

# ON THE Hoeffding's Combinatorial Central Limit Theorem

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## 1. Introduction

Let  $(X_1, X_2, \dots, X_n)$  be a permutation random variable of  $(1, 2, \dots, N)$ , each of which has probability  $1/N!$ , and  $\{c_{ij}^N (1 \leq i \leq N, 1 \leq j \leq N)\}$  a set of  $n^2$  real numbers. Then we shall consider the limit distribution of

$$(1) \quad S_N = \sum c_{iX_i}^N.$$

Many authors have studied this problem (see Wald-Wolfowitz [6], Nother [5], Madow [4], Hoeffding [2], Dwass [1])<sup>1)</sup>. Among them, Hoeffding established this form and proved that under the condition

$$(5) \quad \lim_{N \rightarrow \infty} \frac{\frac{1}{N} \sum d_{ij}^{Nr}}{\left(\frac{1}{N} \sum d_{ij}^{N^2}\right)^{r/2}} = 0 \quad r=3, 4, \dots$$

the distribution of  $S_N$  tends to the normal distribution in law, where

$$(3) \quad d_{ij}^N = c_{ij}^N - \frac{1}{N} \sum_i c_{ij}^N - \frac{1}{N} \sum_j c_{ij}^N + \frac{1}{N^2} \sum_{ij} c_{ij}^N.$$

In this paper we shall prove the following.

**THEOREM 1.** *Under a Lindeberg type condition*

$$(4) \quad \lim_{N \rightarrow \infty} \sum_{|d_{ij}^N/d^{N^2}| > \varepsilon} \left(\frac{d_{ij}^N}{d^{N^2}}\right)^2 = 0$$

*the distribution of  $S_N$  tends to the normal distribution in law, where*

$$d^{N^2} = \frac{1}{N} \sum d_{ij}^{N^2}.$$

## 2. Preliminaries

For simplicity we shall drop the index  $N$  for all coefficients,

<sup>1)</sup> Recently, Dwass has proved in his interesting paper the same result under the non-overlapping condition with that of Hoeffding's. But his condition is of different type from ours.

i.e., we put  $c_{ij}^N = c_{ij}$ ,  $S_N = S$ , etc.

Moreover, we can assume that

i) by adding a suitable constant to  $S$

$$(5) \quad \sum_j c_{ij} = 0, \quad \sum_i c_{ij} = 0 \quad (\text{there for } c_{ij} = d_{ij})$$

ii) by multiplying a suitable constant

$$(6) \quad \frac{1}{N} \sum c_{ij}^2 = 1,$$

iii) by reordering  $i$

$$(7) \quad q_1^2 \geq q_2^2 \geq \dots \geq q_N^2$$

where

$$q_i^2 = \frac{1}{N} \sum_j c_{ij}^2.$$

Then the following lemma is evident.

LEMMA 1. *We have*

$$(8) \quad \sum q_i^2 = 1$$

and

$$\sum_{i>n} q_i^2 \leq \frac{N-n}{N}.$$

Let  $\{y_{ij}\}$ ,  $\{Y_{ij}\}$  be a set of random variables ( $1 \leq i, j \leq N$ ) such that

$$(9) \quad y_{ij} = 1 \quad \text{if } X_i = j,$$

and

$$y_{ij} = 0 \quad \text{if } X_i \neq j,$$

$$(10) \quad Y_{n,j} = \sum_{i \leq n} y_{ij}.$$

Then we have

$$(11) \quad S = \sum c_{ij} y_{ij}.$$

and

LEMMA 2.

