SOME INEQUALITIES RELATING TO THE PARTIAL SUM OF BINOMIAL PROBABILITIES

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1. Introduction.

Uspensky [1, p. 102] gives an inequality relating to the partial sum of binomial probabilities: Let X be a random variable following a binomial distribution B(n, p), arising from n repetitions of an event with probability p. Then it holds that

$$P(|X/n-p| \ge c) < 2e^{-nc^2/2}$$

for any constant c>0 and any p with 0< p<1. Its proof, however, is too tedious, although elementary. In the following we shall give a simplified proof for a somewhat strengthened result (Theorem 1). By the same method we can also obtain some other inequalities which prove to be useful in Matusita's theory [2, 3] of test of fit, two-sample problem, test of independence, etc.

2. Two lemmas

We shall state two lemmas the first of which is a corollary of a theorem given by Chernoff (Theorem 1 in [4]).

LEMMA 1. Let X be a random variable following B(n, p) and x a constant, $0 \le x \le 1$, which may depend on n or p. It holds then that

(i)
$$P(X/n \ge x) \le e^{-n\varphi(x)}$$
 if $x \ge p$, and

(ii)
$$P(X/n \le x) \le e^{-n\varphi(x)}$$
 if $x \le p$,

where

$$\varphi(x) = x \log \frac{x}{p} + (1-x) \log \frac{1-x}{q}$$

and q=1-p.

LEMMA 2. The function $\varphi(x)$ defined in Lemma 1 satisfies the following inequalities:

(a)
$$\varphi(x) \ge 2(x-p)^2$$
 if $0 \le x \le 1$,

(b)
$$\varphi(x) \ge \frac{(x-p)^2}{2pq}$$
 if $p \le x \le 1$, $p \ge \frac{1}{2}$

(c)
$$\varphi(x) \ge 2(\sqrt{x} - \sqrt{p})^2$$
 if $p \le x \le 1$,

(d)
$$\varphi(x) \ge (\sqrt{p} - \sqrt{x})^2$$
 if $0 \le x \le p$,

where the equality sign holds in each case if and only if x=p.

PROOF. First we have

$$\varphi'(x) = \log \frac{x}{p} - \log \frac{1-x}{q} ,$$

$$\varphi''(x) = \frac{1}{x(1-x)} \ge 0 ,$$

consequently

$$\varphi(p) = \varphi'(p) = 0.$$

For the proof of (a), put $\varphi_1(x) = 2(x-p)^2$. Then

$$\varphi_{\mathbf{1}}(p) = \varphi'_{\mathbf{1}}(p) = 0$$

and

$$\varphi_1''(x) = 4 \leq \varphi''(x) \quad \text{if } 0 \leq x \leq 1,$$

where the equality holds at a single point x=1/2. From (2), (3) and (4) we obtain

$$\varphi_1(x) \leq \varphi(x)$$
 if $0 \leq x \leq 1$,

with the equality sign only for x=p.

Concerning (b) and (b'), putting $\varphi_2(x) = \frac{(x-p)^2}{2pq}$, we can prove them similarly.

Re (c). Put
$$\varphi_3(x) = 2(\sqrt{x} - \sqrt{p})^2$$
.

Re (d). The proof of this case is most lengthy. We shall first prove

(5)
$$\varphi(p-\sqrt{p}c) \ge c^2/2 \quad \text{if } 0 \le c \le \sqrt{p}.$$

If $p \le 1/2$, then (b') implies

$$\varphi(p-\sqrt{p}c) \ge \frac{c^2}{2a} \ge \frac{c^2}{2}$$

and if $p \ge 1/2$, then (a) implies

$$\varphi(p-\sqrt{p}c) \geq 2pc^2 \geq c^2 \geq c^2/2$$
.

Then we have (5).

Now we consider two cases:

Case (i) where x satisfies

$$(6) \qquad (\sqrt{2}-1)^2 p \leq x \leq p.$$

Put $c=\sqrt{p}-\sqrt{x}$. Then (6) is equivalent to $0 \le c \le (2-\sqrt{2})\sqrt{p}$, which implies

$$(7) x=(\sqrt{p}-c)^2 \leq p-\sqrt{2p}c.$$

By (1) and (2) $\varphi(x)$ decreases monotonically in the interval $0 \le x \le p$. Therefore (5) and (7) give

$$\varphi(x) \geqq \varphi(p - \sqrt{2p} c) \geqq c^2$$
 ,

which is (d) for the case (i).

