

A NOTE ON UNIFORM APPROXIMATION TO DISTRIBUTIONS
 OF EXTREME ORDER STATISTICS

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Summary

The error bound of an approximation to the distribution of the k th largest order statistic as established by Reiss (1981, *Adv. Appl. Prob.*, 13, 533-547) is improved by making use of an asymptotic expansion of length two in Reiss (the same as above).

1. Introduction

It is well known (see e.g. [1]) that the limit distribution of the k th largest order statistic $Z_{n-k+1:n}$ of a sample of size n is a gamma distribution P_k if the uniform distribution, say, Q on $[0, 1]$ is the underlying probability measure. The classical standardization for this limit law is given by $n(Z_{n-k+1:n} - 1)$. A result of Ikeda and Matsunawa [2] suggests that this result holds true uniformly over all Borel sets if $k \equiv k(n)$ fulfills the condition $k(n)/n \rightarrow 0$ as $n \rightarrow \infty$. Reiss [3] also proves asymptotic expansions arranged in powers of k/n . The asymptotic expansions of length 2 is given by

$$(1.1) \quad \left| Q^n \{n(Z_{n-k+1:n} - 1) \in B\} - \int_B [1 - ((x+k)^2 - k)/(2n - 3k)] dP_k(x) \right| \leq C_2 \left(\frac{k}{n}\right)^2$$

uniformly over all Borel sets B where C_2 is a universal constant and P_k has the Lebesgue-density p_k given by $p_k(x) = (|x|^{k-1}/(k-1)!) e^x$ for $x < 0$. It is immediate from (1.1) that for some universal constant $C_1 > 0$:

$$(1.2) \quad |Q^n \{n(Z_{n-k+1:n} - 1) \in B\} - P_k(B)| \leq C_1 \frac{k}{n}$$

uniformly over all Borel sets B . Moreover, it was proved in [3] that a different standardization leads to a more accurate approximation if

k is large. Combining both results we shall be able to improve the error bound of the second approximation by omitting a term $\exp(-k^{1/2}/4)$. This possibility was already suggested by the numerical computations in [3], p. 542.

2. The result

The main idea of the following approach is to choose constants $a_{n,k}$ and $b_{n,k}$ such that the first two moments of the distribution of $a_{n,k} \cdot (Z_{n-k+1:n} - b_{n,k})$ are accurate approximations to those of the approximating gamma distribution P_k . If $k \geq (4 \log n)^2$ then the following result is an immediate consequence of Theorem 2.9 in [3].

THEOREM 2.1. *There exists a constant $C > 0$ such that for every positive integer n and $k \in \{1, \dots, [n/2]\}$ the following inequality holds true:*

$$\left| Q^n \left\{ \left[\frac{n^{3/2}}{(n-k)^{1/2}} \left(Z_{n-k+1:n} - \frac{n-k}{n} \right) - k \right] \in B \right\} - P_k(B) \right| \leq C \frac{k^{1/2}}{n}$$

uniformly over all Borel sets B .

PROOF. It remains to prove the assertion for $k < (4 \log n)^2$. Set $p_{k,2}(x) = 1 - ((x+k)^2 - k)/(2n - 3k)$. Define $h(x) := (n/(n-k))^{1/2}x + k((n/(n-k))^{1/2} - 1)$ for every real number x . The inverse of h is $g(x) = ((n-k)/n)^{1/2}x - k(1 - ((n-k)/n)^{1/2})$.

Since $n^{3/2}(n-k)^{-1/2}(Z_{n-k+1:n} - (n-k)/n) - k = h(n(Z_{n-k+1:n} - 1))$ we obtain from (1.1) that

$$\begin{aligned} & \left| Q^n \left\{ \left[\frac{n^{3/2}}{(n-k)^{1/2}} \left(Z_{n-k+1:n} - \frac{n-k}{n} \right) - k \right] \in B \right\} - P_k(B) \right| \\ & \leq |P_{k,2}\{h \in B\} - P_k(B)| + O((k/n)^2) \end{aligned}$$

where $P_{k,2}$ is the approximating measure in (1.1) (with P_k -density $p_{k,2}$). By assumption we have $(k/n)^2 \leq Ck^{1/2}/n$. Moreover $|P_{k,2}\{h \in B\} - P_k(B)| \leq 1/2 \int |((n-k)/n)^{1/2}p_k(g(x))p_{k,2}(g(x)) - p_k(x)| dx$.

Notice that $g(x) > 0$ if $x \geq k((n/n-k))^{1/2} - 1$. We split the above integral into two parts, say I_1 and I_2 , namely from $-\infty$ to 0 and from 0 to $k((n/n-k))^{1/2} - 1$. Estimates for I_1 and I_2 will establish the assertion.

To show that $I_2 = O(k^{1/2}/n)$ is trivial and can be left to the reader. Ad I_1 : We have $((n-k)/n)^{1/2} = 1 - k/(2n) + O((k/n)^2)$. Hence we can replace in I_1 the expression $((n-k)/n)^{1/2}$ by $1 - k/(2n)$, $g(x)$ by $(1 - k/(2n))x - k^2/(2n)$ which is equal to $[x - k(x+k)/(2n)]$, and $e^{g(x)}$ and $|g(x)|^{k-1}$ by the

two leading terms in their series representations. The omitted terms are of order $O((k/n)^2)$. Thus we obtain for $k \geq 2$ (the case of $k=1$ is treated in an analogous way):

$$\begin{aligned} I_1 &= \int_{-\infty}^0 \left| \left(1 - \frac{k}{2n} \right) \left(|x|^{k-1} + \frac{(k-1)k(x+k)|x|^{k-2}}{2n} \right) \left(1 - \frac{k(x+k)}{2n} \right) \right. \\ &\quad \cdot \left. \left(1 - \frac{(x+k)^2 - k}{2n - 3k} + \frac{k(x+k)^2}{n(2n - 3k)} \right) - |x|^{k-1} \left| \frac{e^x}{(k-1)!} \right| dx + O\left(\left(\frac{k}{n} \right)^2 \right) \\ &= \frac{1}{2n(k-1)} \int_{-\infty}^0 |x(x+k)^2 + kx(x+k) + (k-1)k(x+k)| \\ &\quad \cdot \frac{|x|^{k-2} e^x}{(k-2)!} dx + O\left(\left(\frac{k}{n} \right)^2 \right). \end{aligned}$$

Thus an application of the Cauchy-Schwarz inequality yields (notice that the terms involving k^4 , k^5 and k^6 vanish):

$$I_1 \leq \frac{(10k^3 + 14k^2 - 24k)^{1/2}}{2n(k-1)} + O\left(\left(\frac{k}{n} \right)^2 \right) = O\left(\frac{k^{1/2}}{n} \right).$$

The proof is complete.

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