r e^{-\theta_i X_i} z_i^{K_i}; X_1 + X_2 + \dots + X_r \leq y\right),$$

 $G(\cdot)$ denotes

$$G(z_1, z_2, \cdots, z_r; \theta_1, \theta_2, \cdots, \theta_r) = E\left(\prod_{i=1}^r e^{-\theta_i X_i} z_i^{K_i}\right),$$

 $B(y;\cdot)$ denotes

$$B(y; z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots \theta_r) = \frac{U(y; z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r)}{G(z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r)},$$

and N^* denotes the number of customers served during a busy period for the queueing system GI/G/1 with the interarrival time distribution as A(t) and the service time distribution as $B(t; z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r)$ (For n=1 $y_1+y_2+\dots+y_t-(v_1+v_2+\dots+v_t)\geq 0$ $(l=1, 2, \dots, n-1)$ is not to be read. Also note that for t<0 A(t) and $U(t; z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r)$ equal 0).

Hence from Finch [1] it follows that the generating function of the Laplace-Stieltjes transforms $E\left(\prod_{i=1}^{r}e^{-\theta_{i}T_{i}}; N_{i}=n_{i} \ (i=1,\,2,\,\cdots,\,r),\,N=n\right)$ is given by

$$(2) E\left\{\left(\prod_{i=1}^{r} e^{-\theta_{i}T_{i}} z_{i}^{N_{i}}\right) z^{N}\right\} = E\left\{\left(zG(z_{1}, z_{2}, \dots, z_{r}; \theta_{1}, \theta_{2}, \dots, \theta_{r})\right)^{N^{*}}\right\}$$

$$= 1 - \exp\left\{-\sum_{n=1}^{\infty} \frac{z^{n}}{n} \int_{t=0}^{\infty} (1 - A_{n}(t)) dB_{n}(t; z_{1}, z_{2}, \dots, z_{r}; \theta_{1}, \theta_{2}, \dots, \theta_{r})\right\}$$

where $A_n(t)$ is the n-fold convolution of A(t) with itself, and

$$egin{aligned} B_n(t\,;\, z_1,\, z_2,\, \, \cdots,\, z_r\,;\, heta_1,\, heta_2,\, \, \cdots,\, heta_r) \ &= E\Big(\prod\limits_{i=1}^r \, e^{- heta_i X_i^{(n)}} z_i^{K_i^{(n)}};\, X_1^{(n)} + X_2^{(n)} + \cdots + X_r^{(n)} \leqq t\Big)\,, \end{aligned}$$

where $K_i^{(n)}$ $(i=1, 2, \dots, r)$ denotes the number of customers of the *i*th type in a group of *n* batches and $X_i^{(n)}$ $(i=1, 2, \dots, r)$ their total service time.

It now readily follows that the generating function of the Laplace-Stieltjes transforms $E(e^{-\theta T}; N_i = n_i \ (i = 1, 2, \dots, r), \ N = n)$, where $T = \sum T_i$ is the length of the busy period, is given by

$$(3) E\left\{\left(\prod_{i=1}^{r} z_{i}^{N_{i}}\right) \exp\left(-\theta T\right) z^{N}\right\}$$

$$=1-\exp\left\{-\sum_{n=1}^{\infty} \frac{z^{n}}{n} \int_{t=0}^{\infty} (1-A_{n}(t)) dB_{n}(t; z_{1}, z_{2}, \dots, z_{r}; \theta, \theta, \dots, \theta)\right\}.$$

3. Some special queueing systems

We shall now consider the following special cases:

(i) Special case with
$$dA(t) = \frac{\lambda^m}{(m-1)!} e^{-\lambda t} t^{m-1} dt$$
 $(t>0)$

From Takács ([5] p. 108) it follows that the generating function

$$\begin{split} E\left\{\left(\prod_{i=1}^{r}e^{-\theta_{i}T_{i}}z_{i}^{N_{i}}\right)z^{N}\right\} \\ =&1-\prod_{k=1}^{m}\left(1-\alpha_{k}(z_{1},z_{2},\,\cdots,\,z_{r};\theta_{1},\,\theta_{2},\,\cdots,\,\theta_{r};\,z)\right) \\ (R(\theta_{i})\geq&0\,(i=1,\,2,\,\cdots,\,r),\,\,|z_{i}|\leq&1\,(i=1,\,2,\,\cdots,\,r),\,|z|<1) \end{split}$$

where $\alpha_k(z_1, z_2, \dots, z_r; \theta_1, \theta_2, \dots, \theta_r; z)$ $(k=1, 2, \dots, m)$ are the m distinct roots in the unit circle $|\alpha| < 1$ of the equation

$$\alpha^m = zG(z_1, z_2, \dots, z_r; \theta_1 + \lambda - \lambda \alpha, \theta_2 + \lambda - \lambda \alpha, \dots, \theta_r + \lambda - \lambda \alpha)$$
.

