

ON THE ROBUSTNESS OF BALANCED FRACTIONAL
 2^m FACTORIAL DESIGNS OF RESOLUTION $2l+1$
IN THE PRESENCE OF OUTLIERS*

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Summary

By use of the algebraic structure, we obtain a simplified expression for the outlier-insensitivity factor for balanced fractional 2^m factorial (2^m -BFF) designs of resolution $2l+1$ derived from simple arrays (S -arrays), whose measure has been introduced by Ghosh and Kipngeno (1985, *J. Statist. Plann. Inference*, 11, 119-129). It is defined by use of the measure suggested by Box and Draper (1975, *Biometrika*, 62 (2), 347-352). As examples, we study the sensitivity of A -optimal 2^m -BFF designs of resolution VII (i.e., $l=3$) given by Shirakura (1976, *Ann. Statist.*, 4, 515-531; 1977, *Hiroshima Math. J.*, 7, 217-285). We observe that these designs are robust in the sense that they have low sensitivities.

1. Introduction

The concept of a balanced array (B -array) was introduced and first studied by Chakravarti [2]. A general connection between a B -array of strength $2l$ and a 2^m -BFF design of resolution $2l+1$ was established by Yamamoto, Shirakura and Kuwada [15]. Furthermore, these authors ([16]) obtained an explicit expression for the characteristic polynomial of the information matrix of a 2^m -BFF design of resolution $2l+1$ by utilizing the decomposition of the triangular multidimensional partially balanced (TMDPB) association algebra into its $l+1$ two-sided ideals. This polynomial includes the results obtained by Srivastava and Chopra [12] as a special case. A - and/or D -optimal 2^m -BFF designs of resolution V and VII have been obtained by several authors (e.g., Shirakura [9] and [10] and Srivastava and/or Chopra [3]-[7], [13] and [14]).

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As a measure of sensitivity in the sense that the design should be insensitive to wild observations, Box and Draper [1] introduced the sum of squares of diagonal elements of the projection matrix, when the number of observations and the number of unknown effects in the model assumed are both fixed. Recently, using the measure suggested by Box and Draper [1], Ghosh and Kipngeno [8] have defined a new measure of robustness of a design with respect to outliers, which is called the "outlier-insensitivity factor", and they have found these values for A -optimal 2^m -BFF designs of resolution V given by Srivastava and/or Chopra [3]–[7], [13] and [14].

In this paper, we obtain a simplified expression for the outlier-insensitivity factor for a 2^m -BFF design of resolution $2l+1$ derived from an S -array, by use of the properties of the TMDPB association scheme and its algebra. As examples, we find its value for A -optimal 2^m -BFF designs of resolution VII given by Shirakura [9] and [10], when $6 \leq m \leq 9$.

2. Measures of sensitivity

Consider a fractional experiment with m factors each at two levels (0 and 1, say). Then the ordinary linear model associated with a fraction T with N assemblies is

$$(2.1) \quad \mathcal{E}[\mathbf{y}(T)] = E_T \boldsymbol{\theta}, \quad \text{Var}[\mathbf{y}(T)] = \sigma^2 I_N, \quad \text{Rank}(E_T) = \nu_l,$$

where $\mathcal{E}[\mathbf{y}]$ stands for the expected value of \mathbf{y} , $\mathbf{y}(T)$ is a vector of N observations, E_T is the $N \times \nu_l$ design matrix, $\boldsymbol{\theta}$ is a vector of unknown effects up to the l -factor interactions, σ^2 is a constant which may or may not be known, and $\nu_l = 1 + \binom{m}{1} + \dots + \binom{m}{l}$. Here $l \leq [m/2]$, where $[x]$ denotes the largest integer not exceeding x . The predicted value of $\mathbf{y}(T)$ is $\hat{\mathbf{y}}(T) = R\mathbf{y}(T)$, where $R = E_T(E_T'E_T)^{-1}E_T'$ which is known as the projection matrix. Here A' denotes the transpose of a matrix A . Suppose that the u -th observation in $\mathbf{y}(T)$ is an outlier in the sense that an unknown aberration c , a fixed constant, is added to it. And we denote the resulting observation vector as $\mathbf{y}^*(T)$ and the corresponding predicted value as $\hat{\mathbf{y}}^*(T) = R\mathbf{y}^*(T)$. Then the quantity $d_u = \{\hat{\mathbf{y}}^*(T) - \hat{\mathbf{y}}(T)\}' \{\hat{\mathbf{y}}^*(T) - \hat{\mathbf{y}}(T)\}$ is a measure of overall discrepancy caused by the effect of c on the u -th observation, and it is equal to $c^2 r_{uu}$, where r_{uu} is the u -th diagonal element of R . Clearly

