# 再生核による指数分布族の構成とその統計的推定への応用

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2008年6月12日 統計数学セミナー

## Outline

- Introduction
- Reproducing kernel Hilbert space and positive definite kernel
- Reproducing kernel exponential manifold (RKEM)
- Application 1: Pseudo-maximum likelihood estimation with RKEM
- Application 2: Statistical asymptotic theory of singular models
- Conclusion

## Introduction

## Maximal Exponential Manifold

- Maximal exponential manifold (Pistone &Sempi 95)
  - Idea: a Banach manifold is defined so that the cumulant generating function is well-defined on a neighborhood of each probability density.  $(\Omega, \mathcal{B}, \mu)$ : probability space

$$f_u = \exp(u - \Psi_f(u))f, \qquad \Psi_f(u) = \log E_f[e^u] < \infty$$

□ Orlicz space L<sub>cosh-1</sub>(f)

$$\begin{split} L_{\cosh-1}(f) &= \left\{ u \mid \exists \alpha > 0 \text{ s.t. } E_f[\cosh(\alpha u)] < \infty \right\} \\ &= \left\{ u \mid \exists \alpha > 0 \text{ s.t. } E_f[e^{\alpha u}] < \infty \text{ and } E_f[e^{-\alpha u}] < \infty \right\} \end{split}$$

This space is (perhaps) the most general to guarantee the finiteness of the cumulant generating functions around a point.

## Estimation with Data

#### Estimation with a finite sample

 A finite dimensional exponential family is suitable for the maximum likelihood estimation (MLE) with a finite sample.

$$X_1,...,X_n$$
: i.i.d.  $\sim f_0 \mu$   $X_n = (X_1,...,X_n)$ 

MLE:  $\theta$  that maximizes

$$\ell_n(\theta; \mathbf{X}_n) = \frac{1}{n} \sum_{i=1}^n \left\{ \sum_{a=1}^m \theta^a u_a(\mathbf{X}_i) - \Psi(\theta) \right\}$$

Is MLE extendable to the maximal exponential manifold?

$$\ell_n(u; X_n) = \frac{1}{n} \sum_{i=1}^n \{ u(X_i) - \Psi_f(u) \}$$

 $\Rightarrow$  But, the function value  $u(X_i)$  is not a continuous functional on u in the exponential manifold.

A small change of u may cause a very different likelihood.

## Reproducing Kernel Hilbert Space and Positive Definite Kernel

## Reproducing kernel Hilbert space

- Reproducing kernel Hilbert space (RKHS)
  - $\ \square$   $\Omega$ : set. A Hilbert space  $\mathcal H$  consisting of functions on  $\Omega$  is called a (real-valued) reproducing kernel Hilbert space (RKHS) if the evaluation functional

$$e_x: \mathcal{H} \to \mathbf{R}, \quad f \mapsto f(x)$$

is continuous for each  $x \in \Omega$ .

□ A Hilbert space  $\mathcal{H}$  consisting of functions on  $\Omega$  is a RKHS if and only if there exists  $k(\cdot, x) \in \mathcal{H}$  (reproducing kernel) for each  $x \in \Omega$  s.t.

$$\langle k(\cdot, x), f \rangle_{\mathcal{H}} = f(x) \quad \forall f \in \mathcal{H}, \ x \in \Omega.$$
 [reproducing property] (by Riesz's lemma)

### Positive definite kernel and RKHS

#### Positive definite kernel

A symmetric function  $k: \Omega \times \Omega \to \mathbb{R}$  is said to be *positive definite*, if for any  $n \in \mathbb{N}$  and  $x_1, \dots, x_n \in \Omega$ , the matrix  $(k(x_i, x_j))$  (Gram matrix) is positive semidefinite, i.e.

$$\sum_{i,j=1}^{n} c_{i} c_{j} k(x_{i}, x_{j}) \ge 0, \quad (\text{for any } c_{1}, \dots, c_{n} \in \mathbf{R}).$$

A reproducing kernel is positive definite.

#### Positive definite kernel and RKHS

Theorem (Moore-Aronszajn)

If  $k: \Omega \times \Omega \to \mathbb{R}$  is positive definite, there uniquely exists a RKHS  $\mathcal{H}_k$  consisting of functions on  $\Omega$  such that

- (1) The linear hull of  $\{k(\cdot, x): \Omega \to \mathbb{R} \mid x \in \Omega\}$  is dense in  $\mathcal{H}_k$ .
- (2)  $k(\cdot, x)$  is a reproducing kernel of  $\mathcal{H}_k$ .