$$(12) \quad y_{ij} = 0 \quad \text{or } 1$$

$$\sum_i y_{ij} = 1, \quad \sum_j y_{ij} = 1$$

$$(13) \quad p_n\{y_{k,j}=1\} = \frac{1 - Y_{nj}}{N - n} \quad \text{for } k > n$$

$$p_n\{y_{k,j}=1, y_{k',j'}=1\} = 0 \quad \text{for } k > n, j \neq j'$$

$$p_n\{y_{k,j}=1, y_{k',j'}=1\} = \frac{(1 - Y_{nj})(1 - Y_{n'j'})}{(N - n)(N - n - 1)} \quad \text{for } k, k' > n, k \neq k', j \neq j'$$

etc.

where  $p_n\{\cdot\}$  means the conditional probability

$$p_n\{\cdot\} = p\{\cdot | y_{i,j}, i \leq n, j = 1, 2, \dots, N\}.$$

### 3. Proof of theorem 1

According to the above convention, Theorem 1 is stated as follows.

**THEOREM 1'.** *If*

$$(14) \quad \frac{1}{N} \sum_{|c_{ij}| > \varepsilon} c_{ij}^2 \rightarrow 0 \quad (N \rightarrow \infty),$$

*the distribution of  $S = \sum c_{ij}y_{ij}$  tends to the normal distribution with 0 mean and unit variance.*

Let

$$(15) \quad C_{nj} = \sum_{i \leq n} c_{ij}$$

$$(16) \quad T_n = \sum_j \left( \sum_{i \leq n} c_{ij}y_{ij} + \frac{C_{nj}Y_{nj}}{N - n} \right) \quad n < N,$$

and for an arbitrarily small  $\eta$  with  $0 < \eta < 1$ , take  $N_0$  such that

$$(17) \quad \eta < \frac{N - N_0}{N} \leq 2\eta^2.$$

Our proof of the theorem is carried out in the three steps, which consist of the proofs of the following three propositions.

**PROPOSITION 1.**

$$(18) \quad E(S - T_{N_0})^2 < 2\eta.$$

**PROPOSITION 2.**

$$(19) \quad |1 - \sigma_{N_0}^2| < 2\eta$$

$$(20) \quad \lim_{N \rightarrow \infty} \left| \sigma_{N_0}^2 - \frac{1}{N} \sum_{n=1}^{N_0} \sigma_n^2 \right| = 0 \quad \text{for fixed } \eta,$$

<sup>2)</sup> This is possible for sufficiently large  $N$ .

where

$$(21) \quad \begin{aligned} \bar{\sigma}_{N_0}^2 &= \frac{1}{N} \sum_j \left( \sum_{i \leq N_0} c_{ij}^2 + \frac{C_{N_0j}^2}{N - N_0} \right) \\ \sigma_n^2 &= \sum_j \left( c_{nj} + \frac{C_{nj}}{N - n} \right)^2. \end{aligned}$$

PROPOSITION 3. *In fixed  $\eta$  we have*

$$(22) \quad |E(e^{T_{N_0} i u}) - e^{-\frac{u^2}{2} \sum_{n=1}^{N_0} \sigma_n^2}| \rightarrow 0 \quad (N \rightarrow \infty)$$

where  $u$  is any (fixed) real number and  $i^2 = -1$ .

When these propositions are proved, it becomes clear that Theorem 1' holds, for  $\eta$  can be chosen arbitrarily small.

PROOF OF PROPOSITION 1.

$$S - T_{N_0} = \sum_j \left( \sum_{i > N_0} c_{ij} y_{ij} - \frac{C_{N_0j} Y_{N_0j}}{N - N_0} \right)$$

Let

$$\bar{C}_j = \sum_{i > N_0} \frac{c_{ij}}{N - N_0},$$

then

$$\bar{C}_j + \frac{C_{N_0j}}{N - N_0} = 0,$$

$$Y_{N_0j} + \sum_{i > N_0} y_{ij} = 1.$$

Therefore

$$S - T_{N_0} = \sum_j \sum_{i > N_0} (c_{ij} - \bar{C}_j) y_{ij},$$

and

$$\begin{aligned} E(S - T_{N_0})^2 &= \frac{1}{N - 1} \sum_j \sum_{i > N_0} (c_{ij} - \bar{C}_j)^2 \leq \frac{1}{N - 1} \sum_j \sum_{i > N_0} c_{ij}^2 \\ &= \frac{N}{N - 1} \sum_{i > N_0} q_i^2 \leq \frac{N - N_0}{N - 1} \leq 2\eta. \end{aligned}$$

