

ON CHARACTERIZATION OF POWER SERIES DISTRIBUTIONS BY A MARGINAL DISTRIBUTION AND A REGRESSION FUNCTION

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Abstract. The conditional distribution of Y given $X = x$, where X and Y are non-negative integer-valued random variables, is characterized in terms of the regression function of X on Y and the marginal distribution of X which is assumed to be of a power series form. Characterizations are given for a binomial conditional distribution when X follows a Poisson, binomial or negative binomial, for a hypergeometric conditional distribution when X is binomial and for a negative hypergeometric conditional distribution when X follows a negative binomial.

Key words and phrases: Characterizations, regression function, Poisson, binomial, negative binomial, hypergeometric.

1. Introduction

Korwar (1975) considered two types of characterization problems for distributions of non-negative integer-valued random variables (r.v.) X and Y , when they follow a Poisson, binomial or negative binomial distribution. He derived characterizations of the marginal distribution of X by the conditional distribution of Y given X , and the linear regression of X on Y and characterizations of the conditional distribution of Y given X by the marginal distribution of X and the linear regression of X on Y .

Extensions and generalizations of Korwar's characterizations by conditional distributions and regression functions were considered by, among others, Dahiya and Korwar (1977), Khatri (1978a, 1978b), Cacoullos and Papageorgiou (1983, 1984), Papageorgiou (1983, 1984 and 1985) and Kyriakoussis (1988).

In this paper we are concerned with characterizing the conditional distribution of Y given $X = x$ by the marginal distribution of X and the regression function $E(X|Y)$ of X on Y , when X has a power series

distribution with parameter θ and $E(X|Y)$ is of a power series form with variable θ . Korwar's characterizations of a conditional binomial when X is a Poisson, binomial or negative binomial, as well as additional characterizations of a conditional hypergeometric or negative hypergeometric, are derived as illustrative examples.

2. The main result

Denote the marginal probability functions of the r.v.'s X and Y by $p(x) = P(X = x)$ and $f(y) = P(Y = y)$, respectively. Also denote the conditional probability function of Y given $X = x$ by $p(y|x)$ and that of X given $Y = y$ by $p(x|y)$.

THEOREM 2.1. *Let X be distributed according to the Power Series Distribution*

$$p(x) = \frac{\alpha(x)\theta^x}{\sum_x \alpha(x)\theta^x}, \quad x = 0, 1, 2, \dots$$

Then if the regression function of X on Y is of the form

$$E(X|Y = y) = m(y; \theta) = \sum_{j=0}^r b_j(y)\theta^j,$$

for all θ , with $b_0(y) = y$, $y = 0, 1, 2, \dots$ and r a positive integer or infinity, the conditional distribution of $Y|X$ can be determined uniquely.

PROOF. We have

$$m(y; \theta) = \sum_x xp(x|y) = \sum_x \frac{xp(y|x)p(x)}{f(y)}.$$

Hence

$$m(y; \theta) \sum_x p(y|x)p(x) = \sum_x xp(y|x)p(x),$$

or from the theorem's assumptions

$$(2.1) \quad m(y; \theta) \sum_x p(y|x)\alpha(x)\theta^x = \sum_x xp(y|x)\alpha(x)\theta^x.$$

Writing

$$(2.2) \quad \sum_x p(y|x) a(x) \theta^x = g(y; \theta),$$

equation (2.1) becomes

$$\frac{g'(y; \theta)}{g(y; \theta)} = \frac{m(y; \theta)}{\theta},$$

where $g'(y; \theta) = (\partial/\partial\theta)g(y; \theta)$. Consequently,

$$g(y; \theta) = c(y) \exp \left\{ \int \frac{m(y; \theta)}{\theta} d\theta \right\},$$

where $c(y)$ is only a function of y . Hence

$$(2.3) \quad g(y; \theta) = c(y) \theta^y \exp \left\{ \sum_{j=1}^r b_j(y) \theta^j / j \right\}.$$

However, from the exponential Bell polynomials, $B_n = B_n(b_1, b_2, \dots, b_n)$, which may be defined by their exponential generating function as

$$\sum_{n=0}^{\infty} B_n \theta^n / n! = \exp [\phi(\theta) - \phi(0)],$$

where

$$(2.4) \quad \phi(\theta) = \sum_{j=0}^{\infty} b_j \theta^j / j!, \quad B_0 = 1,$$

equation (2.3) becomes

$$g(y; \theta) = c(y) \theta^y \sum_{n=0}^{\infty} B_n(y) \theta^n / n!,$$

where

$$B_n(y) = B_n(b_1(y), 1!b_2(y), \dots, (n-1)!b_n(y)), \quad n = 1, 2, \dots, \quad B_0(y) = 1.$$