## Example of positive definite kernel

 $\Box$  Euclidean inner product on  $\mathbf{R}^m$ 

$$k(x, y) = x^T y$$

Polynomial kernel on  $\mathbf{R}^m$ 

$$k(x, y) = (x^T y + c)^d$$
  $(c \ge 0, d \in \mathbb{N})$   
 $\mathcal{H}_k = \{\text{polyn. deg } \le d\}$ 

Gaussian kernel on R<sup>m</sup>

$$k(x, y) = \exp(-\|x - y\|^2 / \sigma^2)$$
 dim  $\mathcal{H}_k = \infty$ 

Laplacian kernel on [0 1]

$$k(x,y) = \exp(-|x-y|)$$

$$\mathcal{H}_k = H^1(0,1) = \left\{ u \in L^2[0,1] \mid \exists u' \in L^2[0,1] \right\} \quad \text{(Sobolev space)}$$

$$\|u\|_{\mathcal{H}_k}^2 = \frac{1}{2} \left\{ u(0)^2 + u(1)^2 \right\} + \frac{1}{2} \int_0^1 \left\{ u(x)^2 + u'(x)^2 \right\} dx$$

## Some properties of RKHS

$$\text{For } f = \sum_{i=1}^n a_i k(\cdot\,,x_i), \ \ g = \sum_{j=1}^m b_j k(\cdot\,,y_j), \\ \left\langle f\,,g\right\rangle_{\mathcal{H}} = \sum_{ij} a_i b_j k(x_i,y_j).$$
 In particular, 
$$\left\|k(\cdot\,,x)\right\|_{\mathcal{H}} = \sqrt{k(x,x)}.$$

□ If a pos. def. kernel k is of class  $C^d$ , so are all the functions in  $\mathcal{H}_k$ . ∵) for  $C^0$ ,

$$|f(x) - f(y)| = |\langle k(\cdot, x) - k(\cdot, y), f \rangle| \le ||k(\cdot, x) - k(\cdot, y)||_{\mathcal{H}_k} ||f||_{\mathcal{H}_k}$$
$$||k(\cdot, x) - k(\cdot, y)||_{\mathcal{H}_k}^2 = k(x, x) - 2k(x, y) + k(y, y)$$

## Reproducing Kernel Exponential Manifold

## Exponential Manifold by RKHS

#### Definitions

Ω: topological space. μ: Borel measure on Ω s.t. support of μ = Ω.

k: continuous pos. def. kernel on  $\Omega$  such that  $\mathcal{H}_k$  contains 1 (constants).

$$\begin{split} M_{\mu}(k) &\coloneqq \left\{ f: \Omega \to \mathbf{R} \, / \, f: \text{continuous}, \, f(x) > 0 \, (\forall x \in \Omega), \, \int f d\mu = 1, \\ &\exists \, \delta > 0, \, \int e^{\delta \sqrt{k(x,x)}} f(x) d\mu(x) < \infty \right\} \end{split}$$

 $M_{ii}(k)$  is provided with a Hilbert manifold structure.

Note: If 
$$\|u\| < \delta$$
,  $E_f[e^{u(X)}] = E_f[e^{\langle u, k(\cdot, X) \rangle}] \le E_f[e^{\|u\| \sqrt{k(X, X)}}] < \infty$ . If  $k$  is bounded, the condition  $E_f[e^{\delta \sqrt{k(x, x)}}] < \infty$  is not needed.

Tangent space

$$T_f := \{ u \in \mathcal{H}_k \mid E_f[u(X)] = 0 \}$$
 closed subspace of  $\mathcal{H}_k$ 

## Exponential Manifold by RKHS (cont'd)

#### Local coordinate

For 
$$f \in M_{\mu}(k)$$
,  $W_f := \left\{ u \in T_f \mid \exists \delta > 0, \ E_f[e^{u(X) + \delta \sqrt{k(X,X)}}] < \infty \right\} \subset T_f$ 

Then, for any  $u \in W_f$ 

$$\begin{split} f_u &\coloneqq \exp(u - \Psi_f(u))f &\in M_{\mu}(k). \\ &\left( :: E_{f_u}[e^{\delta \sqrt{k(X,X)}}] = E_f[e^{\delta \sqrt{k(X,X)}}e^{u(X) - \Psi_f(u)}] < \infty. \right) \end{split}$$

#### Define

#### Lemma

- (1)  $W_f$  is an open subset of  $T_f$ .
- (2)  $g \in \mathcal{E}_f \iff \mathcal{E}_f = \mathcal{E}_g$ .

