

ON THE UNIFORM COMPLETE CONVERGENCE OF
ESTIMATES FOR MULTIVARIATE DENSITY FUNCTIONS
AND REGRESSION CURVES

K. F. CHENG¹ AND R. L. TAYLOR²

(Received Feb. 5, 1979; revised Jan. 28, 1980)

Abstract

Let $(X_1, Y_1), \dots, (X_n, Y_n)$ be a random sample from the $(k+1)$ -dimensional multivariate density function $f^*(\mathbf{x}, y)$. Estimates of the k -dimensional density function $f(\mathbf{x}) = \int f^*(\mathbf{x}, y) dy$ of the form

$$\hat{f}_n(\mathbf{x}) = \frac{1}{nb_1(n) \cdots b_k(n)} \sum_{i=1}^n W\left(\frac{x_1 - X_{i1}}{b_1(n)}, \dots, \frac{x_k - X_{ik}}{b_k(n)}\right)$$

are considered where $W(\mathbf{x})$ is a bounded, nonnegative weight function and $b_1(n), \dots, b_k(n)$ and bandwidth sequences depending on the sample size and tending to 0 as $n \rightarrow \infty$. For the regression function

$$m(\mathbf{x}) = E(Y|X=\mathbf{x}) = \frac{h(\mathbf{x})}{f(\mathbf{x})}$$

where $h(\mathbf{x}) = \int y f^*(\mathbf{x}, y) dy$, estimates of the form

$$\hat{h}_n(\mathbf{x}) = \frac{1}{nb_1(n) \cdots b_k(n)} \sum_{i=1}^n Y_i W\left(\frac{x_1 - X_{i1}}{b_1(n)}, \dots, \frac{x_k - X_{ik}}{b_k(n)}\right)$$

are considered. In particular, uniform consistency of these estimates is obtained by showing that $\|\hat{f}_n(\mathbf{x}) - f(\mathbf{x})\|_\infty$ and $\|\hat{m}_n(\mathbf{x}) - m(\mathbf{x})\|_\infty$ converge completely to zero for a large class of "good" weight functions and under mild conditions on the bandwidth sequences $b_k(n)$'s.

¹ Research supported by National Institute of Environmental Health Sciences Under Grant 5T32 ES07011.

² For this author, revisions made while supported by the Air Force Office of Scientific Research under Contract Number F49620-79-C-0140.

Key words and phrases: Weight function, bandwidth sequences, regression function estimates, and complete convergence.

AMS Classification Number: Primary 62E15 and 62E40, Secondary 60F15.

1. Introduction

Let (\mathbf{X}, Y) be a $(k+1)$ -dimensional random vector with the joint probability density function $f^*(\mathbf{x}, y)$ and let \mathbf{X} be a k -dimensional random vector with the continuous marginal density function $f(\mathbf{x})$. Some modified multivariate density function estimates for f have been discussed by Cacoullos [2] which were based on a random sample $(\mathbf{X}_1, Y_1), (\mathbf{X}_2, Y_2), \dots, (\mathbf{X}_n, Y_n)$ from f^* . In particular, estimates $f_n(\mathbf{x})$ for the multivariate density function $f(\mathbf{x})$ of the form

$$(1.1) \quad f_n(\mathbf{x}) = \frac{1}{nb^k(n)} \sum_{i=1}^n W\left(\frac{\mathbf{x} - \mathbf{X}_i}{b(n)}\right)$$

were considered where $W(\mathbf{x})$ is a bounded, nonnegative, integrable weight function such that

$$\int_{R^k} W(\mathbf{x}) d\mathbf{x} = 1$$

and $b(n)$ is a bandwidth sequence depending on the sample size and tending to 0 as $n \rightarrow \infty$.

For the regression function $m(\mathbf{x})$ of Y on \mathbf{X} ,

$$m(\mathbf{x}) = E(Y | \mathbf{X} = \mathbf{x}) = \frac{h(\mathbf{x})}{f(\mathbf{x})}$$

where $h(\mathbf{x}) = \int y f^*(\mathbf{x}, y) dy$, Watson [9] and Nadaraya [5] independently proposed the following regression function estimates for the case $k=1$:

$$(1.2) \quad m_n(\mathbf{x}) = \frac{h_n(\mathbf{x})}{f_n(\mathbf{x})} = \frac{\sum_{i=1}^n Y_i \frac{W\left(\frac{\mathbf{x} - \mathbf{X}_i}{b(n)}\right)}{\sum_{j=1}^n W\left(\frac{\mathbf{x} - \mathbf{X}_j}{b(n)}\right)}.$$

Thus, the estimates for $h(\mathbf{x})$ are

$$(1.3) \quad h_n(\mathbf{x}) = \frac{1}{nb^k(n)} \sum_{i=1}^n Y_i W\left(\frac{\mathbf{x} - \mathbf{X}_i}{b(n)}\right).$$

An heuristic treatment of $m_n(\mathbf{x})$ as a weighted average of the Y_i 's can be found in Watson [9]. The local properties of (1.1) and (1.2) have been studied extensively (see Rosenblatt [7]), and global measurements of deviations of $m_n(\mathbf{x})$ from $m(\mathbf{x})$ and $f_n(\mathbf{x})$ from $f(\mathbf{x})$ are given by

$$(1.4) \quad \|m_n(\mathbf{x}) - m(\mathbf{x})\|_\infty = \sup_{\mathbf{x} \in R^k} |m_n(\mathbf{x}) - m(\mathbf{x})|$$

and

$$(1.5) \quad \|f_n(\mathbf{x}) - f(\mathbf{x})\|_\infty = \sup_{\mathbf{x} \in R^k} |f_n(\mathbf{x}) - f(\mathbf{x})|.$$