PROOF OF PROPOSITION 2.

$$\begin{aligned} 1 - \sigma_{N_0}^2 &= \frac{1}{N} \sum_{i,j=1}^N c_{ij}^2 - \sigma_{N_0}^2 = \frac{1}{N} \sum_j \left( \sum_{i > N_0} c_{ij}^2 - \frac{C_{N_0j}^2}{N - N_0} \right) \\ &= \frac{1}{N} \sum_j \left( \sum_{i > N_0} c_{ij}^2 - \frac{\bar{C}_j^2}{N - N_0} \right) \end{aligned}$$

$$= \frac{1}{N} \sum_j \sum_{i>N_0} (c_{ij} - \bar{C})^2 .$$

Then

$$0 \leq 1 - \sigma_{N_0}^2 \leq \frac{1}{N} \sum_{i,j>N_0} c_{ij}^2 = \frac{1}{N} \sum_{i>N_0} q_i^2 \leq \frac{N - N_0}{N} \leq 2\eta .$$

Thus (19) is proved.

Before proving (20), we shall state the following remark.

Let  $\epsilon$  be any positive number, and

$$(23) \quad \begin{aligned} \tilde{c}_{ij} &= c_{ij} && \text{when } |c_{ij}| > \epsilon \\ \tilde{c}_{ij} &= 0 && \text{when } |c_{ij}| \leq \epsilon \end{aligned}$$

$$(24) \quad c'_i = c_{ij} - \tilde{c}_{ij} .$$

We can assume by (14)

$$(25) \quad \sum_{i,j=1}^N \tilde{c}_{ij}^2 < \epsilon^2 \quad \text{for sufficiently large } N .$$

Then we get:

LEMMA 3.

$$(26) \quad \begin{aligned} \frac{|C_{nj}|}{N} &\leq 2\epsilon \\ n &= 1, 2, \dots, N \end{aligned}$$

$$(27) \quad \sum_{j=1}^N \frac{C_{nj}^2}{N^2} \leq 1 .$$

PROOF.

$$\frac{|C_{nj}|}{N} \leq \frac{1}{N} \sum_{i=1}^N |c'_{ij}| + \frac{1}{N} \sum_{i=1}^N |\tilde{c}_{ij}| \leq \frac{1}{N} \sum_{i=1}^N \epsilon + \sqrt{\frac{1}{N} \sum_{i=1}^N \tilde{c}_{ij}^2} \leq 2\epsilon$$

$$\sum_j \frac{C_{nj}^2}{N^2} = \sum_j \left( \frac{1}{N} \sum_{i \leq n} c_{ij} \right)^2 \leq \sum_j \frac{1}{N} \sum_{i \leq n} c_{ij}^2 \leq \frac{1}{N} \sum_{i,j} c_{ij}^2 = 1$$

Now we can prove (20) in Proposition 2.

$$\begin{aligned} \bar{\sigma}_{N_0}^2 &= \frac{1}{N} \sum_j \left( \sum_{i=0}^{N_0} c_{ij}^2 + \frac{C_{N_0,j}^2}{N - N_0} \right) \\ &= \frac{1}{N} \sum_{n=1}^{N_0} \sum_{j=1}^N \left( c_{ij}^2 + \frac{C_{nj}^2}{N - n} - \frac{C_{n-1,j}^2}{N - n + 1} \right) \end{aligned}$$

$$= \frac{1}{N} \sum_{n=1}^{N_0} \sum_{j=1}^N \left\{ \left( c_{ij} + \frac{C_{ij}}{N-n} \right)^2 - \left( \frac{c_{nj}^2}{N-n} + \frac{2c_{ij}C_{nj}}{(N-n)(N-n+1)} + \frac{C_{nj}^2}{(N-n)^2(N-n+1)} \right) \right\}.$$