(It may be noted that the usage of the Lagrange expansion establishes that

$$\sum_{k=1}^{m} \log \left(1-\alpha_{k}(z_{1}, z_{2}, \cdots, z_{r}; \theta_{1}, \theta_{2}, \cdots, \theta_{r}; z)\right)$$

$$=\sum_{n=1}^{\infty} \frac{mz^{n}}{(nm)!} \left(\frac{\partial}{\partial x}\right)^{nm-1} \left(G(z_{1}, z_{2}, \cdots, z_{r}; \theta_{1}+\lambda-\lambda x, \theta_{2}+\lambda-\lambda x, \cdots, \theta_{r}+\lambda-\lambda x)^{n} \left(\frac{-1}{1-x}\right)\right)\Big|_{x=0}$$

(For the details see Prabhu ([2] p. 154))). This result can also be obtained by using (2).

It follows immediately that the generating function

(5)
$$E\left\{(e^{-\theta T}z^{N})\prod_{i=1}^{r}z_{i}^{N_{i}}\right\} = 1 - \prod_{k=1}^{m}\left(1 - \alpha_{k}(z_{1}, z_{2}, \dots, z_{r}; \theta, \theta, \dots, \theta; z)\right),$$

$$(R(\theta) \geq 0, |z| < 1, |z_{i}| \leq 1 \ (i = 1, 2, \dots, r)).$$

(ii) Special case with $dA(t) = \lambda \exp(-\lambda t)dt$ (t>0)

It is well known that for a queueing system M/G/1 with the interarrival time mean λ^{-1} and the service time distribution function B(t) the probability that n customers are served during a busy period is given by

$$f_n = \int_0^\infty \exp(-\lambda y) \frac{(\lambda y)^{n-1}}{n!} dB_n(y) , \qquad (n \ge 1)$$

where $B_n(y)$ is the *n*-fold convolution of B(y) with itself (see, for example, Finch [1]). Hence for $n=1, 2, \cdots$ the transform

(7)
$$E\left(\prod_{i=1}^{r} e^{-\theta_{i}T_{i}} z_{i}^{N_{i}}; N=n\right)$$

$$= \int_{y=0}^{\infty} \exp\left(-\lambda y\right) \frac{(\lambda y)^{n-1}}{n!} dB_{n}(y; z_{1}, z_{2}, \dots, z_{r}; \theta_{1}, \theta_{2}, \dots, \theta_{r})$$

$$= \frac{\lambda^{n-1}}{n!} E\left(\left(\prod_{i=1}^{r} e^{-(\theta_{i}+\lambda)X_{i}^{(n)}} z_{i}^{K_{i}^{(n)}}\right) \left(\sum_{i=1}^{r} X_{i}^{(n)}\right)^{n-1}\right)$$

which establishes that for $n=1, 2, \cdots$

(8)
$$P(N_{i}=n_{i}(i=1, 2, \dots, r); T_{i} \leq t_{i} (i=1, 2, \dots, r); N=n)$$

$$= \frac{\lambda^{n-1}}{n!} E\left(e^{-\lambda \sum_{i=1}^{r} X_{i}^{(n)}} \left(\sum_{i=1}^{r} X_{i}^{(n)}\right)^{n-1}; K_{i}^{(n)} = n_{i} (i=1, 2, \dots, r); X_{i}^{(n)} \leq t_{i} (i=1, 2, \dots, r)\right).$$

It may further be seen that for $n=1, 2, \cdots$

(9)
$$P(N_{i}=n_{i} (i=1, 2, \dots, r); T \leq t; N=n)$$

$$= \frac{\lambda^{n-1}}{n!} E(e^{-\lambda X(n)}(X(n))^{n-1}; X(n) \leq t; K_{i}^{(n)}=n_{i} (i=1, 2, \dots, r))$$

where $X(n) = \sum_{i=1}^{r} X_i^{(n)}$.

The reader may note that using the present method the joint behaviour of the number of customers of each type arriving during a busy period initiated by an occupation time of quantity x and their total service time can be studied.

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