In this paper the important large sample properties of the more general estimates $\hat{f}_n(\mathbf{x})$ and $\hat{m}_n(\mathbf{x})$ are explored where

$$(1.6) \quad \hat{f}_n(\mathbf{x}) = \frac{1}{nb_1(n) \cdots b_k(n)} \sum_{i=1}^n W\left(\frac{x_1 - X_{i1}}{b_1(n)}, \dots, \frac{x_k - X_{ik}}{b_k(n)}\right),$$

$$(1.7) \quad \hat{m}_n(\mathbf{x}) = \frac{\hat{h}_n(\mathbf{x})}{\hat{f}_n(\mathbf{x})},$$

and

$$(1.8) \quad \hat{h}_n(\mathbf{x}) = \frac{1}{nb_1(n) \cdots b_k(n)} \sum_{i=1}^n Y_i W\left(\frac{x_1 - X_{i1}}{b_1(n)}, \dots, \frac{x_k - X_{ik}}{b_k(n)}\right).$$

Here, $W(\mathbf{x})$ is a nonnegative, bounded weight function and the bandwidth sequence $b_j(n)$'s depends on the sample size n and tends to zero as $n \rightarrow \infty$. The rate of convergence to zero need not be uniform, and the possible weight functions include more than the uniform kernels. Attention will generally be restricted to a density function $f^*(\mathbf{x}, y)$ with compact support $[a_1, b_1] \times [a_2, b_2] \times [a_3, b_3]$ (for notation convenience only $k=2$) and weight functions $W(\mathbf{x})$ which satisfy

$$\int_{a'_1}^{b'_1} \int_{a'_2}^{b'_2} W(\mathbf{x}) d\mathbf{x} = 1$$

and vanish outside $[a'_1, b'_1] \times [a'_2, b'_2]$. Also, the continuity of $f(\mathbf{x})$ and $h(\mathbf{x})$ on compact support $[a_1, b_1] \times [a_2, b_2]$ and on $[a'_1, b'_1] \times [a'_2, b'_2]$ are assumed. However the compact support of $f(\mathbf{x})$ can be eliminated (Lemma 2) when $f(\mathbf{x})$ has a p th moment, $p > 0$.

The major results of this paper give a new class of "good" weight functions under mild conditions on the bandwidth sequences $b_i(n)$'s. Uniform consistency of the estimates $\hat{f}_n(\mathbf{x})$ and $\hat{m}_n(\mathbf{x})$ is obtained since $\|\hat{f}_n(\mathbf{x}) - f(\mathbf{x})\|_\infty$ and $\|\hat{m}_n(\mathbf{x}) - m(\mathbf{x})\|_\infty$ converge completely to zero (which implies convergence with probability one). The main tools in obtaining these results will be approximating polyhedral functions and sub-Gaussian techniques and will parallel the development in Taylor and Cheng [8].

The modulus of continuity, $\omega_\rho(\delta_1, \delta_2)$, is defined by Billingsley [1] as

$$\omega_\rho(\delta_1, \delta_2) = \sup_{\substack{|t_1 - s_1| \leq \delta_1 \\ |t_2 - s_2| \leq \delta_2}} |g(s_1, s_2) - g(t_1, t_2)|$$

for $\delta_1, \delta_2 > 0$; $(s_1, s_2), (t_1, t_2) \in [a_1, b_1] \times [a_2, b_2]$; and $g \in C([a_1, b_1] \times [a_2, b_2])$, the space of continuous functions with domain on $[a_1, b_1] \times [a_2, b_2]$.

DEFINITION (Chow [3]). A random variable X is said to be *sub-Gaussian* if there exists $\alpha \geq 0$ such that

$$(1.9) \quad \mathbb{E} [\exp(tX)] \leq \exp\left(\frac{\alpha^2 t^2}{2}\right) \quad \text{for all } t \in \mathbb{R}.$$

If X is sub-Gaussian, then let

$$\tau(X) = \inf \{ \alpha \geq 0 : \text{Inequality (1.9) holds} \}.$$

Some basic properties of sub-Gaussian random variables which will be used include:

1. If $\mathbb{P} [|X| \leq K] = 1$ and $\mathbb{E} X = 0$, then

$$\mathbb{E} [\exp(tX)] \leq \exp(K^2 t^2).$$

2. If $\tau(X) = \alpha$, then

$$(1.10) \quad \mathbb{P} [|X| \geq \lambda] \leq 2 \exp(-\lambda^2/2\alpha^2).$$

3. The sum of two independent sub-Gaussian random variables is sub-Gaussian.

A sequence of random variables $\{X_n\}$ is said to *converge completely* to a random variable X if

$$(1.11) \quad \sum_{n=1}^{\infty} \mathbb{P} [|X_n - X| > \varepsilon] < \infty$$

for each $\varepsilon > 0$.

2. Main results

In this section it is shown that $\|\hat{f}_n(\mathbf{x}) - f(\mathbf{x})\|_{\infty}$ and $\|\hat{m}_n(\mathbf{x}) - m(\mathbf{x})\|_{\infty}$ converge completely to zero under conditions on the modulus of continuity of the weight function $W(\mathbf{x})$ and the rate of convergence to zero by the bandwidth sequences $b_j(n)$'s.