$$(2.2) \quad \sum_{u=0}^N d_u = c^2 \sum_{u=1}^N r_{uu} = c^2 \nu_l,$$

because of the idempotency of R . If it is equally likely that c occurs

with any of N observations, giving rise to d_1, \dots, d_N , the average discrepancy is $\bar{d} = c^2 \sum_{u=0}^N r_{uu} / N = c^2 \nu_i / N$. If the d_u ($u=1, \dots, N$) are as small as possible, then a design is said to be insensitive with respect to outliers. Because $\sum d_u$ is fixed as in (2.2), this means that the d_u are as uniformly as possible for a design insensitive to outliers. As one convenient measure of uniformity, Box and Draper [1] introduced the following:

$$V(d) = \sum_{u=1}^N (d_u - \bar{d})^2 / N = c^4 \{r - (\nu_i)^2 / N\} / N,$$

where

$$r = \sum_{u=1}^N (r_{uu})^2.$$

Thus to ensure insensitivity to outliers, $V(d)$ should be made small. Minimization of $V(d)$ is equivalent to minimization of r , when N and ν_i are both fixed. Recently, Ghosh and Kipngeno [8] have defined the outlier-insensitivity factor, E , say, by

$$E = 100 \times (\nu_i)^2 / (Nr),$$

because $r \geq (\nu_i)^2 / N$.

Note that under the model (2.1), for an orthogonal design or $N = \nu_i$, i.e., saturated design, we have $E = 100$.

3. Outlier-insensitivity factors

Under the model (2.1), the expected value of an observation associated with an assembly $(\varepsilon_1, \dots, \varepsilon_m)$ with $\varepsilon_k = 0$ or 1 is given by

$$(3.1) \quad \mathcal{E}[y(\varepsilon_1, \dots, \varepsilon_m)] = \sum_{\eta_1, \dots, \eta_m} d_{\varepsilon_1}(\eta_1) \cdots d_{\varepsilon_m}(\eta_m) \theta(\eta_1, \dots, \eta_m),$$

where the summation is over all binary numbers (η_1, \dots, η_m) with $\eta_k = 0$ or 1 such that $0 \leq \eta_1 + \dots + \eta_m \leq l$, and

$$(3.2) \quad d_0(0) = d_1(0) = d_1(1) = 1 \quad \text{and} \quad d_0(1) = -1.$$

Note that when $\eta_1 + \dots + \eta_m = j$ ($j = 0, 1, \dots, l$), $\theta(\eta_j)$ is called the j -factor interaction, where $\eta'_j = (\eta_1, \dots, \eta_m)$.

When T is a B -array of strength m , size N , m constraints, 2 levels and index set $\{\lambda_0, \lambda_1, \dots, \lambda_m\}$, T is called an S -array, written $SA(m; \lambda_0, \lambda_1, \dots, \lambda_m)$ for brevity (see [9]). For T being an $SA(m; \lambda_0, \lambda_1, \dots, \lambda_m)$, T can be expressed as $T = \|j_{\lambda_i} \otimes T_i\|$ if $\lambda_i \geq 1$ ($i = 0, 1, \dots, m$), where T_i are the $(0, 1)$ matrices of size $\binom{m}{i} \times m$ whose rows denote all distinct

vectors with weight i . Here \mathbf{j}_p and $A \otimes B$ denote, respectively, the $p \times 1$ vector with all unity and the Kronecker product of two matrices A and B , and the weight of a $(0, 1)$ vector means the number of ones in the vector.

