

# A CHARACTERIZATION OF THE NORMAL LAW

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## Summary

If  $X, Y, Z$  are three random observations from a normal population with mean zero then the characteristic function of  $(X/Z, Y/Z)$  is  $\exp(-\sqrt{t^2+u^2})$ . It is shown in this paper that this property characterizes the normal law.

## 1.

Let  $X_1, X_2$  be two independent normal variates with zero mean and common variance. It is then well-known that the quotient  $X_1/X_2$  follows the Cauchy law distributed symmetrically about the origin. It is also well-known that we cannot obtain a characterization of the normal distribution by this property of the quotient [1]. A characterization of the generalized normal law (g.n.l.), that is, a distribution with frequency function

$$(1) \quad f(x) = \frac{\left(\frac{n}{2}\right)^{n/2}}{\Gamma\left(\frac{n}{2}\right)\sigma^n} |x|^{n-1} e^{-x^2/n/2\sigma^2}, \quad -\infty < x < \infty, \quad n \geq 1$$

(and hence the usual normal law) has been obtained in [2] where the following is proved: If  $X, X_1, X_2, \dots$  are independent observations from a population with a distribution function  $F(x)$  assumed to be continuous at  $x=0$  and if the frequency function of  $t_k = X / \sqrt{\frac{1}{k} \left( \sum_{i=1}^k X_i^2 \right)}$  is

$$\frac{\Gamma\left(\frac{n(k+1)}{2}\right)}{\left(\frac{n}{k^2}\right)\Gamma\left(\frac{n}{2}\right)\Gamma\left(\frac{kn}{2}\right)} |x|^{n-1} \left(1 + \frac{x^2}{k}\right)^{-n(k+1)/2} \quad -\infty < x < \infty, \quad n \geq 1$$

then necessarily  $X$  will have the frequency function  $f(x)$  defined in (1).

The object of this paper is to obtain a new characterization of the

g.n.l. in terms of three independent observations from the population. Of course, when  $n=1$  our characterization will be that of the ordinary normal law.

## 2.

LEMMA 1. *If  $X$  follows a g.n.l. then the characteristic function (c.f.) of  $\log |X|$  can never vanish.*

PROOF. If possible, let the c.f. of  $\log |X|$  vanish at a point  $w$ , that is,

$$0 = \int_0^{\infty} \exp \{ iw \log x \} x^{n-1} \exp \left\{ -\frac{nx^2}{2\sigma^2} \right\} dx.$$

The transformation  $x^2 = 2\sigma^2/n \cdot \alpha y$ ,  $\alpha > 0$  yields

$$0 = \int_0^{\infty} \exp \left\{ i \frac{w}{2} \log y \right\} y^{n/2-1} e^{-\alpha y} dy \quad \text{for all } \alpha > 0.$$

This identity in  $\alpha$  contradicts the completeness of the family of frequency functions  $e^{-\alpha x}$ ,  $x \geq 0$ ,  $\alpha > 0$ . Hence the lemma.

COROLLARY. *If  $Z$  follows the generalized Cauchy law (g.c.l.) given by*

$$g(Z) = \frac{\Gamma(n)}{(\Gamma(n/2))^2} |Z|^{n-1} (1+Z^2)^{-n}, \quad -\infty < Z < \infty, \quad n \geq 1$$

*then the c.f. of  $\log |Z|$  does not vanish.*

PROOF. Observe first that if  $X_1, X_2$  are independently distributed according to g.n.l. then their ratio is distributed according to g.c.l. Therefore distributionally  $Z$  can be conceived of as the ratio of two independent g.n.l.'s. Hence the c.f. of  $\log |Z|$  has the form  $|m(t)|^2$  where  $m(t)$  is the c.f. of  $\log |X|$ ,  $X$  being a random variable following g.n.l. That  $m(t)$  does not vanish has been established in Lemma 1. The assertion of the corollary now follows.

LEMMA 2. *If  $X_1, X_2, X_3$  are independent random variables and if the joint c.f. of  $(X_1 - X_3, X_2 - X_3)$  does not vanish, then it determines the distribution of the variables  $X_1, X_2$  and  $X_3$  except for additive constants.*

Note that the assumption that the joint c.f. of  $(X_1 - X_3, X_2 - X_3)$  does not vanish implies and is implied by the assumption that the c.f.'s of the variables  $X_k$ ,  $k=1, 2, 3$  do not vanish.