LEMMA 1. If (i) $nb_1^2(n)b_2^2(n) > n^{\delta}$ for some $\delta > 0$,

(ii) $\omega_W\left(\frac{(b_1+b'_1)-(a_1+a'_1)}{n^{r_1}b_1(n)}, \frac{(b_2+b'_2)-(a_2+a'_2)}{n^{r_2}b_2(n)}\right) = o(b_1(n)b_2(n))$ for some integers $r_1, r_2 > 0$, and (iii) $f(\mathbf{x})$ has a compact support $[a_1, b_1] \times [a_2, b_2]$, then

$$(2.1) \quad \sup_{\mathbf{s} \in \mathbb{R}^2} \left| \hat{f}_n(\mathbf{s}) - \frac{1}{b_1(n)b_2(n)} \mathbb{E} W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$ where $\hat{f}_n(\mathbf{s})$ is defined in (1.6).

PROOF. First, $[a'_1, b'_1] \times [a'_2, b'_2]$ may be expanded to include $(0, 0)$ if

$(0, 0) \notin [a'_1, b'_1] \times [a'_2, b'_2]$. For the positive integers r_1 and r_2 , let $I_{si} = [t_{s,i-1}, t_{si}]$, $s=1$ or 2 , where $t_{si} = (a_s + a'_s) + [(b_s + b'_s) - (a_s + a'_s)]i/n^{r_s}$ and let $I_{ij} = I_{i1} \times I_{2j}$. Thus,

$$[a_1 + a'_1, b_1 + b'_1] \times [a_2 + a'_2, b_2 + b'_2] = \bigcup_{i=1}^{n^{r_1}} \bigcup_{j=1}^{n^{r_2}} I_{ij}.$$

Since W and f vanish outside $[a'_1, b'_1] \times [a'_2, b'_2]$ and $[a_1, b_1] \times [a_2, b_2]$ respectively and $b_1(n) \rightarrow 0$ and $b_2(n) \rightarrow 0$, the sup in (2.1) need only be taken over $[a_1 + a'_1, b_1 + b'_1] \times [a_2 + a'_2, b_2 + b'_2]$. Let $\delta_n^s = \frac{(b_s + b'_s) - (a_s + a'_s)}{n^{r_s} b_s(n)}$ for $s=1$ or 2 and let

$$\tilde{W}_k(s_1, s_2) = W\left(s_1 - \frac{X_{k1}}{b_1(n)}, s_2 - \frac{X_{k2}}{b_2(n)}\right) - E W\left(s_1 - \frac{X_{k1}}{b_1(n)}, s_2 - \frac{X_{k2}}{b_2(n)}\right)$$

for each $k=1, \dots, n$. Thus, $E \tilde{W}_k(\mathbf{s}) = 0$ for each $\mathbf{s} \in [a_1 + a'_1, b_1 + b'_1] \times [a_2 + a'_2, b_2 + b'_2]$ and each k . Furthermore,

$$\begin{aligned} (2.2) \quad \omega_{\tilde{W}_k}(\delta_n^1, \delta_n^2) &= \sup_{\substack{|s_1 - t_1| \leq \delta_n^1 \\ |s_2 - t_2| \leq \delta_n^2}} |\tilde{W}_k(s_1, s_2) - \tilde{W}_k(t_1, t_2)| \\ &\leq \sup_{\substack{|s_1 - t_1| \leq \delta_n^1 \\ |s_2 - t_2| \leq \delta_n^2}} \left| W\left(s_1 - \frac{X_{k1}}{b_1(n)}, s_2 - \frac{X_{k2}}{b_2(n)}\right) \right. \\ &\quad \left. - W\left(t_1 - \frac{X_{k1}}{b_1(n)}, t_2 - \frac{X_{k2}}{b_2(n)}\right) \right| \\ &\quad + \sup_{\substack{|s_1 - t_1| \leq \delta_n^1 \\ |s_2 - t_2| \leq \delta_n^2}} \left| E W\left(s_1 - \frac{X_{k1}}{b_1(n)}, s_2 - \frac{X_{k2}}{b_2(n)}\right) \right. \\ &\quad \left. - E W\left(t_1 - \frac{X_{k1}}{b_1(n)}, t_2 - \frac{X_{k2}}{b_2(n)}\right) \right| \\ &\leq 2\omega_W(\delta_n^1, \delta_n^2). \end{aligned}$$

Hence, $\omega_{\tilde{W}_k}(\delta_n^1, \delta_n^2) \leq 2\omega_W(\delta_n^1, \delta_n^2) = o(b_1(n)b_2(n))$ for each k from Condition (ii). For $\varepsilon > 0$ let

$$\begin{aligned} (2.3) \quad A_n &= \left[\sup_{\substack{a_1 + a'_1 \leq s_1 \leq b_1 + b'_1 \\ a_2 + a'_2 \leq s_2 \leq b_2 + b'_2}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k\left(\frac{s_1}{b_1(n)}, \frac{s_2}{b_2(n)}\right) \right| > \varepsilon \right] \\ &= \left[\max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \sup_{(s_1, s_2) \in I_{ij}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k\left(\frac{s_1}{b_1(n)}, \frac{s_2}{b_2(n)}\right) \right| > \varepsilon \right]. \end{aligned}$$

Hence,

$$(2.4) \quad A_n \subset \left[\max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k\left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)}\right) \right| \right]$$

$$+ \max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \sup_{(s_1, s_2) \in I_{ij}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \left[\tilde{W}_k \left(\frac{s_1}{b_1(n)}, \frac{s_2}{b_2(n)} \right) - \tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) \right] \right| > \varepsilon \Big].$$

