

Supplementary Material

1 Proof of Proposition 1

Notice that $\hat{T}([a, b]; X(n))$ is the sample mean of i.i.d. random variables $Y_i : \Omega \rightarrow \mathbb{R}$ defined as:

$$Y_i = \begin{cases} 1, & \text{if } X_i \cap [a, b] \neq \emptyset, \\ 0, & \text{otherwise.} \end{cases} \quad (36)$$

Therefore, an application of the strong law of large numbers in the classical case yields:

$$\frac{1}{n} \sum_{i=1}^n Y_i \xrightarrow{a.s.} EY_1 = P(X_1 \cap [a, b] \neq \emptyset) = T_{\theta_0}([a, b]), \text{ as } n \rightarrow \infty,$$

$\forall a, b : -\infty < a \leq b < \infty$, and assuming θ_0 is the true parameter value. That is,

$$\hat{T}([a, b]; X(n)) \xrightarrow{a.s.} T_{\theta_0}([a, b]),$$

as $n \rightarrow \infty$. It follows immediately that

$$\left[\hat{T}([a, b]; X(n)) - T_{\theta_0}([a, b]) \right]^2 W(a, b) \xrightarrow{a.s.} 0.$$

Notice that $\forall a, b : -\infty < a \leq b < \infty$, $\left[\hat{T}([a, b]; X(n)) - T_{\theta_0}([a, b]) \right]^2 W(a, b)$ is uniformly bounded by $4C$. By the bounded convergence theorem,

$$\iint_S \left[\hat{T}([a, b]; X(n)) - T_{\theta_0}([a, b]) \right]^2 W(a, b) da db \xrightarrow{a.s.} \iint_S 0 \cdot da db = 0,$$

given any $S \subset \mathbb{R}^2$ with finite Lebesgue measure. This verifies that

$$P_{\theta} \left\{ \omega : \lim_{n \rightarrow \infty} H(X(n); \theta) = 0 \right\} = 1. \quad (37)$$

Similarly, we also get

$$P_{\theta} \left\{ \omega : \lim_{n \rightarrow \infty} H(X(n); \zeta) = \iint_S [T_{\theta}([a, b]) - T_{\zeta}([a, b])]^2 W(a, b) da db \right\} = 1, \quad (38)$$

$\forall \theta, \zeta \in \Theta$. Equations (37) and (38) together imply

$$N(\theta, \zeta) = \iint_S [T_{\theta}([a, b]) - T_{\zeta}([a, b])]^2 W(a, b) da db, \quad \theta, \zeta \in \Theta. \quad (39)$$

By Assumption 2, $T_{\theta}([a, b]) \neq T_{\zeta}([a, b])$, for $\theta \neq \zeta$, except on a Lebesgue set of measure 0. This together with (39) gives

$$N(\theta, \theta) < N(\theta, \zeta), \quad \forall \theta \neq \zeta, \quad \theta, \zeta \in \Theta,$$

which proves that $H(X(n); \theta)$, $\theta \in \Theta$ is a family of contrast functions. To see the equicontinuity of $H(X(n); \theta)$, notice that $\forall \theta_1, \theta_2 \in \Theta$, we have

$$\begin{aligned} & |H(X(n); \theta_1) - H(X(n); \theta_2)| \\ = & \left| \iint_S \left(T_{\theta_1}([a, b]) - \hat{T}([a, b]; X(n)) \right)^2 W(a, b) da db \right. \\ & \left. - \iint_S \left(T_{\theta_2}([a, b]) - \hat{T}([a, b]; X(n)) \right)^2 W(a, b) da db \right| \\ = & \left| \iint_S \left(T_{\theta_1}([a, b]) - T_{\theta_2}([a, b]) \right) \left(T_{\theta_1}([a, b]) + T_{\theta_2}([a, b]) - 2\hat{T}([a, b]; X(n)) \right) W(a, b) da db \right| \\ \leq & 4C \iint_S |T_{\theta_1}([a, b]) - T_{\theta_2}([a, b])| da db, \end{aligned}$$

since, by definition (18), $|W(a, b)|$ is uniformly bounded by C , $\forall a, b : -\infty < a \leq b$. Then the equicontinuity of $H(X(n); \theta)$ follows from the continuity of $T_{\theta}([a, b])$.