Case (ii) where x satisfies

(8)
$$0 \le x \le (\sqrt{2} - 1)^2 p$$
.

If we define the function $\psi(x)$ as

(9)
$$\psi(x) = \varphi(x) - (\sqrt{p} - \sqrt{x})^2,$$

then its first two derivatives are

(10)
$$\psi'(x) = \log \frac{x}{p} - \log \frac{1-x}{q} - \left(1 - \sqrt{\frac{p}{x}}\right),$$

(11)
$$\psi''(x) = \frac{1}{2\sqrt{x^3}(1-x)} \left\{ 2\sqrt{x} - \sqrt{p} (1-x) \right\} .$$

Since the formula in the braces of (11) increases monotonically for $x \ge 0$ and its value at $x = (\sqrt{2} - 1)^2 p$ is easily seen to be non-positive, we obtain for any value of x in the interval (8)

$$\phi''(x) \leq 0$$
.

Since we have from (9) and (10)

$$\psi(0) = -\log q - p \ge 0$$
 and $\psi'(0) = \infty$,

we have only to show

(12)
$$\psi((\sqrt{2}-1)^2p) \ge 0 \quad \text{if } 0 \le p \le 1$$

in order to prove $\psi(x) \ge 0$ for any x in (8). Now, from (9) we have

$$\psi((\sqrt{2}-1)^2p) = [1-(\sqrt{2}-1)^2p] \log \frac{1-(\sqrt{2}-1)^2p}{q} + 2(\sqrt{2}-1)^2[\log (\sqrt{2}-1)-1]p = \zeta(p) \ (say) \ .$$

The function $\zeta(p)$ is defined in $0 \le p \le 1$. Since $\zeta(0) = 0$, in order to prove (12) or $\zeta(p) \ge 0$ it suffices to verify

(13)
$$\zeta'(p) \ge 0 \quad \text{for } 0 \le p \le 1.$$

The derivative of $\zeta(p)$ can be expressed as

(14)
$$\zeta'(p) = (\sqrt{2} - 1)^2 [2 \log (\sqrt{2} - 1) - 3] - (\sqrt{2} - 1)^2 \log \xi(p) + \xi(p)$$
, where

$$\xi(p) = \frac{1 - (\sqrt{2} - 1)^2 p}{q}$$
.

The function $\xi(p)$ defined in $0 \le p \le 1$ is clearly monotone-increasing and therefore

(15)
$$\xi(p) \ge \xi(0) = 1, \qquad 0 \le p \le 1$$
,

which implies

(16)
$$\log \xi(p) \leq \xi(p) - 1, \qquad 0 \leq p \leq 1.$$

Finally (14), (15) and (16) together imply (13), for

$$\zeta'(p) \ge (\sqrt{2} - 1)^2 \left[2 \log (\sqrt{2} - 1) - 3\right] - (\sqrt{2} - 1)^2 \left[\xi(p) - 1\right] + \xi(p)$$

$$\ge (\sqrt{2} - 1)^2 \left[2 \log (\sqrt{2} - 1) - 3\right] + 1 > 0.$$

It will readily be seen that the equality condition in (d) is x=p. This completes the proof of Lemma 2.

3. Theorems

Let X be a binomial variate with B(n,p), 0 , and c a non-negative constant depending possibly on <math>n or p. From Lemmas 1 and 2 in the

preceding section we have readily the following theorems.

THEOREM 1

$$P\left(\frac{x}{n}-p\geq c\right)< e^{-2ncx},$$

(ii)
$$P\left(\frac{x}{n}-p\leq -c\right) < e^{-2nc^2}.$$

THEOREM 2

(i)
$$P\left(\frac{x}{n}-p\geq c\right) < \exp\left(-\frac{nc^2}{2pq}\right)$$
 for $p\geq \frac{1}{2}$,

(ii)
$$P\left(\frac{x}{n} - p \le -c\right) < \exp\left(-\frac{nc^2}{2pq}\right) \qquad for \ p \le \frac{1}{2}.$$

THEOREM 3

$$P\left(\sqrt{\frac{x}{n}} - \sqrt{p} \ge c\right) < e^{-2nc^2}$$
.