However,

$$\begin{aligned} & \max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \sup_{(s_1, s_2) \in I_{ij}} \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \left| \tilde{W}_k \left(\frac{s_1}{b_1(n)}, \frac{s_2}{b_2(n)} \right) - \tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) \right| \\ & \leq \frac{2\omega_W(\delta_n^1, \delta_n^2)}{b_1(n)b_2(n)}. \end{aligned}$$

Since $2\omega_W(\delta_n^1, \delta_n^2) = o(b_1(n)b_2(n))$ by Condition (ii), there exists $N(r_1, r_2)$ such that

$$A_n \subset \left[\max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) \right| > \frac{\varepsilon}{2} \right]$$

for all $n \geq N(r_1, r_2)$. Using the basic properties of sub-Gaussian random variables $\left\{ \left[\tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) : k=1, 2, \dots, n \right] \text{ for each } i, j \right\}$, for each $n \geq N(r_1, r_2)$

$$\begin{aligned} \mathbb{P}[A_n] & \leq \mathbb{P} \left[\max_{\substack{1 \leq i \leq n^{r_1} \\ 1 \leq j \leq n^{r_2}}} \left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) \right| > \frac{\varepsilon}{2} \right] \\ & \leq \sum_{i=1}^{n^{r_1}} \sum_{j=1}^{n^{r_2}} \mathbb{P} \left[\left| \frac{1}{nb_1(n)b_2(n)} \sum_{k=1}^n \tilde{W}_k \left(\frac{t_{1i}}{b_1(n)}, \frac{t_{2j}}{b_2(n)} \right) \right| > \frac{\varepsilon}{2} \right] \\ & \leq n^{r_1+r_2} 2 \exp(-\varepsilon^2/4 \|W\|_\infty^2 B_n) \end{aligned}$$

where

$$\|W\|_\infty = \sup_{(s_1, s_2) \in R^2} |W(s_1, s_2)| \quad \text{and} \quad B_n = \sum_{k=1}^n \left(\frac{1}{nb_1(n)b_2(n)} \right)^2 = \frac{1}{nb_1^2(n)b_2^2(n)}.$$

To obtain the complete convergence of (2.1), consider

$$\begin{aligned} (2.5) \quad \sum_{n=1}^{\infty} \mathbb{P}[A_n] & = \sum_{n=1}^{N(r_1, r_2)} \mathbb{P}[A_n] + \sum_{n=N(r_1, r_2)+1}^{\infty} \mathbb{P}[A_n] \\ & \leq N(r_1, r_2) + \sum_{n=N(r_1, r_2)+1}^{\infty} 2n^{r_1+r_2} \exp \left(\frac{-\varepsilon^2 nb_1^2(n)b_2^2(n)}{4 \|W\|_\infty^2} \right) \\ & \leq N(r_1, r_2) + \sum_{n=N(r_1, r_2)+1}^{\infty} 2n^{r_1+r_2} \exp(-cn^s) \end{aligned}$$

where $c = \varepsilon^2/4 \|W\|_\infty^2$. Thus the series in (2.5) converges by the integral test.

The compact support of the $f(x)$ can be relaxed if any moment $p > 0$ exists. Lemma 2 summarizes this result and again is stated only for $k=2$. The proof of Lemma 2 is similar to the univariate case in Taylor and Cheng [8] and is omitted. Recall that $\|(a, b)\| = (a^2 + b^2)^{1/2}$.

LEMMA 2. If (i) $nb_1^2(n)b_2^2(n) > n^p$ for some $\delta > 0$, (ii) $\int \|x\|^p f(x) dx < \infty$ for some $p > 0$, and (iii) $W\left(\frac{2+(b'_1-a'_1)}{n^{r_1}b_1(n)}, \frac{2+(b'_2-a'_2)}{n^{r_2}b_2(n)}\right) = o(b_1(n)b_2(n))$ for some integers $r_1, r_2 > 1/p$, then

$$\sup_{s \in R^2} \left| \hat{f}_n(s_1, s_2) - \frac{1}{b_1(n)b_2(n)} E W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$ where $\hat{f}_n(s)$ is defined in (1.6).

LEMMA 3. If the underlying density, f , is uniformly continuous, then

$$(2.6) \quad \sup_{(s_1, s_2) \in R^2} \left| \frac{1}{b_1(n)b_2(n)} E W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) - f(s_1, s_2) \right| \rightarrow 0.$$

PROOF. Since f is uniformly continuous on R^2 , given $\epsilon > 0$ there exists $\delta > 0$ such that $|f(s'_1, s'_2) - f(s_1, s_2)| < \epsilon$ whenever $\|(s'_1, s'_2) - (s_1, s_2)\| < \delta$. Let N be sufficiently large so that $\|(b_1(n)y_1, b_2(n)y_2)\| < \delta$ for all $n \geq N$ and all $(y_1, y_2) \in [a'_1, b'_1] \times [a'_2, b'_2]$. Since $W(y_1, y_2) = 0$ for $(y_1, y_2) \notin [a'_1, b'_1] \times [a'_2, b'_2]$,