PROOF. Let  $\phi_k(t)$ ,  $k=1, 2, 3$  denote the c.f. of  $X_k$ ,  $k=1, 2, 3$  and let  $\xi(t_1, t_2)$  denote the joint c.f. of  $(X_1 - X_3, X_2 - X_3)$ :

$$\begin{aligned} \xi(t_1, t_2) &= E\{e^{it_1X_1+it_2X_2+i(-t_1-t_2)X_3}\} \\ &= \phi_1(t_1)\phi_2(t_2)\phi_3(-t_1-t_2), \quad -\infty < t_1, t_2 < \infty. \end{aligned}$$

Let  $Y_1, Y_2, Y_3$  be any other set of independent random variables such that the c.f. of  $(Y_1 - Y_3, Y_2 - Y_3)$  is  $\xi(t_1, t_2)$  defined earlier. Let  $\phi_k(t), k = 1, 2, 3$  be the c.f. of  $Y_k, k = 1, 2, 3$ . Hence  $\phi_1(t_1)\phi_2(t_2)\phi_3(-t_1-t_2) = \phi_1(t_1)\phi_2(t_2)\phi_3(-t_1-t_2)$ . From the assumption that  $\xi(t_1, t_2)$  does not vanish it follows that none of the c.f.'s  $\phi_k(t), \phi_k(t), k = 1, 2, 3$  vanish. Writing  $\phi_k(t) = Q_k(t)\phi_k(t), k = 1, 2, 3$  so that  $Q_k(t), k = 1, 2, 3$  is a complex valued function defined for  $-\infty < t < \infty$ , non-vanishing and satisfying  $Q_k(0) = 1, k = 1, 2, 3$ , we get

$$(2) \quad Q_1(t_1)Q_2(t_2)Q_3(-t_1-t_2) = 1.$$

Putting  $t_1 = t, t_2 = 0$ , we get  $Q_1(t) = 1/Q_3(-t)$ .

Putting  $t_1 = 0, t_2 = t$ , we get  $Q_2(t) = 1/Q_3(-t)$ .

Therefore

$$(3) \quad Q_1(t) = Q_2(t) = \frac{1}{Q_3(-t)}.$$

Substituting in (2) we get

$$Q_3(t_1+t_2) = Q_3(t_1)Q_3(t_2), \quad -\infty < t_1, t_2 < \infty.$$

The most general function  $Q_3(t)$  continuous on the whole line  $-\infty < t < \infty$ , non-vanishing and satisfying the condition  $Q_3(0) = 1$  is the exponential function  $Q_3(t) = e^{bt}, -\infty < t < \infty$  for some  $b$ , a complex number.

From (3) we obtain

$$Q_1(t) = Q_2(t) = Q_3(t) = e^{bt}.$$

Thus we have

$$\phi_k(t) = e^{bt}\phi_k(t), \quad k = 1, 2, 3.$$

Using the property of the c.f.'s  $\phi(-t) = \bar{\phi}(t)$  (complex conjugate), we see that  $\phi_k(t) = e^{ict}\phi_k(t), k = 1, 2, 3$  for some real  $c$ . Hence the assertion.

### 3.

Let  $\eta(t)$  be the c.f. of  $\log |X|$  where  $X$  follows g.n.l. with  $\sigma = 1$ . Recall that  $\eta(t)$  is non-vanishing. Let  $G(x, y)$  be the joint distribution of  $(X_1/X_3, X_2/X_3)$  where  $X_1, X_2, X_3$  are independent observations on  $X$ . Notice that the joint c.f. of  $(\log \{|X_1|/|X_3|\}, \log \{|X_2|/|X_3|\})$  is  $\eta(t)\eta(-t-u)\eta(u)$ .

**THEOREM.** *Let  $Z_1, Z_2, Z_3$  be three independent symmetric (about the*

origin) random variables with distribution functions continuous at zero. Then they all follow the g.n.l. if and only if the bivariate distribution function of  $(Z_1/Z_3, Z_2/Z_3)$  is  $G$ .

PROOF. Observe that the specification of the distribution of  $\log |Z_k|$ ,  $k=1, 2, 3$  determines the distribution of  $|Z_k|$ ,  $k=1, 2, 3$ . Therefore, below we show that under the conditions mentioned in the theorem the distribution of  $\log |Z_k|$ ,  $k=1, 2, 3$  is uniquely determined except for an additive constant which is same for all the variables.

Suppose  $Z_1, Z_2, Z_3$  satisfy the conditions of the theorem. Hence the joint c.f. of  $(\log \{|Z_1|/|Z_3|\}, \log \{|Z_2|/|Z_3|\})$  is  $\eta(t)\eta(-t-u)\eta(u)$ . If  $\theta_k(t)$  is the c.f. of  $\log |Z_k|$ ,  $k=1, 2, 3$ , then we get

$$\theta_1(t)\theta_2(u)\theta_3(-t-u) = \eta(t)\eta(u)\eta(-t-u).$$

Then  $\theta_1, \theta_2, \theta_3$  are non-vanishing. That the claim now follows from Lemma 2 is seen by taking  $\log |Z_k| = X_k$ ,  $k=1, 2, 3$ .

Note 1. If in the theorem the variables  $Z_1, Z_2, Z_3$  are assumed to be identically distributed then the assumption of symmetry of their common distribution is not necessary. This property can be derived as a consequence of the fact that the distribution of  $Z_1/Z_2$  is symmetric about the origin (c.f. [1]).

Note 2. Taking  $n=1$  one observes that the c.f. of  $G(x, y)$  is  $\exp(-\sqrt{t^2+u^2})$ . Thus the joint c.f. of  $(Z_1/Z_3, Z_2/Z_3)$  is  $\exp(-\sqrt{t^2+u^2})$  iff  $Z_1, Z_2, Z_3$  follow the same normal law.

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