2 Lemma 1

Let $H(X(n); \theta)$ be the contrast function defined in (18). Under the hypothesis of Assumption 4,

$$\sqrt{n} \left[\frac{\partial H}{\partial \theta_i} (X(n); \theta_0) \right] \xrightarrow{\mathcal{D}} N(0, \Delta_i), \quad \text{as } n \rightarrow \infty,$$

for $i = 1, \dots, p$, where

$$\begin{aligned} \Delta_i &= 4 \iiint_{S \times S} \{ P(X_1 \cap [a, b] \neq \emptyset, X_1 \cap [c, d] \neq \emptyset) - T_{\theta_0}([a, b]) T_{\theta_0}([c, d]) \} \\ &\quad \times \frac{\partial T_{\theta_0}}{\partial \theta_i}([a, b]) \frac{\partial T_{\theta_0}}{\partial \theta_i}([c, d]) W(a, b) W(c, d) da db dc dd. \end{aligned}$$

Proof. We will write $\frac{\partial T_{\theta_0}([a, b])}{\partial \theta_i} = T_{\theta_0}^i(a, b)$ to simplify notations. Exchanging differentiation and

integration by the bounded convergence theorem, we get

$$\begin{aligned}
& \frac{\partial H}{\partial \theta_i}(X(n); \boldsymbol{\theta}_0) \\
&= \frac{\partial}{\partial \theta_i} \iint_S \left(T_{\boldsymbol{\theta}_0}([a, b]) - \hat{T}([a, b]; X(n)) \right)^2 W(a, b) da db \\
&= \iint_S \frac{\partial}{\partial \theta_i} \left(T_{\boldsymbol{\theta}_0}([a, b]) - \hat{T}([a, b]; X(n)) \right)^2 W(a, b) da db \\
&= \iint_S 2 \left(T_{\boldsymbol{\theta}_0}([a, b]) - \hat{T}([a, b]; X(n)) \right) T_{\boldsymbol{\theta}_0}^i(a, b) W(a, b) da db.
\end{aligned} \tag{40}$$

Define $Y_i(a, b)$ as in (36). Then,

$$\begin{aligned}
(40) &= \iint_S 2 \left(T_{\boldsymbol{\theta}_0}([a, b]) - \frac{1}{n} \sum_{k=1}^n Y_k(a, b) \right) T_{\boldsymbol{\theta}_0}^i(a, b) W(a, b) da db \\
&= \frac{2}{n} \iint_S \sum_{k=1}^n (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) T_{\boldsymbol{\theta}_0}^i(a, b) W(a, b) da db \\
&= \frac{1}{n} \sum_{k=1}^n 2 \iint_S (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) T_{\boldsymbol{\theta}_0}^i(a, b) W(a, b) da db \\
&:= \frac{1}{n} \sum_{k=1}^n R_k.
\end{aligned} \tag{41}$$

Notice that R_k 's are i.i.d. random variables: $\Omega \rightarrow \mathbb{R}$.

Let $\{\Delta s_1, \Delta s_2, \dots, \Delta s_m\}$ be a partition of S , and (a_j, b_j) be any point in Δs_j , $j = 1, \dots, m$. Let $\lambda = \max_{1 \leq j \leq m} \{\text{diam} \Delta s_j\}$. Denote by $\Delta \sigma_j$ the area of Δs_j . By the definition of the double integral,

$$\begin{aligned}
R_k &= 2 \iint_S (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) T_{\boldsymbol{\theta}_0}^i(a, b) W(a, b) da db \\
&= \lim_{\lambda \rightarrow 0} \left\{ \sum_{j=1}^m (T_{\boldsymbol{\theta}_0}([a_j, b_j]) - Y_k(a_j, b_j)) T_{\boldsymbol{\theta}_0}^i(a_j, b_j) W(a_j, b_j) \Delta \sigma_j \right\}.
\end{aligned}$$