Then, since  $N-n \geq N-N_0 \geq N\gamma$ , we have

$$\left| \bar{\sigma}_{N_0}^2 - \frac{1}{N} \sum_{n=1}^{N_0} \sigma_n^2 \right| \leq \sum_{n=1}^{N_0} \sum_{j=1}^N \left( \frac{c_{nj}^2}{N-n} + \frac{(1+c_{nj}^2)|C_{nj}|}{(N-n)(N-n+1)} + \frac{C_{nj}^2}{(N-n)^2(N-n+1)} \right) \\ \leq \frac{1}{N\gamma} + \frac{1}{\gamma^2} \left( 1 + \frac{1}{N} \right) 2\epsilon + \frac{1}{N\gamma^3}.$$

This proves (20) for  $N$  may be arbitrarily large and  $\epsilon$  arbitrarily small.

**PROOF OF PROPOSITION 3.** Assuming (25), and therefore Lemma 3, we divide the proof in four parts.

$$(I) \quad \frac{1}{N} \sum_{n=1}^N \sigma_n^2 \quad \text{is bounded.}$$

This is easily seen by Proposition 2.

$$(II) \quad \sum_{n=0}^{N_0-1} \int_{|\Delta T_n| > 3\delta} (\Delta T_n)^2 dp \rightarrow 0 \quad (N \rightarrow \infty) \quad \text{for any } \delta > 0.$$

where

$$\Delta T_n = T_{n+1} - T_n$$

**PROOF.**

$$\Delta T_n = \sum_{j=1}^N \left( c_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right) y_{n+1,j} + \frac{1}{N-n-1} \left( c_{n+1,j} + \frac{C_{nj}}{N-n} \right) Y_{nj} \\ = u + v + w.$$

where

$$(28) \quad u = \sum_j \tilde{c}_{n+1,j} y_{n+1,j} \\ v = \sum_j \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right) y_{n+1,j} \\ w = \frac{1}{N-n-1} \sum_j \left( c_{n+1,j} + \frac{C_{nj}}{N-n} \right) Y_{nj}.$$

Then we have

$$(29) \quad \int_{|\Delta T_n| > 3\delta} (\Delta T_n)^2 dp$$

$$\begin{aligned}
 &\leq 3 \left( \int u^2 dp + \int w^2 dp + \int_{|u|>\delta} v^2 dp + \int_{|w|>\delta} v^2 dp + \int_{|v|>\delta} v^2 dp \right) \\
 (30) \quad &\int u^2 dp = E(\sum_j \tilde{c}_{n+1,j} y_{n+1,j})^2 = \frac{1}{N} \sum_j \tilde{c}_{n+1,j}^2
 \end{aligned}$$

$$\begin{aligned}
 &\int w^2 dp = \frac{1}{(N-n-1)^2} E \left\{ \sum_j \left( c_{n+1,j} + \frac{C_{nj}}{N-n} \right) Y_{nj} \right\}^2 \\
 (31) \quad &= \frac{1}{(N-n-1)^2} \frac{n(N-n)}{N(N-1)} \sum_j \left( c_{n+1,j} + \frac{C_{nj}}{N-n} \right)^2 \\
 &\leq \frac{2}{N^2 \gamma^2} \sum_j \left( c_{n+1,j}^2 + \frac{C_{nj}^2}{(N-n)^2} \right) \\
 &\leq \frac{2}{\gamma^2 N} q_{n+1}^2 + \frac{1}{\gamma^4 N^2} \quad (\text{by Lemma 3})
 \end{aligned}$$