Let $E_T = \|\mathbf{j}_i \otimes E_i(j)\|$ if $\lambda_i \geq 1$ ($i=0, 1, \dots, m$; $j=0, 1, \dots, l$), where $E_i(j)$ denote the submatrices of E_T corresponding, respectively, to T_i and $\theta(\eta_j)$. Then from (3.1), (3.2) and Appendix, we have

$$E_i(j) = \sum_{\alpha=0}^{\min(i, j, m-i)} (-1)^{j-\min(i, j)+\alpha} A_\alpha^{(i, j)}$$

for $i=0, 1, \dots, m$ and $j=0, 1, \dots, l$,

where $\min(a_1, \dots, a_n)$ denotes the minimum value of integers a_1, \dots, a_n , and $A_\alpha^{(i, j)}$ are the local association matrices which are given in Appendix. Hence, it follows from Appendix that

$$(3.3) \quad E_i(j) = \begin{cases} \sum_{\alpha=0}^{\min(i, j)} (-1)^{j-\min(i, j)+\alpha} A_\alpha^{(i, j)} & \text{for } 0 \leq i \leq [m/2], \\ \sum_{\alpha=0}^{\min(m-i, j)} (-1)^{\min(m-i, j)-\alpha} A_\alpha^{(m-i, j)} & \text{for } [m/2] < i \leq m. \end{cases}$$

Thus, from (3.3) and Appendix, the following is immediate.

LEMMA 3.1. *The submatrices $E_i(j)$ of E_T can be expressed as*

$$E_i(j) = \begin{cases} \sum_{\beta=0}^{\min(i, j)} h_\beta^{(i, j)} A_\beta^{*(i, j)} & \text{for } 0 \leq i \leq [m/2] \text{ and } 0 \leq j \leq l, \\ \sum_{\beta=0}^{\min(m-i, j)} h_\beta^{*(m-i, j)} A_\beta^{(m-i, j)} & \text{for } [m/2] < i \leq m \text{ and } 0 \leq j \leq l, \end{cases}$$

where

$$h_\beta^{(i, j)} = \begin{cases} (-1)^j (-2)^i \left\{ \binom{m-i-\beta}{j-i} / \binom{j-\beta}{j-i} \right\}^{1/2} \\ \quad \times \left\{ \sum_{b=0}^{i-\beta} (-1/2)^b \binom{j-\beta}{i-\beta-b} \binom{m-i-\beta+b}{b} \right\} & \text{if } 0 \leq i \leq j \leq [m/2], \\ 2^j \left\{ \binom{m-j-\beta}{i-j} / \binom{i-\beta}{i-j} \right\}^{1/2} \\ \quad \times \left\{ \sum_{b=0}^{j-\beta} (-1/2)^b \binom{i-\beta}{j-\beta-b} \binom{m-j-\beta+b}{b} \right\} & \text{if } 0 \leq j \leq i \leq [m/2], \end{cases}$$

and $h_\beta^{*(m-i, j)}$ are given by replacing i in $h_\beta^{(i, j)}$ by $m-i$.

For T being an SA $(m; \lambda_0, \lambda_1, \dots, \lambda_m)$, the submatrices $M_{u,v}$ of $E'_T E_T$ corresponding, respectively, to $\{\theta(\eta_u)\}$ and $\{\theta(\eta_v)\}$ can be expressed as

$$M_{u,v} = \sum_{\alpha=0}^{\min(u,v)} \gamma_{|v-u|+2\alpha} A_{\alpha}^{*(u,v)} \quad \text{for } 0 \leq u, v \leq l,$$

where a connection between γ 's and λ 's is given by

$$\gamma_i = \sum_{j=0}^m \sum_{p=0}^i (-1)^p \binom{i}{p} \binom{m-i}{j-i+p} \lambda_j$$

(see [15]). It follows from Appendix that $M_{u,v}$ can be expressed as

$$M_{u,v} = \sum_{\beta=0}^{\min(u,v)} \kappa_{\beta}^{u-\beta, v-\beta} A_{\beta}^{*(u,v)} \quad \text{for } 0 \leq u, v \leq l,$$

where

$$\kappa_{\beta}^{u,v} = \sum_{\alpha=0}^{\beta+u} \gamma_{v-u+2\alpha} z_{\beta\alpha}^{(\beta+u, \beta+v)} \quad \text{for } 0 \leq u \leq v \leq l - \beta$$

and $\kappa_{\beta}^{v,u} = \kappa_{\beta}^{u,v}$ for $u \leq v$ (see [16]). If $E'_T E_T$ is non-singular, then the submatrices $M_{u,v}^*$ of $(E'_T E_T)^{-1}$ corresponding, respectively, to $\{\theta(\eta_u)\}$ and $\{\theta(\eta_v)\}$ is

$$(3.4) \quad M_{u,v}^* = \sum_{\beta=0}^{\min(u,v)} \kappa_{u-\beta, v-\beta}^{\beta} A_{\beta}^{*(u,v)} \quad \text{for } 0 \leq u, v \leq l,$$

where $\|\kappa_{u,v}^{\beta}\| = \|\kappa_{\beta}^{u,v}\|^{-1}$. Note that the order of $\|\kappa_{u,v}^{\beta}\|$ is $l+1-\beta$ ($\beta=0, 1, \dots, l$) and does not depend on m , while the order of $E'_T E_T$ is ν_i . Thus from Lemma 3.1, (3.4) and Appendix, the following is immediate.