## Exponential Manifold by RKHS (cont'd)

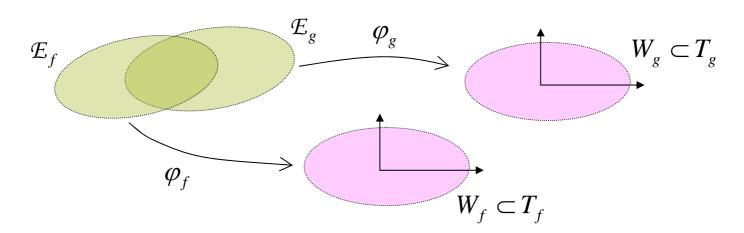
#### Reproducing Kernel Exponential Manifold (RKEM)

Theorem.

The system  $\{(\mathcal{E}_f, \varphi_f)\}_{f \in M_{\mu}(k)}$  is a  $C^{\infty}$ -atlas of  $M_{\mu}(k)$ , i.e.,

(1) 
$$\mathcal{E}_f \cap \mathcal{E}_g \neq \phi \implies \varphi_f(\mathcal{E}_f \cap \mathcal{E}_g)$$
 is open in  $T_f$ 

(2) 
$$\mathcal{E}_f \cap \mathcal{E}_g \neq \phi$$
  
 $\Rightarrow \varphi_g \circ \varphi_f^{-1}|_{\varphi_f(\mathcal{E}_f \cap \mathcal{E}_g)} : \varphi_f(\mathcal{E}_f \cap \mathcal{E}_g) \rightarrow \varphi_g(\mathcal{E}_f \cap \mathcal{E}_g) \text{ is a } C^{\infty}\text{-map.}$ 



## Exponential Manifold by RKHS (cont'd)

- Sketch of the proof
  - (2). Let  $h \in \mathcal{E}_f \cap \mathcal{E}_g$  and  $u = \varphi_f(h)$ , i.e.,  $h = \exp(u \Psi_f(u)) f$ . Then,

$$\varphi_g \circ \varphi_f^{-1}(u) = \varphi_g(h) = \log \frac{\exp(u - \Psi_f(u))f}{g} - E_g \left[ \log \frac{\exp(u - \Psi_f(u))f}{g} \right]$$

$$= u + \log \frac{f}{g} - E_g \left[ u + \log \frac{f}{g} \right]$$

 $u \mapsto E_g[u]$  is affine on  $W_f$ , thus, of  $C^{\infty}$ .

 $\square$  A structure of  $C^{\infty}$  Hilbert manifold is defined on  $M_{\mu}(k)$ .

## Properties of RKEM

#### Properties of RKEM as a Hilbert manifold

- □ The Hilbert manifold  $M_{\mu}(k)$  depends on the choice of a kernel k.
- □ The tangent space at  $f \in M_{\mu}(k)$  is identified with  $T_f$ , which is codimension one in  $\mathcal{H}_k$ .
- □ Log-likelihood  $u(X) \Psi_f(u) + \log f(X)$  is continuous on  $M_{\mu}(k)$ .
- □ Sufficient statistics = k(x,y)

$$\exp(u(x) - \Psi_f(u)) = \exp(\langle u, k(\cdot, x) \rangle - \Psi_f(u))$$

c.f. finite dimensional case:  $\exp(\theta \cdot s(x) - \Psi(\theta))$ 

## Examples of RKEM

- RKEM includes any finite dimensional exponential family.
- $\Omega = \mathbf{R}, \ \mu = N(0,1)$   $k(x,y) = (xy+1)^2. \quad \Rightarrow \quad \mathcal{H}_k = \{\text{polyn. deg} \leq 2\}$   $M_{\mu}(k) = \{N(m, \sigma) \mid m \in \mathbf{R}, \ \sigma > 0 \} \text{ : exponential family of normal distributions.}$

$$\Omega = [0 \ 1], \ \mu = \text{Unif}[0,1]$$

$$k(x,y) = \exp(-|x-y|) \quad \Rightarrow \quad \mathcal{H}_k = H^1(0,1).$$

$$M_{\mu}(k) = \left\{ f : [0,1] \to \mathbf{R} \mid f : \text{continuous}, f > 0, \int_0^1 f(x) dx = 1 \right\}$$

$$\therefore \quad k(x,x) = 1 \quad \Rightarrow \quad E_f[e^{\delta \sqrt{k(X,X)}}] < \infty.$$