$$(2.7) \quad \begin{aligned} & \left| \frac{1}{b_1(n)b_2(n)} E W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) - f(s_1, s_2) \right| \\ &= \left| \frac{1}{b_1(n)b_2(n)} \int_{R^2} W\left(\frac{s_1 - x_1}{b_1(n)}, \frac{s_2 - x_2}{b_2(n)}\right) f(x_1, x_2) dx_1 dx_2 - f(s_1, s_2) \right| \\ &= \left| \int_{R^2} W(y_1, y_2) f(s_1 - b_1(n)y_1, s_2 - b_2(n)y_2) dy_1 dy_2 - f(s_1, s_2) \right| \\ &= \left| \int_{R^2} W(y_1, y_2) [f(s_1 - b_1(n)y_1, s_2 - b_2(n)y_2) - f(s_1, s_2)] dy_1 dy_2 \right| \\ &= \epsilon \int_{R^2} W(y_1, y_2) dy_1 dy_2 = \epsilon \end{aligned}$$

uniformly in (s_1, s_2) for all $n \geq N$. Hence

$$\sup_{s \in R^2} \left| \frac{1}{b_1(n)b_2(n)} E W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) - f(s_1, s_2) \right| \rightarrow 0$$

as $n \rightarrow \infty$.

If the density, f , is continuous on R^2 and has compact support, then it is uniformly continuous and all p th moments exist. Conditions

(ii) and (iii) of Lemma 2 are easily satisfied, and the more general hypotheses of Lemmas 2 and 3 are listed for Theorem 1. First, it should be indicated that the case where f is only known to be continuous on $[a_1, b_1] \times [a_2, b_2]$ is not entirely excluded from consideration.

COROLLARY 1. *If the underlying density function, f , is only known to be continuous on $[a_1, b_1] \times [a_2, b_2]$, then for arbitrarily small $\varepsilon_1, \varepsilon_2 > 0$*

$$\sup_{\substack{a_1 + \varepsilon_1 \leq s_1 \leq b_1 - \varepsilon_1 \\ a_2 + \varepsilon_2 \leq s_2 \leq b_2 - \varepsilon_2}} \left| \frac{1}{b_1(n)b_2(n)} \mathbb{E} W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) - f(s_1, s_2) \right| \rightarrow 0$$

as $n \rightarrow \infty$.

The ε 's in Corollary 1 and Corollary 2 can be considered as functions of n which tend to zero as $n \rightarrow \infty$. The proof of the following theorem is immediate from Lemmas 2 and 3 since for each $\varepsilon > 0$

$$\begin{aligned} (2.8) \quad & \mathbb{P} \left[\sup_{\mathbf{s} \in R^2} \left| \frac{1}{b_1(n)b_2(n)} \sum_{k=1}^n W\left(\frac{s_1 - X_{k1}}{b_1(n)}, \frac{s_2 - X_{k2}}{b_2(n)}\right) - f(s_1, s_2) \right| > \varepsilon \right] \\ & \leq \mathbb{P} \left[\sup_{\mathbf{s} \in R^2} \left| \frac{1}{b_1(n)b_2(n)} \sum_{k=1}^n \left[W\left(\frac{s_1 - X_{k1}}{b_1(n)}, \frac{s_2 - X_{k2}}{b_2(n)}\right) \right. \right. \right. \\ & \qquad \qquad \qquad \left. \left. \left. - \mathbb{E} W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) \right] \right| > \frac{\varepsilon}{2} \right] \\ & \leq \mathbb{P} \left[\sup_{\mathbf{s} \in R^2} \left| \frac{1}{b_1(n)b_2(n)} \mathbb{E} W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) - f(s_1, s_2) \right| > \frac{\varepsilon}{2} \right] \end{aligned}$$

and each of the terms in (2.8) is a convergent series in n . All of the conditions will be stated in Theorem 1 for easy reference, and in particular will be stated for arbitrary k -dimension.

THEOREM 1. *Let $\{X_n\}$ be independent random vectors with the same density function $f(\mathbf{x})$ which is uniformly continuous on R^k . Let $W(\mathbf{x})$ be a nonnegative weight function which is continuous on its compact support $[a'_1, b'_1] \times \dots \times [a'_k, b'_k]$ and integrates to 1. If (a) $nb_1^2(n) \dots b_k^2(n) > n^\delta$ for some $\delta > 0$, (b) $\int \|\mathbf{x}\|^p f(\mathbf{x}) d\mathbf{x} < \infty$ for some $p > 0$, and (c) $\omega_W\left(\frac{2+(b'_1-a'_1)}{n^{r_1}b_1(n)}, \dots, \frac{2+(b'_k-a'_k)}{n^{r_k}b_k(n)}\right) = o(b_1(n) \dots b_k(n))$ for some integers $r_1, \dots, r_k > 1/p$, then*

$$\sup_{\mathbf{s} \in R^k} \left| \frac{1}{nb_1(n) \dots b_k(n)} \sum_{i=1}^n W\left(\frac{s_1 - X_{i1}}{b_1(n)}, \dots, \frac{s_k - X_{ik}}{b_k(n)}\right) - f(s_1, \dots, s_k) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$.

COROLLARY 2. *If all of the conditions of Theorem 1 are satisfied*

except it is only known that f is continuous on its compact support, then for arbitrarily small $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_k > 0$,

$$\sup_{\mathbf{a} + \boldsymbol{\varepsilon} \leq \mathbf{s} \leq \mathbf{b} - \boldsymbol{\varepsilon}} \left| \frac{1}{nb_1(n) \cdots b_k(n)} \sum_{i=1}^n W\left(\frac{s_1 - X_{i1}}{b_1(n)}, \dots, \frac{s_k - X_{ik}}{b_k(n)}\right) - f(s_1, \dots, s_k) \right| \rightarrow 0$$

completely as $n \rightarrow \infty$ where $\mathbf{a} + \boldsymbol{\varepsilon} \leq \mathbf{s} \leq \mathbf{b} - \boldsymbol{\varepsilon}$ indicates that the relationship holds in each coordinate.

