

# ON THE FUNDAMENTAL THEOREM FOR THE DECISION RULE BASED ON DISTANCE || ||

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1. In the theory of decision rule based on distance || ||, which has been developed in the papers [1, 2], the probabilistic inequalities concerning the distance play a fundamental role. We have given in [1, 2] the following inequalities :

*For a given positive number  $\eta$  we have*

$$(I) \quad Pr(||F - S_n|| > \eta) \leq \frac{k-1}{n\eta^2}$$

and

$$(II) \quad Pr(||F - S_n|| > \eta) \leq 2ke^{-\frac{n\eta^4}{2k^2}}$$

where  $F$  denotes the discrete distribution of the random variable under observation,  $k$  the number of events the variable takes on and  $S_n$  the empirical distribution for  $n$  observations.

In this note we intend to give another inequality, which will serve better for a wide class of application than the above two. The comparison of it with the others and its applications will also be given.

2. THEOREM. *For any positive number  $\eta$  we have*

$$(III) \quad Pr(||F - S_n|| > \eta) \leq \frac{k^2 + k - 1}{(n\eta^2)^2} \leq \frac{1.25k^2}{(n\eta^2)^2}$$

*provided that  $k \geq 2$  and  $n > k$ .*

PROOF. Let  $F = \{p_1, p_2, \dots, p_k\}$  and  $S_n = \left\{ \frac{n_1}{n}, \frac{n_2}{n}, \dots, \frac{n_k}{n} \right\}$ .

Further, let

$$p_i > \frac{1}{n} \quad (i=1, 2, \dots, r)$$

$$p_j \leq \frac{1}{n} \quad (j=r+1, r+2, \dots, k)$$

and put  $r+s=k$ ,  $\sum_{i=1}^r p_i = p$ ,  $\sum_{j=r+1}^k p_j = q$ . Then, we have

$$r \geq 1.$$

For, if  $r=0$ , we should have

$$s=k$$

and

$$1 = \sum_{j=1}^k p_j \leq \frac{k}{n} < 1$$

which is a contradiction.

Now we have

$$\begin{aligned} E(\|F - S_n\|^4) &= E\left(\sum_{i=1}^k \left(\sqrt{\frac{n_i}{n}} - \sqrt{p_i}\right)^2\right)^2 \\ &= E\left(\sum_{i=1}^r \left(\sqrt{\frac{n_i}{n}} - \sqrt{p_i}\right)^2 + \sum_{j=r+1}^k \left(\sqrt{\frac{n_j}{n}} - \sqrt{p_j}\right)^2\right)^2 \\ &\leq E\left(\sum_{i=1}^r \frac{\left(\frac{n_i}{n} - p_i\right)^2}{p_i} + \sum_{j=r+1}^k \left(\sqrt{\frac{n_j}{n}}\right)^2\right)^2 \\ &= E\left(\sum_{i=1}^r \frac{\left(\frac{n_i}{n} - p_i\right)^2}{p_i} + \frac{m}{n}\right)^2 \quad \text{where } m = \sum_{j=r+1}^k n_j \\ &= \sum_{i=1}^r \frac{E\left(\frac{n_i}{n} - p_i\right)^4}{p_i^2} + \sum_{i \neq k} \frac{E\left(\frac{n_i}{n} - p_i\right)^2 \left(\frac{n_k}{n} - p_k\right)^2}{p_i p_k} \\ &\quad + 2 \sum_{i=1}^r \frac{E\left(\sum \left(\frac{n_i}{n} - p_i\right)^2 \left(\frac{m}{n} - q\right)\right)}{p_i} + 2q \sum_{i=1}^r \frac{E\left(\frac{n_i}{n} - p_i\right)^2}{p_i} \\ &\quad + E\left(\frac{m}{n} - q\right)^2 + q^2. \end{aligned}$$

Since

$$\begin{aligned} E\left(\left(\frac{n_i}{n} - p_i\right)^2\right) &= \frac{1}{n} p_i (1 - p_i), \\ E\left(\left(\frac{n_i}{n} - p_i\right) \left(\frac{n_j}{n} - p_j\right)^2\right) &= -\frac{1}{n^2} p_i p_j (1 - 2p_j), \end{aligned}$$

$$E\left(\left(\frac{n_i}{n} - p_i\right)^2 \left(\frac{n_j}{n} - p_j\right)^2\right) = \frac{1}{n^2} \left[ p_i p_j \{ (1-p_i)(1-p_j) + 2p_i p_j \} \right. \\ \left. + \frac{1}{n} p_i p_j \{ -1 + 2p_i + 2p_j - 6p_i p_j \} \right],$$