#### Moments in RKEM

□ Mean parameter: for any  $f \in M_{\mu}(k)$ , there uniquely exists  $m_f \in \mathcal{H}_k$  such that

$$E_f[u(X)] = \left\langle u, m_f \right\rangle_{\mathcal{H}_k} \quad \text{for all} \quad u \in \mathcal{H}_k.$$
 
$$m_f(y) = E_f[k(y, X)] \quad \text{: mean of the sufficient statistics } k(\cdot, x)$$

□ Covariance operator: for any  $f \in M_{\mu}(k)$ , there uniquely exists an operator  $\Sigma_f$  on  $\mathcal{H}_k$  such that

$$\langle v, \Sigma_f u \rangle_{\mathcal{H}_k} = \operatorname{Cov}_f[v(X), u(X)]$$
 for all  $u, v \in \mathcal{H}_k$ .

Derivatives of cumulant generating function

For 
$$g = e^{u - \Psi_f(u)} f$$
  $(u \in T_f)$  and  $v_1, v_2 \in T_f$ ,

the derivatives of  $\Psi_f$  at u in the direction of  $v_1$  (and  $v_2$ ) are given by

$$D_{u}\Psi_{f}(v_{1}) = E_{g}[v_{1}(X)] = \langle v_{1}, m_{g} \rangle_{\mathcal{H}_{k}}$$

$$D_{u}\Psi_{f}(v_{1}, v_{2}) = \operatorname{Cov}_{g}[v_{1}(X), v_{2}(X)] = \langle v_{2}, \Sigma_{g}v_{1} \rangle_{\mathcal{H}_{k}}$$
18

## Pseudo-Maximum Likelihood Estimation with RKEM

#### MLE with RKEM

 $m_{u}=\hat{m}^{(n)}$ 

#### Likelihood equation

$$\mathcal{E} \text{: connected component of } M_{\mu}(k). \qquad f_0 \in \mathcal{E} \text{: fixed.}$$
 
$$\mathcal{E} = \{f \in M_{\mu}(k) \mid \exists u \in T_{f_0}, \ f = f_u = \exp(u - \Psi_{f_0}(u)) f_0 \}$$
 
$$f_* = f_{u_*} \text{: true p.d.f. to give i.i.d. samle } X_1, \dots, X_n \sim f_* \mu.$$
 
$$\text{MLE in } \mathcal{E} \text{: } \max_{f \in E} \sum_{i=1}^n \log f(X_i) = \max_{u \in W_0} \sum_{i=1}^n u(X_i) - n \Psi_0(u)$$
 
$$\Longrightarrow \max_{u \in W_0} \left\langle \hat{m}^{(n)}, u \right\rangle - \Psi_0(u) \quad \text{where} \quad \hat{m}^{(n)} = \frac{1}{n} \sum_{i=1}^n k(\cdot, X_i)$$
 
$$\Longrightarrow \left\langle m_u, v \right\rangle_{\mathcal{H}_k} = \left\langle \hat{m}^{(n)}, v \right\rangle_{\mathcal{H}_k} \quad \text{for all} \quad v \in \mathcal{H}_k. \quad \text{-- Moment matching}$$

The ML mean parameter should be  $\hat{m}^{(n)}$ . What is the corresponding p.d.f. element (or natural parameter u) in the RKEM?

## MLE is impossible with RKEM

- Rigorous MLE is impossible for RKEM in general.
  - $\Box$  The mean parameter  $m_f$  uniquely determines the probability for a certain class of kernels (characteristic kernel, Fukumizu et al. 08).

{probability measure on  $\Omega$ }  $\to \mathcal{H}_k$ ,  $P \mapsto m_P$  is injective.

e.g.) Gaussian kernel 
$$k(x, y) = \exp(-\|x - y\|^2/\sigma^2)$$

- Moment matching with the empirical distribution is impossible.
- c.f. For a finite dimensional exponential family, the moments are given by only the finite number of sufficient statistics.
- □ Mean parameter is not a coordinate in general (Pistone & Rogatin 99)  $u \mapsto m_u = D\Psi_0(u)$  does not have a continuous inverse, because  $D^2\Psi_0(u,v) = \langle v, \Sigma_f u \rangle$  and  $\Sigma_f$  can have arbitrary small eigenvalues.