The condition $nb_1^2(n) \cdots b_k^2(n) > n^d$ need not hold for all n but only eventually. Also, the condition could have been stated as

$$\int_d^\infty x^{r_1 + \dots + r_k} \exp(-cxb_1^2(x) \cdots b_k^2(x)) dx < \infty \quad \text{for some } d > 0$$

where $b_i(x)$'s are functions which generate the bandwidth sequences $b_i(1), b_i(2), \dots, i=1, \dots, k$, and $c > 0$ is a constant.

Attention is now directed to establishing the uniform complete convergence of $\|\hat{m}_n(\mathbf{x}) - m(\mathbf{x})\|_\infty$ to zero. Again, for notational convenience $k=2$.

LEMMA 4. *If the regularity conditions in Lemma 1 are satisfied, then*

$$\begin{aligned} \sup_{\mathbf{s} \in R^2} & \left| \frac{1}{nb_1(n)b_2(n)} \sum_{i=1}^n Y_i W\left(\frac{s_1 - X_{i1}}{b_1(n)}, \frac{s_2 - X_{i2}}{b_2(n)}\right) \right. \\ & \left. - \frac{1}{b_1(n)b_2(n)} E Y_1 W\left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)}\right) \right| \rightarrow 0 \end{aligned}$$

completely as $n \rightarrow \infty$.

PROOF. The proof follows similarly to the proof of Lemma 1. First, define

$$\tilde{W}_i^*(s_1, s_2) = Y_i W\left(s_1 - \frac{X_{i1}}{b_1(n)}, s_2 - \frac{X_{i2}}{b_2(n)}\right) - E\left(Y_1 W\left(s_1 - \frac{X_{11}}{b_1(n)}, s_2 - \frac{X_{12}}{b_2(n)}\right)\right)$$

where

$$\begin{aligned} & E\left(Y_1 W\left(s_1 - \frac{X_{11}}{b_1(n)}, s_2 - \frac{X_{12}}{b_2(n)}\right)\right) \\ & = \int_{R^3} y W\left(s_1 - \frac{x_1}{b_1(n)}, s_2 - \frac{x_2}{b_2(n)}\right) f^*(x_1, x_2, y) dx_1 dx_2 dy. \end{aligned}$$

Thus, $E W_i^*(s_1, s_2) = 0$ for all s_1, s_2 and i and $\omega_{W_i^*}(\delta_n^1, \delta_n^2) \leq 2M\omega_W(\delta_n^1, \delta_n^2)$ where $\|Y_i\| \leq M$ a.s. Hence, $\omega_{\tilde{W}_i^*}(\delta_n^1, \delta_n^2) = o(b_1(n)b_2(n))$ a.s., and the remainder of the proof follows from the proof of Lemma 1 with the

new constant being $c' = \frac{\varepsilon^2}{4M^2 \|W\|_\infty^2}$.

A moment condition again could be used to relax the assumption of compact support for Y with the modifications being similar to Lemma 3 but will not be stated. The next result will establish the convergence for the mean.

LEMMA 5. *If $h(\mathbf{x})$ is uniformly continuous on R^2 , then*

$$(2.9) \quad \sup_{\mathbf{s} \in R^2} \left| \frac{1}{b_1(n)b_2(n)} \mathbb{E} Y_1 W \left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)} \right) - h(s_1, s_2) \right| \rightarrow 0$$

as $n \rightarrow \infty$.

PROOF. First,

$$(2.10) \quad \begin{aligned} & \frac{1}{b_1(n)b_2(n)} \mathbb{E} Y_1 W \left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)} \right) \\ &= \frac{1}{b_1(n)b_2(n)} \int_{R^2} y W \left(\frac{s_1 - x_1}{b_1(n)}, \frac{s_2 - x_2}{b_2(n)} \right) f^*(x_1, x_2, y) dx_1 dx_2 dy \\ &= \frac{1}{b_1(n)b_2(n)} \int_{R^2} \mathbb{E} \left(Y_1 W \left(\frac{s_1 - x_1}{b_1(x)}, \frac{s_2 - x_2}{b_2(x)} \right) \middle| X_{11} = x_1, X_{12} = x_2 \right) \\ & \quad \cdot f(x_1, x_2) dx_1 dx_2 \\ &= \frac{1}{b_1(n)b_2(n)} \int_{R^2} W \left(\frac{s_1 - x_1}{b_1(n)}, \frac{s_2 - x_2}{b_2(n)} \right) h(x_1, x_2) dx_1 dx_2 \\ &= \int_{R^2} W(y_1, y_2) h(s_1 - b_1(n)y_1, s_2 - b_2(n)y_2) dy_1 dy_2. \end{aligned}$$

Next, N can be chosen large enough so that $\|(b_1(n)y_1, b_2(n)y_2)\| < \delta$ for all $(y_1, y_2) \in [a_1, b_1] \times [a_2, b_2]$ when $n \geq N$. Hence, from the uniform continuity of h and (2.10),

$$\begin{aligned} & \left| \frac{1}{b_1(n)b_2(n)} \mathbb{E} Y_1 W \left(\frac{s_1 - X_{11}}{b_1(n)}, \frac{s_2 - X_{12}}{b_2(n)} \right) - h(s_1, s_2) \right| \\ &= \left| \int_{R^2} W(y_1, y_2) [h(s_1 - b_1(n)y_1, s_2 - b_2(n)y_2) - h(s_1, s_2)] dy_1 dy_2 \right| \\ &\leq \varepsilon \int_{R^2} W(y_1, y_2) dy_1 dy_2 = \varepsilon \text{ uniformly in } (s_1, s_2) \end{aligned}$$

for all $n \geq N$.