For

$$\begin{aligned}
 |v| &= \left| \sum_j \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n} \right) y_{n+1,j} \right| \leq \sup_j \left| c_{n+1,j} + \frac{C_{n+1,j}}{N-n} \right| \\
 &\leq \varepsilon + \frac{2\varepsilon}{\gamma} = \varepsilon \left( 1 + \frac{2}{\gamma} \right) \quad (\text{by Lemma 3})
 \end{aligned}$$

and

$$\begin{aligned}
 \text{Prob. } (|u|>\delta) &\leq \frac{1}{\delta^2 N} \sum_j \tilde{c}_{n+1,j}^2 \\
 \text{Prob. } (|w|>\delta) &\leq \frac{2}{\delta^2 N} \left( \frac{q_{n+1}^2}{\gamma^2} + \frac{1}{\gamma^4 N} \right) \\
 &\quad (\text{by (30), (31) and Chebyshev's inequality}),
 \end{aligned}$$

$$(32) \quad \int_{|u|>\delta} v^2 dp + \int_{|w|>\delta} v^2 dp \leq \frac{1}{\delta^2 N} \varepsilon \left( 1 + \frac{2}{\gamma} \right) \left( \sum_j \tilde{c}_{n+1,j}^2 + \frac{2q_{n+1}^2}{\gamma^2} + \frac{1}{\gamma^4 N} \right)$$

Finally, we have

$$\begin{aligned}
 &\int_{|v|>\delta} v^2 dp \leq \frac{1}{\delta} \int |v|^3 dp \leq \frac{1}{\delta} E \left\{ \sum_j \left| c'_{n+1} + \frac{C_{n+1,j}}{N-n-1} \right| y_{n+1,j} \right\}^3 \\
 &= \frac{1}{\delta N} \sum_j \left| c'_{n+1} + \frac{C_{n+1,j}}{N-n-1} \right|^3 \\
 (33) \quad &\leq \frac{1}{\delta N} \sup_j \left| c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right| \sum_j \left( c'_{n+1} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \\
 &\leq \frac{2\varepsilon}{\delta} \left( 1 + \frac{2}{\gamma} \right) \left( q_{n+1}^2 + \frac{1}{\gamma^2 N} \right) \quad (\text{Lemma 3})
 \end{aligned}$$

By (29), (30), (31), (32) and (33) we obtain

$$\sum_{n=0}^{N_0-1} \int_{|\Delta T_n| > 3\delta} \Delta T_n^2 dp \leq 3 \left\{ \varepsilon + \frac{2}{N\gamma^2} + \frac{1}{N\gamma^4} + \frac{\varepsilon}{\delta^2} \left(1 + \frac{2}{\gamma}\right) \right. \\ \left. \cdot \left( \varepsilon^2 + \frac{2}{N\gamma^2} + \frac{1}{N\gamma^4} \right) + \frac{2\varepsilon}{\delta} \left(1 + \frac{2}{\gamma}\right) \left(1 + \frac{1}{\gamma^2}\right) \right\},$$

which proves (II).

$$(III) \quad E_n(\Delta T_n) = 0,$$

where

$$E_n(\dots) = E(\dots | y_{ij}, \quad j=1, 2, \dots, N, \quad i \leq n)$$

PROOF.

$$E_n(\Delta T_n) = E_n \left\{ \sum_j \left( c_{n+,j} + \frac{C_{n+,j}}{N-n-1} \right) y_{n+,j} \right\} \\ + \frac{1}{N-n-1} \sum_j \left( c_{n+,j} + \frac{C_{n_j}}{N-n} \right) Y_{n_j} \\ = \sum_j \left( c_{n+,j} + \frac{C_{n+,j}}{N-n-1} \right) \frac{1-Y_{n_j}}{N-n} + \frac{1}{N-n-1} \sum_j \left( c_{n+,j} \right. \\ \left. + \frac{C_{n_j}}{N-n} \right) Y_{n_j} = 0.$$

$$(IV) \quad \sum_{n=0}^{N_0-1} E \left| E_n(\Delta T_n)^2 - \frac{1}{N} \sigma_{n+1}^2 \right| \rightarrow 0 \quad (N \rightarrow \infty)$$