## Asymptotics of mean parameter

□ Theorem ( $\sqrt{n}$ -consistency of the ML mean parameter) ( $\Omega$ ,  $\mathcal{B}$ , P): probability space.

*k*: positive definite kernel on  $\Omega$  s.t.  $E_p[k(X,X)] < \infty$ .

Proof) 
$$E\|\hat{m}^{(n)} - m_P\|_{\mathcal{H}_k}^2 = \frac{1}{n} \{E[k(X,X)] - E[k(X,\widetilde{X})]\},$$
 where  $\widetilde{X}$  is an independent copy of  $X$ .  $q.e.d.$ 

Theorem implies the uniform law of large numbers;

$$\sup_{f \in \mathcal{H}_k, \, \|f\|_{\mathcal{H}_k} \le 1} \left| \frac{1}{n} \sum_{i=1}^n f(X_i) - E_P[f(X)] \right| = O_p(1/\sqrt{n}).$$

#### Pseudo-MLE with RKEM

#### Pseudo-MLE by regularization

 $\{\mathcal{H}^{(\ell)}\}_{\ell=1}^{\infty} \text{: sequence of finite dim. subspaces in } \mathcal{H}_{k} \text{ such that } \mathcal{H}^{(\ell)} \subset \mathcal{H}^{(\ell+1)} \\ \text{and the inclusions } \mathcal{H}^{(\ell)} \to \mathcal{H}^{(\ell+1)}, \ \mathcal{H}^{(\ell)} \to \mathcal{H}_{k} \text{ are continuous.}$ 

$$T_f^{(\ell)} = T_f \cap \mathcal{H}^{(\ell)}, \quad W_f^{(\ell)} = W_f \cap \mathcal{H}^{(\ell)}$$

Pseudo-MLE: 
$$\hat{u}^{(\ell)} := \underset{u \in W_f^{(\ell)}}{\operatorname{arg max}} \left[ \left\langle \hat{m}^{(n)}, u \right\rangle - \Psi_0(u) \right]$$

Assumptions

(A-1) For 
$$u \in W_f$$
, let  $u_*^{(\ell)} \coloneqq \min_{w \in W_f^{(\ell)}} KL(f_u \parallel f_w)$ . Then, 
$$\left\| u - u_*^{(\ell)} \right\|_{\mathcal{H}_k} \to 0 \ (\ell \to \infty).$$
 (approximation)

$$(\mathsf{A-2}) \ \exists \, \delta > 0, \ \exists (\ell_n)_{n=1}^{\infty} \subset \mathbf{N} \ \text{ s.t.}$$
 
$$\lambda^{(\ell)} := \inf_{u \in \mathcal{H}_k, ||u-u_*|| \leq \delta} \inf_{v \in T_{f_u}^{(\ell)}, ||v|| \leq 1} \left\langle v, \Sigma_{f_u} v \right\rangle \quad \text{satisfies} \quad \lim_{n \to \infty} \sqrt{n} \ \lambda^{(\ell_n)} = +\infty.$$
 (stability) 
$$_{23}$$

## Consistency of Pseudo-MLE

Theorem (Fukumizu, IGAIA2005)

$$\hat{f}_n = \exp(\hat{u}^{(n)} - \Psi_0(\hat{u}^{(n)})) f_0.$$

Under the assumptions (A-1) and (A-2),

$$KL(f_* \parallel \hat{f}_n) \to 0 \quad (n \to \infty)$$
 in probability.