THEOREM 4

$$P\left(\sqrt{\frac{x}{n}} - \sqrt{p} \leq -c\right) < e^{-nc^2}$$
.

We note that the equality signs for c=0 which are to be present in these formulas in applying Lemmas 1 and 2 are absent there. This is justified by the direct consideration of properties of the binomial distribution, where we restrict p in the open interval 0 .

4. Application to Matusita's multionomial distance.

Let F be a multinomial distribution with k classes and a set of probabilities (p_1, \dots, p_k) , $p_i > 0$, $\sum p_i = 1$, and let S_n be an empirical distribution with relative frequencies $(n_1/n, \dots, n_k/n)$, $(\sum n_i = n)$. Matusita [2], [3] defined the distance between S_n and F by the formula

(17)
$$||S_n - F||^2 = \sum_{i=1}^k \left(\sqrt{\frac{n_i}{n}} - \sqrt{p_i} \right)^2.$$

which we shall refer to as Matusita's multinomial distance. He and M. Motoo [5] proved that

(18)
$$P(||S_n - F||^2 \ge \eta^2) \le \frac{k^2 + k - 1}{(n\eta^2)^2}$$

for any positive constant η . Now we obtain from Theorems 3 and 4

the following

THEOREM 5

(19)
$$P(||S_n - F||^2 \ge \eta^2) < k \left\{ \exp\left(-\frac{2n\eta^2}{k}\right) + \exp\left(-\frac{n\eta^2}{k}\right) \right\}.$$

PROOF. Clearly

$$P(||S_n - F||^2 \ge \eta^2) \le \sum_{i=1}^k P\left(\left|\sqrt{\frac{n_i}{n}} - \sqrt{p_i}\right| \ge \frac{\eta}{\sqrt{k}}\right).$$

Since for each i the random variable n_i is distributed according to $B(n, p_i)$, we have from Theorems 3 and 4

$$P\!\!\left(\sqrt{rac{n_i}{n}}\!-\!\sqrt{p_i}\!\geq\!\!rac{\gamma}{\sqrt{k}}\!
ight)\!\!<\!\exp\left(-rac{2n\eta^2}{k}
ight)$$
 ,

$$P\!\!\left(\sqrt{\frac{n_i}{n}} - \sqrt{p_i} \le -\frac{\eta}{\sqrt{k}}\right) < \exp\left(-\frac{n\eta^2}{k}\right)$$
,

whence the required inequality follows.

We shall compare our result (19) with that of Matusita and Motoo (18), that is, we shall ask which of

$$A = \frac{k^2 + k - 1}{(n\eta^2)^2}$$
 and $D = k(e^{-2n\eta^2/k} + e^{-n\eta^2/k})$

is better (smaller in value). If we put $A' = k^2/(n\eta^2)^2$, which is better than A, it holds identically

$$D = k(e^{-2/\sqrt{A'}} + e^{-1/\sqrt{A'}})$$
.

Now we mention two examples of the comparison of A and D: For A'=1/25

$$D = k(e^{-10} + e^{-5}) \le 1/25 = A' \le A$$
 if $k \le 5$,

and for A' = 1/100

$$D=k(e^{-20}+e^{-10})\leq 1/10=A'\leq A$$
 if $k\leq 220$.

Though the comparison depends on k, the number of classes, if A is around or below 0.01, then D is seen to be better than A in almost all practical cases.

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- estimation, Ann. Math. Stat., Vol. 26 (1955), pp. 631-640.
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ERRATA

These Annals Vol. IX, No. 3

- P. 204, in the determinant of the second member in formula (9): read " $1-\alpha_{nn}$ " instead of " $-\alpha_{nn}$ ".
- P. 207, 1st line: read "quantitative" instead of "quantitive".
- P. 211, the last line: read " $\cdots + a_{nk}^0 q_n + \frac{R'_k}{X'_k P_k^0}$ " instead of

"
$$\cdots + a_{nk}^0 q_n \frac{R_k'}{X_k' P_k^0}$$
".

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P. 33, Theorems $1\sim 4$: read "X" instead of "x".