$$E\left(\left(\frac{n_i}{n} - p_i\right)^4\right) = \frac{1}{n^2} \left\{ 3\left(1 - \frac{1}{n}\right) p_i^2 (1-p_i)^2 \right. \\ \left. + \frac{1}{n} p_i (1-p_i) (1-3p_i + 3p_i^2) \right\} \\ = \frac{1}{n^2} \left\{ 3p_i^2 (1-p_i)^2 + \frac{1}{n} p_i (1-p_i) (1-6p_i + 6p_i^2) \right\},$$

we obtain

$$n^2 \left( \sum_{i=1}^r \frac{\left(\frac{n_i}{n} - p_i\right)^4}{p_i^2} + \sum_{i \neq j} \frac{E\left(\frac{n_i}{n} - p_i\right)^2 \left(\frac{n_j}{n} - p_j\right)^2}{p_i p_j} \right) \\ = \sum_{i=1}^r 3(1-p_i)^2 + \frac{1}{np_i} + \frac{1}{n} \{ -7 + 12p_i - 6p_i^2 \} \\ + \sum_{i \neq j} \left\{ (1-p_i)(1-p_j) + 2p_i p_j + \frac{1}{n} - 1 + 2p_i + 2p_j - 6p_i p_j \right\} \\ \leq (r-p)^2 + 2r - 4p + 2p^2 + r + \frac{1}{n} \{ -r^2 - 7r + 12p - 6p^2 + 4p(r-1) \} \\ = r^2 + r(3-2p) - 4p + 3p^2 + \frac{1}{n} \{ -r^2 - 7r + 4pr + 8p - 6p^2 \} \\ = r^2 + r - 1 + 2qr - 2q + 3q^2 + \frac{1}{n} \{ -r^2 - 7r + 4pr + 8p - 6p^2 \} \\ \leq r^2 + r - 1 + 2qr - 2q + 3q^2$$

$$n^2 \left( 2 \sum_{i=1}^r \frac{E\left(\frac{n_i}{n} - p_i\right)^2 \left(\frac{m}{n} - q\right)}{p_i} \right) = -2q \sum_{i=1}^r (1-2p_i) \\ = -2qr + 2pq = -2qr + 2q - 2q^2,$$

$$n^2 \left( 2q \sum_{i=1}^r \frac{E\left(\frac{n_i}{n} - p_i\right)^2}{p_i} \right) = 2nq \sum_{i=1}^r (1-p_i) \leq 2rs,$$

$$n^2 \left( E\left(\frac{m}{n} - q\right)^2 + q^2 \right) = nq(1-q) + n^2 q^2 \leq nq + (n^2 - n)q^2$$

$$= s + s^2 - \frac{s^2}{n}.$$

Therefore, we have

$$\begin{aligned} E(\|F - S_n\|^4) &\leq \frac{1}{n^2} \left\{ r^2 + 2rs + s^2 + r + s - 1 + q^2 - \frac{s^2}{n} \right\} \\ &\leq \frac{1}{n^2} \{(r+s)^2 + (r+s) - 1\} \\ &= \frac{1}{n^2} (k^2 + k - 1). \end{aligned}$$

As  $k \geq 2$ , the last term is less than or equal to  $1.25k^2$ . We thus obtain

$$P_r\{\|F - S_n\| > \eta\} \leq \frac{k^2 + k - 1}{(n\eta^2)^2} \leq \frac{1.25k^2}{(n\eta^2)^2}$$

3. *Comparison with the previous inequalities.* Now, we compare the inequality (III) with (I) and (II).

Put

$$A = \frac{k^2 + k - 1}{(n\eta^2)^2}, \quad B = \frac{k - 1}{n\eta^2}, \quad C = 2ke^{-n\eta^4/2k^2}$$

and denote by  $\alpha$  the upper bound which we want to set on  $P_r\{\|F - S_n\| > \eta\}$ . Actually we evaluate an upper bound of  $P_r\{\|F - S_n\| > \eta\}$  by  $A$ ,  $B$  or  $C$ . Therefore, for a given  $\alpha$  we take  $n$  such that at least one of  $A$ ,  $B$  and  $C$  becomes less than  $\alpha$ .

First we have :

$$\begin{aligned} A \geq B &\Leftrightarrow n\eta^2 \leq \frac{k^2 + k - 1}{k - 1} \\ &\Leftrightarrow B \geq \frac{(k - 1)^2}{k^2 + k - 1} \end{aligned}$$

and

$$\frac{(k-1)^2}{k^2+k-1} \begin{cases} = \frac{1}{5} & \text{when } k=2, \\ = \frac{4}{11} & \text{when } k=3, \\ \rightarrow 1 & \text{when } k \rightarrow \infty. \end{cases}$$

Therefore, when we take  $\alpha$  less than or equal to 0.2, or 0.36 if  $k \geq 3$ , and when we make  $A$  or  $B$  less than  $\alpha$ , it always holds that

$$A \leq B,$$

which means (III) is preferable to (I). When  $\alpha > 0.2$ ,  $A \geq B$  does not necessarily hold. We have  $A \leq B$  almost always for sufficiently large  $k$ . For example, we have  $A \leq B$  for  $k \geq 10$  and  $\alpha \leq 0.73$ .