*Sketch of the Proof)* 

$$KL(f_* \parallel \hat{f}_n) = KL(f_* \parallel f_{u_*^{(\ell_n)}}) + \Psi_0(\hat{u}^{(n)}) - \Psi_0(u_*^{(\ell_n)}) - E_{f_*}[\hat{u}^{(n)} - u_*^{(\ell_n)}]$$

- (i)  $KL(f_* || f_{u_*^{(\ell_n)}}) \to 0$  by (A-1).
- (ii) For the rest terms, it suffices to show  $\Pr(\|\hat{u}^{(n)} u_*^{(\ell_n)}\| \ge \varepsilon) \to 0$  for an arbitrary  $\varepsilon > 0$ .

$$\Pr\left(\left\|\hat{u}^{(\ell_n)} - u_*^{(\ell_n)}\right\| \ge \varepsilon\right) \\
\leq \Pr\left(\sup_{u \in W^{(\ell_n)}, \|u - u_*^{(\ell_n)}\| \ge \varepsilon} \left\langle u, \hat{m}^{(n)} \right\rangle - \Psi_0(u) \ge \left\langle u_*^{(\ell_n)}, \hat{m}^{(n)} \right\rangle - \Psi_0(u_*^{(\ell_n)})\right) \equiv \mathbf{P}_n$$

## Consistency of Pseudo-MLE (cont'd)

For any  $u \in W^{(\ell_n)}$ ,

$$\begin{aligned}
& \left\langle u, \hat{m}^{(n)} \right\rangle - \left\langle u_{*}^{(\ell_{n})}, \hat{m}^{(n)} \right\rangle - \Psi_{0}(u) + \Psi_{0}(u_{*}^{(\ell_{n})}) \\
&= \left\langle u - u_{*}^{(\ell_{n})}, \hat{m}^{(n)} - m_{f_{*}} \right\rangle - \left\langle u - u_{*}^{(\ell_{n})}, m_{f_{*}} - m_{u_{*}^{(\ell_{n})}} \right\rangle \\
&+ \left\langle u - u_{*}^{(\ell_{n})}, m_{u_{*}^{(\ell_{n})}} \right\rangle - \Psi_{0}(u) + \Psi_{0}(u_{*}^{(\ell_{n})}) \\
&= \left\langle u - u_{*}^{(\ell_{n})}, \hat{m}^{(n)} - m_{f_{*}} \right\rangle - \left\langle \Psi_{0}(u) - \Psi_{0}(u_{*}^{(\ell_{n})}) - \left\langle u - u_{*}^{(\ell_{n})}, m_{u_{*}^{(\ell_{n})}} \right\rangle \right\} \quad (*)
\end{aligned}$$

By convexity of  $\Psi_0$ , the supremum can be considered in a neighborhood. By (A-2),

$$(*) \leq ||u - u_{*}^{(\ell_{n})}|| ||\hat{m}^{(n)} - m_{f_{*}}|| - \frac{1}{2}\lambda^{(\ell_{n})}||u - u_{*}^{(\ell_{n})}||^{2}$$

$$\Rightarrow \mathbf{P}_{n} \leq \Pr(||\hat{m}^{(n)} - m_{f_{*}}|| \geq \frac{1}{2}\lambda^{(\ell_{n})}\varepsilon) \to 0.$$

$$q.e.d.$$

## Remarks on pseudo-MLE

#### Remarks

- □ If  $\mathcal{H}_k$  is finite dimensional, the Pseudo-MLE is equal to the ordinary MLE.
- Another way of regularization Tikhonov regularization.
   (Canu&Smola06)

## Statistical Asymptotic Theory of Singular Models

## Singular Submodel of exponential family

#### Standard asymptotic theory

Statistical model  $\{f(x;\theta) | \theta \in \Theta\}$  on a measure space  $(\Omega, \mathcal{B}, \mu)$ .

Θ: (finite dimensional) manifold.

"True" density:  $f_0(x) = f(x; \theta_0)$   $(\theta_0 \in \Theta)$   $X_1, \dots, X_n$ : i.i.d.  $\sim f_0 \mu$ 

Maximum likelihood estimator (MLE)

$$\hat{\theta}_n = \arg\max_{\theta \in \Theta} \sum_{i=1}^n \log f(X_i; \theta)$$

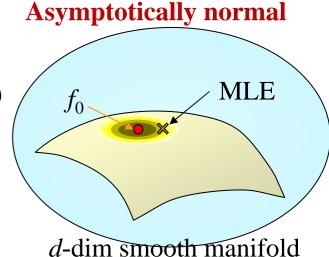
Under some regularity conditions,

$$\sqrt{n}(\hat{\theta}_n - \theta_0) \Rightarrow N(0, I(\theta_0)^{-1}) \text{ in law } (n \to \infty)$$