LEMMA 3.2. *The diagonal submatrices R_{ii} of R corresponding to T_i can be expressed as*

$$R_{ii} = \begin{cases} \sum_{j=0}^l \sum_{k=0}^l \left\{ \sum_{\beta=0}^{\min(i,j,k)} h_{\beta}^{(i,j)} h_{\beta}^{(i,k)} \kappa_{j-\beta, k-\beta}^{\beta} A_{\beta}^{*(i,i)} \right\} & \text{if } 0 \leq i \leq [m/2], \\ \sum_{j=0}^l \sum_{k=0}^l \left\{ \sum_{\beta=0}^{\min(m-i, j, k)} h_{\beta}^{*(m-i, j)} h_{\beta}^{*(m-i, k)} \kappa_{j-\beta, k-\beta}^{\beta} A_{\beta}^{*(m-i, m-i)} \right\} & \text{if } [m/2] < i \leq m. \end{cases}$$

Since $A_0^{(i,i)} = I_{\binom{m}{i}}$ for $i=0, 1, \dots, m$, it follows from Appendix that the coefficients of $A_0^{(i,i)}$ in $A_{\beta}^{*(i,i)}$ are given by

$$(3.5) \quad z_{(i,i)}^{\beta 0} = \phi_{\beta} z_{\beta 0}^{(i,i)} / \left\{ \binom{m}{i} \binom{i}{0} \binom{m-i}{0} \right\} = \phi_{\beta} / \binom{m}{i},$$

where $z_{\beta\alpha}^{(u,v)}$, $z_{(u,v)}^{\beta\alpha}$ and ϕ_{β} are given in Appendix. Hence, from Lemma 3.2 and (3.5), the diagonal elements r_{ii} of R_{ii} are given by

$$(3.6) \quad r_{ii} = \begin{cases} \left[\sum_{j=0}^l \sum_{k=0}^l \left\{ \sum_{\beta=0}^{\min(i,j,k)} \phi_{\beta} h_{\beta}^{(i,j)} h_{\beta}^{(i,k)} \kappa_{j-\beta, k-\beta}^{\beta} \right\} \right] / \binom{m}{i} & \text{if } 0 \leq i \leq [m/2], \\ \left[\sum_{j=0}^l \sum_{k=0}^l \left\{ \sum_{\beta=0}^{\min(m-i,j,k)} \phi_{\beta} h_{\beta}^{*(m-i,j)} h_{\beta}^{*(m-i,k)} \kappa_{j-\beta, k-\beta}^{\beta} \right\} \right] / \binom{m}{i} & \text{if } [m/2] < i \leq m. \end{cases}$$

Thus the following is immediate.

LEMMA 3.3. For T being an SA $(m; \lambda_0, \lambda_1, \dots, \lambda_m)$, under the model (2.1), r is given by

$$r = \sum_{i=0}^m \lambda_i \binom{m}{i} (r_{ii})^2,$$

where r_{ii} are given by (3.6).

From Lemma 3.3, we have the main results of this paper as follows.

THEOREM 3.1. Let T be an SA $(m; \lambda_0, \lambda_1, \dots, \lambda_m)$. Then under the model (2.1), the outlier-insensitivity factor E is given by

$$E = 100 \times (\nu_i)^2 / (Nr),$$

where r is given in Lemma 3.3.