The following theorem on the complete uniform consistency of \hat{h}_n can be proved immediately from Lemmas 4 and 5.

THEOREM 2. *If the regularity conditions of Lemmas 4 and 5 hold, then*

$$(2.11) \quad \sup_{\mathbf{s} \in R^2} |\hat{h}_n(s_1, s_2) - h(s_1, s_2)| \rightarrow 0$$

completely as $n \rightarrow \infty$.

COROLLARY 3. *If the regularity conditions of Lemma 4 are assumed and $h(s_1, s_2)$ is continuous on its compact support $[a_1, b_1] \times [a_2, b_2]$ only, then for arbitrarily small $\varepsilon_1, \varepsilon_2 > 0$*

$$\sup_{(s_1, s_2) \in [a_1 + \varepsilon_1, b_1 - \varepsilon_1] \times [a_2 + \varepsilon_2, b_2 - \varepsilon_2]} |\hat{h}_n(s_1, s_2) - h(s_1, s_2)| \rightarrow 0$$

completely as $n \rightarrow \infty$.

The stage is now set to obtain the complete uniform consistency of \hat{m}_n to m .

THEOREM 3. *If the regularity conditions of Theorem 1, Lemma 4 and Lemma 5 are satisfied and if there exist $\varepsilon_0^1, \varepsilon_0^2 > 0$ such that in $[a_1 + \varepsilon_0^1, b_1 - \varepsilon_0^1] \times [a_2 + \varepsilon_0^2, b_2 - \varepsilon_0^2] = C$ $\inf_{s \in C} f(s_1, s_2) = \mu > 0$ and $\sup_{s \in C} |m(s_1, s_2)| = v < \infty$, then*

$$(2.12) \quad \sup_{s \in C} |\hat{m}_n(s_1, s_2) - m(s_1, s_2)| \rightarrow 0$$

completely as $n \rightarrow \infty$.

PROOF. First

$$\begin{aligned} (2.13) \quad & \sup_{s \in C} |\hat{m}_n(s_1, s_2) - m(s_1, s_2)| \\ & \leq \sup_{s \in C} \left| \frac{\hat{h}_n(s_1, s_2)}{\hat{f}_n(s_1, s_2)} - \frac{h(s_1, s_2)}{f(s_1, s_2)} \right| + \sup_{s \in C} \left| \frac{h(s_1, s_2)}{\hat{f}_n(s_1, s_2)} - \frac{h(s_1, s_2)}{f(s_1, s_2)} \right| \\ & \leq (\inf_{s \in C} \hat{f}_n(s_1, s_2))^{-1} |\hat{h}_n(s_1, s_2) - h(s_1, s_2)| + \sup_{s \in C} \left| \frac{h(s_1, s_2)}{\hat{f}_n(s_1, s_2)} - \frac{h(s_1, s_2)}{f(s_1, s_2)} \right| \\ & \leq (\inf_{s \in C} \hat{f}_n(s_1, s_2))^{-1} \{ |\hat{h}_n(s_1, s_2) - h(s_1, s_2)| \\ & \quad + \sup_{s \in C} |m(s_1, s_2)| \sup_{s \in C} |\hat{f}_n(s_1, s_2) - f(s_1, s_2)| \}. \end{aligned}$$

Since $\inf_{s \in C} f(s_1, s_2) = \mu > 0$ and $\|\hat{f}_n - f\|_\infty \rightarrow 0$ completely by Theorem 1, the result easily follows from Lemmas 4 and 5 and (2.13).

3. Comparisons and useful weight functions

In this part, a few brief comments on Nadaraya's [6] conditions and on useful weight functions which satisfy the results of this paper are listed for comparisons.

To obtain a strong law rather than uniform consistency in probability, the conditions on the weight function and bandwidth sequences are expected to be more stringent. For example in the case $k=1$, let $f(s)$ and $m(x)$ be unknown continuous density and regression functions on R . If

(N1) $W(x)$ is a function of bounded variation such that

$$\begin{aligned} \sup_{x \in R} |W(x)| < \infty, \quad \lim_{x \rightarrow \pm\infty} |xW(x)| = 0, \\ \int |W(x)| dx < \infty, \quad \int W(x) dx = 1; \end{aligned}$$

(N2) $-\infty < A \leq Y \leq B < \infty$ with probability one and

$$\min_{-\infty < a \leq x \leq b < \infty} f(x) = \mu > 0; \text{ and}$$

(N3) $\sum_{n=1}^{\infty} \exp(-\gamma nb^2(n))$ exists for each $\gamma > 0$, then

$\|\hat{m}_n(x) - m(x)\|_{\infty} \rightarrow 0$ with probability one (Nadaraya [6]).