Likelihood ratio

$$2\ell_n(\hat{\theta}_n) = 2\sum_{i=1}^n \log \frac{f(X_i; \hat{\theta}_n)}{f(X_i; \theta_0)} \implies \chi_d^2$$

in law  $(n \to \infty)$ 



28

## Singular Submodel of exponential family (cont'd)

Singular submodel in ordinary exponential family

Finite dimensional exponential family 
$$M$$
:  $f(x;\theta) = \exp(\theta^T u(x) - \Psi(\theta))$   
Submodel  $S = \{f(x;\theta) \in M \mid \theta \in \Theta_S\}$   $(\theta \in \Theta)$ 

#### Tangent cone:

$$C_{f_0}S = \{ \xi^T u(x) \in T_{f_0}M \mid \exists \{\theta_n\} \subset \Theta_S, \exists \lambda_n > 0 \text{ s.t. } \lambda_n(\theta_n - \theta_0) \to \xi \quad (n \to \infty) \}$$

Under some regularity conditions,

$$\ell_{n}(\hat{\theta}_{n}) = \sum_{i=1}^{n} \log \frac{f(X_{i}; \hat{\theta}_{n})}{f(X_{i}; \theta_{0})}$$

$$= \frac{1}{2} \sup_{\xi^{T} u \in C_{f_{0}}S, \ E_{f_{0}} \mid \xi^{T} u \mid^{2} = 1} \underbrace{\left\{ \xi^{T} \left( \frac{1}{n} \sum_{i=1}^{n} u(X_{i}) \right) \right\}^{2} + o_{p}(1)}_{\text{projection of empirical} \ (n \to \infty) \text{ mean parameter}}$$

More explicit formula can be derived in some cases.

## Singular submodel in RKEM

- Submodel of an infinite dimensional exponential family
  - The tangent cone of a model defined by a finite number of parameters may not be in a finite dimensional space.
  - Interesting parametric models are
    - not embeddable into a finite dimensional exponential family,
    - but can be embedded into an infinite dimensional RKEM.

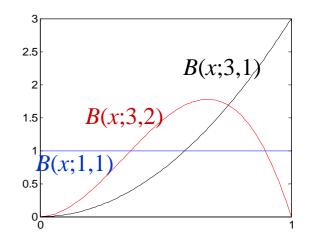
#### Mixture of Beta distributions

Mixture of Beta distributions (on [0,1])

S: 
$$f(x;\alpha,\beta) = \alpha B(x;\beta,1) + (1-\alpha)B(x;1,1)$$

$$B(x; \beta, \gamma) = \frac{\Gamma(\beta + \gamma)}{\Gamma(\beta)\Gamma(\gamma)} x^{\beta - 1} (1 - x)^{\gamma - 1}$$

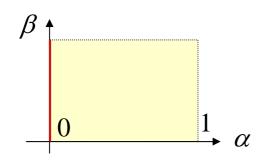
: Beta distribution

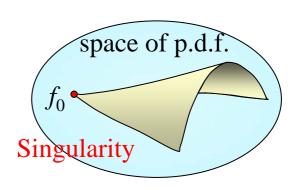


□ Singularity at  $f_0(x) = f(x;0,\beta) = B(x;1,1)$ 

If  $\alpha = 0$ ,  $\beta$  is not identifiable.

→ singularity in the space of probability densities.





## Mixture of Beta distributions (cont'd)

□  $\mathcal{H}_k$  = Sobolev space  $H^1(0,1)$  defined by  $k(x,y) = \exp(-|x-y|)$ .

Fact: 
$$\log f(x; \alpha, \beta) \in H^1(0,1)$$
 for  $0 \le \alpha < 1, \beta > 3/2$ .

□ RKEM with the Sobolev space Connected component including  $f_0 = \text{Unif}[0,1]$ :

$$\mathcal{E}_{f_0} = \{ g \in M_{\mu}(k) \mid \exists u \in T_{f_0}, g = \exp(u - \Psi_{f_0}(u)) f_0 \}$$

$$\begin{aligned} u_{\alpha,\beta}(x) &\coloneqq \log f(x;\alpha,\beta) - E_{f_0}[\log f(x;\alpha,\beta)] \in T_{f_0} \\ S &= \{ f(\cdot;\alpha,\beta) = \exp(u_{\alpha,\beta} - \Psi_f(u_{\alpha,\beta})) f_0 \mid 0 \le \alpha < 1, \beta > 3/2 \} \end{aligned}$$

- $\Rightarrow$   $f_0$  is a singularity of S.
- $\Box$  Tangent cone at  $f_0$  is not finite dimensional.