4. Calculation of E for Shirakura's designs

In this section, we study the sensitivity of A -optimal 2^m -BFF designs of resolution VII (i.e., $l=3$) given by Shirakura [9] and [10]. It follows from Lemma 3.1 that

$$\begin{aligned} h_0^{(i,0)} &= \left\{ \binom{m}{i} \right\}^{1/2} & \text{for } 0 \leq i \leq [m/2], \\ h_0^{(0,j)} &= (-1)^j \left\{ \binom{m}{j} \right\}^{1/2} & \text{for } j=1, 2, 3, \\ h_0^{(i,1)} &= -(m-2i) \left\{ \binom{m-1}{i-1} / i \right\}^{1/2} & \text{for } 1 \leq i \leq [m/2], \\ h_1^{(i,1)} &= 2 \left\{ \binom{m-2}{i-1} \right\}^{1/2} & \text{for } 1 \leq i \leq [m/2], \\ h_{\beta}^{(i,j)} &= (-1)^{j-1} h_{\beta}^{(j,1)} & \text{for } j=2, 3 \text{ and } \beta=0, 1, \\ h_0^{(i,2)} &= \left\{ 4 \binom{i}{2} - 2(m-1)i + \binom{m}{2} \right\} \left\{ \binom{m-2}{i-2} / \binom{i}{2} \right\}^{1/2} & \text{for } 2 \leq i \leq [m/2], \end{aligned}$$

$$\begin{aligned}
 h_1^{(i,2)} &= -2(m-2i) \left\{ \binom{m-3}{i-2} / (i-1) \right\}^{1/2} && \text{for } 2 \leq i \leq [m/2], \\
 h_2^{(i,2)} &= 4 \left\{ \binom{m-4}{i-2} \right\}^{1/2} && \text{for } 2 \leq i \leq [m/2], \\
 h_\beta^{(2,3)} &= -h_\beta^{(3,2)} && \text{for } \beta = 0, 1, 2, \\
 h_0^{(i,3)} &= \left\{ 8 \binom{i}{3} - 4(m-2) \binom{i}{2} + 2 \binom{m-1}{2} i - \binom{m}{3} \right\} \left\{ \binom{m-3}{i-3} / \binom{i}{3} \right\}^{1/2} && \text{for } 3 \leq i \leq [m/2], \\
 h_1^{(i,3)} &= \left\{ 8 \binom{i-1}{2} - 4(m-3)(i-1) + 2 \binom{m-2}{2} \right\} \left\{ \binom{m-4}{i-3} / \binom{i-1}{2} \right\}^{1/2} && \text{for } 3 \leq i \leq [m/2], \\
 h_2^{(i,3)} &= -4(m-2i) \left\{ \binom{m-5}{i-3} / (i-2) \right\}^{1/2} && \text{for } 3 \leq i \leq [m/2], \\
 h_3^{(i,3)} &= 8 \left\{ \binom{m-6}{i-3} \right\}^{1/2} && \text{for } 3 \leq i \leq [m/2],
 \end{aligned}$$

and $h_\beta^{*(m-t, j)}$ are given by replacing i in $h_\beta^{(i, j)}$ by $m-i$. Thus for T being an SA $(m; \lambda_0, \lambda_1, \dots, \lambda_m)$, we have

$$E = 100 \times (\nu_3)^2 / (Nr),$$

where

$$r = \sum_{i=0}^m \lambda_i \binom{m}{i} (r_{ii})^2$$

and

$$r_{ii} = \begin{cases} \left[\sum_{j=0}^3 \sum_{k=0}^3 \left\{ \sum_{\beta=0}^{\min(t, j, k)} \phi_\beta h_\beta^{(t, j)} h_\beta^{(t, k)} \kappa_{j-\beta, k-\beta}^\beta \right\} \right] / \binom{m}{i} & \text{if } 0 \leq i \leq [m/2], \\ \left[\sum_{j=0}^3 \sum_{k=0}^3 \left\{ \sum_{\beta=0}^{\min(m-t, j, k)} \phi_\beta h_\beta^{*(m-t, j)} h_\beta^{*(m-t, k)} \kappa_{j-\beta, k-\beta}^\beta \right\} \right] / \binom{m}{i} & \text{if } [m/2] < i \leq m. \end{cases}$$

For $6 \leq m \leq 9$, all A -optimal 2^m -BFF designs of resolution VII given by Shirakura [9] and [10] except for the designs corresponding to $m=8$ and $N=127b, 128b$ are S -arrays. In Tables 1, 2, 3 and 4, the values of the outlier-insensitivity factor E of A -optimal 2^m -BFF designs of resolution VII are presented for $m=6$ and $42 \leq N \leq 64$, $m=7$ and $64 \leq N \leq 90$, $m=8$ and $93 \leq N \leq 128$, and $m=9$ and $130 \leq N \leq 150$, respectively.

