## AN EXAMPLE OF THE TWO-SIDED WILCOXON TEST WHICH IS NOT UNBIASED

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The purpose of this note is to answer the question stated in Lehmann ([1], p. 240), of whether the two-sided Wilcoxon test is unbiased against the two-sided alternatives or not. Although every one-sided Wilcoxon test is unbiased against the one-sided alternatives, we shall show by an example that the two-sided Wilcoxon test is not necessarily unbiased against some special two-sided translation alternatives.

Let  $X_1, \dots, X_m$  and  $Y_1, \dots, Y_n$  be two random samples drawn from the distributions F(x) and G(x), respectively. These distributions are assumed to be absolutely continuous with respect to the Lebesgue measure. To test the hypothesis H: F(x) = G(x), consider the following test function  $\phi(X_1, \dots, X_m; Y_1, \dots, Y_n)$ :

(1) 
$$\phi(X_1, \dots, \hat{X}_m; Y_1, \dots, Y_n) = \begin{cases} 1 & \text{when } X_1, \dots, X_m < Y_1, \dots, Y_n \\ & \text{or } X_1, \dots, X_m > Y_1, \dots, Y_n, \\ 0 & \text{otherwise.} \end{cases}$$

Then it is clear that the test  $\phi$  is a two-sided Wilcoxon test with the level  $\alpha = 2(m!n!)/(m+n)!$ . The power of the test is given by

$$(2) \quad \beta = P\{X_1, \dots, X_m < Y_1, \dots, Y_n\} + P\{X_1, \dots, X_m > Y_1, \dots, Y_n\}$$

$$= P\{\max_{1 \le i \le m} X_i < \min_{1 \le j \le n} Y_j\} + P\{\min_{1 \le i \le m} X_i > \max_{1 \le j \le n} Y_j\}$$

$$= \int_{-\infty}^{\infty} F(x)^m d[1 - (1 - G(x))^n] + \int_{-\infty}^{\infty} G(x)^n d[1 - (1 - F(x))^m]$$

$$= n \int_{-\infty}^{\infty} F(x)^m (1 - G(x))^{n-1} dG(x) + m \int_{-\infty}^{\infty} G(x)^n (1 - F(x))^{m-1} dF(x).$$

Let us now specify the distribution functions F(x) and G(x) as follows:

$$F(x) = \begin{cases} 1 - e^{-x} & x \ge 0 \\ 0 & x < 0 \end{cases},$$

$$G(x) = F(x - \Delta),$$

and test the hypothesis  $H:\Delta=0$  against the two-sided alternatives  $K:\Delta\neq 0$ .

The power function,  $\beta_{m,n}(\Delta)$ , of the test  $\phi$  for this special case is obtained from (2):

$$(4) \qquad \beta_{m, n}(\Delta) = n \int_{\Delta}^{\infty} (1 - e^{-x})^{m} e^{-n(x - \Delta)} dx + m \int_{\Delta}^{\infty} (1 - e^{-(x - \Delta)})^{n} e^{-mx} dx$$

$$= \frac{m! n!}{(m + n)!} e^{-m\Delta} + n e^{n\Delta} \int_{0}^{e^{-\Delta}} t^{n-1} (1 - t)^{m} dt$$

$$= \frac{m! n!}{(m + n)!} e^{-m\Delta} + n \int_{0}^{1} x^{n-1} (1 - e^{-\Delta}x)^{m} dx$$

for  $\Delta \ge 0$  and similarly

(5) 
$$\beta_{m,n}(\Delta) = \frac{m! n!}{(m+n)!} e^{n\Delta} + m \int_0^1 x^{m-1} (1-e^{\Delta}x)^n dx$$

for  $\Delta < 0$ . From (4) and (5) we have

$$\beta_{m, n}(\Delta) = \beta_{n, m}(-\Delta).$$

Differentiating (4) with respect to  $\Delta$ , we have

(7) 
$$h(\Delta) = e^{\Delta} \beta'_{m,n}(\Delta)$$

$$= -\frac{m! n!}{(m+n)!} m e^{-(m-1)\Delta} + m n \int_0^1 x^n (1 - e^{-\Delta} x)^{m-1} dx,$$

and

(8) 
$$k(0) = \frac{m! n!}{(m+n)!} (n-m).$$

Since  $h(\Delta)$  is strictly increasing for  $\Delta \ge 0$  whenever m > 1, (a) the equation  $h(\Delta) = 0$  has a unique solution in the interval  $(0, \infty)$  when n < m, (b) h(0) = 0, when m = n > 1, (c)  $h(\Delta) > 0$  for  $\Delta \ge 0$  when n > m > 1. In the case that m = 1,  $\beta'_{1, n}(\Delta) = (n-1)e^{-\Delta}/(n+1)$  and hence  $\beta'_{1, n}(\Delta)$  is always positive for  $\Delta \ge 0$  when n > m = 1.

Considering the symmetric property of  $\beta_{m,n}(\Delta)$  as in (6), we can conclude that in all cases except when m=n=1,  $\beta'_{m,n}(\Delta)=0$  has a unique solution  $\Delta=\Delta_0$ ,  $\beta_{m,n}(\Delta)$  is strictly decreasing for  $\Delta<\Delta_0$  and strictly increasing for  $\Delta>\Delta_0$ , and further that  $\Delta_0>0$ ,  $\Delta_0=0$  and  $\Delta_0<0$  according as m>n, m=n>1 and n>m, respectively. In other words, the test  $\phi$  defined by (1) is not unbiased against the two-sided alternatives  $\Delta\neq 0$  when  $m\neq n$  and is unbiased when m=n>1. From (4) and (5) it can easily be seen that

$$\lim_{\Delta \to +\infty} \beta_{m, n}(\Delta) = 1.$$

We shall give two numerical examples which are obtained from (4) and (5).

(i) In case m=19 and n=1:  $\alpha=0.10$ 

$$eta_{_{19,1}}\!(\Delta)\!=\!\left\{egin{array}{ll} rac{1}{20}\!\left[e^{_{_{_{19,1}}}}\!\!+\!e^{_{\Delta}}\!\left\{1\!-\!(1\!-\!e^{_{_{_{_{_{_{19,1}}}}}}}\!
ight\}
ight] & (\Delta\!\geq\!0) \ 1\!-\!rac{9}{10}e^{_{\Delta}} & (\Delta\!<\!0) \;. \end{array}
ight.$$

(ii) In case m=n=3:  $\alpha=0.10$ 

$$eta_{3,3}(\Delta) = \left\{egin{array}{ll} 1 - rac{9}{4}e^{-\Delta} + rac{9}{5}e^{-2\Delta} - rac{9}{20}e^{-3\Delta} & (\Delta \geqq 0) \ 1 - rac{9}{4}e^{\Delta} + rac{9}{5}e^{2\Delta} - rac{9}{20}e^{3\Delta} & (\Delta < 0) \end{array}
ight..$$

The following figure shows the power function  $\beta_{m,n}(\Delta)$  for the cases (i) and (ii). In both cases the level of the test  $\phi$  is equal to 0.10. In case of (i) the power function attains the minimum value 0.062 at about  $\Delta=0.2$  and is less than 0.10 for  $0<\Delta<0.7$ .

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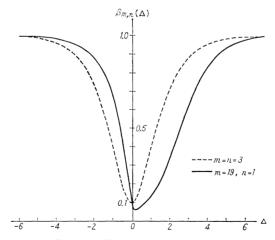


Fig. 1. The power of the test

## REFERENCE

[1] E. L. Lehmann, Testing Statistical Hypotheses, John Wiley and Sons Inc., 1959.