$$\frac{\log f(\cdot;\alpha,\beta)}{\alpha} \to w_{\beta} := \beta x^{\beta-1} - 1 \quad (\alpha \downarrow 0) \text{ in } H^1(0,1)$$

## Asymptotics on singular submodel

#### General theory of singular submodel

$$M_{\mu}(k)$$
: RKEM.  $f \in M_{\mu}(k)$ ,

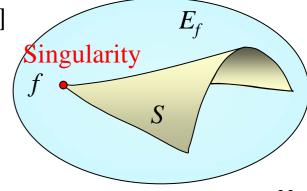
Submodel 
$$S \subset E_f$$
 defined by  $\varphi : K \times [0,1] \to T_f$  
$$S = \{ \exp(u - \Psi_f(u)) f \in E_f \mid u \in \varphi(K \times [0,1]) \}$$

#### such that

- (1) K: compact set
- (2)  $\varphi(a,t) = 0 \iff t = 0$
- (3)  $\varphi(a,t)$ : Frechet differentiable w.r.t. t and  $\partial \varphi$

 $\frac{\partial \varphi}{\partial t}(a,t)$  is continuous on  $K \times [0,1]$ 

(4)  $\min_{a \in K} \left\| \frac{\partial \varphi}{\partial t}(a, t) \right|_{t=0} \right\| > 0$ 



## Asymptotics on singular submodel (cont'd)

#### Lemma (tangent cone)

$$C_f S = \mathbf{R}_{\geq} \left\{ \frac{\partial \varphi}{\partial t} (a, t) \big|_{t=0} \; \middle| \; a \in K \right\}$$

#### **Theorem**

$$\sup_{g \in S} \sum_{i=1}^{n} \log \frac{g(X_i)}{f(X_i)} = \frac{1}{2} \sup_{w \in C_f S, E_f |w|^2 = 1} \frac{\langle w, \hat{m}_n \rangle^2}{\text{projection of empirical mean parameter}}$$

$$\Rightarrow \frac{1}{2} \sup_{w \in C_f S, E_f |w|^2 = 1} G_w^2 \qquad G_w: \text{ Gaussian process}$$

- Analogue to the asymptotic theory on a submodel in a finite dimensional exponential family.
- The same assertion holds without assuming exponential family, but the sufficient conditions and the proof are much more involved.

## Conclusion

- Reproducing kernel exponential manifold are defined as a Hilbert manifold.
  - It is an extension of ordinary finite dimensional exponential family.
  - The model depends on the choice of kernel; the dimension is either finite or infinite.
  - It allows estimation for finite sample, since the likelihood is a continuous functional.
- The pseudo-MLE based on a series of finite dimensional subspaces is proposed, and proved to be consistent.
- It can be used for the asymptotic theory of singular models. The theoretical discussion is easier than general cases.
- Future works:
  - Application to expectation propagation.
  - Dual geometry on reproducing kernel exponential manifolds.

#### References

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## Appendix: Relation with maximal exponential manifold

#### **Proposition**

Let 
$$f\in M_{\mu}(k)$$
, and 
$$A_f=\inf\left\{\alpha>0\,\middle|\, E_f\left[\exp\left(\sqrt{k(X,X)}\middle/\alpha\right)\right]\leq 2\right\}.$$
 Then, 
$$\mathcal{H}_k\subset L_{\cosh-1}(f)\quad\text{and}\quad \left\|u\right\|_{L_{coh}^{-1}(f)}\leq A_f\left\|u\right\|_{\mathcal{H}_k}\quad\text{for any }\;u\in\mathcal{H}_k.$$

Proof. 
$$E_f[\cosh(u(X)/\alpha)-1] = \frac{1}{2}E_f[e^{u(X)/\alpha} + e^{-u(X)/\alpha}] - 1$$
  
 $\leq E_f[e^{|u(X)|/\alpha}] - 1 \leq E_f[\exp(\frac{||u||_{\mathcal{H}_k}}{\alpha}\sqrt{k(X,X)})] - 1.$ 

Thus, 
$$\|u\|_{\mathcal{H}_k}/\alpha < 1/A_f \implies E_f[\cosh(u(X)/\alpha)-1] \le 1$$
.  $q.e.d$ .

 $\square$   $M_{\mu}(k)$  is a subset of the exponential manifold proposed by Pistone and Sempi (1995)