These values are greater than 93, 87, 92 and 91 for $m=6, 7, 8$ and 9, respectively. Therefore we conclude that the sensitivities of the Shirakura's designs to outliers are low.

Table 1.		Table 2.		Table 3.		Table 4.	
N	E	N	E	N	E	N	E
42	100.00000	64	100.00000	93	100.00000	130	100.00000
43	99.81403	65a	99.23683	94	99.47096	131	99.61979
44	98.68351	65b	99.59870	95	99.05598	132	98.99250
45	97.42432	66	98.81355	96	98.19897	133	98.31154
46	95.70471	67a	97.59471	97	97.51840	134	97.96954
47	99.30583	67b	97.86155	98	96.61199	135	97.28181
48	99.28371	68	96.67693	99	95.83656	136	96.78959
48	98.83269	69a	95.40590	100	94.93192	137	96.10858
49	98.80295	69b	95.56592	101	96.21760	138	95.55457
50	97.79658	70	98.29494	102	95.81256	139	96.87867
50	97.70547	71a	97.54619	103	95.06137	140	96.57143
51	96.72022	71b	95.35560	104	94.23622	141	96.01454
52	95.31230	72a	97.25246	105	93.58704	142	95.40207
52	95.27303	72b	94.81481	106	92.75629	143	94.77291
53	93.92315	73a	96.33422	107	98.81635	144	94.13993
54	95.26166	73b	94.09180	108	98.50349	145	93.62473
55	95.05672	74a	95.54002	109	98.27383	146	93.00120
56	94.22506	74b	93.06986	110	97.57208	147	92.46373
57	98.92164	75a	94.48201	111	97.05163	148	92.08806
58	99.10209	75b	92.16546	112	96.28676	149	91.84034
59	98.91176	76a	93.52295	113	95.65440	150	91.35135
60	98.30057	76b	91.07877	114	94.87960		
61	97.22206	77	94.42167	115	96.66927		
62	99.83649	78	93.82685	116	96.45581		
63	99.89680	79	93.24751	117	96.00049		
64	100.00000	80	92.47121	118	95.37221		
		81	91.66177	119	94.79438		
		82	90.74376	120	94.10029		
		83	89.85773	121	93.47427		
		84	88.90973	122	92.76839		
		85a	95.80994	123	95.96798		
		85b	90.07147	124	95.48760		
		86a	95.59750	125	95.07700		
		86b	89.86398	126	94.45884		
		87a	95.32026	127a	99.89345		
		87b	89.95140	128a	100.00000		
		88a	94.71190				
		88b	89.20927				
		89a	93.95288				
		89b	88.75775				
		90a	93.18285				
		90b	87.91516				

5. Concluding remarks

Let T^* be an SA $(7; 3, 0, 0, 1, 1, 0, 1, 3)$, which is also a B -array of strength 6 and size 83. Then we have $\det(\|\kappa_1^{v_i, v_j}\|) = 129466368$, $\det(\|\kappa_1^{v_i, v_j}\|) = 73728$, $\det(\|\kappa_2^{p_i, p_j}\|) = 3072$ and $\det(\|\kappa_3^{0_i, 0_j}\|) = 128$, where $\det(A)$ denotes the determinant of a matrix A . Hence, T^* is a 2^7 -BFF design of resolution VII. It follows from Section 4 that $E = 96.91502$ for T^* , which is the most insensitive design to outliers in the class of balanced designs derived from S -arrays, while, from Table 2, we have $E = 89.85773$

for an A -optimal 2^l -BFF design T of resolution VII. On the other hand, it follows from Theorem 2.1 of Shirakura [9] that

$$\text{tr} \{(E_T' E_T)^{-1}\} = 1.50710 \quad \text{for } T^*$$

and

$$\text{tr} \{(E_T' E_T)^{-1}\} = 0.88119 \quad \text{for } T.$$

This implies that in the restricted class mentioned above, the most insensitive design to outliers is not always good design in some sense. It, however, is worth to calculate the values of E for some optimal designs with respect to the popular criteria (e.g., A -, D - and E -optimal).