Therefore, by the Lebesgue dominated convergence theorem,

$$\begin{aligned}
& ER_k \\
&= 2E \lim_{\lambda \rightarrow 0} \left\{ \sum_{j=1}^m (T_{\theta_0}([a_j, b_j]) - Y_k(a_j, b_j)) T_{\theta_0}^i(a_j, b_j) W(a_j, b_j) \Delta\sigma_j \right\} \\
&= 2 \lim_{\lambda \rightarrow 0} \left\{ \sum_{j=1}^m [E(T_{\theta_0}([a_j, b_j]) - Y_k(a_j, b_j))] T_{\theta_0}^i(a_j, b_j) W(a_j, b_j) \Delta\sigma_j \right\} \\
&= 2 \lim_{\lambda \rightarrow 0} \left\{ \sum_{j=1}^m 0 \right\} = 0.
\end{aligned}$$

Moreover,

$$\begin{aligned}
& Var(R_k) = ER_k^2 \\
&= 4E \left\{ \lim_{\lambda \rightarrow 0} \left\{ \sum_{j=1}^m (T_{\theta_0}([a_j, b_j]) - Y_k(a_j, b_j)) T_{\theta_0}^i(a_j, b_j) W(a_j, b_j) \Delta\sigma_j \right\} \right\}^2 \\
&= 4E \lim_{\lambda_1 \rightarrow 0} \lim_{\lambda_2 \rightarrow 0} \left\{ \sum_{j_1=1}^{m_1} (T_{\theta_0}([a_{j_1}, b_{j_1}]) - Y_k(a_{j_1}, b_{j_1})) T_{\theta_0}^i(a_{j_1}, b_{j_1}) W(a_{j_1}, b_{j_1}) \Delta\sigma_{j_1} \right\} \\
&\quad \left\{ \sum_{j_2=1}^{m_2} (T_{\theta_0}([a_{j_2}, b_{j_2}]) - Y_k(a_{j_2}, b_{j_2})) T_{\theta_0}^i(a_{j_2}, b_{j_2}) W(a_{j_2}, b_{j_2}) \Delta\sigma_{j_2} \right\} \\
&= 4E \lim_{\lambda_1 \rightarrow 0} \lim_{\lambda_2 \rightarrow 0} \sum_{j_1=1}^{m_1} \sum_{j_2=1}^{m_2} (T_{\theta_0}([a_{j_1}, b_{j_1}]) - Y_k(a_{j_1}, b_{j_1})) (T_{\theta_0}([a_{j_2}, b_{j_2}]) - Y_k(a_{j_2}, b_{j_2})) \\
&\quad T_{\theta_0}^i(a_{j_1}, b_{j_1}) T_{\theta_0}^i(a_{j_2}, b_{j_2}) W(a_{j_1}, b_{j_1}) W(a_{j_2}, b_{j_2}) \Delta\sigma_{j_1} \Delta\sigma_{j_2} \\
&= 4 \lim_{\lambda_1 \rightarrow 0} \lim_{\lambda_2 \rightarrow 0} \sum_{j_1=1}^{m_1} \sum_{j_2=1}^{m_2} E(T_{\theta_0}([a_{j_1}, b_{j_1}]) - Y_k(a_{j_1}, b_{j_1})) (T_{\theta_0}([a_{j_2}, b_{j_2}]) - Y_k(a_{j_2}, b_{j_2})) \\
&\quad T_{\theta_0}^i(a_{j_1}, b_{j_1}) T_{\theta_0}^i(a_{j_2}, b_{j_2}) W(a_{j_1}, b_{j_1}) W(a_{j_2}, b_{j_2}) \Delta\sigma_{j_1} \Delta\sigma_{j_2} \\
&= 4 \lim_{\lambda_1 \rightarrow 0} \lim_{\lambda_2 \rightarrow 0} \sum_{j_1=1}^{m_1} \sum_{j_2=1}^{m_2} Cov(Y_k(a_{j_1}, b_{j_1}), Y_k(a_{j_2}, b_{j_2})) \\
&\quad T_{\theta_0}^i(a_{j_1}, b_{j_1}) T_{\theta_0}^i(a_{j_2}, b_{j_2}) W(a_{j_1}, b_{j_1}) W(a_{j_2}, b_{j_2}) \Delta\sigma_{j_1} \Delta\sigma_{j_2} \\
&= 4 \iiint\limits_{S \times S} Cov(Y_k(a, b), Y_k(c, d)) T_{\theta_0}^i(a, b) T_{\theta_0}^i(c, d) W(a, b) W(c, d) dadbdcd \\
&= 4 \iiint\limits_{S \times S} \{P(X_k \cap [a, b] \neq \emptyset, X_k \cap [c, d] \neq \emptyset) - T_{\theta_0}([a, b]) T_{\theta_0}([c, d])\} \\
&\quad T_{\theta_0}^i(a, b) T_{\theta_0}^i(c, d) W(a, b) W(c, d) dadbdcd.
\end{aligned}$$