PROOF. By using the inequality

$$(34) \quad |(a+b)^2 - a^2| \leq pa^2 + \left(1 + \frac{1}{p}\right)b^2 \quad \text{where } p > 0,$$

we get for  $p = \frac{1}{\sqrt{N}}$

$$E \left\{ \left| E_n(\Delta T_n^2) - E_n \left\{ \sum_j \left( c_{n+,j} + \frac{C_{n+,j}}{N-n-1} \right) y_{n+,j} \right\}^2 \right| \right\} \\ \leq \frac{1}{\sqrt{N}} E \left\{ \sum_j \left( c_{n+,j} + \frac{C_{n+,j}}{N-n-1} \right) y_{n+,j} \right\}^2 \\ (35) \quad + \frac{2\sqrt{N}}{(N-n-1)^2} E \left\{ \sum_j \left( c_{n+,j} + \frac{C_{n_j}}{N-n} \right) Y_{n_j} \right\}^2$$



$$\begin{aligned} &\leq \frac{1}{\sqrt{N^3}} \sigma_{n+1}^2 + \frac{4}{\sqrt{N^3} \gamma^2} \sum_j \left\{ c_{n+1,j}^2 + \frac{C_{n,j}^2}{(N-n)^2} \right\} \\ &\leq \frac{1}{\sqrt{N^3}} \sigma_{n+1}^2 + \frac{4}{\sqrt{N} \gamma^2} \left( q_{n+1}^2 + \frac{1}{N \gamma^2} \right) \end{aligned}$$

and

$$\begin{aligned} (36) \quad E \left\{ \left| E_n \left\{ \sum_j \left( c_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right) y_{n+1,j} \right\}^2 - \frac{1}{N} \sigma_{n+1}^2 \right| \right\} \\ &= E \left\{ \left| \sum_j \left( c_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \frac{1-Y_{nj}}{N-n} - \frac{1}{N} \sigma_{n+1}^2 \right| \right\} \\ &= E \left\{ \left| \frac{1}{N-n} \sum_j \left( \frac{n}{N} - Y_{nj} \right) \left( c_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \right| \right\} \\ &\leq \frac{1}{N\gamma} \sum_{j: |c_{n+1,j}| < \varepsilon} \left( \tilde{c}_{n+1,j}^2 + \frac{C_{n+1,j}}{N-n-1} \right)^2 \\ &\quad + \frac{1}{N\gamma} E \left\{ \left| \sum_{j: |c_{n+1,j}| \geq \varepsilon} \left( \frac{n}{N} - Y_{nj} \right) \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \right| \right\} \end{aligned}$$

But, for  $j$ , with  $|c_{n+1,j}| > \varepsilon$  and

$$\frac{|C_{n+1,j}|}{N-n-1} \leq \frac{\varepsilon}{\gamma} \leq \frac{2|c_{n+1,j}|}{\gamma} = \frac{2|\tilde{c}_{n+1,j}|}{\gamma},$$

we have

$$(37) \quad \frac{1}{N\gamma} \sum_{j: |c_{n+1,j}| > \varepsilon} \left( \tilde{c}_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \leq \frac{1}{N\gamma} \sum_j \left( 1 + \frac{2}{\gamma} \right)^2 \tilde{c}_{n+1,j}^2.$$