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REFERENCES

- [1] Box, G. E. P. and Draper, N. R. (1975). Robust designs, *Biometrika*, **62**(2), 347-352.
- [2] Chakravarti, I. M. (1956). Fractional replication in asymmetrical factorial designs and partially balanced arrays, *Sankhyā*, **17**, 143-164.
- [3] Chopra, D. V. (1975). Balanced optimal 2^8 fractional factorial designs of resolution V, $52 \leq N \leq 59$, *A Survey of Statistical Design and Linear Models* (ed., J. N. Srivastava), North-Holland, Amsterdam, 91-100.
- [4] Chopra, D. V. and Srivastava, J. N. (1973a). Optimal balanced 2^7 fractional factorial designs of resolution V with $N \leq 42$, *Ann. Inst. Statist. Math.*, **25**, 587-604.
- [5] Chopra, D. V. and Srivastava, J. N. (1973b). Optimal balanced 2^7 fractional factorial designs of resolution V, $49 \leq N \leq 55$, *Commun. Statist.*, **2**, 59-84.
- [6] Chopra, D. V. and Srivastava, J. N. (1974). Optimal balanced 2^8 fractional factorial designs of resolution V, $37 \leq N \leq 51$, *Sankhyā*, **A36**, 41-52.
- [7] Chopra, D. V. and Srivastava, J. N. (1975). Optimal balanced 2^7 fractional factorial designs of resolution V, $43 \leq N \leq 48$, *Sankhyā*, **B37**, 429-447.
- [8] Ghosh, S. and Kipngeno, W. A. K. (1985). On the robustness of the optimum balanced 2^m factorial designs of resolution V (given by Srivastava and Chopra) in the presence of outliers, *J. Statist. Plann. Inference*, **11**, 119-129.
- [9] Shirakura, T. (1976). Optimal balanced fractional 2^m factorial designs of resolution VII, $6 \leq m \leq 8$, *Ann. Statist.*, **4**, 515-531.
- [10] Shirakura, T. (1977). Contributions to balanced fractional 2^m factorial designs derived from balanced arrays of strength $2l$, *Hiroshima Math. J.*, **7**, 217-285.
- [11] Shirakura, T. and Kuwada, M. (1976). Covariance matrices of the estimates for balanced fractional 2^m factorial designs of resolution $2l+1$, *J. Japan Statist. Soc.*, **6**, 27-31.
- [12] Srivastava, J. N. and Chopra, D. V. (1971a). On the characteristic roots of the information matrix of 2^m balanced factorial designs of resolution V, with applications, *Ann. Math. Statist.*, **42**, 722-734.
- [13] Srivastava, J. N. and Chopra, D. V. (1971b). Balanced optimal 2^m fractional factorial

designs of resolution V, $m \leq 6$, *Technometrics*, **13**, 257-269.

[14] Srivastava, J. N. and Chopra, D. V. (1974). Balanced trace-optimal 2^l fractional factorial designs of resolution V, with 56 to 68 runs, *Utilitas Math.*, **5**, 263-279.

[15] Yamamoto, S., Shirakura, T. and Kuwada, M. (1975). Balanced arrays of strength $2l$ and balanced fractional 2^m factorial designs, *Ann. Inst. Statist. Math.*, **27**, 143-157.

[16] Yamamoto, S., Shirakura, T. and Kuwada, M. (1976). Characteristic polynomials of the information matrices of balanced fractional 2^m factorial designs of higher $(2l+1)$ resolution, *Essays in Probability and Statistics* (eds., S. Ikeda et al.), Shinko Tsusho, Tokyo, 73-94.

Appendix

Let $(\epsilon_1, \dots, \epsilon_m)$ ($=\epsilon'_i$, say) be a $(0, 1)$ vector with weight i . Further let $S_i = \{\epsilon_i\}$ ($i=0, 1, \dots, m$). Then $|S_i| = \binom{m}{i}$ ($=n_i$, say), where $|S|$ denotes the cardinality of a set S . Suppose a relation of association defined among the sets S_i in such a way that $\epsilon_i (\in S_i)$ and $\epsilon_j (\in S_j)$ are the α -th associates if

$$(A.1) \quad \epsilon'_i \epsilon_j = \min(i, j) - \alpha,$$

where $\alpha = 0, 1, \dots, \min(i, m-i, j, m-j)$. Let $\tilde{\epsilon}_k = j_m - \epsilon_k$. Then if ϵ_i is the α -th associate of ϵ_j , then (A.1) shows that ϵ_i is the α^* -th associate of $\tilde{\epsilon}_j$, $\tilde{\epsilon}_i$ is the α^{**} -th associate of ϵ_j , and $\tilde{\epsilon}_i$ is the α^{***} -th associate of $\tilde{\epsilon}_j$, where $\alpha^* = \min(i, m-j) - i + \min(i, j) - \alpha$, $\alpha^{**} = \min(m-i, j) - j + \min(i, j) - \alpha$, and $\alpha^{***} = \min(m-i, m-j) - m + i + j - \min(i, j) + \alpha$.