For the results of this paper, the weight function $W(x)$ was required to be nonnegative and continuous on its compact support along with a smoothness condition. The condition on the bandwidths sequence for the case $k=1$ reduces to $nb^2(n) > n^\delta$ for some $\delta > 0$. For useful weight functions, Epanechnikov [4] considered multivariate density function estimates of the form

$$f_n(s_1, \dots, s_k) = \frac{1}{n} \sum_{i=1}^n \prod_{j=1}^k \frac{1}{b_j(n)} W_j\left(\frac{s_j - X_{ij}}{b_j(n)}\right).$$

Setting $b_j(n) = b(n)$ and $W_j(s) = W(s)$ for $j=1, \dots, k$, the optimal weight function $W_0(s)$ was found to be

$$W_0(s) = \begin{cases} \frac{3}{4(5)^{1/2}} - \frac{3s^2}{20(5)^{1/2}} & \text{for } |s| \leq (5)^{1/2} \\ 0 & \text{otherwise} \end{cases}$$

in minimizing the relative global error. For this case, let $a = -(5)^{1/2}$ and $b = (5)^{1/2}$. Then, $W_0(s_1, s_2) = W_0(s_1)W_0(s_2)$, and $|W_0(s_1, s_2) - W_0(t_1, t_2)| \leq c_1|s_1 - t_1| + c_2|s_2 - t_2|$ for constants c_1 and c_2 . Moreover, $\omega_{w_0}(\delta_n^1, \delta_n^2) \leq \frac{c'_1}{b_1(n)n^{r_1}} + \frac{c'_2}{b_2(n)n^{r_2}}$ for constants c'_1 and c'_2 , and the optimal weight function easily satisfies the smoothness condition of this paper.

If the weight function $W(s_1, s_2)$ satisfies a Lipschitz condition of order α , then

$$\begin{aligned} |W(s_1, s_2) - W(t_1, t_2)| &\leq M \| (s_1, s_2) - (t_1, t_2) \|^\alpha \text{ and} \\ \omega_w\left(\frac{2(b_1 - a_1)}{b_1(n)n^{r_1}}, \frac{2(b_2 - a_2)}{b_2(n)n^{r_2}}\right) &\leq M \left(\frac{(2b_1 - 2a_1)^2}{b_1^2(n)n^{2r_1}} + \frac{(2b_2 - 2a_2)^2}{b_2^2(n)n^{2r_2}} \right)^{\alpha/2} \end{aligned}$$

for some $M > 0$. Hence, bandwidth sequences $b_1(n)$ and $b_2(n)$ are to be

chosen so that

$$\frac{1}{b_1(n)b_2(n)} \left[\frac{1}{n^{r_1}b_1(n)} + \frac{1}{n^{r_2}b_2(n)} \right]^\alpha \rightarrow 0$$

as $n \rightarrow \infty$ for some integers $r_1, r_2 > 0$. Let $b_1(n) = n^{-p_1}$ and $b_2(n) = n^{-p_2}$ for $p_1 > p_2 > 0$ and let $r_1 = r_2 = r$, then

$$(3.1) \quad \frac{1}{n^{-(p_1+p_2)}} \left(\frac{1}{n^r n^{-p_1}} + \frac{1}{n^r n^{-p_2}} \right)^\alpha \leq \frac{2^\alpha}{n^{-(p_1+p_2)} n^{\alpha r} n^{-\alpha p_1}}.$$

If $\alpha > 0$ or $\alpha < -(p_1 + p_2)/p_1$, then a positive integer r can be chosen so that (3.1) converges to zero as $n \rightarrow \infty$.

For the bandwidth sequences $b_j(n) = b(n)$, Epanechnikov [4] found the optimum bandwidth sequence (minimizing the asymptotic relative global error) to be

$$b_0(n) \sim \left(\frac{kL^k}{nM_0} \right)^{1/(k+4)}$$

where k is the dimension, $L = \int_R W^2(x)dx$, and

$$M_0 = \int \dots \int \left[\sum_{i=1}^k \frac{\partial^2 f(x_1, \dots, x_k)}{\partial x_i^2} \right]^2 dx_1 \dots dx_k.$$

When L and M_0 are bounded and $M_0 \neq 0$, then there is no difficulty in showing that the optimum bandwidth sequence satisfies the conditions of this paper. Finally, it should be noted that a weight function $W(x)$ exists satisfying a Lipschitz condition of order $0 < \alpha < 1$ on $[a, b]$ but which is not of bounded variation.

THE FLORIDA STATE UNIVERSITY
UNIVERSITY OF SOUTH CAROLINA

REFERENCES

[1] Billingsley, P. (1968). *Convergence of Probability Measures*, Wiley, New York.
 [2] Cacoullos, T. (1966). Estimation of multivariate density, *Ann. Inst. Statist. Math.*, **18**, 179-189.
 [3] Chow, Y. S. (1966). Some convergence theorems for independent random variable, *Ann. Math. Statist.*, **37**, 1482-1493.
 [4] Epanechnikov, V. A. (1969). Nonparametric estimates of multivariate probability density, *Theory Prob. Appl.*, **14**, 153-158.
 [5] Nadaraya, E. A. (1964). On estimating regression, *Theory Prob. Appl.*, **9**, 141-142.
 [6] Nadaraya, E. A. (1970). Remarks on nonparametric estimates for density function and regression curve, *Theory Prob. Appl.*, **15**, 134-137.
 [7] Rosenblatt, M. (1971). Curve estimates, *Ann. Math. Statist.*, **42**, 1815-1842.
 [8] Taylor, R. L. and Cheng, K. F. (1978). On the uniform complete convergence of density function estimates, *Ann. Inst. Statist. Math.*, **30**, A, 397-406.
 [9] Watson, G. S. (1964). Smooth regression analysis, *Sankhyā*, **A**, **26**, 359-372.