Explicit expressions for $B_n = B_n(b_1, b_2, \dots, b_n)$, as functions of b_1, b_2, \dots, b_n , are given in Kendall and Stuart ((1969), p. 69) and David and Barton ((1962), p. 42) where the role of the cumulants k_i is played here by the b_i as defined in (2.4) and the Bell polynomials themselves are the moments μ_i' . From equation (2.2), we have

$$\sum_{x=0}^{\infty} p(y|x) \alpha(x) \theta^x = c(y) \sum_{x=y}^{\infty} B_{x-y}(y) \theta^x / (x-y)! .$$

Equating coefficients of θ^x , we obtain

$$(2.5) \quad \begin{aligned} p(y|x) &= c(y) \frac{1}{\alpha(x)} \frac{B_{x-y}(y)}{(x-y)!}, & y = 0, 1, 2, \dots, x, \quad x = 0, 1, 2, \dots, \\ p(y|x) &= 0 & \text{for } x < y. \end{aligned}$$

To determine the coefficients $c(y)$, note first that, from (2.5), for a given x , y ranges from zero to x , so that in turn

$$\sum_{i=0}^x p(i|x) = 1, \quad x = 0, 1, 2, \dots ;$$

equivalently,

$$p(y|y) = 1 - \sum_{i=0}^{y-1} p(i|y) .$$

From the relation (2.5), we have

$$(2.6) \quad c(y) = \alpha(y) - \sum_{i=0}^{y-1} c(i) \frac{B_{y-i}(i)}{(y-i)!}, \quad y = 1, 2, \dots ,$$

and

$$(2.7) \quad c(0) = \alpha(0) ,$$

because $p(0|0) = 1$.

From (2.6) and (2.7) the remaining coefficients $c(1), c(2), \dots$ can be determined uniquely.

Conversely, if the conditional distribution of $Y|X=x$ and the marginal distribution of X are known, then the joint distribution of (X, Y) is determined; therefore the marginal distribution of Y , the conditional distribution of $X|Y$ and the regression $E(X|Y)$ can also be determined.

Characterizations of some well-known distributions are given as illustrative examples when the r.v. X follows a Poisson, binomial or negative binomial distribution. For clarity these results are summarized in Table 1.

Table 1. Bell polynomials and conditional distributions.

Marginal Distribution $P(X = x)$	Regression $E(X Y = y) = m(y; \theta)$	Exponential Bell Polynomials $B_n(y)$	Conditional Distribution $P(Y X = x)$
$e^{-\theta} \theta^x / x!$	$y + \theta q$	q^n	$\binom{x}{y} p^y q^{x-y}$
$\theta > 0$	for all $\theta > 0, y = 0, 1, 2, \dots$ and some $0 < q < 1$	$n = 0, 1, 2, \dots$	$y = 0, 1, 2, \dots, x, p = 1 - q$
$\binom{y}{x} \alpha^x (1 - \alpha)^{y-x}$	$(y + \nu q \theta) / (1 + q \theta)$	$n! \binom{y-\nu}{n} q^n$	$\binom{x}{y} p^y q^{x-y}$
$0 < \alpha < 1, \nu = 0, 1, 2, \dots$	for all $\theta = \alpha / (1 - \alpha),$ $y = 0, 1, 2, \dots$ and some $0 < q < 1 - \alpha$	$n = 0, 1, 2, \dots, y - \nu,$ $y \leq \nu$	$y = 0, 1, 2, \dots, x, p = 1 - q$
$\binom{-\nu}{x} (-1)^x \alpha^x (1 - \alpha)^{-x}$	$(y + \nu q \theta) / (1 - q \theta)$	$n! \binom{-y-\nu}{n} (-1)^n q^n$	$\binom{x}{y} p^y q^{x-y}$
$0 < \alpha < 1, \nu > 0$	for all $\theta = 1 - \alpha,$ $y = 0, 1, 2, \dots$ and some $0 < q < 1$	$n = 0, 1, 2, \dots$	$y = 0, 1, 2, \dots, x, p = 1 - q$
$\binom{N}{x} p^x (1 - p)^{N-x}$	$y + (N - \nu) \theta / (\theta + 1)$	$n! \binom{N-\nu}{n}$	$\binom{x}{y} \binom{N-x}{y-y} / \binom{N}{y}$
$0 < p < \frac{1}{2}, N = 0, 1, \dots$	for all $\theta = p / (1 - p),$ $y = 0, 1, \dots$ and some positive integer $\nu < N$	$n = 0, 1, 2, \dots, N - \nu$	$y = 0, 1, 2, \dots, \min\{x, \nu\}$
$\binom{-N}{x} (-1)^x p^x (1 - p)^{-x}$	$y + (N - \nu) \theta / (1 - \theta)$	$n! \binom{-N+\nu}{n} (-1)^n$	$\binom{-\nu}{y} \binom{-N+\nu}{x-y} / \binom{-N}{x}$
$0 < p < 1, N > 0$	for all $\theta = 1 - p,$ $y = 0, 1, 2, \dots$ and for some positive integer $\nu < N$	$n = 0, 1, 2, \dots$	$y = 0, 1, 2, \dots, x$

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