Let $A_\alpha^{(u,v)}$ ($=A_\alpha^{(v,u)'}$) be the $n_u \times n_v$ local association matrix of the TMDPB association scheme, where $0 \leq u \leq v \leq [m/2]$ and $\alpha = 0, 1, \dots, l$ (e.g., [15]). Further let $A_\beta^{*(u,v)}$ ($=A_\beta^{*(v,u)'}$) ($0 \leq u \leq v \leq [m/2]$; $\beta = 0, 1, \dots, l$) be the $n_u \times n_v$ matrices which are linearly linked with $A_\alpha^{(u,v)}$ as follows (e.g., [11] and [16]).

$$(A.2) \quad A_\alpha^{(u,v)} = \sum_{\beta=0}^u z_{\beta\alpha}^{(u,v)} A_\beta^{*(u,v)} \quad \text{for } 0 \leq \alpha \leq u \leq v \leq [m/2]$$

and

$$(A.3) \quad A_\beta^{*(u,v)} = \sum_{\alpha=0}^u z_{\beta\alpha}^{(u,v)} A_\alpha^{(u,v)} \quad \text{for } 0 \leq \beta \leq u \leq v \leq [m/2],$$

where

$$z_{\beta\alpha}^{(u,v)} = \sum_{b=0}^{\alpha} (-1)^{\alpha-b} \frac{\binom{u-\beta}{b} \binom{u-b}{u-\alpha} \binom{m-u-\beta+b}{b} \left\{ \binom{m-u-\beta}{v-u} \binom{v-\beta}{v-u} \right\}^{1/2}}{\binom{v-u+b}{b}}$$

and

$$z_{\beta\alpha}^{(u,v)} = \phi_\beta z_{\beta\alpha}^{(u,v)} / \left\{ \binom{m}{u} \binom{u}{\alpha} \binom{m-u}{v-u+\alpha} \right\} \quad \text{for } 0 \leq u \leq v \leq [m/2].$$

Here

$$\phi_\beta = \binom{m}{\beta} - \binom{m}{\beta-1} \quad \text{for } \beta=0, 1, \dots, \min(u, v).$$

The matrices $A_\beta^{*(u,v)}$ have the following properties,

$$\begin{aligned} \sum_{\beta=0}^u A_\beta^{*(u,u)} &= I_{n_u}, \\ A_\beta^{*(u,w)} A_\tau^{*(w,v)} &= \delta_{\beta,\tau} A_\beta^{*(u,v)}, \\ \text{Rank}(A_\beta^{*(u,v)}) &= \phi_\beta, \end{aligned}$$

where $\delta_{a,b}$ denotes Kronecker's delta, i.e., $\delta_{a,b}=1$ or 0 according as $a=b$ or not. As mentioned above, we have

$$A_\alpha^{(i,j)} = \begin{cases} A_\alpha^{(i,j)} & \text{if } 0 \leq i, j \leq [m/2], \\ A_{\alpha^*}^{(i,m-j)} & \text{if } 0 \leq i \leq [m/2] < j \leq m, \\ A_{\alpha^{**}}^{(m-i,j)} & \text{if } 0 \leq j \leq [m/2] < i \leq m, \\ A_{\alpha^{***}}^{(m-i,m-j)} & \text{if } [m/2] < i, j \leq m. \end{cases}$$

Thus $A_\alpha^{(i,j)}$ ($0 \leq i, j \leq m; 0 \leq \alpha \leq \min(i, m-i, j, m-j)$) can be expressed as the linear combinations of $A_\beta^{*(*,*)}$ as in (A.2) and (A.3).

It is to be noted that the importance of the TMDPB association algebra \mathcal{A} generated by the ordered association matrices $D_\alpha^{(u,v)}$ and also generated by $D_\beta^{*(u,v)}$ has been discussed in the works of Yamamoto, Shirakura and Kuwada [16], and others. A few references are given above; for further information the readers are requested to see the references therein.