Secondly, we have :

$$A \geq C \Leftrightarrow A \geq 2ke^{-\frac{k^2+k-1}{2nAk^2}} \Leftrightarrow A \geq 2ke^{-\frac{\sqrt{k^2+k-1}}{2k^2\sqrt{A}}\eta^2}$$

From the second relation it follows that  $A \leq \frac{1}{n \log 2k}$ , and from the last relation it can be seen that

$$(1) \quad 1 \geq \frac{4}{A} e^{-\frac{\sqrt{5}}{4\sqrt{A}}} \quad \text{for } k \geq 2 \text{ and } \eta^2 < 2,$$

$$(2) \quad 1 \geq \frac{4}{A} e^{-\frac{\sqrt{5}}{8\sqrt{A}}} \quad \text{for } \eta^2 \leq 1,$$

$$(3) \quad 1 \geq \frac{6}{A} e^{-\frac{\sqrt{11}}{9\sqrt{A}}} \quad \text{for } k \geq 3.$$

Now, when  $1 \geq A \geq 1/100$ , (1) does not hold, when  $1 \geq A \geq 1/500$  (2) does not hold and when  $1 \geq A \geq 1/400$  (3) does not hold. Therefore, (III) is always preferable to (II) for  $\alpha \geq 0.01$ , that is, when  $A$  or  $C$  can be greater than or equal to 0.01, although at least one of them remains less than  $\alpha$ . (III) is preferable to (II) for  $\eta^2 \leq 1$ ,  $\alpha \geq 0.002$  or for  $\alpha \geq 0.0025$  when  $k \geq 3$ . Even when  $\alpha < 0.01$ ,  $A \geq C$  does not necessarily hold. As  $k$  becomes larger and  $\eta$  smaller, the case  $A \leq C$  happens more frequently. For example, when  $\eta^2 < 1/2$ , we have  $A \leq C$  for  $\alpha > 1/5000$ . On the other hand, when we fix  $k$  and  $\eta$ , and make  $n$  large, (accordingly  $\alpha$  small), we have  $A \geq C$ . For example, when  $k \leq 10$  and  $\eta \geq 0.2$ , we have always  $A \geq C$  for  $\alpha \leq 2.5 \times 10^{-5}$ .

4. *Application.* Let  $\omega$  be a set of distributions which are defined on the same  $k$  events. Further, let  $F_0$  be the distribution of the random variable defined on the same events under observation, and  $\delta_n$  the empirical distribution on  $n$  observations of the variable. The problem then is to decide whether  $F_0$  is contained in  $\omega$  or  $F_0$  lies apart by  $\epsilon$  ( $> 0$ ) from  $\omega$ . This problem has been treated in various forms in [1, 2]. The decision is, however, made more precisely in a wide class of cases by

employing (III) than by employing (I) or (II). This will be illustrated by the examples below.\*

Let  $d$  denote  $\inf_{F \in \omega} ||F - S_n||$ . Then we have:

	$n$	$k$	$d^2$	$\frac{k-1}{nd^2}$	$\frac{k^2+k-1}{n^2d^4}$
1	253	9	0.0214	1.47	3.03
2	497	9	0.0699	0.2302	0.0737
3	300	9	0.410	0.656	0.00588
4	300	9	0.238	0.2738	0.0788
5	300	9	0.426	0.0626	0.00545
6	300	6	0.294	0.057	0.00398
7	300	6	0.032	0.521	0.0336
8	300	6	0.416	0.040	0.00199
9	270	9	0.296	0.100	0.0114
10	270	9	0.178	0.166	0.0316
11	270	9	0.238	0.124	0.0177
12	270	9	0.402	0.074	0.0062

Thus, we can decide with risk 0.05 that  $F_0$  is contained in  $\omega$  in examples 1, 2, 4, and  $F_0$  is not contained in  $\omega$ , that is, lies by  $2\sqrt{\frac{k^2+k-1}{n^2(0.05)^2}}$  apart from  $\omega$  in examples 3, 5, 6, 7, 8, 9, 10, 11 and 12.

For the above examples it can be seen at a glance that (III) is more precise than (I). It can also be seen easily that (III) is more precise than (II) for the above examples.

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#### REFERENCE

- [1] Matsusita, Kameo, Decision rules, based on the distance, for the problems of fit, two samples, and estimation, *Ann. Math. Stat.*, Vol. 26 (1955), pp. 631-640.
- [2] Matsusita, Kameo and Hirotugu AKAIKE: Decision rules, based on the distance, for the problems of independence, invariance and two samples, *Ann. Inst. Stat. Math.*, Vol. VII (1956), pp. 67-80.

\* As to these examples see the examples in [2].