On the other hand,

$$\begin{aligned} (38) \quad &\frac{1}{N\gamma} E \left\{ \left| \sum_{j: |c_{n+1,j}| \leq \varepsilon} \left( \frac{n}{N} - Y_{nj} \right) \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \right| \right\} \\ &\leq \frac{1}{N\gamma} \sqrt{E \left\{ \sum_{j: |c_{n+1,j}| \leq \varepsilon} \left( \frac{n}{N} - Y_{nj} \right) \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2 \right\}^2} \\ &\leq \frac{1}{N\gamma} \sqrt{\sum_j \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^4 \frac{n(N-n)}{N(N-1)}} \\ &\leq \frac{1}{N\gamma} \sup \left| c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right| \sqrt{\sum_j \left( c'_{n+1,j} + \frac{C_{n+1,j}}{N-n-1} \right)^2} \\ &\leq \frac{\varepsilon}{N\gamma} \left( 1 + \frac{2}{\gamma} \right) \sigma_{n+1} \leq \frac{\varepsilon}{2N\gamma} \left( 1 + \frac{2}{\gamma} \right) (1 + \sigma_{n+1}^2). \end{aligned}$$

From (35), (36), (37) and (38) it follows that

$$\begin{aligned} & \sum_{n=0}^{N_0-1} E \left\{ \left| E_n (\Delta T_n)^2 - \frac{1}{N} \sigma_{n+1}^2 \right| \right\} \\ & \leq \frac{1}{\sqrt{N}} \left\{ \frac{1}{N} \sum_{n=0}^{N_0-1} \sigma_{n+1}^2 + \frac{4}{\eta^2} \left( 1 + \frac{1}{\eta} \right) \right\} + \frac{1}{\eta} \left( 1 + \frac{2}{\eta} \right)^2 \varepsilon^2 \\ & \quad + \frac{\varepsilon}{2\eta} \left( 1 + \frac{2}{\eta} \right) \left( 1 + \frac{1}{N} + \sum_{n=0}^{N_0-1} \sigma_{n+1}^2 \right) \end{aligned}$$

which proves IV.

Noting that  $T_n$  is  $\{y_{ij}, i \leq n, j=1, 2, \dots, N\}$  measurable, (I), (II), (III) and (IV) altogether show that the condition of Loève's theorem is satisfied for  $T_{N_0} = \Delta T_0 + \Delta T_1 + \dots + \Delta T_{N_0-1}$  (Theorem C: page 337 in [3]), and Proposition 3 is proved.

#### 4. Relation with Hoeffding's condition

$$\text{THEOREM 2. If } \frac{\frac{1}{N} \sum |d_{ij}|^r}{\left( \frac{1}{N} \sum d_{ij}^2 \right)^{r/2}} \rightarrow 0 \quad (N \rightarrow \infty) \quad \text{for } r > 2,$$

then the condition of Theorem 1 is satisfied.

PROOF. Normalizing  $c_{ij}$ , it is sufficient to prove that, if

$$\frac{1}{N} \sum |c_{ij}|^r \rightarrow 0, \quad \text{then} \quad \frac{1}{N} \sum_{|c_{ij}| > \varepsilon} c_{ij}^2 \rightarrow 0.$$

This is easily seen from the fact:

$$\frac{1}{N} \sum_{|c_{ij}| > \varepsilon} c_{ij}^2 \leq \frac{1}{\varepsilon^{r-1} N} \sum |c_{ij}|^r.$$

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#### REFERENCES

- [1] M. Dwass, On the asymptotic normality of some statistics used in non parametric tests. *Ann. Math. Stat.* vol. 26 (1955).
- [2] W. Hoeffding, Combinatorial central limit theorem. *Ann. Math. Stat.* vol. 22 (1951).
- [3] M. Loève, *Probability theory*. New York (1955).
- [4] W.G. Madow, On the limiting distribution of estimates based on samples from finite universe. *Ann. Math. Stat.* vol. 20 (1949).
- [5] G.E. Noether, On the theorem of Wald and Wolfowitz. *Ann. Math. Stat.* vol. 20 (1949).
- [6] A. Wald and J. Wolfowitz, Statistical tests based on permutations of the observations. *Ann. Math. Stat.* vol. 15 (1944).