From the central limit theorem for i.i.d. random variables, the desired result follows. \square

3 Proof of Proposition 2

By the Cramér-Wold device, it suffices to prove

$$\sqrt{n} \sum_{i=1}^p \lambda_i \frac{\partial H}{\partial \theta_i} (X(n); \boldsymbol{\theta}_0) \xrightarrow{\mathcal{D}} N \left(0, \sum_{1 \leq i, j \leq p} \lambda_i \lambda_j \Xi(i, j) \right), \quad (42)$$

for arbitrary real numbers $\lambda_i, i = 1, \dots, p$. It is easily seen from (41) in the proof of Lemma 1 that

$$\begin{aligned} & \sum_{i=1}^p \lambda_i \frac{\partial H}{\partial \theta_i} (X(n); \boldsymbol{\theta}_0) \\ &= \frac{1}{n} \sum_{k=1}^n \left(2 \sum_{i=1}^p \lambda_i \iint_S (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) \frac{\partial T_{\boldsymbol{\theta}_0}}{\partial \theta_i}([a, b]) W(a, b) da db \right) \\ &:= \frac{1}{n} \sum_{k=1}^n \left(2 \sum_{i=1}^p \lambda_i Q_k^i \right). \end{aligned}$$

By Lemma 1,

$$E \left(2 \sum_{i=1}^p \lambda_i Q_k^i \right) = 2 \sum_{i=1}^p \lambda_i \cdot 0 = 0.$$

In view of the central limit theorem for i.i.d. random variables, (42) is reduced to proving

$$Var \left(2 \sum_{i=1}^p \lambda_i Q_k^i \right) = \sum_{1 \leq i, j \leq p} \lambda_i \lambda_j \Xi(i, j). \quad (43)$$

By a similar argument as in Lemma 1, together with some algebraic calculations, we obtain

$$\begin{aligned} & Var \left(2 \sum_{i=1}^p \lambda_i Q_k^i \right) \\ &= 4 \sum_{1 \leq i, j \leq p} \lambda_i \lambda_j Cov \left(Q_k^i, Q_k^j \right) \\ &= 4 \sum_{1 \leq i, j \leq p} \lambda_i \lambda_j E \left(\iint_S (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) \frac{\partial T_{\boldsymbol{\theta}_0}}{\partial \theta_i}([a, b]) W(a, b) da db \right) \\ & \quad \left(\iint_S (T_{\boldsymbol{\theta}_0}([a, b]) - Y_k(a, b)) \frac{\partial T_{\boldsymbol{\theta}_0}}{\partial \theta_j}([a, b]) W(a, b) da db \right) \\ &= 4 \sum_{1 \leq i, j \leq p} \lambda_i \lambda_j \iiint_{S \times S} \{ P(X_1 \cap [a, b] \neq \emptyset, X_1 \cap [c, d] \neq \emptyset) - T_{\boldsymbol{\theta}_0}([a, b]) T_{\boldsymbol{\theta}_0}([c, d]) \} \\ & \quad \frac{\partial T_{\boldsymbol{\theta}_0}}{\partial \theta_i}([a, b]) \frac{\partial T_{\boldsymbol{\theta}_0}}{\partial \theta_j}([c, d]) W(a, b) W(c, d) da db dc dd. \end{aligned}$$

This validates (43), and hence finishes the proof.